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- New procedures were developed to maximize EEG comparability across testing sites and platforms
- Test-retest reliability was examined for three EEG measures with 39 healthy adults in a multisite project
- Good-to-excellent test-retest reliability was observed across testing sites

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Demonstrating Test-Retest Reliability of Electrophysiological Measures for Healthy Adults in a Multisite Study of Biomarkers of Antidepressant Treatment Response

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Abstract

Growing evidence suggests that loudness dependency of auditory evoked potentials (LDAEP) and resting EEG alpha and theta may be biological markers for predicting response to antidepressants. In spite of this promise, little is known about the joint reliability of these markers, and thus their clinical applicability. New, standardized procedures were developed to improve the compatibility of data acquired with different EEG platforms, and used to examine test-retest reliability for the three electrophysiological measures selected for a multisite project—Establishing Moderators and Biosignatures of Antidepressant Response for Clinical Care (EMBARC). Thirty nine healthy controls across four clinical research sites were tested in two sessions separated by about one week. Resting EEG (eyes-open and eyes-closed conditions) was recorded and LDAEP measured using binaural tones (1000 Hz, 40 ms) at five intensities (60-100 dB SPL). Principal components analysis (PCA) of current source density (CSD) waveforms reduced volume conduction and provided reference-free measures of resting EEG alpha and N1 dipole activity to tones from auditory cortex. Low Resolution Electromagnetic Tomography (LORETA) extracted resting theta current density measures corresponding to rostral anterior cingulate (rACC), which has been implicated in treatment response. There were no significant differences in posterior alpha, N1 dipole or rACC theta across sessions. Test-retest reliability was .84 for alpha, .87 for N1 dipole, and .70 for theta rACC current density. The demonstration of good-to-excellent reliability for these measures provides a template for future EEG/ERP studies from multiple testing sites, and an important step for evaluating them as biomarkers for predicting treatment response.

Keywords: EEG, evoked potentials, Surface Laplacian, LORETA, biomarkers, reliability

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Electrophysiological Markers for Predicting Antidepressant Treatment Response

Despite the availability of pharmacologic treatments for major depressive disorder (MDD), high failure rates for specific treatments can introduce significant delays before relief is obtained from depression. Fortunately, there is growing evidence that electrophysiological measures of brain function show potential value as biological markers for predicting subsequent clinical response to antidepressants (Bruder et al., 2013). Of clinical relevance, measures such as the electroencephalogram (EEG) and evoked or event-related potentials (ERPs) provide the advantages of being noninvasive, widely applicable and economical, while providing information about neuronal generator patterns at scalp on a millisecond scale.

Resting EEG. Resting measures of spontaneous brain activity in the alpha and theta bands have shown particular promise as predictors of response to a range of antidepressants (see Alhaj et al., 2011; Bruder et al., 2013 for reviews). Greater alpha power prior to treatment, particularly identifiable at posterior scalp locations, is more likely to be observed in patients who subsequently respond to antidepressants than in nonresponders (Bruder et al., 2008; Prichep et al., 1993; Tenke et al., 2011; Ulrich et al., 1986). Some studies have also found that responders to a selective serotonin reuptake inhibitor (SSRI) differ from nonresponders in pretreatment alpha asymmetry (Arns et al., 2015a; Bruder et al., 2001; 2008), although this is not a universal finding (Tenke et al., 2011). Greater alpha over right compared to left frontal (Arns et al., 2015a) or across frontal, central and parietal regions (Bruder et al., 2001) was found in women who responded to an SSRI compared to nonresponders, and Bruder et al. (2008) found this difference in alpha asymmetry between SSRI responders and nonresponders over occipital locations. Greater alpha in SSRI responders, particularly over right posterior regions, may be indicative of reduced cortical arousal, which has been hypothesized to be associated with depression (Heller et al., 1995; 1997).

EEG alpha. The EEG alpha rhythm is a posterior oscillation at 8-13 Hz that is characteristic of a relaxed, wakeful state, and blocked (desynchronized) when visual processes are engaged by opening the eyes. However, the specific topography that is observed depends to no small degree on the chosen EEG reference, the impact of which may be irreversible (Fig. 1 of Tenke and Kayser, 2005; Tenke and Kayser, 2015). Feige et al. (2005) reported an inverse association between posterior alpha and the functional magnetic resonance imaging blood oxygenation level-dependent response in cortical visual regions. Alpha is also generated within the ventral visual stream, but its organization differs across regions (Bollimunta et al., 2008).

The stability of resting EEG alpha is consistent with a trait characteristic (Allen et al., 2004; Bruder et al., 2008; Hagemann et al., 2005; Smit et al., 2005; Stewart et al., 2014). While alpha differs between depressed patients and healthy control subjects, these differences persist following antidepressant treatment (Bruder et al., 2008; Pollock and Schneider, 1989). Condition-dependent posterior alpha (i.e., greater for eyes-closed than eyes-open) has also been observed to be greatest in individuals with a strong familial risk for depression (i.e., both parents having MDD; Bruder et al., 2005). In view of the success of antidepressants with serotonergic mechanisms, it is noteworthy that the inverse association between posterior alpha and physiological or emotional arousal (Heller et al., 1995; 1997) parallels the association between serotonergic activity and behavioral arousal (Jacobs and Azmitia, 1992).

EEG theta. In contrast to alpha, EEG theta is classically linked to limbic activity. This association is clearest in nonhuman rodent models, in which a highly regular theta rhythm is observed during active exploration (Vanderwolf, 1969), when it is synchronized to vibrissae movements (Semba and Komisaruk, 1984). However, midline frontal theta has also been studied in humans in demanding cognitive tasks and has been shown to be reliable across testing sessions (Iramina et al., 1996; McEvoy et al., 2000). In the resting EEG, frontal midline theta may also appear in close association with posterolateral low-frequency alpha (Tenke and Kayser, 2005).

For EEG theta, early conflicting reports on the direction of the difference predictive of a favorable treatment response have been supplemented by more consistent findings obtained using a model-dependent inverse, called Low Resolution Electromagnetic Tomography (LORETA; Pascual-Marqui et al., 1994), which has been used to infer current density through the rostral anterior cingulate cortex (rACC). Using this measure, patients who eventually responded to antidepressants showed increased pretreatment theta when compared to nonresponders (Korb et al., 2009; Mulert et al., 2007; Pizzagalli et al., 2001; although see Arns et al., 2015b). In addition, a multisite study (Leuchter et al., 2009) reported that an Antidepressant Treatment Response (ATR) index derived from a weighted combination of alpha and theta obtained before and one week after treatment onset may allow a differential prediction of response to a SSRI antidepressant as opposed to a noradrenaline-dopamine reuptake inhibitor (bupropion). However, this particular index is proprietary, and is only derived from forehead and earlobe electrodes, leaving its biophysical origins unknown.

Loudness dependency of auditory evoked potentials (LDAEP). In addition to resting EEG markers, ERP measures of brain activity elicited during sensory or cognitive processing (e.g., N1 or P3) have also been linked to clinical response to antidepressants (Bruder et al., 2013). The most replicated finding has been for the LDAEP, which refers to the monotonic increase in amplitude of N1 or P2 potentials with increasing tone intensity. These components predominantly reflect processing in modality specific cortical regions (Kayser and Tenke, 2006a,b; Tenke and Kayser, 2012; Tenke et al., 2010; Vaughan and Ritter, 1970). The change in component amplitude across intensities is viewed as an index of the gain of the auditory system to loudness (Hegerl and Juckel, 1993), thereby serving as a more selective measure of responsivity than that provided by resting alpha desynchronization.

Hegerl and Juckel (1993) reviewed evidence showing that the slope of the function relating tone loudness to the amplitude difference between successive component

peaks N1 and P2 provides an indicator of serotonergic activity. In this model, serotonergic neurons originating in dorsal raphé modulate activity in auditory cortex: a low firing rate in dorsal raphé is associated with a strong loudness dependency (steep LDAEP function), whereas a high firing rate is related to weak loudness dependency (shallow LDAEP function; Juckel et al., 1999). Depressed patients with pronounced LDAEP (putatively low serotonergic activity) prior to treatment responded better to a SSRI compared to patients with weak LDAEP (Gallinat et al., 2000; Hegerl et al., 2001; Lee et al., 2005; Page et al., 1994). Studies have not, however, found LDAEP to be related to severity of current depressive symptoms and improvement of depression following treatment was not associated with a change in LDAEP, which suggests that it is not state-dependent (Gallinat et al., 2000; Linka et al., 2009). Likewise, the specificity of LDAEP for predicting response to SSRI as opposed to antidepressants with a different mode of action is still in question. As a result, LDAEP is a promising predictor of response even to non-serotonergic treatments (O'Neill et al., 2008). Some studies have suggested that LDAEP may differentially predict clinical response to a SSRI as opposed to a noradrenaline reuptake inhibitor (reboxetine) (Linka et al., 2004, 2005; Mulert et al., 2007), but further study is needed concerning its specificity for SSRI antidepressants.

Reliability of Electrophysiological Measures

Resting EEG. The gross morphology and local topography of EEG waveforms are useful for applications in clinical neurology, but more nuanced neurologic and psychopathologic applications require equally nuanced quantitative methods and measures (e.g., Duffy et al., 1994). For a marker of clinical response to be viable, the measure must also have good test-retest reliability. Classifications of EEG spectral patterns have been reported to be stable at 12 – 40 month retest (Napfli et al., 2007), and high test-retest correlations have been reported for broad band spectral amplitude measures ($r = .92$ at 5 min, $.84$ at 12-14 wk; Salinsky et al., 1991). Resting EEG alpha and theta power at frontal and more posterior electrodes have shown high test-retest reliability (between $.82$ and $.97$) in both healthy adults (Smit et al., 2005;

Tomarken et al., 1992) and depressed patients (Bruder et al., 2008). In a report of alpha as a predictor of antidepressant treatment response, alpha amplitude and asymmetry were unchanged following treatment (Bruder et al., 2008). Retest reliability of alpha asymmetry has, however, been found to be lower (between .41 and .76) in both healthy adults (Debener et al., 2000; Hagemann et al., 2002; Tomarken et al., 1992) and depressed patients (Bruder et al., 2008; Allen et al., 2004). Inasmuch as the measures for the present study are based on EEG amplitudes, asymmetry measures are of secondary interest and will therefore be presented as Supplementary Material (Table S1).

LDAEP. Hegerl, Gallinat & Mrowinski (1994) quantified N1-P2 as the strength of the tangential equivalent dipole corresponding to superior temporal cortex within the Sylvian fissure (cf. Tenke & Kayser, 2012). The loudness dependency of this measure showed high reliability ($r = 0.88$) when retested after three weeks. A more recent study (Hensch et al., 2008) reported comparable high reliability for nose-referenced vertex AEP measures. For 62 healthy adults, test-retest reliability of the N1/P2 peak-to-peak amplitude ranged from .59 to .89, and N1/P2 LDAEP slope showed a reliability between .78 and .87. Beaudeau et al. (2000) reported that the use of temporal PCA (tPCA) to derive ERP component measures for N1 and P2 at mid-central sites improved test-retest reliability over a 2-4 week interval when compared to baseline-to-peak ERP measures (for N1, .42 to .52 vs. .06 to .38; for N1/P2, .76 to .80 vs. .59 to .77).

LORETA rACC theta. Cannon et al. (2012) reported high test-retest reliabilities after 30 days for total resting EEG power and coherence in traditional EEG bands. In the same study, reliabilities for LORETA-based measures across frequency bands and participants was also high, including measures for left rACC (BA 32; Cronbach's alpha = 0.65 and 0.98 for eyes-closed and eyes-open, respectively).

Rationale for present study and selection of measures. The aim of this study was to evaluate the test-retest reliability of the electrophysiological predictors used in an ongoing multisite study of antidepressant treatment response Establishing

Moderators and Biosignatures of Antidepressant Response for Clinical Care [EMBARC]; Trivedi et al., 2016). It was important to first establish the test-retest reliability of all of the electrophysiological measures in healthy adults, in order to preclude the possible impact of change in clinical state of patients over time (spontaneous or treatment-related). Moreover, EEG predictors in at least one report (Leuchter et al., 2009) rely on EEG changes between baseline and week one of treatment, making the test-retest reliability at week one of critical importance.

For resting EEG alpha and LDAEP, current source density (CSD) measures were derived to avoid problems associated with the choice of a recording reference. This approach reduces volume conduction from distant locations, while representing the strength of the current generators underlying the topography (Kayser and Tenke, 2006a,b; Tenke and Kayser, 2005, 2012).

For EEG alpha, the corresponding CSD topographies are predominantly posterior (Tenke et al., 2011), without the computational bias that causes various reference schemes to misallocate it to anterior regions (Tenke and Kayser, 2015). These CSD measures were then quantified using frequency-based PCA (fPCA) methods. In a similar manner, temporal PCA (tPCA) provided component measures for the LDAEP paradigm, yielding measures with larger effect sizes and increased reliability when compared to peak or time window estimates (Beauducel et al., 2000; Kayser et al., 1998; Kayser and Tenke, 2015b).

In addition to CSD-fPCA measures, LORETA was also applied to the resting EEG data to derive measures of current density attributable to the region of the rACC. Although different versions of the LORETA algorithm have been described (Pascual-Marqui, 2002; Pascual-Marqui, 2007), we elected to use the one originally described (Pascual-Marqui et al., 1994), since this was the method used in prior studies linking rACC theta current density to antidepressant response (Korb et al., 2009; Mulert et al., 2007; Pizzagalli et al., 2001).

Data Unification Across EEG Platforms

Multisite studies pose a number of difficult challenges for efforts to pool or

equate data across testing sites, which may have different EEG acquisition hardware systems and which may rely on distinct software and technical methods to acquire, quantify and analyze the EEG data. As a result, findings in the literature are frequently embedded in hardware-specific domains, with little or no effort to generalize them across platforms. This oversight imposes implicit limits on the clinical applicability of reported findings. The present multisite study addressed these challenges by developing : (1) a standardized EEG procedure manual to maximize the comparability of data collected at four laboratories across the United States with different EEG recording montages and platforms; (2) data interpolation to a common montage and sample rate; and (3) implementation of a single, standardized analysis pipeline to process these data.

Methods

Participants

A total of 39 healthy adults (24 female) were tested as part of the EMBARC project (Trivedi et al., 2016), with participants locally recruited and tested at each of four research testing sites: Columbia University Medical Center (CU) in New York ($n = 10$, 6 female), University of Texas Southwestern Medical Center (TX) in Dallas ($n = 10$, 5 female), Massachusetts General Hospital (MG) in Boston ($n = 10$, 6 female)ⁱ, and University of Michigan (UM) in Ann Arbor ($n = 9$, 7 female). The participants at these testing sites did not differ significantly in mean age ($F[3,35] = 1.00$, $p = .40$) or gender ratio. Participants were recruited using advertisements in local newspapers or online, flyers or posters. After a telephone screening, adults aged 18-65 of all races and ethnicities were invited to participate. A trained rater administered the Structured Clinical Interview for DSM-IV Axis I Disorders, Non-patient Edition (First et al., 1996), obtained information about psychiatric and medical history, reviewed eligibility, explained study procedures, and answered questions about the study. Blood samples were drawn from eligible participants to test for

hematology, liver, thyroid and kidney function, and urinalysis was used as a drug screen, and as a pregnancy test for women of child-bearing age. Participants also completed self-rating scales, including the Quick Inventory of Depressive Symptoms (QIDS-SR; Rush et al., 2003), Antidepressant Treatment History Questionnaire (ATRQ; Chandler et al., 2010), and Edinburgh Inventory (Oldfield, 1971) to assess handedness. The study was approved by the Institutional Review Board at each testing site and all participants signed an informed consent form.

Inclusion criteria included: 18-65 years old, QIDS-SR score of less than 8, fluent in English, and capacity to understand the nature of the study and provide written informed consent. Exclusion criteria included: (1) Current or lifetime history for major depression, bipolar disorder, schizophrenia or other Axis I psychotic disorders; (2) any current Axis I or Axis II diagnoses except for nicotine or caffeine dependence; (3) meeting DSM-IV criteria for substance dependence in last 6 months (except for nicotine) or substance abuse in the last 2 months; (4) positive urine drug screen at evaluation; (5) any current history of an unstable general medical condition deemed to be clinically significant; (6) epilepsy or other conditions requiring an anticonvulsant; (7) any clinically significant abnormal laboratory results. The 39 participants (24 female/15 male) who met the inclusion criteria had a mean age of 37.6 years (SD= 14.8) and mean education of 15.7 years (SD= 4.4). Their mean handedness laterality quotient (LQ) on the Edinburgh Handedness Inventory (Oldfield, 1971) was 70.6 (SD= 51.4), with all but 5 being right-handed (LQ > 0).

Procedures

Although waveshape distortion caused by the use of filters with different properties are primary concerns for ERP, its impact can be mitigated by coupling broad band acquisition methods with offline filtering using narrow parameters sufficient for the final measures. However, resting EEG measures may also be influenced by the signal-to-noise properties of the disparate platforms. Since one of the aim of EMBARC is to increase sample sizes by pooling across testing sites, we developed methods to uniquely counter or account for differences between testing sites. This approach

thereby maximizes the impact of individual differences and participant groupings across consecutive EEG recordings.

The most fundamental distinction between recording systems is between those capable of recording low frequency EEG activity down to 0 Hz (i.e., DC) and those that cannot. These differences are not merely reflections of amplifier construction, since they require unique electrodes as well (e.g., Ag-AgCl instead of tin). Such fundamental distinctions affect the low-frequency content of EEG spectra by their differential sensitivity to slow drifts, both physiological (e.g., skin conductance) and artifactual (e.g., electrode polarization) in origin. The present generation of EEG systems adds to this the distinction between acceptable electrode-to-scalp impedances required of high- and low-impedance amplifiers, as well as passive vs. active electrodes (i.e., sources followers, with or without balancing currents). Even though recording systems have considerably improved over time, it is clearly preferable to completely eliminate testing site-based noise differences wherever possible. For resting EEG alpha, this goal is facilitated by the use of testing site-independent methods, coupled with the use of testing site control factor. In contrast to resting EEG, signal averaging alone is generally sufficient for LDAEP.

Test-retest sessions. Electrophysiological tests were administered in a baseline session of 1-2 hours and the tests were repeated after 5 to 16 days (mean= 7.6 ± 2.5 days). An effort was made to retest individuals at about the same time of day as session 1. The average local time of day when subjects were tested in session 1 ($1:14\text{pm} \pm 2\text{hr}, 22\text{min}$) did not differ from session 2 ($12:52\text{pm} \pm 2\text{hr}, 24\text{min}$; $t[38] = 1.10, p=.28$).

Resting EEG. EEG was recorded while participants sat quietly during four 2-min periods, half with eyes open (O) and half with eyes closed (C) in a counterbalanced order (OCCO). Participants were instructed to remain still and inhibit blinks or eye movements during each period. During the eyes open condition, participants fixated on a central cross on the monitor.

LDAEP. Participants sat quietly with their eyes open, fixating on a central cross

during each of 5 blocks of 100 trials (about 5 min per block), while binaural tones (1000 Hz, 40 ms duration with 10 ms rise and decay time) were presented at five intensities (60, 70, 80, 90, 100 dB SPL) in a pseudorandomized order with inter-stimulus intervals ranging from 1600 – 2100 ms using Presentation (Neurobehavioral Systems, 2007). Each stimulus intensity was repeated 100 times for a total of 500 trials. Calibration of the output of headphones (in dB SPL) for the five tone intensities in the LDAEP paradigm was done using a sound level meter with a coupler appropriate for the headphones (CU, TX, UM) or ear inserts (MG).

EEG Acquisition

Inter-site standardization. All personnel responsible for administering the above tests used an EEG Procedure Manual designed to standardize test administration, including a set of instructions to participants at scheduling of the tests (e.g., emphasizing to have a good night's sleep and avoid drinking alcohol the night before, minimizing caffeine and nicotine on test day, and making sure hair was clean and dry), instructions to participants prior to each test, and detailed instructions to experimenters concerning the administration of EEG and LDAEP tests. Each of the experimenters at all testing sites required certification by the Columbia lab after demonstrating EEG cap placement and task instruction via video conference and submitting satisfactory EEG data acquired from a volunteer.

The continuous EEG data were acquired using different recording equipment at each of the four research testing sites, with acquisition filters set to broadband cutoffs to approximate DC-50Hz (or greater). The acquisition methods will be described for CU, followed by variations for each of the other testing sites. To enhance inter-site comparability, the location of the recording electrode montage was optimized in all cases using direct measurements of electrode locations corresponding to landmarks of the 10-20 system (nasion, inion, auditory meati, vertex). Feedback was provided to each testing site to identify and minimize artifacts, bad channels, and electrode bridging as soon as possible to allow for correction of technical errors.

CU acquisition methods. The electrode montage consisted of 72 expanded 10-20 system scalp channels (Pivik et al., 1993) on a Lycra stretch electrode cap (Electro Cap International, Inc.) including 12 midline locations (Nose, Nz to Iz) and 30 homologous pairs over the left and right hemisphere, extending laterally to include the inferior temporal lobes. Signals from the Ag/AgCl electrodes were recorded using an active reference (ActiveTwo EEG system) at electrode locations PPO1 (common mode sense) and PPO2 (driven right leg), but monitored using a nose reference. The scalp placements were prepared using a conventional water soluble electrolyte gel and the interface was verified by the acquisition software (ActiView), with additional care taken to avoid electrolyte bridges (Tenke and Kayser, 2001; Alschuler et al., 2014). Continuous EEG was acquired at 256 samples/s (bandwidth: DC - 51.3 Hz at 3 dB attenuation; -20.5 dB at 128 Hz) using the 24-bit Biosemi system and the raw data files were saved in the native (.bdf) format. Amplifier calibration was accomplished through saline between each active electrode and CMS-DRL using a 100 μ V, 100 ms square pulse (2 s ISI).

MG acquisition methods. The electrode montage consisted of a 128-channel geodesic net (24-bit, Electrical Geodesics, Inc.; EGI), including 10 midline locations (Nz to Iz) and 52 homologous pairs over the left and right hemisphere, extending laterally to include the two mastoids (below the 10-20 landmarks) and recorded using a Cz reference (a nose channel was not included). The montage also included 2 electrodes below each ear and 5 on each side of the face. The scalp electrodes were prepared using a saline solution, with scalp connectivity verified by the 24-bit acquisition software (Netstation), with additional care taken to avoid electrolyte bridges. To facilitate subsequent interpolation to a common montage, particular care was taken to optimize the montage based on landmarks of the 10-20 system (nasion,inion, auditory meatus, vertex). Continuous EEG was acquired from Ag/AgCl electrodes at 250 samples/s and the raw data files were saved in the native (.raw) format. A 60 Hz notch filter was used with a DC – 100 Hz bandpass. Amplifiers were calibrated using a 20 Hz, 5000 μ V sine wave into the amplifier input.

TX acquisition methods. The electrode montage consisted of 62 expanded 10-20 system scalp channels on a Lycra stretch electrode cap including 8 midline locations (Nz to Iz) and 27 homologous pairs over the left and right hemisphere, extending laterally to include the two mastoids, recorded using a nose reference. Continuous EEG was acquired from Ag/AgCl electrodes at 250 samples/s using the 32-bit Neuroscan Synamp system and the raw data files were saved in the native (.cnt) format. Data were recorded at DC – 100 Hz with a 60 Hz notch. Amplifier calibration used a 20 Hz, 50 μ V sine wave into the amplifier input.

UM acquisition methods. The electrode montage consisted of 60 expanded 10-20 system scalp channels on a Lycra stretch electrode cap including 8 midline locations (FPz to Oz) and 26 homologous pairs over the left and right hemisphere, extending laterally to include the two mastoids, recorded using a nose reference. Continuous EEG was acquired from tin electrodes at 250 samples/s using the 32-bit Neuroscan Synamp system and the raw data files were saved in the native (.cnt) format. Data were recorded at .5 – 100 Hz with a 60 Hz notch. Amplifier calibration used a 20 Hz, 50 μ V sine wave into the amplifier input.

Preprocessing Pipeline for Continuous EEG

The preprocessing strategy for continuous resting EEG and LDAEP data is shown in Fig 1. In the first step, data were converted from their native formats to BDF format using EEGLAB (Delorme and Makeig, 2004) and a custom Matlab script to preserve the original data acquisition gain. Bad channels were then identified from runtime notes and visual inspection of the continuous data, as was the overall integrity of the data, taking particular note of block transitions, missing or unusable periods within blocks, or nonstandard blocks or files owing to technical errors. In the third step, PolyRex (Kayser, 2013a) was used to remove DC offsets, remove drift across each block via a polynomial filter, re-reference to a nose-tip reference, optimize data scaling if data representation of the native acquisition format exceeded the range for a common 16-bit A/D conversion, and convert to 16-bit Neuroscan (.cnt) format that included the CU 72-channel montage.

(insert Figure 1 about here)

After the channel montage has been created, missing or bad channels, including those identified as bridged (Tenke and Kayser, 2001; Alschuler et al., 2014), were replaced by spherical spline interpolation (Perrin et al., 1989) from the remaining electrode locations (step 4). Interpolated channels included all CU channels that were not in the UM and TX montages, and all 72 channels for the overlapping 128-channel MG data. EEG data were not further analyzed if 20% or more of the electrodes were bad. Note that the backward paths in Fig. 1 indicate possible breaks in the processing stream for iterative interactive data handling by the technician in the case of recording errors that required the removal of data between blocks or the flagging of additional channels as bad due to excessive drift or intermittent contact.

The final preprocessing step for continuous EEG data was blink-correction using a spatial, singular value decomposition (NeuroScan, 2003). Bipolar EOG recordings (horizontal: lateral to outer canthi; vertical: above and below right eye), were interpolated using spherical splines (Perrin et al., 1989) as an aid in identifying blinks and eye movements during visual inspection and validation of rejected artifacts (see footnote 1 in Kayser and Tenke, 2015b).

Data Segmentation and Processing of Resting EEG and LDAEP Epochs

Blink-corrected data were segmented into 2-s epochs (75% overlap) for the resting EEG, or into stimulus-locked epochs (-200 to 1000 ms) for the LDAEP. Resting EEG data were band-passed at 1-60 Hz (24 dB/octave), and LDAEP data low-passed at 50 Hz (24 dB/octave). Channels containing artifacts or noise for any given epoch were identified using a semiautomated reference-free approach to identify isolated EEG channels containing amplifier drift, residual eye activity, muscle or movement-related artifacts on a trial-by-trial basis (Kayser and Tenke, 2006d). If 25% or more of all channels were identified as containing artifact, the trial was

rejected. Otherwise channels containing artifact were replaced by spline interpolation (Perrin et al., 1989). For LDAEP, ERP averages were computed for all accepted trials, baseline corrected, and low-passed at 12.5 Hz (12 dB/octave). If required, epoched data were adjusted to 256 samples/s using a temporal spline interpolation. For the resting EEG, an additional automated step was included to reject any remaining epochs exceeding a 100 μ V threshold on any channel (including uncorrected EOG channels), thereby removing from consideration epochs containing well defined blinks.

CSD. EEG epochs and ERPs were transformed into reference-free CSD estimates (μ V/cm²) using a spherical spline surface Laplacian ($m = 4$; $\lambda = 10^{-5}$; Kayser and Tenke, 2006a; Perrin et al., 1989; Tenke et al., 2011). CSD estimates represent the magnitude of the radial current flow entering and leaving the skull and scalp from the subjacent dura (Nunez, 1981; Nunez and Srinivasan, 2006), and thereby identify the direction, location and intensity of current generators underlying a surface potential topography (Mitzdorf, 1985; Nicholson, 1973; Tenke and Kayser, 2012). CSD is a true reference-free technique in that any EEG reference scheme provides identical CSD estimates, which resolves the ubiquitous problem of arbitrarily choosing a reference (Kayser and Tenke, 2010, 2015a).

Principal Components Analysis (PCA). The averaged EEG/CSD spectra and ERP/CSD waveforms were separately submitted to frequency (spectra) or temporal (waveforms) PCA derived from the covariance matrix, followed by unrestricted Varimax rotation of the covariance loadings (Kayser and Tenke, 2003, 2006a; Tenke and Kayser, 2005). This approach determines common sources of variance in the original EEG/ERP data or their reference-free transformations in the form of distinctive PCA components (factor loadings) and corresponding weighting coefficients (factor scores), and thereby provides a concise, efficient simplification of the spectral or temporal pattern and spatial distribution of surface potentials (EEG/ERP) or their neuronal generators (CSD). PCA-based estimates provide superior measures (e.g., larger effect sizes, increased internal consistency, better

test-retest reliability) when compared to peak-to-peak amplitudes (Beauducel et al., 2000; Beauducel and Debener, 2003) or integrated time window amplitudes (Kayser et al., 1997; 1998; Kayser and Tenke, 2015b).

The correspondence between the spectral pattern or time course and topography of the extracted orthogonal factors, in conjunction with the observed CSD spectra or waveforms, allows identification and measurement of complex, physiologically-relevant CSD components for further analysis (i.e., only a limited number of meaningful, high variance CSD factors are retained for further statistical analysis; for complete rationale, see Kayser and Tenke, 2003; 2005; 2006a; 2006c). At the same time, the CSD-PCA approach provides additional protection against artifacts (i.e., extracting EMG and EOG as distinct components), reduce the impact of noise and eliminate reference-related errors (e.g., reversed local asymmetries with weak rhythmicity; Tenke and Kayser, 2005).

CSD-fPCA for resting EEG. Data from one participant were eliminated because of topographic distortion owing to excessive electrolyte bridging, and another one due to abnormal EEG spectra. Data from two additional participants were eliminated for poor EEG quality (excessive artifact) in one or both of the two conditions (eyes-closed, eyes-open). For the remaining 35 participants, the total number of epochs was: Session 1: eyes-closed = 331.8 ± 73.0 ; eyes-open = 377.8 ± 86.1 ; Session 2: eyes-closed = 330.1 ± 82.4 ; eyes-open = 373.7 ± 89.1 . The 2-s CSD epochs were tapered using a 50% Hanning window and padded with zeros (1 s at each end) to yield a FFT-transformed spectral resolution of .25 Hz. This is consistent with the resolution of Tenke et al. (2011; 1-s epochs padded to 4 s; 1024 points/epoch), but relies on less spectral interpolation. Mean power spectra were then computed for accepted trials.

Our CSD-fPCA implementation uses CSD amplitude (Root Mean Squared [RMS] power) spectra to obtain factors with an alpha structure that simply subdivided the alpha band, while preserving a linear relationship to the amplitudes of the underlying current generators (Tenke and Kayser, 2005). Due to the likelihood that the unique

characteristics of the different acquisition systems at each testing site would degrade the comparability of the spectra across testing site, additional steps were taken to eliminate site differences for alpha. First, since low frequency drift, high frequency noise, and differences in the bandpass properties were known to differ, the spectral data were limited to 1-40 Hz (data points 5-161). However, for the present sample from each testing site, there were noticeable differences between testing sites in the alpha band, with MG showing lowest and TX the largest amplitude (see Supplementary Material, Fig. S1)ⁱⁱ. Although this is of less concern when the sample is large, as would be the case in the upcoming analysis of patient data, it was important to preclude the possibility that differential variance contributions from each testing site might lead to a bias in the description and quantification of alpha by the PCA. For this reason, scale factors were computed for each testing site to equate the standard deviation of the CSD amplitude spectrum waveforms in alpha (8-12 Hz) for all participants and recording sessions at 19 posterior locations where alpha is greatest (P9/10, P7/8, P5/6, P3/4, P1/2, Pz, PO7/8, PO3/4, POz, O1/2, Oz). The complete CSD amplitude spectra for the full montage were then scaled, thereby matching posterior CSD alpha across testing sites without distorting the association between activity at each electrode for the rest of the spectra.

After scaling, the CSD amplitude spectra (1-40 Hz) were submitted to unrestricted fPCA based on the covariance matrix, with Varimax rotation of the covariance loadings (Tenke and Kayser, 2005; Tenke et al., 2011), yielding three dominant factors representing EEG alpha (49.1% variance of amplitude spectra). These factors were used as a spectral filter (Tenke et al., 2011) to reconstruct the alpha amplitude spectra without the influence of high-variance noise (e.g., EMG: broad peak at 28Hz, 20.6%; EOG: 1.25 Hz peak, 20.3%) or other low-variance activity (remaining factors < 2.5%). A second PCA with Varimax rotation was confined to 1-20 Hz (see Supplementary Material, Fig. S2), resulting in a low-alpha factor (47.9%), a high-alpha factor (36.1%), and a residual factor including low beta (16.0%).

Based on prior EEG studies evaluating posterior alpha power as a marker of

antidepressant response (Tenke et al., 2011), estimates of low-frequency alpha were computed as regional means across three posterolateral locations on each hemisphere (P7/8, P9/10, PO7/PO8). The high-frequency alpha factor score topography likewise included posteromedial and midline locations, although the variability across testing sites required a broader region for confident quantification (P7/8, P5/6, P3/4, P1/2, PO7/8, PO3/4, POz, O1/2, Oz). The resulting estimates were then averaged to examine the reliability of overall posterior alpha (mean of eyes-closed and eyes-open). For comparison with the alpha asymmetry literature, these posterior measures were supplemented by medial and lateral parietal (P3/4; P7/8) and frontal (F3/4; F7/8) electrodes to examine the reliability of alpha amplitude (mean of homologous electrodes) and asymmetry (right hemisphere minus left hemisphere) at these locations.

CSD-tPCA for LDAEP. Acceptable LDAEP averages were available for 38 participants. To optimize the identification and quantification of N1, latency jitter (Möcks, 1986) was eliminated between participants by temporally adjusting CSD waveforms for N1 sink peak latency (Kayser et al., 2012). This was accomplished by computing mean CSD waveforms across all five intensities, pooling them across 8 medial frontocentral locations (FC1/2, FC3/4, C1/2, C3/4) and across 4 lateral temporoparietal locations (TP7/8, P9/10) to provide an optimized estimate for N1 sink activity at frontocentral locations and its opposite (i.e., source) side of the underlying N1 dipole at temporoparietal locations. The most negative deflection of the corresponding difference waveform (i.e., frontocentral minus temporoparietal pooled CSDs) was determined between 0 and 200 ms post stimulus onset, resulting in N1 sink peak latencies between 90 and 195 ms (188 ± 15 ms). These individual N1 sink peak latencies were used to jointly align all 72 CSD waveforms for each stimulus intensity.

The optimized CSD waveforms were submitted to unrestricted tPCA as described above for fPCA (Kayser & Tenke, 2003, 2006a), in order to determine common sources of variance related to N1 sink activity and to quantify its amplitude. The input

matrices consisted of 257 variables (samples between -101 and 898 ms) and 27,360 observations stemming from 38 participants, 2 tests, 5 intensities and 72 electrode locations. Because this approach provides a concise, efficient simplification of the temporal pattern and spatial distribution of neuronal generators (Kayser and Tenke, 2003, 2006a), the present analysis focused on the PCA factor representing N1 sink.

To further minimize the problem of spatial component jitter between participants, bihemispheric N1 sink maxima and minima were determined from the individual N1 sink topographies (i.e., mean PCA factor scores across all five intensities for each participant). The most negative location within an array of 12 frontocentral and centroparietal locations (i.e., locations for the left hemisphere were F1, F3, F5, FC1, FC3, FC5, C1, C3, C5, CP1, CP3, and CP5; homologous locations were used for the right hemisphere) and the most positive location within an array of 7 lateral frontotemporal and temporoparietal locations (i.e., FT7, FT9, T7, TP7, TP9, P7, and P9 for the left hemisphere, and homologous locations for the right hemisphere) were determined, and these locations were then used to compute an estimate of N1 sink dipole strength for each hemisphere (i.e., difference between maximum and minimum) and intensity (see Fig. 3A in Tenke and Kayser, 2012, p. 2335, for a comparison between ERP and CSD topographies of N1 during LDAEP). For the present report, N1 dipoles computed for left and right hemisphere were averaged to obtain a single estimate for the tangentially-oriented N1 dipole in the vicinity of primary auditory cortex (Hegerl et al., 2001; highly similar but less robust reliabilities were observed for other quantifications of N1 amplitude, including PCA-based N1 amplitudes measured at C3 and C4 only).

The number of artifact-free trials included in the computation of the LDAEP averages did not differ between session 1 (Mean \pm SD, 90.0 \pm 13.7) and session 2 (88.7 \pm 6.8), yielding similar means across intensities (range from 87.6 \pm 7.8 to 88.3 \pm 7.6) and more than sufficient means for each testing site (CU: 92.6 \pm 3.1; MG: 87.9 \pm 3.3; TX: 86.5 \pm 5.9; UM: 84.2 \pm 12.1), despite a marginally significant difference

between testing sites ($F[3,34] = 2.36, p = 0.09$). However, there were no significant interactions between session, intensity, or testing site (all $p > 0.29$).

LORETA analysis of resting EEG. Although LORETA data were processed in parallel with those described for the resting EEG, only the eyes-closed condition was used, in line with prior studies linking rACC theta current density to treatment response (e.g., Pizzagalli et al., 2001). Acceptable data were available for 37 participants. Consecutive 2-s, nose-referenced EEG epochs, precisely matching those subjected to CSD-fPCA, were processed using LORETA (Pascual-Marqui et al., 1999) following the elimination of overlapping data (i.e., one out of four epochs retained). This approach mimics analyses from prior LORETA studies implicating rACC theta current density in predicting antidepressant response (e.g., Mulert et al., 2007; Pizzagalli et al., 2001).

LORETA computed the three-dimensional intracerebral current density distribution of EEG theta (6.5-8 Hz) based on the assumption that similar levels of activation characterize neighboring neurons, but with no assumptions about the number of generating sources. LORETA partitions the solution space into 2,394 cubic “voxels” (voxel dimension: 7 mm^3) limited to cortical gray matter and hippocampi, according to the digitized MNI probability atlases available from the Montreal Neurologic Institute (MNI; Montreal, Quebec, Canada). This distributed source localization technique has received cross-modal validation from studies combining LORETA with functional MRI (fMRI) (Mulert et al., 2004; Vitacco et al., 2002), structural MRI (Cannon et al., 2011; Worrell et al., 2000), intracranial EEG recordings (Zumsteg et al., 2006) and PET (Pizzagalli et al., 2004; Zumsteg et al., 2005; but see Gamma et al., 2004). Given that prior research has implicated theta current density in the rACC as a predictor of treatment response to antidepressant medication (Korb et al., 2009; Mulert et al., 2007; Pizzagalli et al., 2001; Rentzsch et al., 2014), analyses were restricted to this band (6.5-8 Hz) and a predefined rACC region-of-interest involving 13 voxels (Korb et al., 2009; Pizzagalli et al., 2001).

For the baseline session, the mean number of artifact-free epochs included was

83.4 ± 16.5 – amounting to an average of 170.8 ± 33.7 s of artifact-free EEG data available for analyses. For the Week 1 session, 82.2 ± 16.5 artifact-free epochs were available (168.3 ± 37.7 s). No significant differences emerged across testing sites or across sessions with respect to the number of artifact-free EEG epochs available for the LORETA analyses, all $ps > 0.45$. Consistent with established procedures (e.g., Pizzagalli et al., 2004), LORETA activity was normalized to a total power of 1 before statistical analyses. To minimize variations in signal-to-noise ratios across testing sites, over-smoothing was used (option TM04 in the LORETA transformation matrix module).

Results

Resting EEG alpha

CSD-fPCA of the resting EEG yielded expected low- and high-frequency alpha factors, identifiable by their factor loadings spectra, their distinct posterior topographies, and their condition-dependency (greater alpha for eyes-closed than for eyes open conditions; see Supplementary Material, Fig. S2). A residual alpha factor primarily reflected beta, and was showed the opposite condition-dependency (maximal for eyes-open). Fig. 2 shows the resulting mean alpha factor score topographies obtained at each testing site, showing similar posterior topographies and condition dependencies. Since previous studies have not identified differences of interest between the two alpha factors, they have been combined.

(insert Figure 2 about here)

A three-way ANOVA including *Testing Site* (CU, UT, MG, UM), *Session* (baseline, retest) and *Condition* (eyes open, eyes closed) yielded the expected *Condition* effect with posterior alpha (averaged across low and high alpha factors; 8-12 Hz) being greater with eyes closed than eyes open at each testing site ($F[1, 31] = 30.80, p <$

.001, partial eta squared = .50). Fig. 3 illustrates this effect for each testing site, and supports the impression by Fig. 2 of greater alpha for CU than the other sites. However, the only significant *Testing Site* effect was an overall *Testing Site* by *Session* interaction ($F[3, 31] = 3.93, p = .02$, partial eta squared = .275).

(insert Figure 3 about here)

Fig. 4 shows a scatterplot of overall posterior alpha in session 1 and 2 for each participant at the four testing sites. Although two CU cases showed appreciably greater alpha at session two than session one, the overall correlation showed high test-retest reliability of alpha across testing sites ($r = .84, p < .0001$) and ranged from $r = .74$ to $r = .99$ across testing sites.ⁱⁱⁱ Cases from each testing site were appropriately distributed along the overall regression line, and no other ANOVA effects were observed. Supplementary analyses of alpha asymmetries indicated lower reliability than for amplitude, particularly at frontal electrodes (Supplementary Material, Table S1).

(insert Figure 4 about here)

LDAEP

Fig. 5 shows grand mean CSD waveforms for three stimulus intensities in the LDAEP. The expected N1 topographies and loudness-dependency were observed, including the sink-to-source transition across the Sylvian fissure (Tenke and Kayser, 2012). These topographies are simplified in Fig. 6, showing waveforms at selected left central (C3) and left inferior-parietal (P9) sites for all five loudness intensities with the CSD-tPCA loadings waveform for the factor corresponding to N1 sink. The corresponding factor score topographies are shown in Fig. 7 for each of the five loudness intensities for both sessions (week 1 and 2).

(insert Figure 5 about here)

(insert Figure 6 about here)

(insert Figure 7 about here)

As summarized in Fig. 8, all testing sites showed the expected monotonic increase in N1 dipole amplitude with increasing tone intensity. A repeated measures ANOVA, including *Testing Site*, *Session* and *Intensity*, yielded a significant effect of *Intensity*, $F(12, 136) = 79.5, p < 0.0001$, $\epsilon = 0.56$), but no significant difference in N1 dipole across sessions or testing sites, and no interactions involving these variables. Fig. 9 shows the scatterplot of N1 dipole amplitude (averaged across intensity) for individual participants at each testing site in the two sessions. The test-retest reliability of N1 across testing sites was high ($r = .87, p < .0001$) and ranged from $r = .70$ to $r = .98$ for the individual sites.

(insert Figure 8 about here)

(insert Figure 9 about here)

LORETA Measure of rACC Theta.

Based on the findings of prior studies (Korb et al., 2009; Mulert et al., 2007; Pizzagalli et al., 2001), we computed theta current density for the rACC. Although preliminary analyses calculated current density measures for three different levels of spatial smoothing, higher spatial smoothing yielded greatest consistency across sites. A repeated measures ANOVA, including *Testing Site* and *Session*, revealed a

main effect of *Testing Site* ($F[3,33] = 8.27, p < .001$, partial eta squared = 0.429), owing to higher rACC current density at MG than the other testing sites (post-hoc unpaired *t*-tests, all^{iv} $p < .01$). As evident in the scatterplot of Fig. 10, current density was greater for MG in both sessions, and there was no significant difference in rACC current density across sessions. Although the test-retest correlation attained statistical significance ($p < .05$) at all levels of spatial smoothing, it was largest with the highest smoothing ($r = 0.70, p < .0001$), ranging from $r = .29$ to $r = .84$ across testing sites.

(insert Figure 10 about here)

Discussion

Overview

To our knowledge, this is the first study to examine the test-retest reliability of three electrophysiological measures that show promise as markers for antidepressant response. The current study in healthy controls, who were tested at four different research sites in the United States using different EEG acquisition systems, was conducted in preparation for the multisite EMBARC project, which will examine the value of biomarkers for differential prediction of response to antidepressants.

EEG Alpha

Most prior studies of test-retest reliability of EEG have used scalp potential measures in standard spectral bands. As in our prior study, in which posterior alpha predicted antidepressant treatment response in patients (Tenke et al., 2011), reference-free CSD was used for sharper, reference-independent topographies, and PCA provided measures for more robust, empirically-derived alpha bands. Reliability was examined for alpha CSD measures (integrated across low and high alpha factors) at posterior locations where alpha is maximal. Test-retest reliability of alpha was high ($r = .84$) and consistent across testing sites, which agrees with the reliability coefficients reported for scalp potential measures of alpha recorded at a single testing site (Allen

et al., 2004; Bruder et al., 2008; Smit et al., 2005).

LDAEP

Although studies have found LDAEP predicts response to SSRI antidepressants, there is less agreement on the best way of measuring it. A variety of different methods have been used to measure LDAEP, including scalp potential, dipole source analysis, or LORETA measures of N1, P2 or N1-P2 difference waveforms. The model of Hegerl and Juckel (1993) related LDAEP of N1/P2 to serotonergic neurotransmission in primary auditory cortex. The tangentially-oriented N1 dipole within the superior temporal gyrus in the vicinity of primary auditory cortex is thought to be uniquely important (Hegerl and Juckel, 1993; Hegerl et al., 1994), and Gallinat et al. (2000) found evidence that LDAEP of the tangential dipole of N1/P2 predicts response to a SSRI better than LDAEP scalp potentials from a single electrode (Cz). Simultaneous measurement of EEG and fMRI showed a high correlation of loudness dependence of activity in primary auditory cortex between fMRI and LORETA measures (Mulert et al., 2005). Both dipole source analysis and LORETA measures of LDAEP were found to predict response to an SSRI to the same degree, but were not highly correlated (Mulert et al., 2002). Moreover, Beauducel et al. (2000) found that the using tPCA-based LDAEP measures provided superior test-retest reliabilities compared to baseline-to-peak LDAEP measures.

Our CSD-tPCA dissociates the tangential N1 generator from a radially oriented, temporal lobe subcomponent of N1 (e.g., Kayser and Tenke, 2006a, 2006b; Tenke and Kayser, 2012), and our findings suggest that this direct, overall amplitude measure of the tangential N1 spanning the Sylvian fissure may provide an improved measure of serotonergic activation related to auditory intensity processing (see also Manjarrez et al., 2005). CSD-tPCA measures of the N1 dipole showed the expected monotonic increase with increasing tone intensity, which did not differ across testing sites or sessions. Overall, N1 amplitude (averaged over intensity) showed high test-retest reliability ($r = .87$) and was consistent across testing sites. In a prior study, the average amplitude of the N1 dipole in the LDAEP paradigm was strongly correlated

with the slope of LDAEP function and was predictive of response to antidepressants including a serotonergic agent (Kayser, 2013b). The use of N1 amplitude at only one or two intensities as an alternative to the slope of LDAEP function over a broad range of intensities has been suggested (Hensch et al., 2008) and could be more feasible for application in clinical settings. In the EMBARC study, we will use CSD-tPCA measures of the N1 dipole and evaluate whether overall N1 amplitude or slope of the LDAEP function is the best predictor of response to a SSRI antidepressant.

LORETA Measure of rACC Theta

Current density of theta, as localized by LORETA to the region of the rACC, has been reported to predict antidepressant response (Korb et al., 2009; Mulert et al., 2007; Pizzagalli et al., 2001), but has not previously been evaluated for test-retest reliability. Extensive neuropsychological and neuroimaging evidence has implicated the rACC in both the pathophysiology of depression and putative mechanisms of treatment response (for a review, see Pizzagalli, 2011). In particular, the rACC has been hypothesized to be implicated in treatment outcomes by supporting adaptive self-referential processing and recalibrating relationships between the default network and a 'task-positive network' spanning dorsolateral prefrontal and dorsal cingulate cortices. Animal data have also demonstrated an independent generator of theta oscillations in ACC (Feenstra and Holsheimer, 1979; Holsheimer, 1982), a finding also confirmed in various human neuroimaging studies (e.g., Asada et al., 1999; Pizzagalli et al., 2003). The convergence of these independent lines of evidence supported our *a priori* focus on theta activity in the rACC.

When using a high degree of spatial smoothing to minimize differences across testing sites, there was no difference in rACC theta across test sessions. The overall reliability coefficient ($r = .70$) was somewhat less than seen for CSD measures of posterior alpha. These properties suggest that LORETA solutions restricted to rACC theta may be subject to greater interindividual variability than the scalp-based CSD measures. However, it should be noted that the N1 dipole measure that was computed directly accounted for the spatial variability between subjects, suggesting

that the equivalent LORETA measure might require the identification of individual maxima within the rACC region.

There was a significant difference in rACC levels across testing sites, with one site (MG) being greater than the others. Although this testing site differed from the others in using a 129-channel EGI system, which relies on Hydrocel Geodesic nets rather than electrode caps with an extended 10-20 coordinate system, all sites relied on the identical 72-channel interpolated montage for the inverse computation. In planned analyses of EMBARC data, it will therefore be necessary to include research testing site as a covariate or to implement additional normalizations across the four testing sites. It is also important to note that the prescaling strategy applied to alpha was not used for the LORETA measure, which was computed directly from the eyes-closed EEG epochs, rather than from eyes-open and eyes-closed CSD amplitude spectra. The theta measure was also delimited by an *a priori* band.

Standardization Across Acquisition Sites and Platforms: Strengths and Limitations

The use of different EEG systems across testing sites poses a unique challenge that must be dealt with if the neurophysiological predictors determined in the EMBARC study are to be applied in real world clinical settings. Considerable efforts were made to standardize the training of testers and administration of the EEG across testing sites. The main purpose of testing healthy controls in this study was to establish sufficient reliability of the potential predictors of treatment outcome. Although limited by the small number of participants at each testing site, the results show that retest reliability across testing sites was high for alpha power and LDAEP despite differences in EEG systems. Reliability of rACC measures obtained with LORETA was lower, but still acceptable, and not affected by site differences, suggesting that it may be a property of the measure itself.

Another limitation of this study is that there was no control of the mental state or wakefulness of individuals during the EEG assessments. In particular, no diary was

obtained of prior night sleep or daily activities. Global brain states, such as CNS arousal or vigilance levels^v, can impact resting EEG measures (Hegerl et al. 2012; Olbrich et al., 2012), and circadian phase and sleep pressure during wakefulness also affect resting EEG (Aeschbach et al., 1997). Although time of day of EEG tests did not differ significantly across test and retest sessions, lack of control of individual's wakefulness or vigilance during these sessions could have increased variance of alpha and theta measures and reduced retest reliability. This will, however, be the case in real world applications of EEG tests and despite the lack of control of these variables, good-to-excellent retest reliability was obtained for each of the EEG measures in the EMBARC study, which represents a clear strength of the current findings and analytic approaches.

One potential source of variance between measures obtained at the different testing sites is the different calibration strategies used for different recording systems or preferred by different laboratories. It might be supposed that the use of a single calibration signal at all testing sites would be sufficient to assure comparability across sites. Unfortunately, there is no common mechanism for introducing the signal into all systems. Although the Neuroscan and EGI systems are equipped to introduce a calibration signal directly into the amplifier, this approach implicitly ignores the contributions of the electrode-scalp interface, including the different properties of Ag/AgCl and tin electrodes. Moreover, the electrodes of the Active2 system are all active, and its native recording reference is a common-mode sense electrode combined with a driven right leg balance electrode (CMS-DRL), which makes measurements through saline preferable for calibration. Following this line of reasoning one step further, the optimal common calibration signal for a study of alpha might be a 10 Hz sinusoid recorded through each system through the recording electrodes. Calibration across a wider range of frequencies (e.g., 1-20 Hz used for the final PCA) would either require a series of sinusoids or a variable frequency sweeping across the frequency range, resulting in a site-specific correction for EEG spectra. The same approach might also provide better comparability of LDAEP waveforms

across sites than using rectangular pulses of appropriate durations for signals (as used at CU). However, further consideration of these alternatives is well beyond the scope of the present study.

The resting CSD spectra were prescaled to protect against the possibility that alpha amplitude differences between testing sites might differentially bias the contribution of each site to the final PCA solution. In the case of small samples of healthy controls, such as used in the present report, this approach also redistributed cases from the different testing sites along the across-site regression line in Fig. 4, suggesting its applicability as a more general method for enhancing the consistency of alpha across testing sites quite apart from the rest of the EEG spectrum. This approach clearly has face validity for evaluating stability over time, but it does not provide a universal method for pooling across testing sites. Since healthy controls show considerable variability in overall resting alpha and task-related prestimulus alpha (Tenke et al., 2015), it is not impossible for even large samples of patients to differ in alpha amplitudes. It is therefore mandatory to include testing site as a control factor in all analyses that might distinguish between subgroups based on means (e.g., repeated measures ANOVA; etc.).

Conclusion

In summary, this multisite study demonstrated good test-retest reliability of CSD measures of resting EEG alpha and N1 dipole measures of LDAEP, and adequate test-retest reliability of LORETA measures of the activity in rACC, all of which have shown promise as predictors of clinical response to antidepressants. This report also details standardized procedures for improving compatibility of EEG and ERP data across testing sites using different EEG platforms and electrode montages, which should be highly relevant in other research contexts. This report is therefore both a critical step in evaluating the usefulness of electrophysiological measures as biomarkers for predicting clinical response to antidepressants as well as a template to guide future EEG/ERP studies derived from multiple testing sites, as is the current

trend in government-funded research.

Accepted Article

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Figure Captions

Figure 1. Flowchart of the preprocessing pipeline for continuous EEG. Data acquisition from the four testing sites differed in electrode composition, recording montages, broadband cutoffs and acquisition hardware and software. 1) Raw data files were unified to *bdf* format using EEGlab routines. 2) Raw data were then evaluated for data integrity and channel exclusion based on runtime notes and preliminary visual inspection. 3) Data were then preprocessed using Polyrex to include the common 72-channel montage (CU), eliminate baseline drifts using a polynomial filter, and scale the data to optimize the range of the resulting file in *cnt* format. 4) Data were then interpolated from all good electrodes in the original montage using a spherical spline following tests for electrode bridging. If additional electrodes are identified as bad, or if the performance of the polynomial filter is degraded by recording errors (e.g., extraneous data between blocks), raw data will be reevaluated (step 2) and corrected. 5) Following successful data interpolation, electrodes that differ from the common 72-channel montage are eliminated, bipolar eye channels created by interpolation, and the EEG channels are blink-corrected.

Figure 2. Mean alpha factor score topographies obtained at each testing site, for net (eyes-closed minus eyes-open) and overall (mean of eyes-closed and eyes-open) alpha. Means are across low- and high-frequency factors for both test sessions. Alpha topographies have similar posterior topographies at all testing sites, particularly for overall alpha.

Figure 3. Mean and SE of posterior alpha factor scores for eyes-open and eyes-closed conditions at each of the four testing sites. Posterior alpha is averaged across electrode regions for low- and high-frequency factors for both test sessions. ANOVA results identified a test-retest difference at CU as the origin of the apparent difference in overall alpha.

Figure 4. Scatter plot of overall posterior alpha in session 1 and 2 for individual

participants at the four testing sites. The correlation across testing sites showed high test-retest reliability for overall alpha ($r=.84$, $p<.0001$), with cases from each testing site distributed along the regression line. Apparent differences in overall amplitude differences for CU in Figs. 2-3 reflect two cases with greater alpha at session two than session one.

Figure 5. Grand mean ($N = 38$) current source density (CSD) [$\mu\text{V}/\text{cm}^2$] waveforms (-100 to 900 ms, 100 ms pre-stimulus baseline) comparing stimuli of low (60 dB), medium (80 dB), and high (100 dB) loudness intensity (pooled across testing site and test-retest session) at all 72 scalp recording locations. CSDs had been individually adjusted for N1 sink peak latency (see text).

Figure 6. Enlargements (-100 to 900 ms; cf. Fig. 5) of current source density (CSD) [$\mu\text{V}/\text{cm}^2$] waveforms at selected left central (C3) and left inferior-parietal (P9) sites comparing all five loudness intensities. The loadings of factor 121 corresponding to N1 sink are shown for comparison on the same scale. The inset shows CSDs for 100 dB between 60 and 220 ms to highlight a peak latency shift of 45 ms that differentiates N1 sink at site C3 from temporal N1 sink at site T7, the latter corresponding to a separate CSD-PCA factor. Note that these distinct latency shifts can not be appreciated from a cursory review of Fig. 5.

Figure 7. Topographies of N1 sink for five loudness intensities for both test-retest session (week 1 and 2). All topographies are two-dimensional representations of spherical spline interpolations ($m = 2$; $\lambda = 0$) derived from the mean factors scores ($N = 38$) for each recording site at each test session and each intensity.

Figure 8. Mean and SE for N1 dipole amplitude at five tone intensities show LDAEP function for the four testing sites.

Figure 9. Scatter plot of N1 dipole amplitude in session 1 and 2 for individual participants at the four testing sites. The correlation across testing sites showed high test-retest reliability ($r=.87$, $p<.0001$).

Figure 10. Theta current density (eyes-closed) localized to rACC in session 1 and 2 for individual participants at the four testing sites. The correlation across testing sites showed high test-retest reliability ($r=.70$, $p<.0001$).

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Endnotes

ⁱ EEG for MG participants was collected at McLean Hospital in Belmont

ⁱⁱ Although TX and UM share a higher peak frequency than the other sites in Fig. S1, they had widely different recording environments, owing to the distinction between DC with Ag/AgCl electrodes vs. a .5 Hz filter with tin electrodes.

ⁱⁱⁱ For comparison purposes, test-retest correlations for overall alpha amplitude and asymmetry are shown for medial and lateral parietal and frontal electrodes in the Supplementary Material, Table S1.

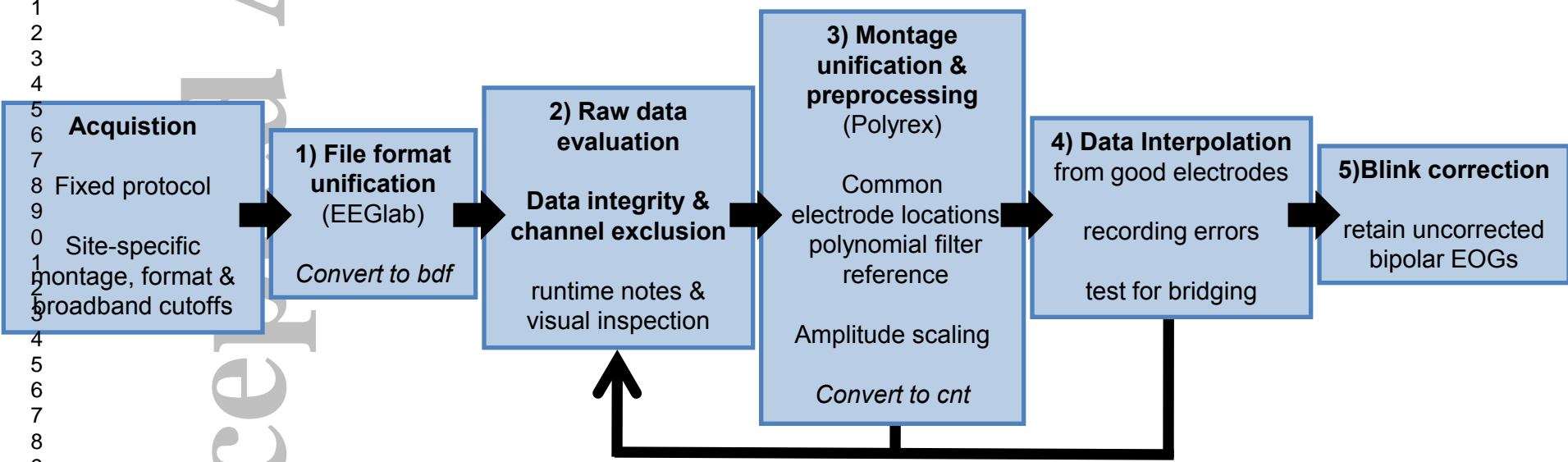
^{iv} The statistical results for the post-hoc tests are as follows: MG vs. CU: $t(17) = 5.35$, $p < .001$; MG vs. TX: $t(17) = 4.82$, $p < .001$; MG vs. UM: $t(15) = 3.82$, $p < .003$.

^v Analyses are underway examining the application of a vigilance algorithm on the EEG of patients and controls in this study (Hegerl et al. 2012; Olbrich et al., 2012), which are extensive and will be reported separately.

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Preprocessing Pipeline for Continuous EEG

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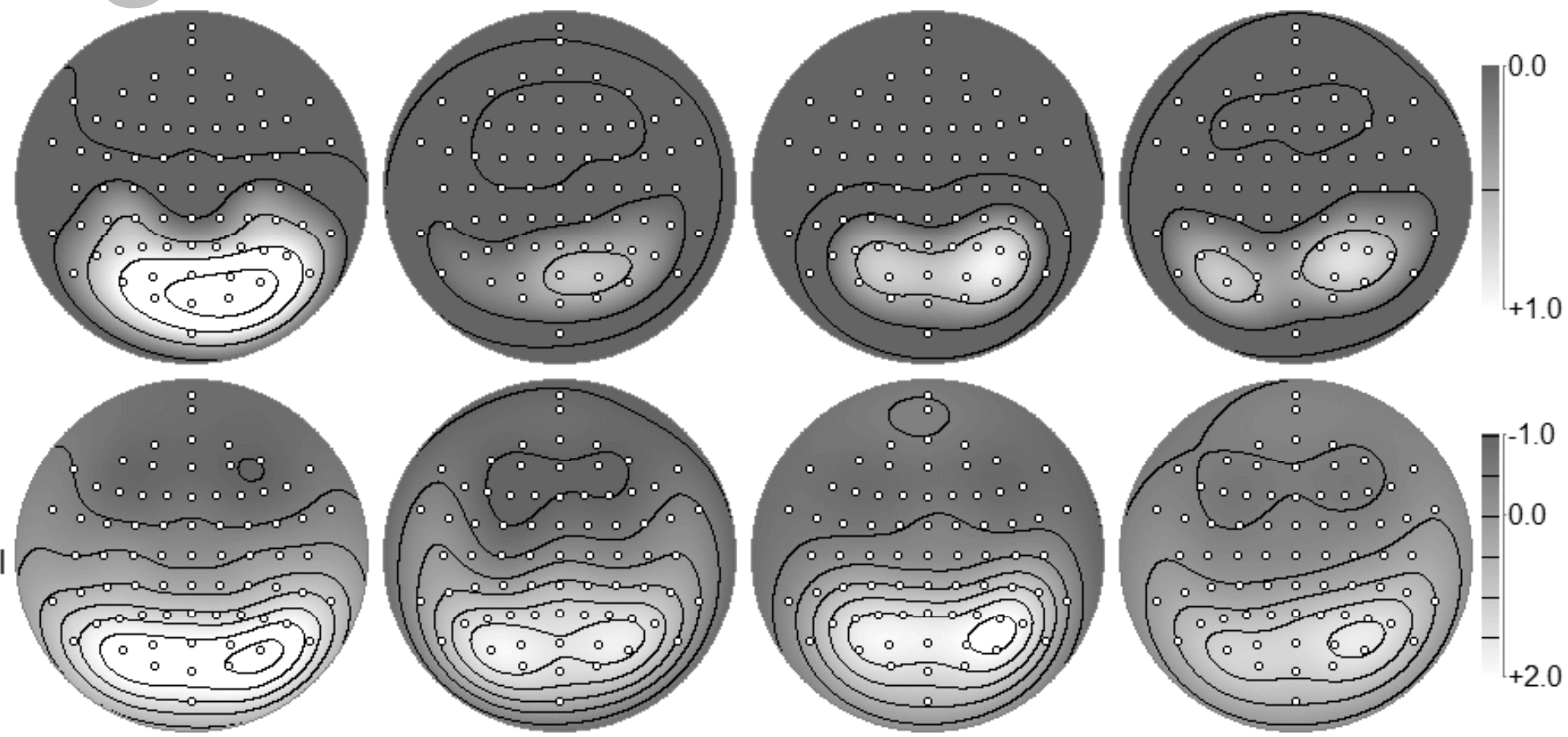
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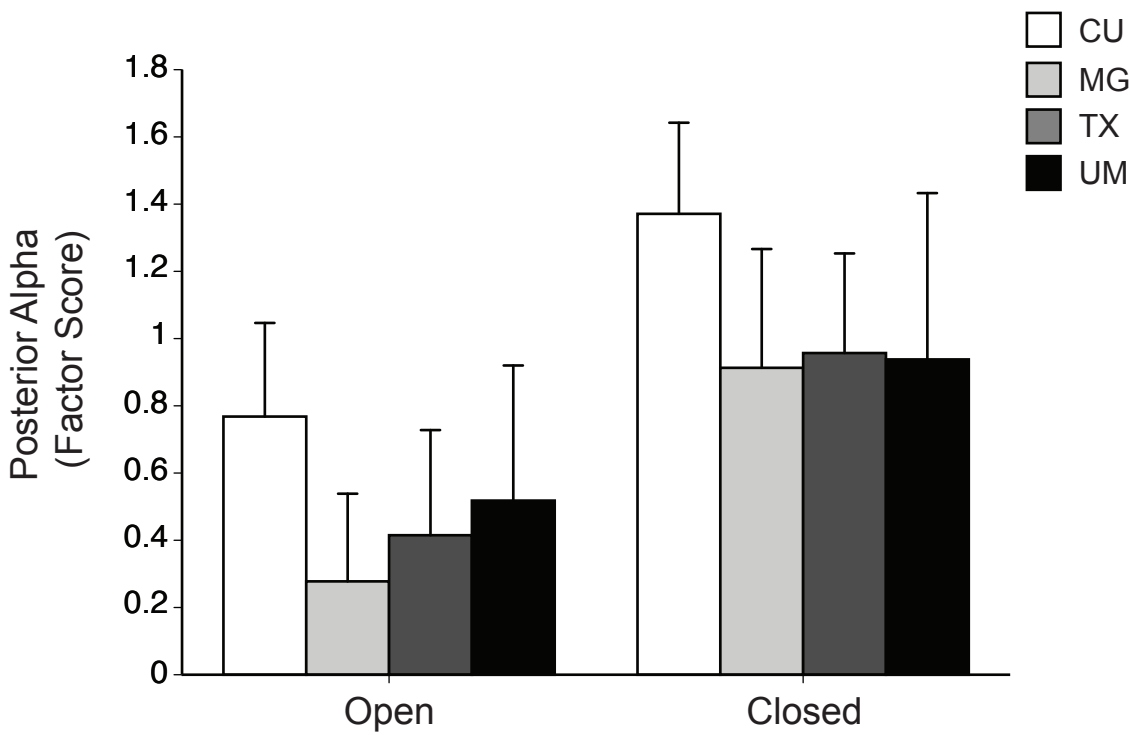
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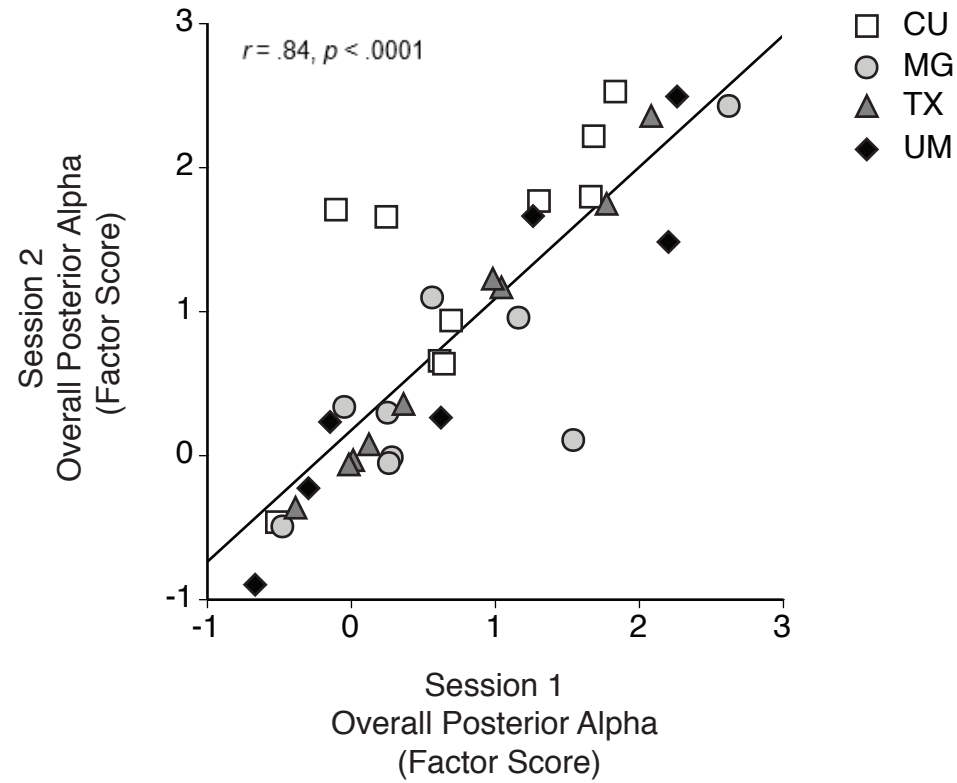
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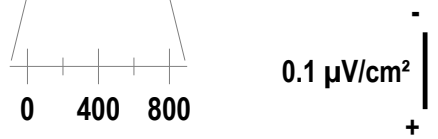
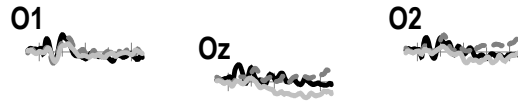
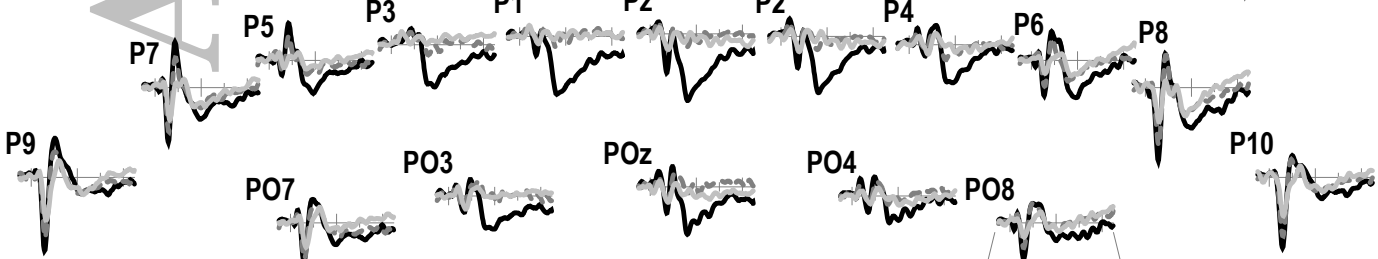
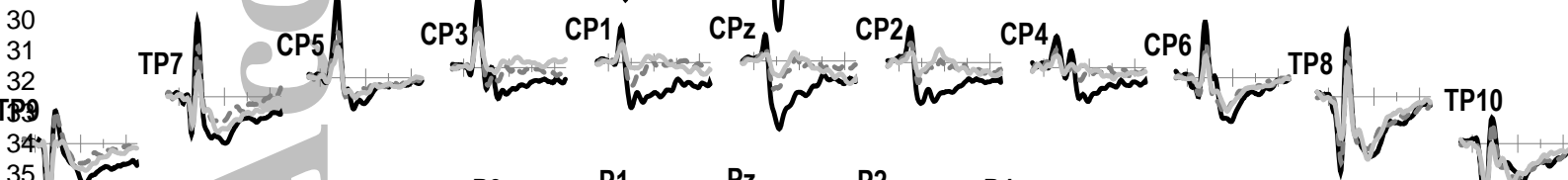
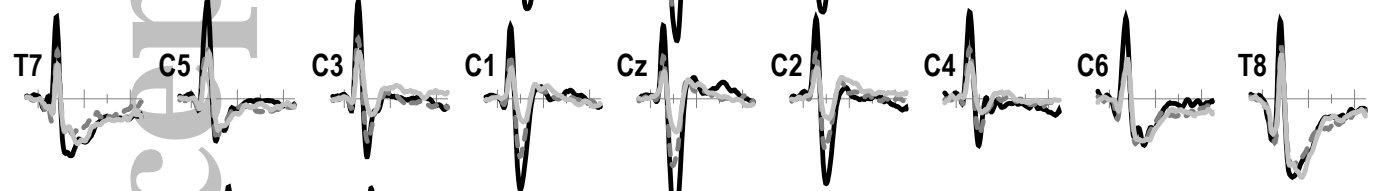
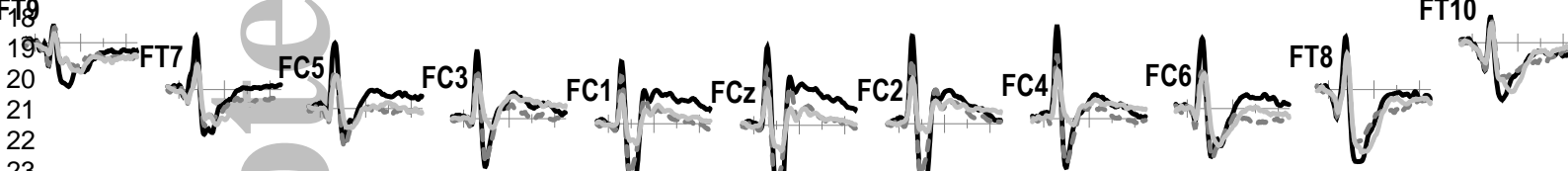
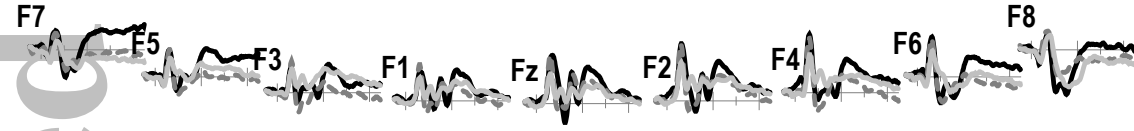
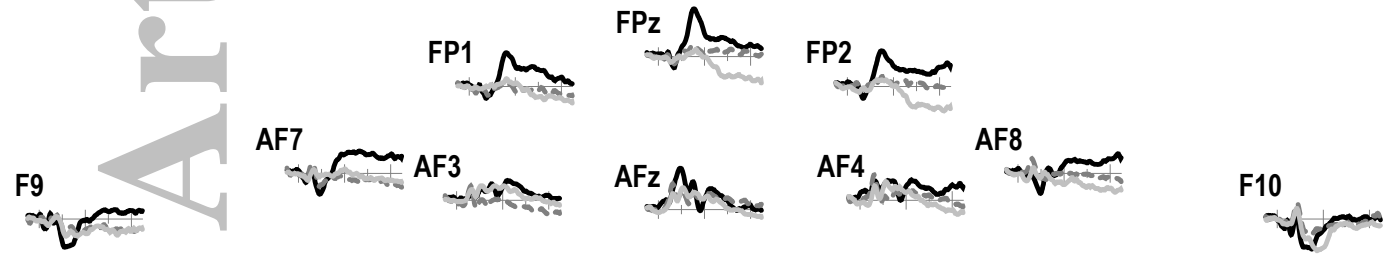
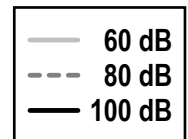
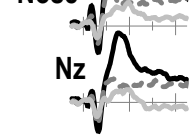




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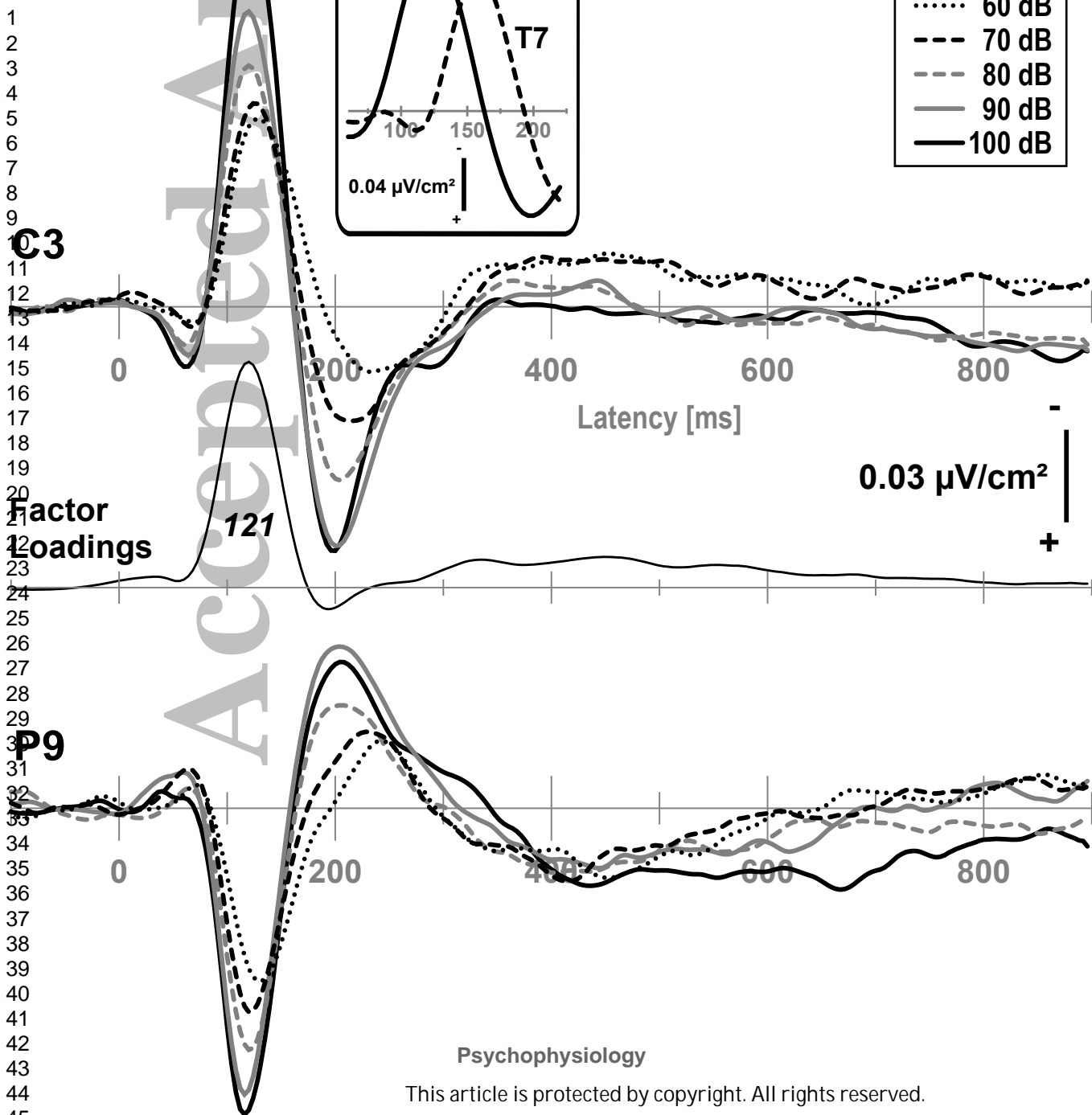
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Psychophysiology



Psychophysiology



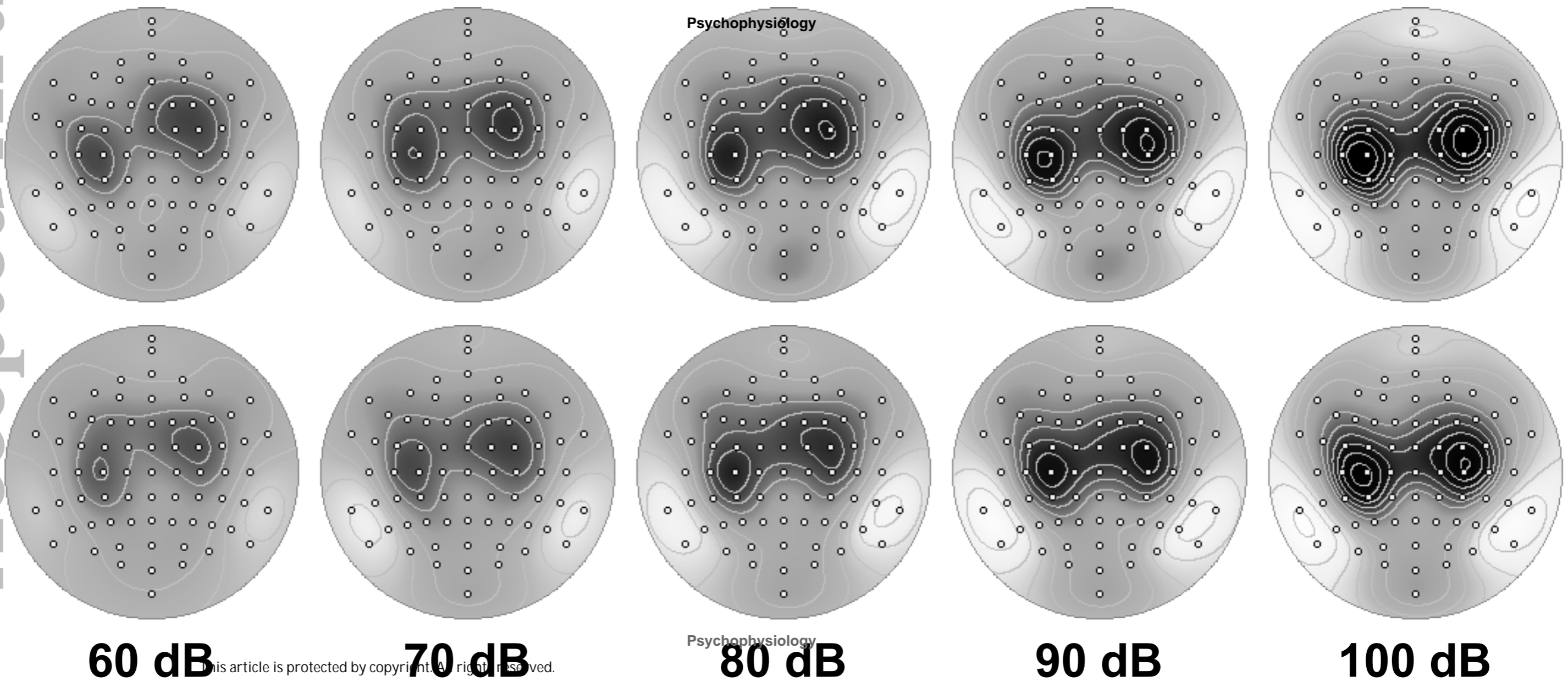


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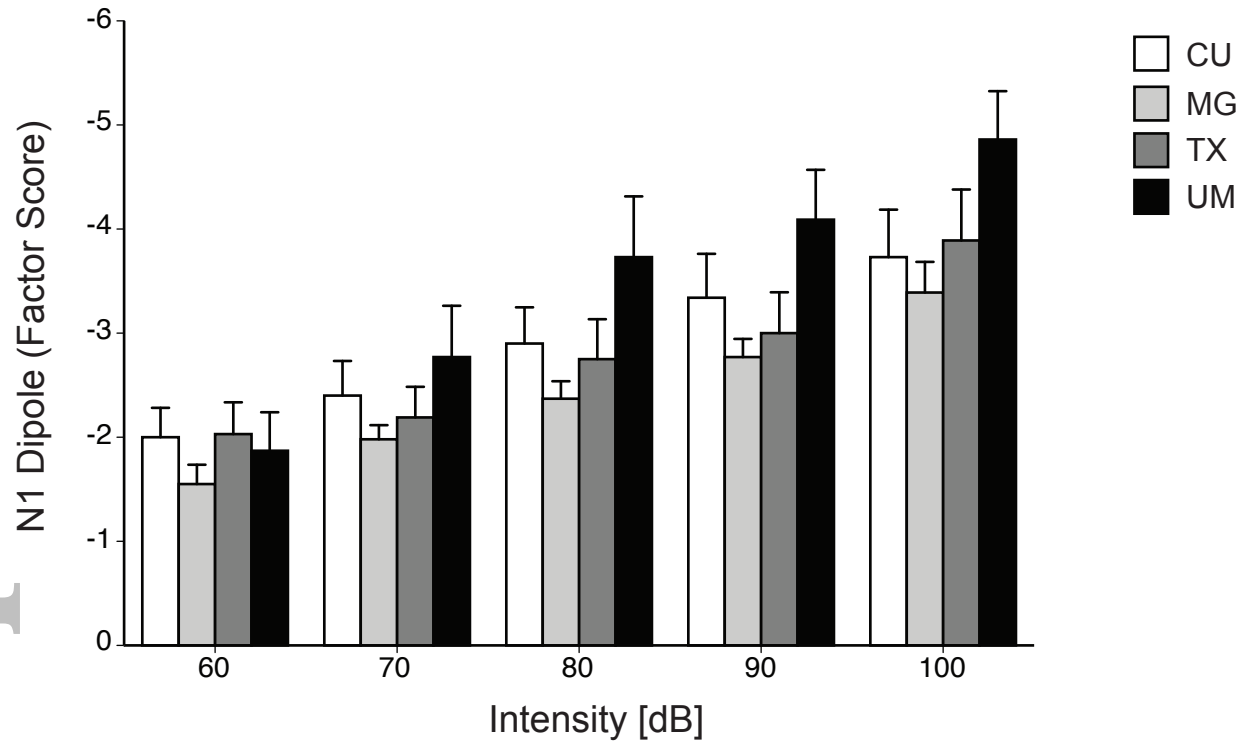
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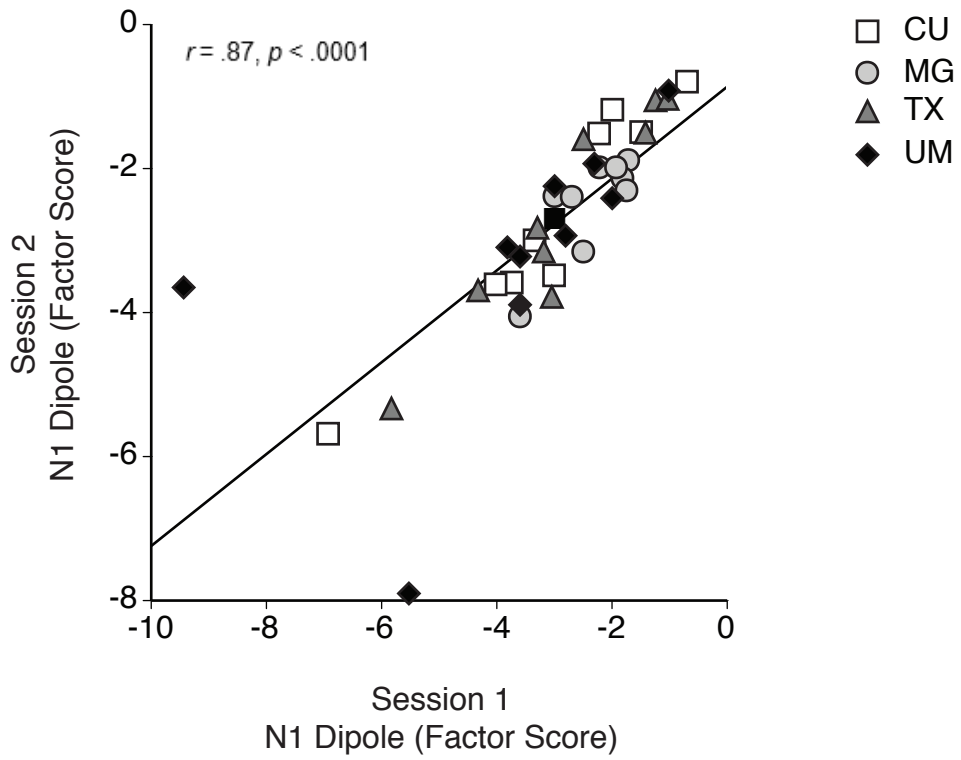
Psychophysiology
80 dB

90 dB

100 dB



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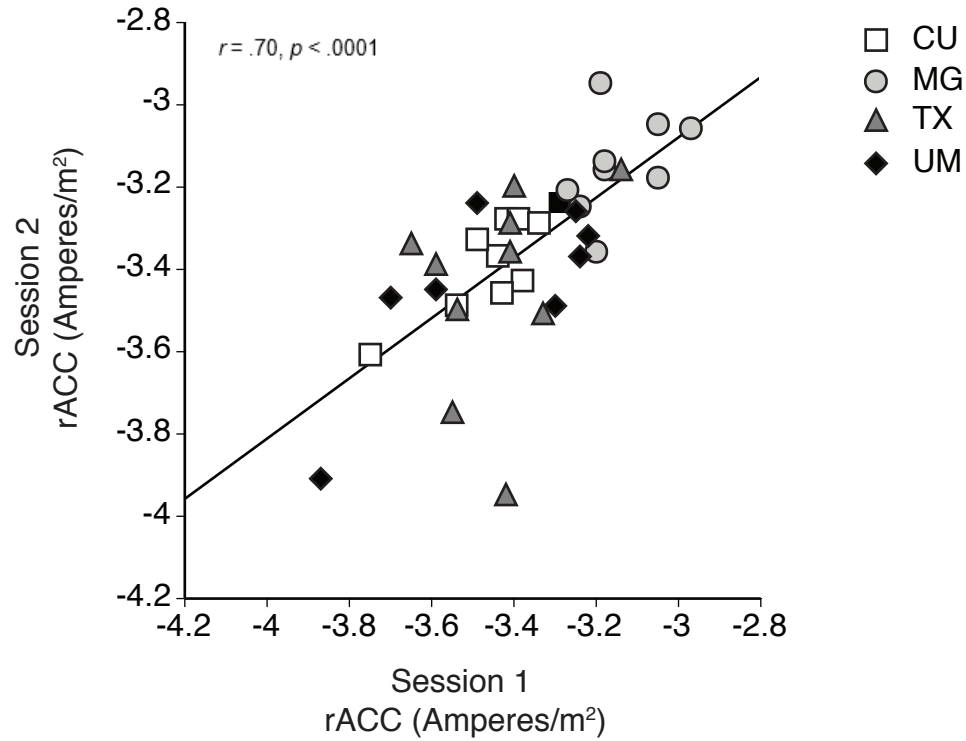


Figure S1. Grand mean CSD amplitude spectra for the alpha band (8 - 12 Hz) across 19 posterior electrodes for all participants and recordings, separated by testing site. **A.** Unscaled amplitude spectra from which scale factors were computed to equate their standard deviations across testing sites (CU, MG, TX, UM). **B.** The corresponding scaled CSD amplitude spectra. Equating the variance of alpha between testing sites in this way protected against the disproportionate representation of testing sites in the extraction of CSD-fPCA factors.

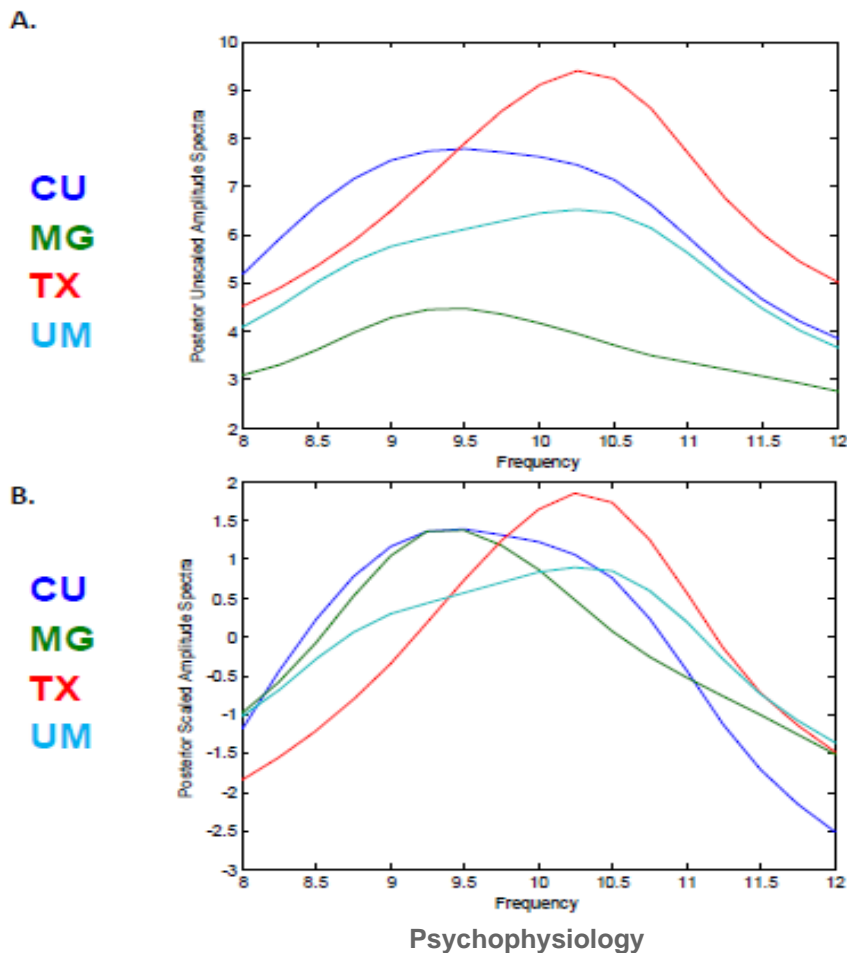
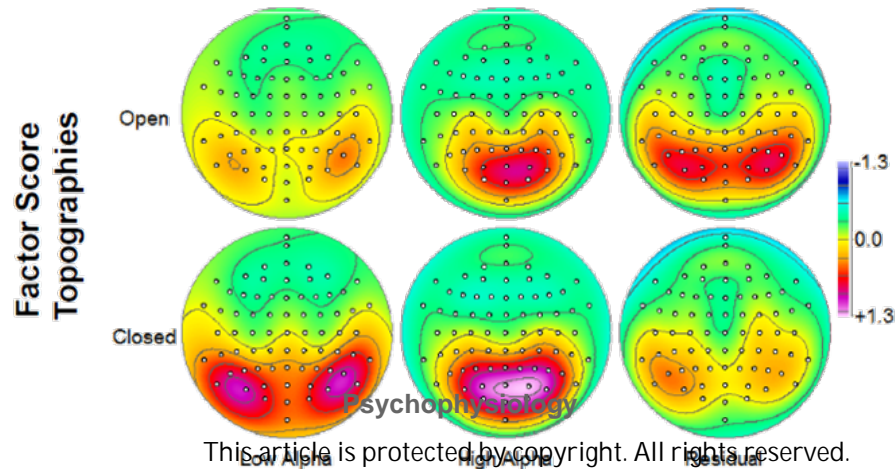
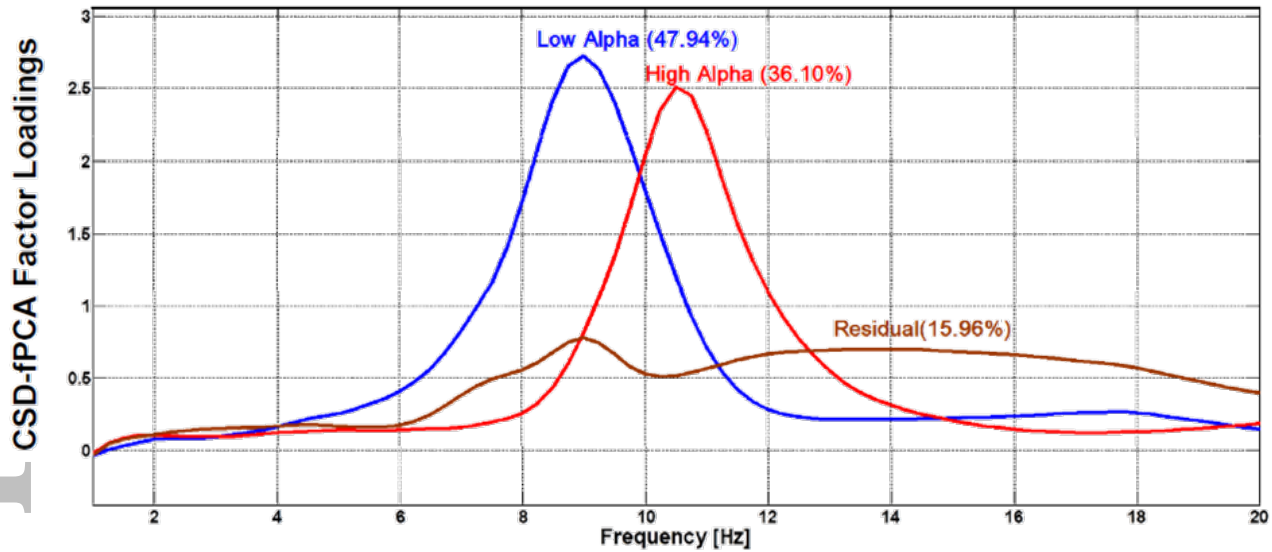


Figure S2. CSD-fPCA of prefiltered CSD amplitude spectra (Tenke et al., 2011). **Top.** Factor loadings spectra yielded a low-frequency alpha factor and a high-frequency alpha factor, as well as a residual factor including low beta. **Bottom.** Factor score topographies for low-frequency alpha and high-frequency alpha factors had the expected posterior topographies and condition dependency (greater alpha for eyes-closed than for eyes-open condition). The residual factor did not show a condition dependency consistent with alpha.



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Table S1. Test-Retest Correlations of Overall Alpha at Parietal and Frontal Electrodes

	Overall Alpha Amplitude				Overall Alpha Asymmetry			
	Parietal		Frontal		Parietal		Frontal	
	Medial (P3/4)	Lateral (P7/8)	Medial (F3/4)	Lateral (F7/8)	Medial (P4-P3)	Lateral (P8-P7)	Medial (F4-F3)	Lateral (F8-F7)
Across Testing-Sites (N=35)	.837 ***	.867 ***	.728 ***	.883 ***	.485 **	.700 ***	.371 *	.593 ***
CU (n=10)	.788 **	.757 *	.802 **	.578	.313	.947 ***	.656 *	.471
MG (n=9)	.678 *	.902 **	.610	.981 ***	.493	-.333	-.098	.798 **
TX (n=9)	.966 ***	.991 ***	.955 ***	.936 ***	.576	.824 **	.428	.197
UM (n=7)	.905 **	.899 **	.344	.322	.596	-.008	.354	.432

* $p < .05$ ** $p < .01$ *** $p < .001$

Test-retest correlations for overall posterior alpha (across conditions and factors) indicated high test-retest reliability across testing sites ($r = .84$). Although limited by small sample sizes, this correlation was also evidenced for the individual testing sites, ranging from $r = .74$ to $r = .99$. While the present study focused on regional measures of posterior alpha as a potential marker for treatment response, the comparison of corresponding amplitude and asymmetry correlations at frontal and parietal electrodes provides the opportunity to compare them to the broader literature on alpha asymmetries. Table S1 shows that parietal amplitude correlations were comparable to those for the overall posterior regional measures. Despite the low amplitudes of frontal alpha (cf. factor score topographies, Fig. S2), frontal amplitude correlations were also robust, albeit more variable across testing sites. The test-retest correlations for overall alpha asymmetry were smaller than those for alpha amplitude, particularly at frontal locations, and were extremely variable across testing sites or regions (medial, lateral).