

## Original Article

## A reduced risk of stroke with lithium exposure in bipolar disorder: a population-based retrospective cohort study

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**Objectives:** The risk of stroke is increased in patients with bipolar disorder. Lithium exhibits neuroprotective effects but the association between lithium use and the risk of stroke is unknown.

**Methods:** A population-based retrospective cohort study was conducted by utilizing the National Health Insurance Research Database in Taiwan. Subjects who had first been diagnosed with bipolar disorder between 2001 and 2006 were identified. A propensity score (PS) for receiving lithium was calculated with variables of age, gender, and comorbidities. The patients with bipolar disorder receiving lithium within the period from diagnosis through to December 2011 were designated as the lithium group ( $n = 635$ ). A 1:2 ratio was used to select PS-matched subjects with bipolar disorder without lithium use ( $n = 1,250$ ). Multivariate Cox proportional hazards regression models were used to explore the association, rather than causal inference, of lithium exposure and the risk of stroke.

**Results:** Of the 1,885 subjects, 86 (4.6%) experienced stroke, including 2.8% of the lithium group and 5.4% of the non-lithium group. Lithium use was associated with a significantly reduced risk of stroke [hazard ratio (HR) = 0.39, 95% confidence interval (CI): 0.22–0.68]. Reduced risks of stroke were also associated with the highest cumulative lithium dose  $\geq 720$  defined daily dose (DDD), HR = 0.25, 95% CI: 0.10–0.59], the longest cumulative exposure period ( $\geq 720$  days, HR = 0.20, 95% CI: 0.06–0.64), and the highest exposure rate ( $\geq 2$  DDD/day, HR = 0.39, 95% CI: 0.21–0.70).

**Conclusions:** Lithium use was significantly related to a reduced risk of stroke in patients with bipolar disorder.

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Stroke is one of the most severe medical conditions and it has been reported to account for 5.5 million deaths and 44 million disability-adjusted life-years lost worldwide annually (1). Previous studies have demonstrated an increased risk of stroke of up to

twofold among patients with bipolar disorder (2–4). However, the possible effects of medication on the risk of stroke were not further analyzed in these studies. Lithium was first suggested as a treatment for ‘psychotic excitement’ in 1949 (5)

and its efficacy in the management of bipolar disorder was established in the 1970s (6). In addition, lithium has been shown to increase neuronal density and gray matter volume in patients with bipolar disorder (7, 8).

*In vitro* and *in vivo* animal studies have demonstrated multiple neurotrophic and neuroprotective effects of lithium (9). The principal mechanisms underlying lithium's neuroprotective effects are as follows. (i) Inhibition of glycogen synthase kinase-3 (GSK-3), an enzyme linked to apoptosis induced by various neural insults (10). In animal stroke models, GSK-3 $\beta$  was found to be upregulated, and its inhibition by lithium was implicated as an important neuroprotective mechanism (11). (ii) Induction of brain-derived neurotrophic factor (BDNF), which plays a key role in neuronal development, plasticity, and survival (12). (iii) Increased expression of vascular endothelial growth factor (VEGF), which induces neuronal proliferation and neurovascular remodeling (13). (iv) Inhibition of *N*-methyl-D-aspartate (NMDA) receptor-mediated calcium influx and downstream apoptotic pathways, resulting in decreased glutamate excitotoxicity (14). (v) Inhibition of phosphoinositol phosphatase and regulation of autophagy (15), a critical process for neuronal survival and functioning.

Animal studies have also shown that lithium reduces neuronal injury after various insults, including ischemia (16–18), hemorrhage (19), and trauma (20). However, only one human study has investigated the neuroprotective effects of lithium after stroke, and the results were inconclusive, in that only the subgroup of patients with cortical ischemic stroke showed enhanced motor recovery with lithium treatment (21).

It is still unclear whether lithium has any effect on the susceptibility to cerebrovascular disease. The progression from atherosclerotic plaque formation to plaque rupture and ischemic stroke is driven by inflammation, extracellular matrix degradation, and neovascularization (22). Macrophages play an important role in the inflammatory process of atherosclerosis. In an animal study, lithium chloride induced macrophage apoptosis via inhibition of inositol monophosphatase (23). Lithium also reduced the expression of vascular cell adhesion molecule (VCAM)-1 and the infiltration of macrophages into the atherosclerotic plaque (24). This, in turn, decreased the macrophage load within the atherosclerotic plaques and implicated a plaque-stabilizing effect exhibited by lithium. Metalloproteinases participated in the initiation, progression, and rupture of atherosclerotic plaque and the formation of thrombus (25). Matrix

metalloproteinase (MMP)-9 was associated with abnormal degradation of the extracellular matrix when its production by macrophages was deregulated. Lithium treatment was found to suppress MMP-9 gene expression and reduce the level of this metalloproteinase (26), thereby reducing plaque instability. Vascular smooth muscle cell (VSMC) proliferation and migration in the intima of the arterial wall also play an important role in atheroma progression (27). VSMCs are also important mediators of abnormal neovascularization in the atherosclerotic plaque which expedites the process toward plaque rupture (28). Lithium inhibited VSMC proliferation and migration in an animal study (29). It also reduced blood glucose and cholesterol levels in mice fed a high-fat diet (24) and this may alleviate the detrimental effects of diabetes or hyperlipidemia on stroke events. The above mechanisms contribute to the beneficial effect of lithium on atherosclerosis and cerebral ischemia.

Lithium treatment has been demonstrated dramatically to reduce atherosclerotic plaque formation in mice (24). Another animal study discovered that the combination of low-dose lithium with captopril prevented stroke and improved survival in salt-loaded, stroke-prone spontaneously hypertensive rats, an animal model for cerebrovascular disease (30). However, no study has focused on the effect of lithium on the risk of stroke in human populations. If lithium were consistently to demonstrate neurotrophic and neuroprotective effects, this could be translated into clinical benefits for central nervous system disorders such as stroke, and would have significant implications for patients at risk of stroke and other cerebral insults.

We conducted a retrospective cohort study by analyzing a population-based dataset from the National Health Insurance Research Database in Taiwan. The aim was to determine whether lithium is related to the susceptibility to stroke and its occurrence in patients with bipolar disorder.

## Methods

### Database

The National Health Research Institutes (NHRI), Taiwan, maintain the National Health Insurance Research Database (NHIRD) for research purposes. We obtained a subset of the NHIRD with 1,000,000 random subjects, which accounted for about 5% of all enrollees in the National Health Insurance program. The random sampling process from the NHIRD to yield the 1,000,000-subject database was performed by the NHRI and

including the following main steps: (i) a serial number was assigned to each of the 23,251,700 subjects in the original population; (ii) a random number generator was used to generate 1,000,000 random numbers; (iii) the subjects, with their serial number matching the generated random numbers, were selected to form the database subset (31). There were no statistically significant differences in age, gender, or costs between the sample group and all other enrollees. The database contained information about medical claims for ambulatory care, inpatient care, dental services, and prescription drugs as well as registration files of the enrollees, insured from January 1996 to December 2011. The International Classification of Diseases, Ninth Revision (ICD-9) diagnostic system was incorporated into the data since the year 2000.

### Study design

The current study was an 11-year retrospective cohort study from January 1, 2001 to December 31, 2011. Subjects newly diagnosed with bipolar disorder (i.e., ICD-9 codes: 296.0, 296.1, 296.4–296.9 in three or more outpatient visits or any inpatient hospitalization records) within the interval of January 1, 2001 through to December 31, 2006 were first identified. The subjects were excluded if they had any diagnosis of stroke before the diagnosis of bipolar disorder. The patients with bipolar disorder were further divided into the lithium group and the non-lithium group, depending on lithium exposure from the prescription record. The lithium group and the non-lithium group, comprised the unmatched cohort in our study.

As the patients with bipolar disorder in the study were not randomly assigned to lithium treatment, we further utilized a propensity score (PS) to reduce potential confounding and selection biases. A multivariate logistic regression model was generated to predict the probability of receiving lithium, given the following set of covariates: age, gender, and comorbidities including diabetes mellitus, hypertension, hyperlipidemia, coronary artery disease, and chronic kidney disease. A PS was calculated for each subject and subsequently utilized to match the lithium group with the non-lithium group at a ratio of 1:2. Both the unmatched cohort and the PS-matched cohort were analyzed. The endpoints of the follow-up were the occurrences of the first stroke event (after lithium exposure in the lithium group), death, or end of study, on December 31, 2011.

To clarify further the association between lithium exposure and the risk of stroke in patients with bipolar disorder, we utilized three indexes –

the cumulative lithium dose, the cumulative lithium exposure period, and the lithium exposure rate – to describe these specific characteristics in the PS-matched cohort. The cumulative lithium dose of each subject was represented by the number of cumulative defined daily dose (DDD) of lithium. This value was acquired by dividing the total lithium dosage during the study period by the DDD of lithium (one DDD equals 24 mmol of lithium). The cumulative lithium dose was stratified into five subgroups: no use, <60 DDD, 60–359 DDD, 360–719 DDD, and  $\geq 720$  DDD. The cumulative lithium exposure period was the sum of the number of days for which each subject was prescribed lithium; it was stratified into five subgroups: no use, <60 days, 60–359 days, 360–719 days, and  $\geq 720$  days. The lithium exposure rate was the ratio of the cumulative lithium dose divided by the cumulative lithium exposure period; it was stratified into four subgroups: 0 DDD/day, <1 DDD/day, 1–2 DDD/day,  $\geq 2$  DDD/day.

### Definition of stroke and comorbidities

The stroke diagnosis was defined by ICD-9 codes 430–438 (ischemic: 433–437, hemorrhagic: 430–432) in any inpatient or three or more outpatient treatments during the study period. Patients who had a diagnosis of transient ischemic attack (ICD-9 code: 435) were excluded. Comorbidities in the 12-month period prior to the diagnosis of bipolar disorder, including diabetes mellitus (ICD-9 code: 250), hypertension (code: 401–405), hyperlipidemia (code: 272), chronic kidney disease (code: 585), coronary artery disease (code: 410–413), and congestive heart failure (code: 428), were confirmed by any inpatient or two or more outpatient treatments for the diagnoses in the NHI records.

### Statistical analysis

Distributions of lithium users and non-users in terms of age, gender, clinical comorbidities, use of typical antipsychotic agents, use of atypical antipsychotic agents, use of mood stabilizers including valproate and carbamazepine, and stroke were examined using Student *t*-tests for continuous variables and chi-square tests for categorical variables. Multivariate Cox proportional hazards regression models were used in both the unmatched cohort and the PS-matched cohort to explore the relationship between lithium exposure and the occurrence of stroke, adjusted for age, gender, comorbidities, use of antipsychotic agents, and use of mood stabilizers. Three additional Cox proportional hazards regression models were

implemented in the PS-matched cohort. *Model 1* was adjusted for age, gender, comorbidities, use of antipsychotic agents, use of mood stabilizers, and the cumulative lithium dose. *Model 2* was adjusted for age, gender, comorbidities, use of antipsychotic agents, use of mood stabilizers, and the cumulative lithium exposure period. *Model 3* was adjusted for age, gender, comorbidities, use of antipsychotic agents, use of mood stabilizers, and the lithium exposure rate. The proportional hazard assumptions were satisfied by testing graphically and by including the interaction of time with each covariate in the Cox regression models. The stroke-free survival curves of the lithium group and non-lithium group of both the unmatched cohort and the PS-matched cohort were plotted with the product-limit survival estimators for stroke. All statistical tests were two-sided, conducted at a significance level of 0.05, and reported using p values and 95% confidence intervals. All analyses were performed using the SAS software, version 9.1 (SAS Institute, Cary, NC, USA).

## Results

Data for a total of 3,681 subjects in the unmatched cohort and 1,885 subjects in the PS-matched cohort during the observation period from 2001 to 2011 were analyzed. The lithium group of the unmatched cohort was younger, consisted of more male subjects, and had significantly lower rates of diabetes mellitus, hypertension, hyperlipidemia, and coronary artery disease compared to the non-lithium group and had no cases of chronic kidney disease (Table 1). By contrast, the PS-matching process successfully removed these differences in the PS-matched cohort. The lithium group in the PS-matched cohort was not significantly different in terms of age, gender, and comorbidities compared to the non-lithium group (Table 1). Of note, none of the subjects in the PS-matched cohort had chronic kidney disease. The rates of use of typical antipsychotic agents, atypical antipsychotic agents, and mood stabilizers were all significantly lower in the non-lithium group compared to the lithium group in both the unmatched and the PS-matched cohort.

Among the total sample of 3,681 subjects in the unmatched cohort, 325 (8.8%) had stroke, including 18 (2.8%) among the lithium group and 307 (10.1%) among non-users. Among the 1,885 subjects in the PS-matched cohort, 86 (4.6%) had stroke, including 18 (2.8%) among the lithium group and 68 (5.4%) among non-users.

As shown in Table 2, after adjusting for possible confounders by using the multivariate Cox

proportional hazards regression model, lithium users in the unmatched cohort were associated with a significantly lower risk of stroke compared to non-users [hazard ratio (HR) = 0.43,  $p = 0.0010$ ]. The age-related risk of stroke was 1.06 times per decade older (HR = 1.06,  $p < 0.0001$ ). An increased risk of stroke was also associated with male gender (HR = 1.32,  $p = 0.0130$ ), diabetes mellitus (HR = 1.87,  $p < 0.0001$ ), hypertension (HR = 1.45,  $p = 0.0055$ ), and mood-stabilizer use (HR = 1.54,  $p = 0.008$ ). Neither typical (HR = 0.95,  $p = 0.6784$ ) nor atypical (HR = 1.05,  $p = 0.6992$ ) antipsychotic agent use was significantly associated with the risk of stroke.

In the PS-matched cohort, the lithium group also had a significantly reduced risk of stroke compared to non-users (HR = 0.39,  $p = 0.0010$ ). The age-related risk of stroke was 1.07 times per decade older (HR = 1.07,  $p < 0.0001$ ). No comorbidities were shown to be associated with the risk of stroke in the PS-matched cohort. Neither typical (HR = 0.81,  $p = 0.4392$ ) nor atypical (HR = 1.32,  $p = 0.2792$ ) antipsychotic agent use was significantly associated with the risk of stroke. The use of mood stabilizers was associated with an increased risk of stroke in this cohort (HR = 2.23,  $p = 0.0017$ ).

The risk of stroke associated with different levels of lithium exposure among the PS-matched cohort is listed in Table 3. A significantly reduced risk of stroke was demonstrated in the subgroup with a cumulative lithium dose  $\geq 720$  DDD (HR = 0.25,  $p < 0.01$ ), the subgroups with a cumulative lithium exposure period of 60–359 days (HR = 0.35,  $p < 0.05$ ) and  $\geq 720$  days (HR = 0.20,  $p < 0.01$ ), and the subgroup with the lithium exposure rate of  $\geq 2$  DDD/day (HR = 0.39,  $p < 0.05$ ).

The stroke-free survival curves of the lithium and non-lithium groups in the unmatched and PS-matched cohorts are shown in Figures 1 and 2, respectively. The stroke-free survival probabilities were significantly lower in the non-lithium group in both the unmatched cohort and the PS-matched cohort ( $p < 0.0001$  and  $p = 0.0080$ , respectively). The HRs within each subsequent year during the study period are also marked in the Figures 1 and 2. The HRs for stroke associated with lithium use reached statistical significance in the seventh and eighth year of follow-up in the unmatched cohort and the PS-matched cohort, respectively.

## Discussion

The current study showed that lithium treatment is significantly associated with a reduced risk of stroke in patients with bipolar disorder. The

Table 1. Characteristics of study participants

| Variables                            | Unmatched cohort               |                           |         | PS-matched cohort              |                           |         |
|--------------------------------------|--------------------------------|---------------------------|---------|--------------------------------|---------------------------|---------|
|                                      | Without lithium<br>(n = 3,046) | With lithium<br>(n = 635) | p-value | Without lithium<br>(n = 1,250) | With lithium<br>(n = 635) | p-value |
| Age, years, mean ± SD                | 44.5 ± 17.2                    | 36.0 ± 13.7               | <0.0001 | 37.0 ± 14.3                    | 36.0 ± 13.7               | 0.1610  |
| Gender, n (%)                        |                                |                           |         |                                |                           |         |
| Female                               | 1,800 (59.1)                   | 335 (52.8)                | 0.0030  | 692 (55.4)                     | 335 (52.8)                | 0.2830  |
| Male                                 | 1,246 (40.9)                   | 300 (47.2)                |         | 558 (44.6)                     | 300 (47.2)                |         |
| Diabetes mellitus, n (%)             |                                |                           |         |                                |                           |         |
| No                                   | 2,849 (93.5)                   | 624 (98.3)                | <0.0001 | 1,232 (98.6)                   | 624 (98.3)                | 0.6260  |
| Yes                                  | 197 (6.5)                      | 11 (1.7)                  |         | 18 (1.4)                       | 11 (1.7)                  |         |
| Hypertension, n (%)                  |                                |                           |         |                                |                           |         |
| No                                   | 2,561 (84.1)                   | 602 (94.8)                | <0.0001 | 1,185 (94.8)                   | 602 (94.8)                | 0.9980  |
| Yes                                  | 485 (15.9)                     | 33 (5.2)                  |         | 65 (5.2)                       | 33 (5.2)                  |         |
| Hyperlipidemia, n (%)                |                                |                           |         |                                |                           |         |
| No                                   | 2,886 (94.7)                   | 621 (97.8)                | 0.0010  | 1,221 (97.7)                   | 621 (97.8)                | 0.8740  |
| Yes                                  | 160 (5.3)                      | 14 (2.2)                  |         | 29 (2.3)                       | 14 (2.2)                  |         |
| Chronic kidney disease, n (%)        |                                |                           |         |                                |                           |         |
| No                                   | 3,017 (99.1)                   | 635 (100)                 | 0.0140  | –                              | –                         | –       |
| Yes                                  | 29 (0.9)                       | 0 (0)                     |         | –                              | –                         | –       |
| Coronary artery disease, n (%)       |                                |                           |         |                                |                           |         |
| No                                   | 2,882 (94.6)                   | 625 (98.4)                | <0.0001 | 1,233 (98.6)                   | 625 (98.4)                | 0.7110  |
| Yes                                  | 164 (5.4)                      | 10 (1.6)                  |         | 17 (1.4)                       | 10 (1.6)                  |         |
| Congestive heart failure, n (%)      |                                |                           |         |                                |                           |         |
| No                                   | 3,004 (98.6)                   | 630 (99.2)                | 0.2270  | 1,243 (99.4)                   | 630 (99.2)                | 0.5520  |
| Yes                                  | 42 (1.4)                       | 5 (0.8)                   |         | 7 (0.6)                        | 5 (0.8)                   |         |
| Typical antipsychotic agents, n (%)  |                                |                           |         |                                |                           |         |
| No                                   | 882 (29.0)                     | 72 (11.3)                 | <0.0001 | 338 (27.0)                     | 72 (11.3)                 | <0.0001 |
| Yes                                  | 2,164 (71.0)                   | 563 (88.7)                |         | 912 (73.0)                     | 563 (88.7)                |         |
| Atypical antipsychotic agents, n (%) |                                |                           |         |                                |                           |         |
| No                                   | 1,855 (60.9)                   | 150 (23.6)                | <0.0001 | 727 (58.2)                     | 150 (23.6)                | <0.0001 |
| Yes                                  | 1,191 (39.1)                   | 485 (76.4)                |         | 523 (41.8)                     | 485 (76.4)                |         |
| Mood stabilizers, n (%)              |                                |                           |         |                                |                           |         |
| No                                   | 1,949 (64.0)                   | 180 (28.3)                | <0.0001 | 742 (59.4)                     | 180 (28.3)                | <0.0001 |
| Yes                                  | 1,097 (36.0)                   | 455 (74.7)                |         | 508 (40.6)                     | 455 (74.7)                |         |
| Incident stroke events, n (%)        |                                |                           |         |                                |                           |         |
| No                                   | 2,739 (89.9)                   | 617 (97.2)                | 0.0030  | 1,182 (94.6)                   | 617 (97.2)                | 0.0100  |
| Yes                                  | 307 (10.1)                     | 18 (2.8)                  |         | 68 (5.4)                       | 18 (2.8)                  |         |

PS = propensity score; SD = standard deviation.

association appeared to be most prominent with a greater amount or longer duration of lithium treatment. Significantly reduced risks of stroke were observed in the subgroups with the highest cumulative lithium dose, the longest cumulative lithium exposure period, and the highest lithium exposure rate. This indicated that the dose, duration, and rate of lithium exposure might be related to the risk of stroke in patients with bipolar disorder.

Around 93% of the stroke cases in the PS-matched cohort were of ischemic origin. Additional analysis revealed that the HR of ischemic stroke associated with lithium use in the PS-matched cohort was 0.40 [95% confidence interval (CI): 0.22–0.71, p = 0.0019] after adjusting for age, gender, and related confounders. Lithium might alter the process of atherosclerosis via its effect on

inflammation, extracellular matrix degradation, and neovascularization.

Our findings that age, male gender, diabetes mellitus, and hypertension are associated with an increased risk of stroke in patients with bipolar disorder were compatible with two previous studies similarly focusing on patients with bipolar disorder (2, 3). Of note, our finding that female gender was associated with a reduced risk of stroke was consistent with these two studies but the underlying mechanism is not well understood. Lin et al. (3) also pointed out that renal disease was a risk factor for stroke. However, lithium treatment is generally avoided in patients with severe renal disease, owing to a heightened risk of toxicity (32). None of the 635 patients with bipolar disorder in the lithium group in the current study had chronic kidney

Table 2. Multivariate Cox proportional analyses for hazard ratio of stroke

| Variables                     | Unmatched cohort      |         | PS-matched cohort     |         |
|-------------------------------|-----------------------|---------|-----------------------|---------|
|                               | Hazard ratio (95% CI) | p-value | Hazard ratio (95% CI) | p-value |
| Lithium                       | 0.43 (0.26–0.71)      | 0.0010  | 0.39 (0.22–0.68)      | 0.0010  |
| Age (10 years)                | 1.06 (1.05–1.07)      | <0.0001 | 1.07 (1.06–1.09)      | <0.0001 |
| Gender (male/female)          | 1.32 (1.06–1.65)      | 0.0130  | 1.50 (0.98–2.32)      | 0.0639  |
| Diabetes mellitus             | 1.87 (1.37–2.54)      | <0.0001 | 2.60 (0.87–7.78)      | 0.0872  |
| Hypertension                  | 1.45 (1.12–1.89)      | 0.0055  | 1.56 (0.86–2.84)      | 0.1458  |
| Hyperlipidemia                | 0.85 (0.57–1.26)      | 0.4139  | 0.70 (0.20–2.41)      | 0.5659  |
| Chronic kidney disease        | 0.96 (0.42–2.19)      | 0.9125  | –                     | –       |
| Coronary artery disease       | 1.15 (0.82–1.61)      | 0.4114  | 1.29 (0.49–3.39)      | 0.6002  |
| Congestive heart failure      | 0.62 (0.27–1.42)      | 0.2599  | 0.59 (0.08–4.40)      | 0.6082  |
| Typical antipsychotic agents  | 0.95 (0.74–1.22)      | 0.6784  | 0.81 (0.48–1.37)      | 0.4392  |
| Atypical antipsychotic agents | 1.05 (0.82–1.34)      | 0.6992  | 1.32 (0.80–2.19)      | 0.2792  |
| Mood stabilizers              | 1.54 (1.19–1.97)      | 0.0008  | 2.23 (1.35–3.67)      | 0.0017  |

CI = confidence interval; PS = propensity score.

disease. This reflected the extreme precaution taken by doctors when prescribing lithium to patients with bipolar disorder and renal dysfunction.

By implementing PS matching, we successfully removed the differences in comorbidities between the original groups of patients with bipolar disorder. In the unmatched cohort, the subjects in the non-lithium group were significantly older and had a higher prevalence of diabetes mellitus, hypertension, hyperlipidemia, coronary artery disease, and chronic kidney disease, each of which alone was a risk factor for stroke (33, 34). This potentially may have caused an overestimation of the association between lithium exposure and the reduced risk of stroke as the lithium group had significant fewer risk factors at baseline. The finding that the risk of stroke remained significantly reduced in association with the lithium group in the PS-matched cohort strengthened our findings.

Psychiatric disorders such as depression (35) and schizophrenia (36) are also associated with a higher risk of stroke. Lithium may be used as an adjunctive treatment for subjects with mood swings or those at risk of suicide in these psychiatric patient groups. There is a possibility that lithium alters the risk of stroke in such patients but our results could not be extrapolated to subjects beyond the bipolar disorder group.

Various psychotropic agents may also alter the risk of stroke. No previous studies have focused on the association between medication use and the risk of stroke in patients with bipolar disorder. In our analyses of the unmatched cohort and the PS-matched cohort, we adjusted for the use of both typical and atypical antipsychotic agents, which were not associated with the risk of stroke in patients with bipolar disorder. From studies

targeting other disease populations, such as those with schizophrenia and dementia, the risk of stroke associated with the use of antipsychotic agents was found to be equivocal. Some studies reported no association between the use of either typical or atypical antipsychotic agents and the risk of stroke (37), whereas others reported that the use of typical antipsychotic agents was associated with an increased risk of stroke (38–40), and still others reported this association for atypical antipsychotic agents (41). This finding of no significant association may also be related to the distinctive effects of different individual antipsychotic agents on the risk of stroke. In addition, we found that the use of mood stabilizers was associated with an increased risk of stroke. No previous studies have evaluated the association between the risk of stroke and use of mood stabilizers in patients with bipolar disorder. However, in previous epilepsy studies, in which the mood stabilizers we used in the current study were used as anticonvulsants, a similarly increased risk of stroke associated with anticonvulsants use was demonstrated (42, 43). Further studies, with a broader range of subjects and a detailed differentiation of prescribed medication, are needed to fully delineate the effect of lithium treatment on the risk of stroke.

#### Strengths and limitations

One of the strengths of the current study was the use of the NHIRD, which included diagnostic coding and records of prescribed medication in detail. Lithium is a medication with clear clinical indications; we therefore considered the records of lithium prescriptions to be valid and reliable. The database also provided a sample with a large

## Lithium reduces risk of stroke in bipolar disorder

Table 3. Adjusted hazard ratios (HR) of lithium dosage or duration of stroke in PS-matched cohort

| Variables                          | n     | Model 1     |                        | Model 2     |                        | Model 3     |                        |
|------------------------------------|-------|-------------|------------------------|-------------|------------------------|-------------|------------------------|
|                                    |       | Adjusted HR | 95% CI                 | Adjusted HR | 95% CI                 | Adjusted HR | 95% CI                 |
| Age (10 years)                     |       | 1.07        | 1.06–1.09 <sup>a</sup> | 1.07        | 1.06–1.09 <sup>a</sup> | 1.07        | 1.06–1.09 <sup>a</sup> |
| Gender                             |       | 1.51        | 0.98–2.33              | 1.52        | 0.99–2.35              | 1.48        | 0.96–2.27              |
| Male                               | 858   |             |                        |             |                        |             |                        |
| Female                             | 1,027 |             |                        |             |                        |             |                        |
| Cumulative lithium dose            |       |             |                        |             |                        |             |                        |
| No use                             | 1,250 | 1.00        | Reference              |             |                        |             |                        |
| <60 DDD                            | 86    | 0.61        | 0.19–1.98              |             |                        |             |                        |
| 60–359 DDD                         | 151   | 0.49        | 0.19–1.24              |             |                        |             |                        |
| 360–719 DDD                        | 75    | 0.58        | 0.20–1.63              |             |                        |             |                        |
| ≥720 DDD                           | 323   | 0.25        | 0.10–0.59 <sup>b</sup> |             |                        |             |                        |
| Cumulative lithium exposure period |       |             |                        |             |                        |             |                        |
| No use                             | 1,250 |             |                        | 1.00        | Reference              |             |                        |
| <60 days                           | 179   |             |                        | 0.70        | 0.33–1.48              |             |                        |
| 60–359 days                        | 181   |             |                        | 0.35        | 0.14–0.88 <sup>c</sup> |             |                        |
| 360–719 days                       | 79    |             |                        | 0.36        | 0.09–1.51              |             |                        |
| ≥720 days                          | 196   |             |                        | 0.20        | 0.06–0.64 <sup>b</sup> |             |                        |
| Lithium exposure rate              |       |             |                        |             |                        |             |                        |
| 0 DDD/day                          | 1,250 |             |                        |             |                        | 1.00        | Reference              |
| <1 DDD/day                         | 24    |             |                        |             |                        | 1.05        | 0.24–4.58              |
| 1–2 DDD/day                        | 67    |             |                        |             |                        | 0.19        | 0.03–1.41              |
| ≥2 DDD/day                         | 544   |             |                        |             |                        | 0.39        | 0.21–0.70 <sup>c</sup> |
| Diabetes mellitus                  |       | 2.65        | 0.88–7.98              | 2.62        | 0.87–7.87              | 2.70        | 0.89–8.13              |
| Yes                                | 29    |             |                        |             |                        |             |                        |
| No                                 | 1,856 |             |                        |             |                        |             |                        |
| Hypertension                       |       | 1.54        | 0.85–2.81              | 1.54        | 0.85–2.82              | 1.53        | 0.84–2.80              |
| Yes                                | 98    |             |                        |             |                        |             |                        |
| No                                 | 1,787 |             |                        |             |                        |             |                        |
| Hyperlipidemia                     |       | 0.66        | 0.19–2.28              | 0.67        | 0.19–2.29              | 0.68        | 0.19–2.39              |
| Yes                                | 43    |             |                        |             |                        |             |                        |
| No                                 | 1,842 |             |                        |             |                        |             |                        |
| Coronary artery disease            |       | 1.32        | 0.50–3.48              | 1.28        | 0.49–3.38              | 1.26        | 0.48–3.33              |
| Yes                                | 27    |             |                        |             |                        |             |                        |
| No                                 | 1,858 |             |                        |             |                        |             |                        |
| Congestive heart failure           |       | 0.54        | 0.07–4.09              | 0.52        | 0.07–4.03              | 0.52        | 0.07–3.98              |
| Yes                                | 12    |             |                        |             |                        |             |                        |
| No                                 | 1,873 |             |                        |             |                        |             |                        |
| Typical antipsychotic agent        |       | 0.82        | 0.49–1.38              | 0.81        | 0.48–1.37              | 0.81        | 0.48–1.37              |
| Yes                                | 1,475 |             |                        |             |                        |             |                        |
| No                                 | 410   |             |                        |             |                        |             |                        |
| Atypical antipsychotic agent       |       | 1.31        | 0.79–2.18              | 1.34        | 0.81–2.21              | 1.32        | 0.79–2.19              |
| Yes                                | 1,008 |             |                        |             |                        |             |                        |
| No                                 | 877   |             |                        |             |                        |             |                        |
| Mood stabilizers                   |       | 2.27        | 1.38–3.72 <sup>b</sup> | 2.28        | 1.39–3.74 <sup>b</sup> | 2.23        | 1.35–3.67 <sup>b</sup> |
| Yes                                | 963   |             |                        |             |                        |             |                        |
| No                                 | 922   |             |                        |             |                        |             |                        |

CI = confidence interval; DDD = defined daily dose; PS = propensity score.

<sup>a</sup>p < 0.001.

<sup>b</sup>p < 0.01.

<sup>c</sup>p < 0.05.

number of subjects, allowing for analysis with strong statistical power. We focused on lithium use among people with bipolar disorder, thereby avoiding the bias of use by indication. PS matching and Cox proportional hazards regression models minimized possible confounding effects. In addition,

the present study was the first to take into consideration of the use of psychotropic medications when analyzing the risk of stroke in patients with bipolar disorder.

There were several limitations in the current study. One of the most important of these was that

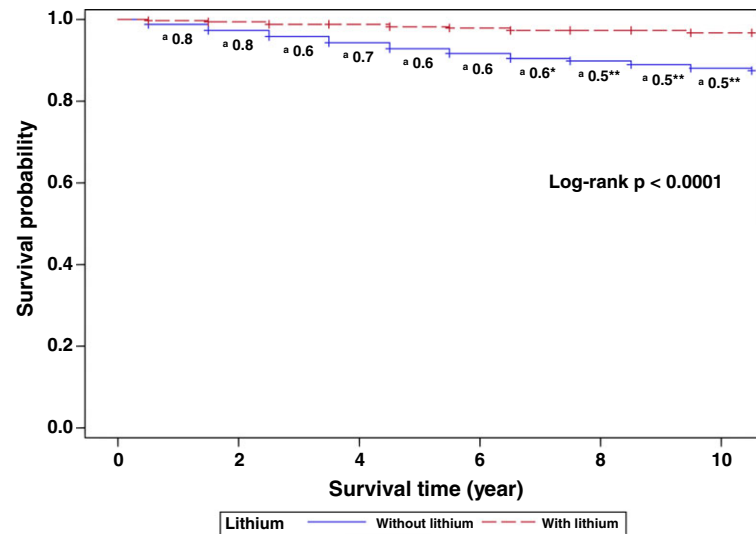


Fig. 1. The product-limit survival estimates for stroke in the unmatched cohort. <sup>a</sup>Hazard ratio. \*p < 0.50, \*\*p < 0.01.

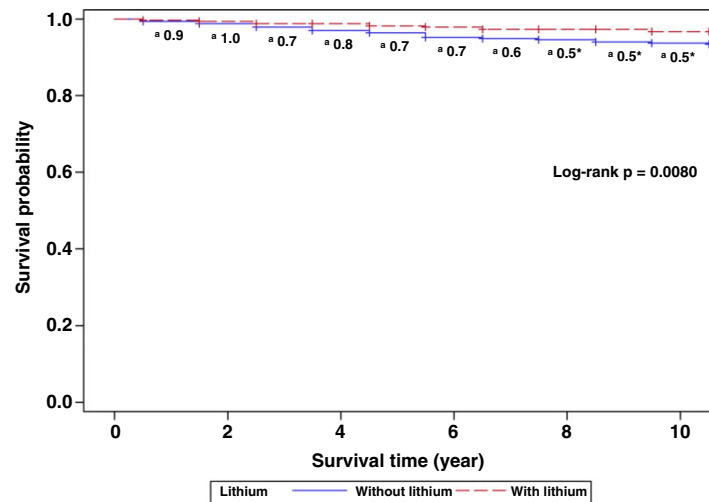


Fig. 2. The product-limit survival estimates for stroke in the propensity score-matched cohort. <sup>a</sup>Hazard ratio. \*p < 0.50.

the study design precluded causal inference, so the associations between lithium use and reduced risk of stroke should be interpreted with caution. In addition, because of the lack of information in NHIRD records of other possible risk factors for stroke, such as body mass index, cigarette smoking, and alcohol consumption, the current study could not control for these confounders. Due to the fact that lithium was rarely prescribed for patients with severe renal dysfunction, and that the lithium group in our cohorts contained no patients with chronic kidney disease, the results could not be extrapolated to patients with bipolar disorder with chronic kidney disease. Finally, the results should not be generalized to patients receiving lithium treatment for conditions other than bipolar disorder.

### Conclusions

Lithium use may be associated with a reduced risk of stroke in patients with bipolar disorder. This association may be strongest among patients with bipolar disorder receiving a higher dose, longer duration, and higher rate of lithium exposure. In clinical practice, the prevention of stroke by monitoring and early intervention for relevant cerebrovascular risk factors is of paramount importance for patients with bipolar disorder as this condition is itself a risk factor for stroke. Lithium treatment in patients with bipolar disorder may be a reasonable choice if the patient does not have chronic kidney disease. The duration or timing of discontinuation of lithium treatment may influence the association with the risk of stroke. As



the design of the current study precluded causal inference, future studies are indicated to clarify whether lithium exerts a protective effect on stroke and the possible underlying mechanisms if such a protective effect exists.

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