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Tip rates, phylogenies, and diversification: what are we estimating, and how good are the estimates?

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Abstract

1. Species-specific diversification rates, or “tip rates”, can be computed quickly from phylogenies and are widely used to study diversification rate variation in relation to geography, ecology, and phenotypes. These tip rates provide a number of theoretical and practical advantages, such as the relaxation of assumptions of rate homogeneity in trait-dependent diversification studies. However, there is substantial confusion in the literature regarding whether these metrics estimate speciation or net diversification rates. Additionally, no study has yet compared the relative performance and accuracy of tip rate metrics across simulated diversification scenarios.
2. We compared the statistical performance of three model-free rate metrics (inverse terminal branch lengths; node density metric; DR statistic) and a model-based approach

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30 (BAMM). We applied each method to a large set of simulated phylogenies that had been
31 generated under different diversification processes. We summarized performance in
32 relation to the type of rate variation, the magnitude of rate heterogeneity and rate regime
33 size. We also compared the ability of the metrics to estimate both speciation and net
34 diversification rates.

- 35 3. We show decisively that model-free tip rate metrics provide a better estimate of the rate
36 of speciation than of net diversification. Error in net diversification rate estimates
37 increases as a function of the relative extinction rate. In contrast, error in speciation rate
38 estimates is low and relatively insensitive to extinction. Overall, and in particular when
39 relative extinction was high, BAMM inferred the most accurate tip rates and exhibited
40 lower error than non-model-based approaches. DR was highly correlated with true
41 speciation rates but exhibited high error variance, and was the best metric for very small
42 rate regimes.
- 43 4. We found that, of the metrics tested, DR and BAMM are the most useful metrics for
44 studying speciation rate dynamics and trait-dependent diversification. Although BAMM
45 was more accurate than DR overall, the two approaches have complementary strengths.
46 Because tip rate metrics are more reliable estimators of speciation rate, we recommend
47 that empirical studies using these metrics exercise caution when drawing biological
48 interpretations in any situation where the distinction between speciation and net
49 diversification is important.

50
51 Keywords: tip rates, diversification, terminal branch length, node density, DR, BAMM, trait-
52 dependent diversification

53 54 55 Introduction

56 Rates of speciation and extinction vary through time and among lineages (Nee, Mooers &
57 Harvey 1992; Sanderson & Donoghue 1996; Etienne & Haegeman 2012; Jetz et al. 2012; Moen
58 & Morlon 2014; Alfaro et al. 2018), contributing to dramatic heterogeneity in species richness
59 across the tree of life (Alfaro et al. 2009; Jetz et al. 2012; Barker et al. 2013). By characterizing
60 variation in rates of speciation and extinction, we can better understand the dynamics of

61 biological diversity through time, across geographic and environmental gradients (Zink, Klicka
62 & Barber 2004; Ricklefs 2006; Mittelbach et al. 2007; Silvestro, Schnitzler & Zizka 2011;
63 Rabosky, Title & Huang 2015), and in relation to traits and key innovations (FitzJohn, Maddison
64 & Otto 2009; Near et al. 2012; Beaulieu & O’Meara 2016). Consequently, there has been great
65 interest in statistical methods for inferring rates of speciation and extinction from molecular
66 phylogenies.

67 Although rates of diversification have traditionally been quantified for clades, there has
68 been a growing interest in estimating species-specific rates of diversification, which we refer to
69 here as “tip rates”. Tip rates are increasingly used to describe patterns of geographic and trait-
70 associated variation in diversification (Freckleton, Phillimore & Pagel 2008; Jetz et al. 2012;
71 Kennedy et al. 2016; Harvey & Rabosky 2017; Quintero & Jetz 2018; Rabosky et al. 2018). It
72 may seem strange to view evolutionary rates as a property of individual lineages, but such rates
73 emerge naturally from the birth-death model we typically use to conceptualize the diversification
74 process (Nee, Mooers & Harvey 1992; Nee, May & Harvey 1994). Under the birth-death
75 process, individuals (species) are characterized by per-lineage rates of species origination
76 (speciation, λ) and extinction (μ). For the purposes of inference, these rates are typically
77 assumed to be constant among contemporaneous members of a focal clade. However, tip rates
78 can be viewed as our best estimate of the present-day rate of speciation or extinction for an
79 individual lineage, conditional on past (usually recent) evolutionary history. As such, they
80 provide information about the expected amount of time that will elapse before a lineage splits or
81 becomes extinct.

82 A number of approaches have been used to estimate tip rates, including both model-based
83 and non-model-based approaches (i.e., models that are parameterized with speciation and
84 extinction rates, vs metrics that simply rely on branch lengths and splitting events). These
85 approaches vary in terms of how much information they derive from a focal species (i.e., a
86 terminal branch) relative to the amount of information they incorporate from other regions of the
87 phylogeny. On one end of the spectrum, tree-wide estimates (i.e., one rate for the entire
88 phylogeny) of speciation and extinction rates under a constant-rate birth-death (CRBD) model
89 provide tip rates that are maximally auto-correlated (identical) across species in the clade; such
90 rates for any given species are not independent of rates for any other species in the group of
91 interest. On the other end of the spectrum, terminal branch lengths can be used to derive a

92 censored estimate of the rate of speciation that is minimally autocorrelated with rates for other
93 species in the focal clade. Terminal branch lengths are largely unique to each species (rates
94 might be identical only for sister taxa), but provide a noisy measure of speciation, due to the
95 stochasticity inherent in the diversification process (Nee, May & Harvey 1994), and they have
96 been employed as a summary statistic in assessing model adequacy (Bromham, Hua & Cardillo
97 2016; Gomes, Sorenson & Cardoso 2016). In contrast to single (terminal) branch estimates, tree-
98 wide estimates should be less susceptible to stochastic noise, because they incorporate
99 information from the entirety of the tree (e.g., multiple branches are used in the estimates). Of
100 course, the tree-wide estimate necessarily assumes that all tips share a common underlying
101 diversification process. Other tip rate metrics fall somewhere between these two extremes,
102 incorporating some tree-wide information but relaxing the assumption of homogeneous rates
103 across all lineages (node density metric: Freckleton, Phillimore & Pagel 2008; DR: Jetz et al.
104 2012). The estimation of tip-specific rates thus entails a tradeoff between the precision of
105 individual estimates and the stochastic error associated with those estimates.

106 BMM (Bayesian Analysis of Macroevolutionary Mixtures, Rabosky 2014) is a model-
107 based approach that can accommodate heterogeneity in the rate of diversification through time
108 and among lineages. BMM simulates a posterior distribution of macroevolutionary rate shift
109 configurations given a phylogeny of interest; marginal rates of speciation and extinction for
110 individual taxa can then be extracted from this distribution. In this framework, the correlation in
111 rates between any pair of species is a function of the posterior probability that they share a
112 common macroevolutionary rate regime (Rabosky et al. 2014). If the tree-wide posterior
113 probability of rate variation is low, the marginal rates estimates for individual species will be
114 similar across the entire tree, as under a CRBD model. Likewise, any pair of taxa that are
115 consistently assigned to the same macroevolutionary rate regime will necessarily have identical
116 tip rates.

117 Tip rates are best suited to a host of questions and hypotheses where the diversification
118 dynamics over the evolutionary history of a group are either less relevant, or no more relevant,
119 than the rates of diversification closer to the present day. For example, many hypotheses
120 involving trait-dependent diversification implicitly assume a time-homogeneous, or constant
121 through time, effect of the trait on diversification rate (Coyne & Orr 2004; Kay et al. 2006;
122 Jablonski 2008; FitzJohn 2010; Claramunt et al. 2011). Harvey & Rabosky (2017) found that the

123 use of tip rates for assessing correlations between continuous traits and diversification has good
124 performance across a range of diversification scenarios. Furthermore, hypotheses pertaining to
125 non-historical geographic patterns of diversity are also better addressed with recent rates of
126 diversification. For example, many hypotheses for the latitudinal diversity gradient propose time-
127 homogeneous effects of particular environmental factors (temperature, energy, geographic area)
128 on rates of diversification (Mittelbach et al. 2007; Kennedy et al. 2014; Rabosky, Title & Huang
129 2015; Schluter 2016; Rabosky et al. 2018). Put simply, if such time-homogeneous processes
130 have shaped the latitudinal diversity gradient (e.g., correlation between speciation and
131 temperature: Rohde 1992), then the effect should be manifest in the distribution of present-day
132 evolutionary rates.

133 At present, there is substantial confusion in the literature over what quantity various tip
134 rate metrics actually measure. The DR statistic (Jetz et al. 2012) was originally described as a
135 measure of the “species-level lineage diversification rate”. While supplemental analyses and
136 subsequent work suggested that DR was a better measure of speciation rate than net
137 diversification (Jetz et al. 2012, Belmaker & Jetz 2015, Quintero & Jetz 2018), many studies
138 have nonetheless continued to describe DR as an estimate of the lineage-level net diversification
139 rate (Marin & Hedges 2016; Oliveira et al. 2016; Cai et al. 2017; and many others). The node
140 density metric of Freckleton, Phillimore & Pagel (2008) has also been described as a measure of
141 net diversification. Whether these metrics more accurately measure speciation or net
142 diversification is critically important for interpreting biodiversity patterns (e.g., two regions
143 might differ dramatically in speciation rate, but net diversification rates in each might
144 nonetheless be zero). An objective of our study is thus to compare the ability of DR, node
145 density, and other metrics to estimate speciation and net diversification rates.

146 Despite the potential utility of tip rates in geographic and trait-based analyses of
147 speciation rate heterogeneity (Jetz et al. 2012; Belmaker & Jetz 2015; Oliveira et al. 2016;
148 Quintero & Jetz 2018), there has yet been no comprehensive comparative assessment of the
149 accuracy and precision of the estimates, save for supplemental analyses in Jetz et al. (2012) and
150 Quintero & Jetz (2018). BAMM has low power to infer small rate regimes (Rabosky, Mitchell &
151 Chang 2017; Meyer & Wiens 2017), leading to the possibility that other approaches might
152 perform better for smaller phylogenies or when the variation in rates among clades is subtle.
153 However, DR and related methods will always identify variation in tip rates, even when none

154 exists, provided there is stochastic variation in branch lengths. A goal of this study is therefore to
155 evaluate the trade-off between the stochastic noise inherent in non-model-based approaches, and
156 the conservative but less noisy estimates from model-based metrics. We compare the
157 performance of these metrics across a range of simulation scenarios, which include both discrete
158 and continuous variation in rates.

159

160 **Methods**

161 Tip rate metrics

162 We assessed the accuracy of four tip rate metrics in this study at quantifying rates of
163 speciation. As we demonstrate below (see also Supplementary Figure 5 in Jetz et al. 2012;
164 extended Figure 5 in Quintero & Jetz 2018; Belmaker & Jetz 2015), these metrics are estimators
165 of speciation rate and not net diversification rate, and we refer to them as such throughout. The
166 first metric is the inverse of the equal splits measure (Redding and Mooers 2006), also called the
167 *DR statistic* (Jetz et al. 2012), *DivRate* (Belmaker & Jetz 2015; Oliveira et al. 2016), *ES* (Harvey
168 & Rabosky 2017) or *tip DR* (Quintero & Jetz 2018), which we denote in this study as λ_{DR} . This
169 species-specific measure incorporates the number of splitting events and the internode distances
170 along the root-to-tip path of a phylogeny, while giving greater weight to branches closer to the
171 present (Redding & Mooers 2006; Jetz et al. 2012). λ_{DR} is computed as:

$$\lambda_{DR_i} = \sum_{j=1}^{N_i} b_j \frac{1}{2^{j-1}}$$

172 where λ_{DR_i} is the tip rate for species i , N_i is the number of branches between species i and the
173 root, b_j is the length of branch j , starting at the terminal branch ($j = 1$) and ending with the root.
174 Jetz et al. (2012) demonstrated that, for trees deriving from a Yule process, and with mild
175 extinction, the mean λ_{DR} across tips converges on the true speciation rate.

176 We also considered a simpler metric, node density (Freckleton, Phillimore & Pagel 2008;
177 denoted by λ_{ND}). This is simply the number of splitting events along the path between the root
178 and tip of a phylogeny, divided by the age of the phylogeny. While λ_{DR} down-weights the
179 contribution of branch lengths that are closer to the root, λ_{ND} equally weights the contributions
180 of all branches along a particular root-to-tip path, regardless of where they occur in time. Under a

181 pure-birth model ($\mu = 0$), both λ_{DR} and λ_{ND} should yield unbiased estimates of the rate of
182 speciation.

183 The third measure we considered is the inverse of the terminal branch lengths (λ_{TB}).
184 Rapid speciation rates near the present should be associated with proportionately shorter terminal
185 branches; smaller values of λ_{TB} should thus characterize species with faster rates of speciation.
186 This measure has recently been used as a summary statistic to assess model adequacy in trait-
187 dependent diversification studies (Bromham, Hua & Cardillo 2016; Gomes, Sorenson & Cardoso
188 2016; Harvey & Rabosky 2017). Following Steel & Mooers (2010), we note that the terminal
189 branch lengths can be used to derive an estimate of the speciation rate; this follows from the fact
190 that interior and terminal branches have the same expected value under the Yule process (Steel &
191 Mooers 2010). The corresponding estimator for the i 'th tip, λ_{TB} is approximately $1 / 2b$ where b
192 is the length of a given terminal branch (Steel & Mooers 2010). To our knowledge, λ_{TB} has not
193 been used to explicitly estimate tip rates as we do here, but given its utility as a summary statistic
194 and general theoretical properties (Steel & Mooers 2010), we see value in comparing the
195 performance of this metric to others currently in use.

196 Finally, we considered a Bayesian, model-based approach to estimating tip rates. BAMM
197 (Rabosky 2014) assumes that phylogenies are generated by set of discrete diversification
198 regimes. Using MCMC, the program simulates a posterior distribution of rate shift regimes, from
199 which marginal posterior rate distributions can be extracted for each tip in the phylogeny. Priors
200 for BAMM analyses were set using default settings from the `setBAMMpriors` function from
201 `BAMMtools` (Rabosky et al. 2014). The prior parameterizations specified by this function ensure
202 that the prior density on relative rate changes across the tree is invariant to the scale of the tree
203 (e.g., multiplying branch lengths by 10^6 will not change inferences about relative rates across the
204 tree). We denote BAMM tip speciation rates (mean of the marginal posterior) as λ_{BAMM} . As
205 BAMM also estimates extinction rates for each regime, we also calculated tip-specific net
206 diversification rate as $\lambda_{BAMM} - \mu_{BAMM}$, denoted as r_{BAMM} .

207

208 Tip rate metrics estimate speciation, not net diversification

209 As suggested previously (Belmaker & Jetz 2015; supplemental analyses in Jetz et al.
210 2012), DR and presumably other tip-based measurements, more accurately estimate the rate of
211 speciation than the rate of net diversification. However, numerous studies continue to refer to DR

212 as a measure of net diversification (Marin & Hedges 2016; Oliveira et al. 2016; Cai et al. 2017;
213 Quintero & Jetz 2018; and many others). This is incorrect and it is straightforward to
214 demonstrate that λ_{TB} , λ_{ND} and λ_{DR} are more reliable measures of speciation rates and not net
215 diversification rates, at least when extinction is moderate to high.

216 To illustrate this property of the metrics, we applied all approaches to constant-rate birth-
217 death phylogenies simulated across a range of extinction fractions ($\varepsilon = \mu / \lambda$), including pure-
218 birth trees ($\varepsilon = 0$) as well as trees exhibiting very high turnover ($\varepsilon = 1$). To evaluate accuracy of
219 speciation estimates as a function of ε , we generated 1000 phylogenies with 100 tips each, where
220 λ and ε were drawn from uniform distributions (λ : [0.05, 0.3]; ε : [0, 1]). Importantly, when λ is
221 sampled uniformly with respect to ε , the distribution of r is not uniform: the mean, range and
222 variance in r decrease dramatically as ε increases. To evaluate the accuracy of r as a function of
223 ε , we thus generated a second set of trees by sampling r and ε from uniform distributions (r :
224 [0.05, 0.3], ε [0, 1]). As a result, λ has constant mean and variance with respect to ε in the first
225 set of simulations, and the same is true for r in the second set of simulations (Figure S1). All
226 phylogeny simulations were conducted with the TreeSim package in R (Stadler 2011).

227 We compared tip rate metrics to true speciation rates λ_{TRUE} (with the first simulation set)
228 and to true net diversification rates r_{TRUE} (with the second simulation set). We evaluated mean
229 per-tip accuracy of the tip rate metrics with three measures of error:

230 mean absolute error = $\sum_{i=1}^{N_i} |\lambda_i - \lambda_{TRUE_i}| / N$

231 RMSE = $\sqrt{\sum_{i=1}^{N_i} (\lambda_i - \lambda_{TRUE_i})^2 / N}$

232 mean proportional error = $\sum_{i=1}^{N_i} \frac{\lambda_i - \lambda_{TRUE_i}}{\lambda_{TRUE_i}} / N$

233 where λ_i is the estimated tip rate for species i out of N total species, λ_{TRUE} is the true tip rate.
234 Mean absolute error and root mean square error capture the magnitude in error in tip rates, and
235 mean proportional error quantifies the bias in tip rates, as a function of the true tip rates
236 (Rabosky et al. 2014). In analyses below, all error summaries yield generally congruent results;
237 results for mean absolute error are presented in the main text, and others in supplement.

238

239 Assessment of tip rate metrics

240 We tested the performance of the metrics by compiling publicly-available datasets from a
241 number of simulation-based studies (Table 1). By focusing on simulations from previously-
242 published work, we thus ensured that the simulation process itself was effectively blinded to the
243 objectives of the present study. We further note that our trial datasets included several studies
244 that were critical of BAMM (Moore et al. 2016, Meyer & Wiens 2017). These simulated trees
245 include rate heterogeneity in time and across lineages. Together, these phylogenies present a
246 wide range of tree sizes and diversification rate shifts, providing an ideal comparative dataset for
247 our purposes. To more easily distinguish between these tree types in the text, we refer to the
248 BAMM-type, multi-regime time-constant phylogenies simply as “multi-regime”, and the multi-
249 regime diversity-dependent phylogenies simply as “diversity-dependent”, even though discrete
250 rate shifts are present in both types of trees. In addition to discrete-shift scenarios (e.g., BAMM-
251 type process), we simulated phylogenies under an “evolving rates” model of diversification
252 (Rabosky 2010; as corrected in Beaulieu & O’Meara 2015) to explore performance of tip rate
253 metrics when diversification rates change continuously and independently along branches, as
254 might occur if diversification rates are correlated with an underlying continuous trait (FitzJohn
255 2010). In these simulations, we allowed the logarithm of λ to evolve across the tree under a
256 Brownian motion process, while holding ε constant. The magnitude of rate heterogeneity among
257 branches is controlled by the diffusion parameter σ , where greater values lead to greater
258 heterogeneity in speciation rates. Although published phylogenies with rate data were
259 unavailable for this simulation scenario, we used simulation code and parameters taken directly
260 from Beaulieu & O’Meara (2015) to generate trees with similar statistical properties to those in
261 their study. Simulations were performed with the following parameters: $\lambda = 0.078, 0.103, 0.145,$
262 0.249 and $\varepsilon = 0.0, 0.25, 0.50, 0.75$. We simulated 100 phylogenies for each (λ, ε) pair, and for
263 three values of σ ($\sigma = 0.03, 0.06, 0.12$). We evaluated tip rate accuracy by comparing estimated
264 to true tip rates, using the absolute and proportional error metrics described above. We also
265 examined the correlation between true and estimated tip rates, combining tip rates from all
266 phylogenies generated under the same class of diversification process, and visualizing these data
267 as density scatterplots, generated with the LSD package in R (Schwalb et al. 2018), where colors
268 indicate the density of points.

269 Size of diversification rate regimes might be an important factor in a tip rate metric’s
270 ability to accurately estimate rates. For example, BAMM’s statistical power in detecting a shift

271 to a new rate regime is a function of the number of taxa in that rate regime, and tip rates for taxa
272 from small regimes will more likely be parameterized according to the larger parent regime or
273 the tree-wide average rate (Rabosky, Mitchell & Chang 2017); this is the expected behavior
274 when BAMM fails to identify a rate shift. However, non-model-based approaches such as those
275 examined in this study might be more accurate for small regimes. To explore how rate regime
276 size influences the accuracy of tip rate metrics, we calculated the mean tip rate for each true rate
277 regime from all multi-regime phylogenies (simulation datasets from Moore et al. 2016; Rabosky,
278 Mitchell & Chang 2017; Meyer & Wiens 2017; Mitchell, Etienne & Rabosky 2018). We then
279 calculated the Pearson correlation coefficient and the slope of a linear model between true and
280 estimated mean regime rates. We explored the performance of all metrics with respect to regime
281 sample size, as in Rabosky, Mitchell & Chang (2017: Figure 13). For comparison, we repeated
282 all performance summaries on tip rates estimated by applying a simple constant-rate birth-death
283 (CRBD) process to each simulated phylogeny. This exercise is an important control, because it
284 indicates how much error we would expect for each simulated phylogeny under the simplifying
285 (incorrect) assumption that rates are constant among lineages and through time for each dataset.

286

287 **Results**

288 **Speciation or net diversification?**

289 As expected, the tip rate metrics examined in this study are more accurate estimators of
290 the rate of speciation (λ) and not the net rate of species diversification (r). Mean absolute error
291 increased exponentially with respect to the extinction fraction ε (Figure 1). However, mean
292 absolute error in speciation rate was largely invariant with respect to ε (0.95 quantile of r -based
293 and λ -based mean absolute error for λ_{DR} : 2.28 and 0.17, respectively). Nearly identical patterns
294 were found with RMSE (Figure S2). Note that r and λ for these simulations were drawn from
295 identical uniform distributions, and absolute error in the rates is thus comparable. Proportional
296 error generally exhibited the same pattern, and in terms of λ versus r , differences in speciation-
297 based error varied across ε (Figure S3). There was a weak but significant trend towards
298 progressively greater underestimation of speciation rates with increasing values of relative
299 extinction (linear model slopes: -0.08, -0.014, -0.011 for λ_{ND} , λ_{DR} and λ_{BAMM} , respectively).
300 Overall, error was highest for λ_{TB} by an order of magnitude (Figure S4), and decreased

301 progressively with λ_{ND} and λ_{DR} , with the lowest overall error in λ_{BAMM} . BAMM estimates of net
302 diversification rate were relatively accurate, except at the highest values of ε (Figures 1, S2, S3).

303

304 Tip rate accuracy across rate-variable phylogenies

305 Tip rates estimated with BAMM were consistently more accurate than those obtained
306 using the other methods across all diversification scenarios considered, including multi-regime,
307 diversity-dependent and evolving rates trees (Figure 2). λ_{DR} was the second-most accurate
308 metric, although its relationship with true rates was substantially weaker than λ_{BAMM} . λ_{ND} and
309 λ_{TB} were correlated with true rates but performed relatively poorly overall. However, λ_{TB}
310 performed better than λ_{ND} , and just as well as λ_{DR} at estimating speciation rates for diversity-
311 dependent trees (Figure 2, S5). All metrics performed best for multi-regime trees, followed by
312 evolving rates and diversity-dependent trees, respectively. For diversity-dependent trees, λ_{ND}
313 rates are effectively uncorrelated with the true rates (Figure 2). Additionally, the performance of
314 the different tip rate metrics for multi-regime phylogenies is not sensitive to the source of the
315 simulated phylogenies (Figure S6). We found that BAMM substantially outperformed all other
316 metrics on datasets from studies that independently assessed BAMM's performance (Figure S6:
317 Moore et al. 2016; Meyer & Wiens 2017). Tip rates were also generally but more weakly
318 correlated with true net diversification rates, with the exception of λ_{ND} , which was uncorrelated
319 with true rates for diversity-dependent trees, presumably because this metric equally weights the
320 full depth of the tree (Figure S7).

321 In terms of mean per-tip error, λ_{BAMM} consistently outperformed the other metrics for
322 multi-regime, diversity-dependent and evolving rates trees (Figures 3, S8). Error in λ_{BAMM}
323 increased as a function of rate heterogeneity for evolving rate phylogenies, but was largely
324 independent of the magnitude of rate heterogeneity for the other scenarios. λ_{DR} generally
325 exhibited greater error than λ_{BAMM} , and this error increased as a function of the level of
326 heterogeneity for both the evolving rates and multi-regime trees. Error in λ_{DR} was generally
327 invariant to the number of rate regimes for the diversity-dependent scenarios. However, λ_{DR}
328 tended to have greater error than tip estimates from a simple model that assumes no variation in
329 rates through time or among lineages (λ_{CRBD} ; all tips assigned the tree-wide CRBD rate). λ_{ND}
330 performed somewhat similarly to λ_{DR} for constant-rate and evolving rates trees, but worse for

331 diversity-dependent trees. Error in λ_{TB} increased with increasing rate heterogeneity for constant-
332 rate and evolving rates trees, but was relatively unaffected by rate heterogeneity in diversity-
333 dependent trees (Figure S9). However, error for this metric was far greater than for all other tip
334 metrics.

335

336 Effects of regime size on performance

337 Both metrics of performance assessment – the Pearson correlation and OLS slope –
338 generally increased with increasing regime size (Figure 4). This was found to be true for all tip
339 rate metrics, although λ_{TB} and λ_{ND} never achieved high performance. λ_{DR} tended to perform
340 better than other metrics when small rate regimes were included (e.g., 10 tips or fewer);
341 however, the slope between estimated and true rates was greater than 1 across the majority of
342 minimum regime sizes, indicating that λ_{DR} overestimates speciation rates (see also Figure S3).
343 Similar patterns were observed for net diversification rates with λ_{DR} , but the magnitude of the
344 overestimation was greater than for speciation (Figure S10). λ_{BAMM} , in contrast, approached a
345 slope of 1 when estimating speciation rates and slightly underestimated net diversification rates
346 (regimes with > 30 tips: OLS slope = 0.96 for λ , 0.87 for r).

347 Absolute error in regime mean tip rates was lowest for λ_{DR} and λ_{BAMM} , regardless of the
348 size of the rate regime (Figure 5). BAMM's ability to accurately estimate tip rates improved with
349 regime size, whereas absolute error was relatively consistent across regime sizes for λ_{DR} for
350 regimes greater than 10 species. We also found that λ_{DR} slightly outperformed λ_{BAMM} for small
351 rate regimes.

352 Note that, in Figures 4 and 5, each rate regime is treated as a single data point. Rate
353 regimes of sizes 1000, 100, and 1 tip are equivalent under this method of error assessment.
354 Figure 4 assesses how well these methods estimate rates for individual regimes, regardless of the
355 size of those regimes. In contrast, Figures 1-3 ask how well these methods perform at estimating
356 rates for a given tip.

357

358 **Discussion**

359 We assessed several tip rate metrics and confirmed that these are more accurate
360 estimators of the rate of speciation, rather than net diversification (Figures 1, 4, S7, S10). This
361 distinction was especially pronounced at high extinction fractions, where the rate of lineage

362 turnover is high, and rates of speciation and net diversification have the potential to be more
363 divergent. These results are consistent with supplemental analyses performed in Jetz et al. 2012.
364 It is also important to note that recent extinction will have a much greater influence on these
365 metrics than extinction events deeper in time (Quental & Marshall 2011). Net diversification rate
366 is a critical determinant of species richness, yet this quantity is potentially independent of the
367 underlying rate of speciation. Misinterpretation of tip rate metrics could therefore lead to highly
368 misleading perspectives on large-scale diversity dynamics. As we demonstrate (Figures 1, S2,
369 S3), tip rate metrics (λ_{ND} , λ_{DR}) provide relatively little information about net diversification, and
370 high values of these metrics are fully consistent with equilibrational models of speciation where the
371 true net diversification rate is zero. Thus, λ_{DR} and λ_{ND} should not be used to support claims
372 about the dynamics of species richness or net diversification *per se* without independent
373 evidence bearing on plausible levels of extinction.

374 In terms of accuracy, we found that BAMM performed better than non-model-based
375 metrics across all datasets we considered: estimated tip rates were most highly correlated with
376 true tip rates, and mean per-tip error in rates was lower across a range of rate-variable simulation
377 scenarios. This performance is likely to be at least partially due to the inclusion of extinction in
378 the BAMM inference model. BAMM is expected to perform well for phylogenies with discrete
379 shifts in diversification rates as this type of rate variation is most consistent with BAMM's
380 assumptions (Rabosky 2014; Mitchell & Rabosky 2016; Rabosky, Mitchell & Chang 2017;
381 Mitchell, Etienne & Rabosky 2018). However, BAMM performed surprisingly well for the
382 evolving rates phylogenies, which conform poorly to the assumptions of the inference model. In
383 these trees, the rate of speciation changes continuously under a diffusion process, and as a result,
384 the phylogeny exhibits rate heterogeneity without discrete rate shifts.

385 On evolving rates phylogenies, λ_{BAMM} performed better than λ_{DR} (Figure 2; Spearman's
386 ρ for $\lambda_{BAMM} = 0.83$, ρ for $\lambda_{DR} = 0.62$), despite the fact that λ_{DR} does not rely on the detection of
387 distinct rate regimes to estimate tip rates (Figure 5). λ_{BAMM} also exhibited the lowest mean per-
388 tip error across varying levels of rate heterogeneity (Figure 3).

389 Why do λ_{BAMM} and λ_{DR} exhibit such striking differences in performance across the
390 simulation scenarios considered here? To illustrate the differences between inference under these
391 metrics, we compared true tip rates to λ_{BAMM} and to λ_{DR} on a simulated birth-death tree with a
392 single rate shift (Figure 6), as well as on one evolving rates tree simulated for this study (Figure

393 7). It is clear that if BAMM has the statistical power to detect true rate shifts, then it will perform
394 well under rate shift scenarios. In contrast, λ_{DR} tracks true rate shifts but exhibits high sample
395 variance. With an evolving rates tree (Figure 7), the simulation model is very different from the
396 inference model in BAMM. However, it conservatively places rate shifts in order to
397 accommodate rate heterogeneity that is spread across the phylogeny under a rather different
398 model of rate variation. λ_{DR} also broadly tracks the overall pattern of the true rates, but the
399 variance in the corresponding estimates is so high that performance is negatively affected. If we
400 calculate mean (absolute) per-tip error in λ_{BAMM} and λ_{DR} , the error is relatively similar between
401 λ_{BAMM} and λ_{DR} , but the variance in per-tip error for λ_{DR} is higher. Overall, BAMM exhibited
402 substantially lower error than λ_{DR} under precisely this scenario (Figure 3).

403 Thus, although BAMM is conservative in the estimation of tip rates relative to λ_{DR} , the
404 method exhibits lower overall error. It appears that λ_{DR} can recover more subtle rate
405 heterogeneity relative to BAMM (see Rabosky, Mitchell & Chang 2017 for discussion of power
406 in BAMM), but this apparent power advantage comes at the cost of increased variance (error) in
407 the resulting estimates. Remarkably, on a per-tip basis, we find that a simple constant-rate birth-
408 death process (λ_{CRBD}) frequently yields tip estimates with lower median error and less error
409 variance than those obtained with λ_{DR} (Figure 3), despite the simplifying (and incorrect)
410 assumption that rates are identical across all tips in a given tree. For example, across all multi-
411 regime simulations (Figure 3), λ_{CRBD} point estimates were more accurate than the corresponding
412 λ_{DR} point estimates for 84% of trees in the simulations; for λ_{BAMM} , the λ_{CRBD} estimates were
413 more accurate for a much smaller fraction of the total (36%). Similar results were noted for
414 diversity-dependent (λ_{CRBD} more accurate than 98% of λ_{DR} estimates, versus 15% of λ_{BAMM}
415 estimates) and evolving rates trees (λ_{CRBD} more accurate than 93% of λ_{DR} estimates, versus 36%
416 of λ_{BAMM} estimates). Given that λ_{DR} can and does track true heterogeneity in speciation rate
417 (Figures 6, 7), this pattern suggests that the metric is especially sensitive to the stochastic
418 variation in branch lengths that can emerge even when all tips have the same underlying
419 speciation rate.

420 Regardless of the performance summaries presented in this article, important questions
421 remain with respect to how well tip rate metrics can estimate the true rate of speciation from
422 empirical phylogenies. The phylogenies analyzed in this study were simulated under idealized

423 processes and neglect potential biases and sources of uncertainty that are present in real datasets.
424 For example, if the process of speciation takes time to complete, as is generally believed to be
425 the case (i.e., the protracted speciation process; Rosindell et al. 2010; Etienne and Rosindell
426 2012), then the most recent speciation events may still be on-going at the present and typical
427 species-level molecular phylogenies may fail to recognize these events. This will lead to an
428 overestimation of terminal branch lengths, as some terminal branches potentially include
429 incipient species. A related bias might arise due to incomplete taxon sampling, which
430 disproportionately affects the length of terminal (or otherwise recent) branch lengths (Pybus &
431 Harvey 2000). Likewise, variation in taxonomic practice across a phylogeny might lead to
432 spurious rate variation, particularly if different species concepts are used, or if some clades in the
433 phylogeny – but not others – have been subject to population genetic analysis or screens for
434 cryptic species diversity. Additionally, it has been shown that BAMM and other methods may
435 fail to infer accurate speciation rate dynamics if the phylogeny is in diversity decline – that is,
436 when extinction rates increase towards the present and ultimately exceed speciation rates
437 (Quental & Marshall 2011; Burin et al. 2018). A major, if obvious, caveat in the interpretation of
438 tip rates is that they apply to recent speciation rates and are necessarily limited with respect to
439 inferences about historical variation in speciation rate.

440 The greater the importance of the terminal branches in tip rate metrics, the greater the
441 impact these biases might have on tip rate estimates. On one end of the spectrum, metrics such as
442 λ_{TB} will be very sensitive to such biases as they rely exclusively on terminal branch lengths.
443 Such approaches may retain utility as summary statistics (e.g., Bromham, Hua & Cardillo 2016),
444 but we found that λ_{TB} exhibited the greatest amount of error in estimating speciation rates. On
445 the other end of the spectrum, a metric like λ_{ND} would be minimally impacted as this metric is
446 attempting to capture an average speciation rate over an entire root-to-tip path and does not
447 upweight the contribution of recent branch lengths. λ_{DR} is likely somewhere in the middle of this
448 spectrum, as it gives decreasing weight to branches towards the root. λ_{BAMM} is potentially
449 sensitive to such issues as well, although it may be possible to analytically correct for some
450 biases in the mechanics of the model itself (e.g., Rosindell et al. 2010; Etienne and Rosindell
451 2012).

452 Potential empirical biases aside, tip rates present a number of practical advantages in the
453 study of diversification rate variation. First, tip rates can be summarized and compared across

454 non-monophyletic assemblages of species (Jetz et al. 2012; Kennedy et al. 2016; Belmaker &
455 Jetz 2015; Oliveira et al. 2016; Quintero & Jetz 2018; Rabosky et al. 2018), making it possible to
456 summarize rate characteristics of entire communities or regional assemblages of species. Second,
457 estimation of rates at the present should be more robust to the influence of extinction, as
458 extinction can erase the history of lineage splitting deeper in the phylogeny (Nee et al. 1994;
459 Nee, May & Harvey 1994; Rabosky & Lovette 2008). Third, tip-specific rates can be paired with
460 species-specific trait values or geographic attributes in order to test potential trait- or geography-
461 dependent speciation rates (Freckleton, Phillimore & Pagel 2008; Jetz et al. 2012, Rabosky &
462 Goldberg 2017; Harvey & Rabosky 2017). Tip rates make it possible to relax strong assumptions
463 of rate homogeneity within character states, which are inherent to certain trait-dependent models,
464 including BiSSE and GeoSSE (Maddison, Midford & Otto 2007; Goldberg, Lancaster & Ree
465 2011; Ng & Smith 2014). Recent work has provided a conceptually rich and robust interpretive
466 framework for SSE models that does not assume rate-constancy within character states (Beaulieu
467 & O'Meara 2016; Caetano, O'Meara & Beaulieu 2018), but tip rates nonetheless can provide an
468 important check on results obtained with SSE models by providing a direct means of visualizing
469 the relationship between branch lengths and character states (Bromham, Hua & Cardillo 2016;
470 Hua & Bromham 2016; Harvey & Rabosky 2017). Visual inspection of data in this fashion has
471 the potential to reduce false positives by calling attention to potential outliers and other sources
472 of model inadequacy (Maddison & FitzJohn 2014; Rabosky & Goldberg 2015). A final
473 advantage for non-model-based tip rates, especially λ_{DR} , is that they can profitably be applied to
474 extremely large phylogenies: there are few computational limits to using them on phylogenies
475 with tens of thousands of tips or more, in contrast to formal model-based approaches for which
476 BAMM, HiSSE (Hidden State Speciation and Extinction; Beaulieu & O'Meara 2016), and other
477 methods are poorly suited. This computational efficiency also lends itself to more readily
478 accounting for phylogenetic uncertainty, because tip rate metrics can rapidly be computed across
479 posterior distributions of phylogenies and averaged (for example, see Jetz et al. 2012; Rabosky et
480 al. 2018).

481 In summary, tip rates offer a number of theoretical and practical advantages, particularly
482 in the study of associations between traits and diversification. We found that λ_{BAMM}
483 outperformed other metrics evaluated in this study and proved to be relatively accurate, even
484 under diversification scenarios that depart from the BAMM inference model. λ_{DR}

485 underperformed in comparison to λ_{BAMM} , but in many cases still did reasonably well, particularly
486 for small rate regimes. Despite our performance results, λ_{DR} is likely to remain a useful tool in
487 the study of trait- and geography-dependent diversification (Rabosky & Goldberg 2017; Harvey
488 & Rabosky 2017).

489

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498

499 **Data accessibility**

500 All trees, data and code necessary to repeat the analyses have been deposited in the Dryad
501 repository: <https://datadryad.org/resource/doi:10.5061/dryad.5hr25nv> (Title & Rabosky 2019).

502 **Author contributions**

503 P.O.T. and D.L.R. designed the project. P.O.T. assembled the datasets and performed all
504 analyses. P.O.T. and D.L.R. wrote the manuscript. Both authors contributed critically to
505 subsequent drafts and approved the final publication.

506

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667 **Figures/Captions**

668

669 **Table 1.** Summary of simulated phylogenies used in this study.

simulation model	number		regime	source
	of trees	tree size	number	
single-regime, constant-rate birth-death	100	100	1	Mitchell & Rabosky 2016
single- and multi-regime, constant-rate birth-death	100	51-148	1-6	Moore et al. 2016
single- and multi-regime, constant-rate birth-death	400	10-4296	1-67	Rabosky, Mitchell & Chang 2017
multi-regime, constant-rate birth-death	20	939-3708	11	Meyer & Wiens 2017
single- and multi-regime, constant-rate birth-death	188	4-3955	1-73	Mitchell, Etienne & Rabosky 2018
single-regime, constant-rate birth-death, lambda uniform	1000	100	1	this study
single-regime, constant-rate birth-death, net diversification uniform	1000	100	1	this study
pure birth root regime, 1-4 discrete shifts to diversity-dependent regimes	1200	54-882	1-5	Rabosky 2014; Mitchell & Rabosky 2016
speciation rate evolves via diffusion process	1200	25-1208	1	Rabosky 2010; Beaulieu & O'Meara 2015; Rabosky 2016; this study

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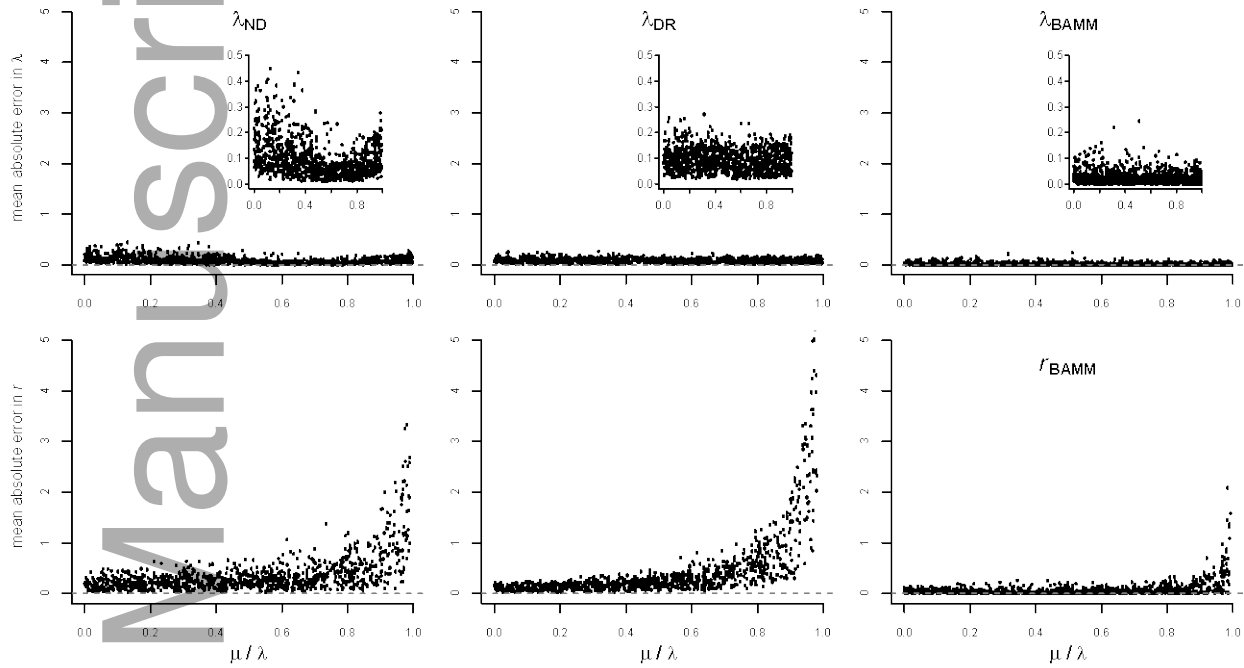
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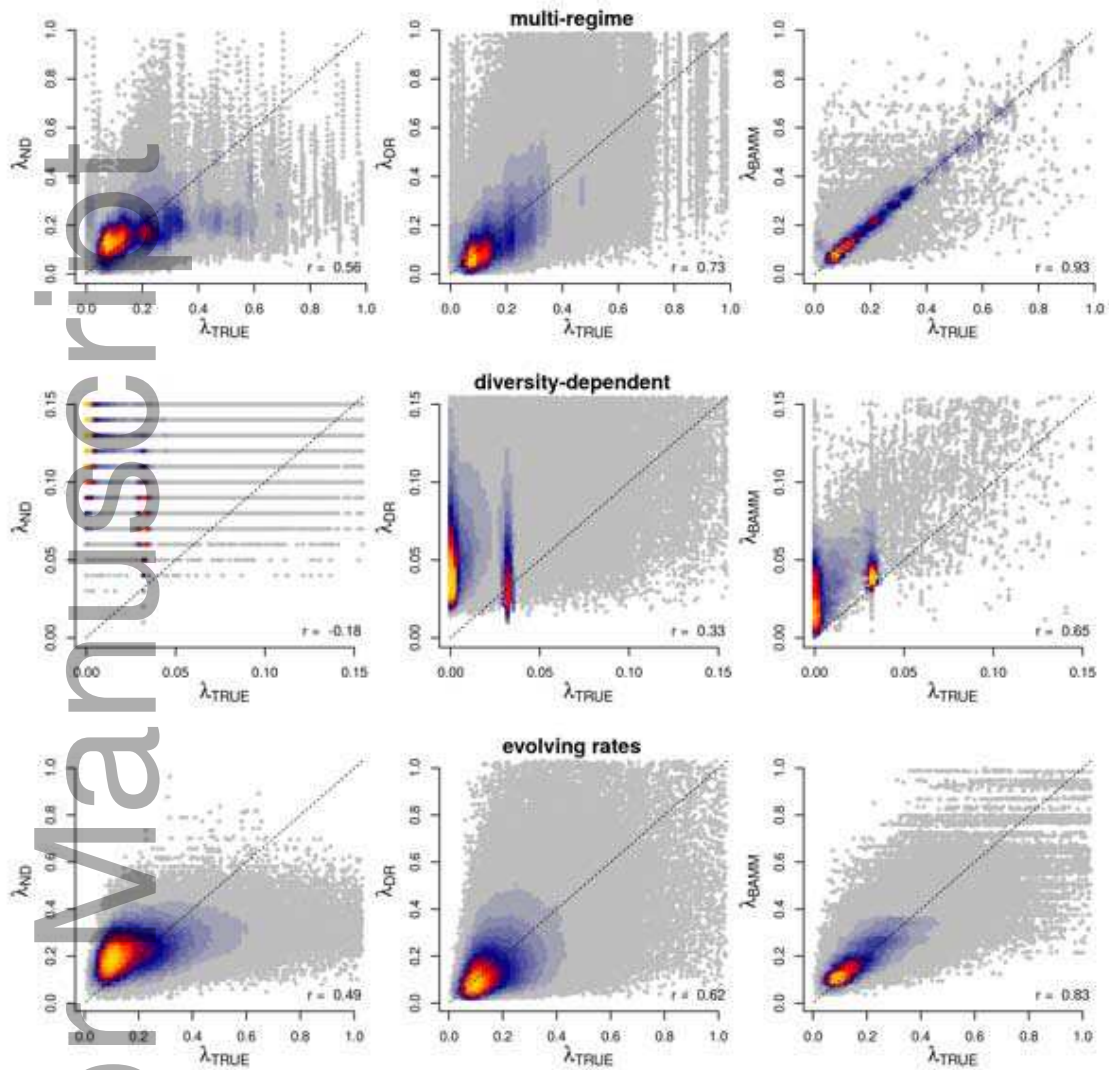
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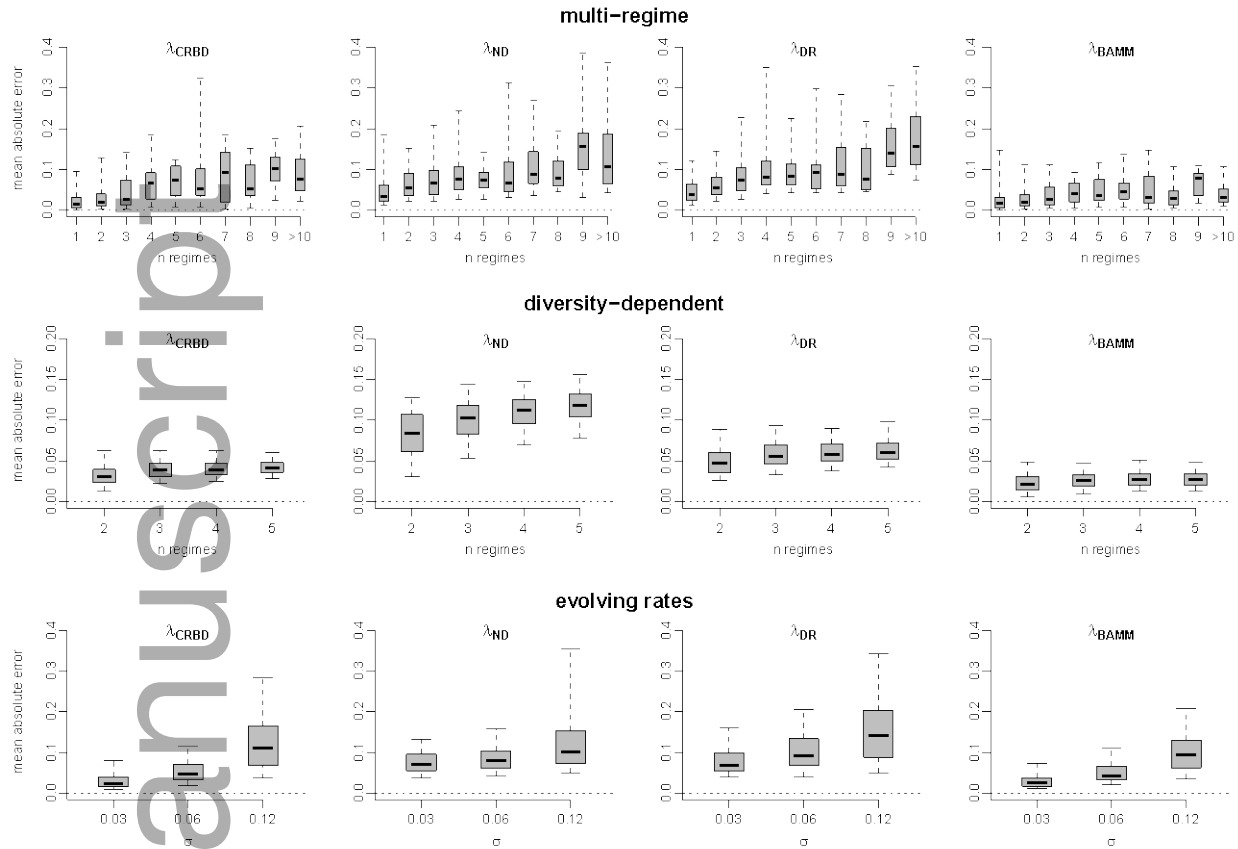
681
682 **Figure 1.** Mean absolute error in λ (top) and r (bottom) for three different tip rate metrics, across
683 a range of relative extinction rates. For BAMM, the estimated speciation and net diversification
684 rates are presented in the top and bottom panels, respectively. Absolute error of zero implies
685 perfect accuracy. Inset plots show error in λ with truncated y-axis scale to facilitate comparison
686 among metrics. All tip rate metrics track λ more accurately than they track r . See Figure S4 for
687 λ_{TB} , which performed much worse than the other metrics.

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694 **Figure 2.** True tip rates (λ_{TRUE}) in relation to estimated tip rates. Tip rates were compared
 695 separately for different major categories of phylogeny simulations (rows) and are plotted
 696 separately by inference method (columns). Plotting region is restricted to the 99th percentile of
 697 true rates, but Spearman correlations between true and estimated rates (lower right of each figure
 698 panel) are based on the full range of the data. Colors indicate the density of points in the scatter
 699 plots. The horizontal gaps in λ_{ND} for diversity-dependent trees are an artefact of all trees having
 700 the same crown age. λ_{BAMM} exhibited the strongest correlation with true rates for all simulation
 701 categories.



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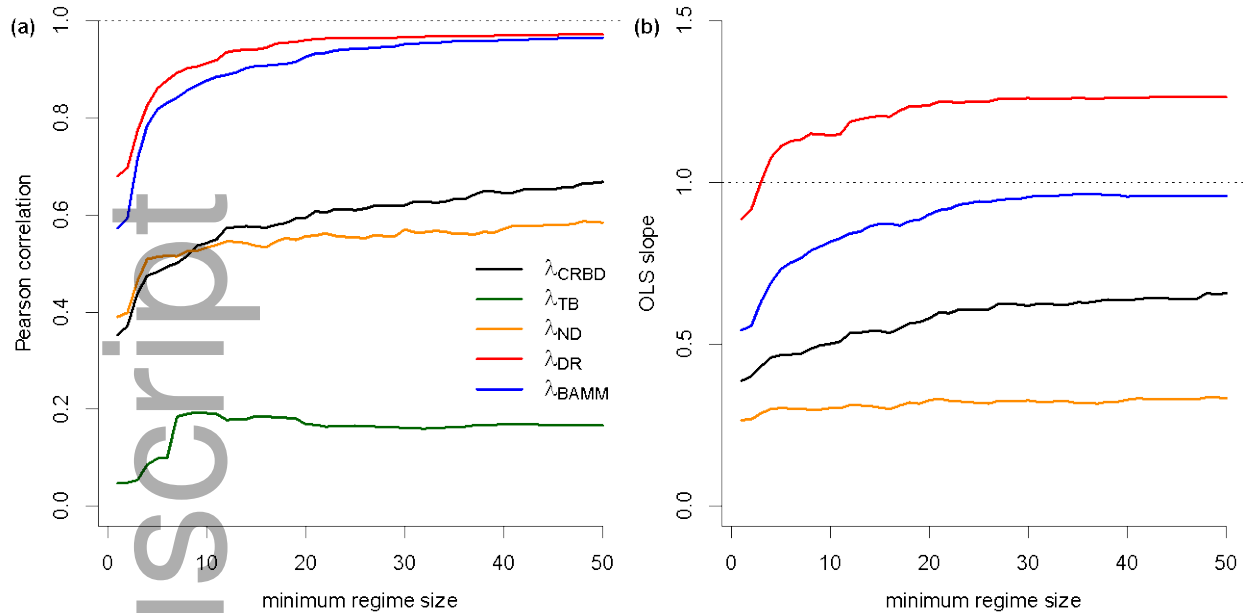
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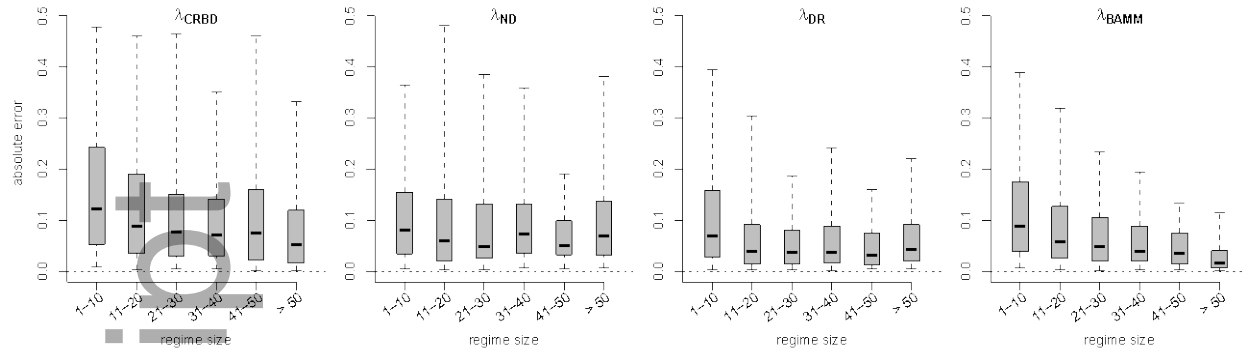
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Figure 3. Mean per-tip absolute error in speciation rates as a function of the magnitude of rate heterogeneity in each simulated phylogeny. Results are presented separately for different categories of rate variation (Table 1); left column shows estimates from a constant-rate birth-death model for reference. The boxes and whiskers represent the 0.25 – 0.75, and the 0.05 – 0.95 quantile ranges, respectively. In some cases, λ_{ND} and λ_{DR} had more error than a simple CRBD model with no variation in tip rates. λ_{BAMM} had the least amount of error across all amounts of rate heterogeneity. See Figure S9 for λ_{TB} .



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Figure 4. Performance of tip rate metrics as a function of regime size, including Pearson correlation (a) and OLS regression slope (b) for mean rates with respect to λ_{TRUE} . λ_{DR} and λ_{BAMM} outperform the other metrics when summarized in this fashion, although λ_{DR} tends to overestimate the rate of speciation. The x-axis denotes the minimum regime size across which performance was summarized. For example, $x = 20$ corresponds to the correlations and slopes computed for all regimes with 20 or more tips; a value of $x = 1$ is the corresponding results for all regimes. The OLS slope for λ_{TB} is not visible as it ranges between 7 and 9.



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727 **Figure 5.** Mean per-regime absolute error in relation to true rate regime size, as binned into 10
 728 size categories. The boxes and whiskers represent the 0.25 – 0.75, and the 0.05 – 0.95 quantile
 729 ranges, respectively. Perfectly estimated rates have an error of zero. λ_{DR} and λ_{BAMM} exhibit the
 730 least error when averaged by regimes, and λ_{DR} does slightly better for small clades (10-clade
 731 median error 0.07 for λ_{DR} , and 0.08 for λ_{BAMM}).

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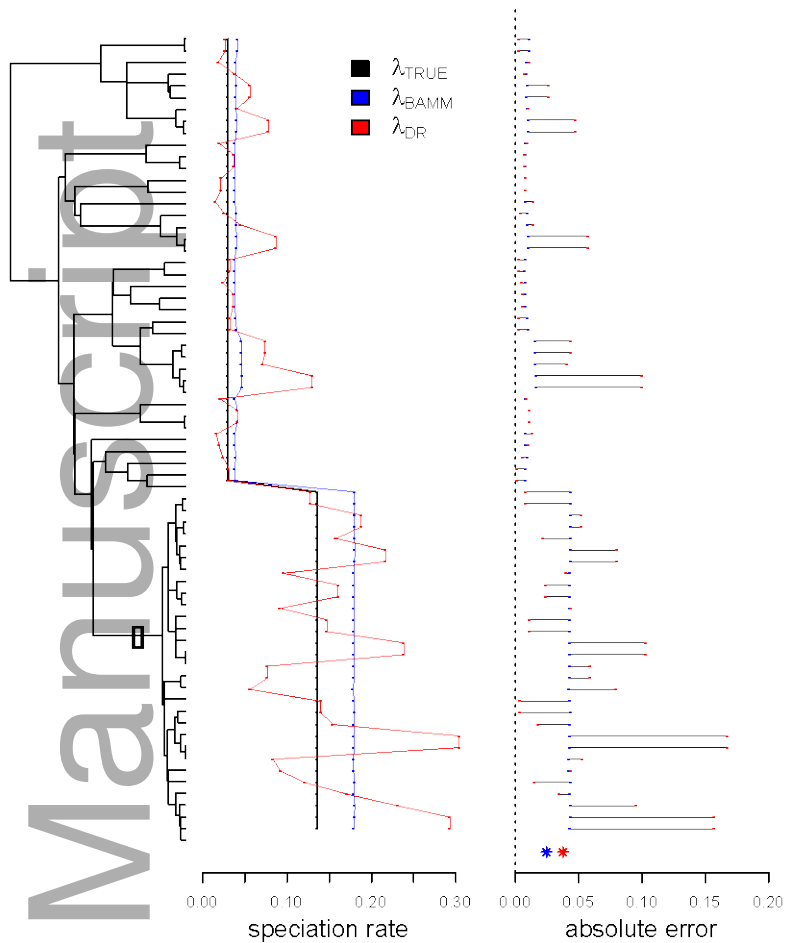
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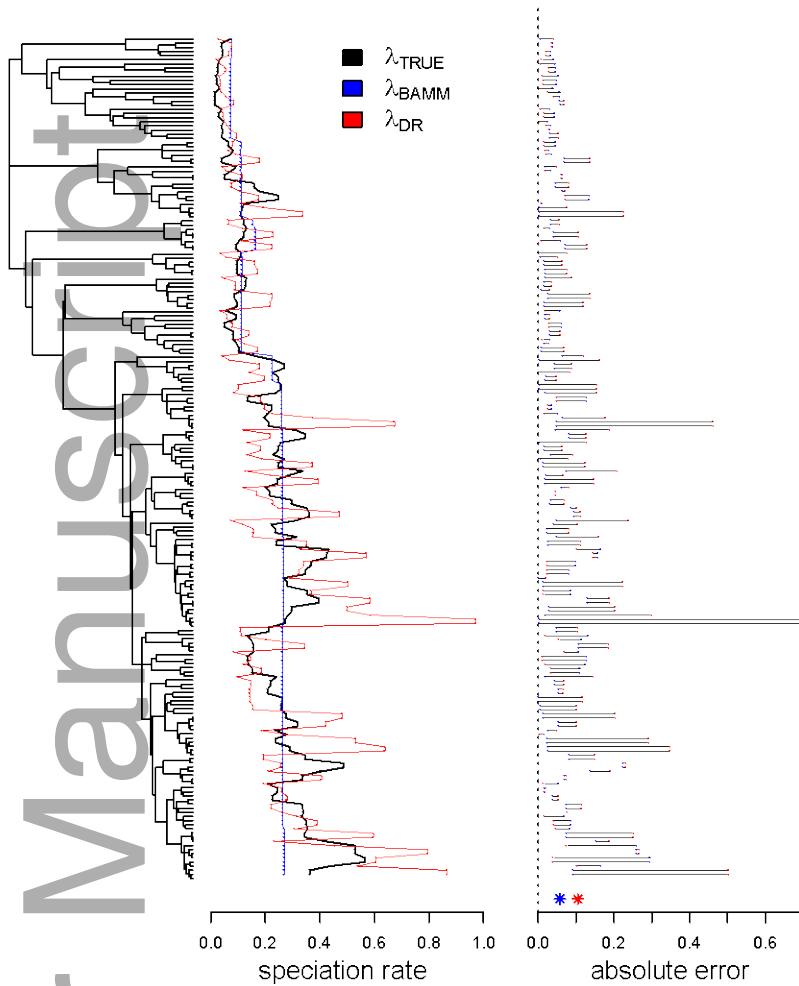
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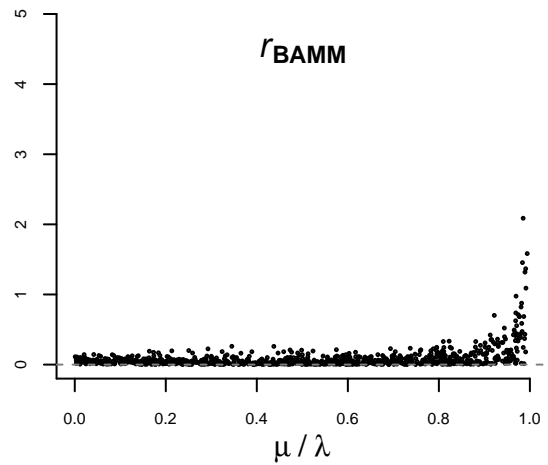
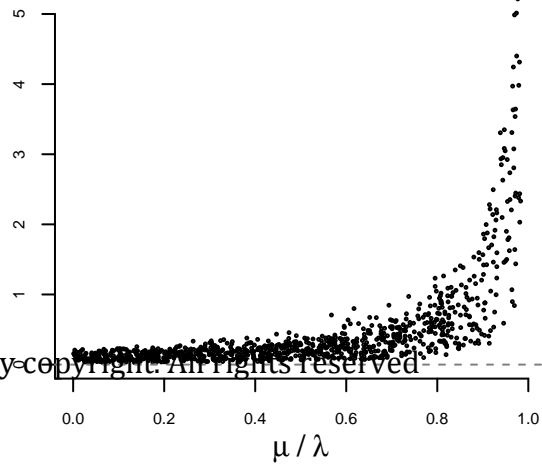
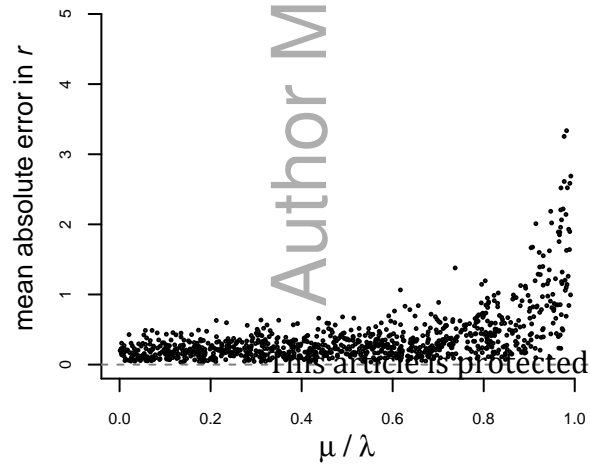
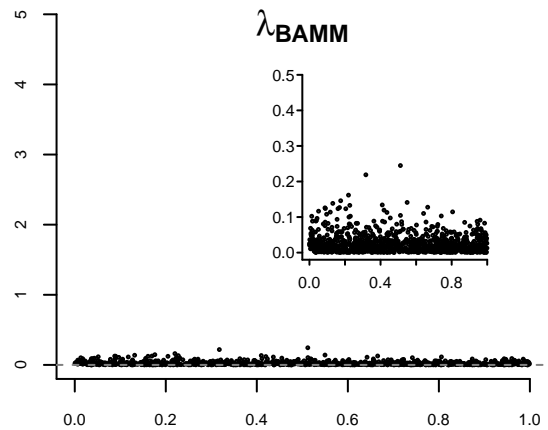
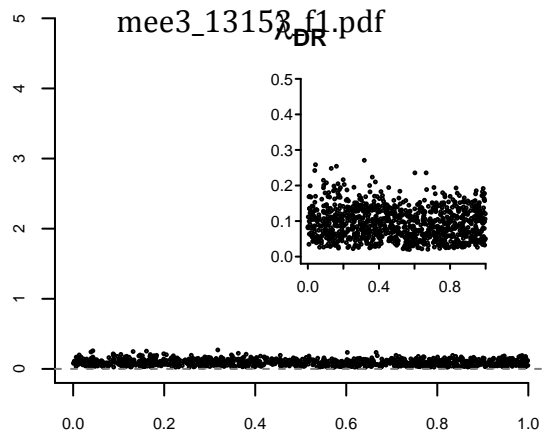
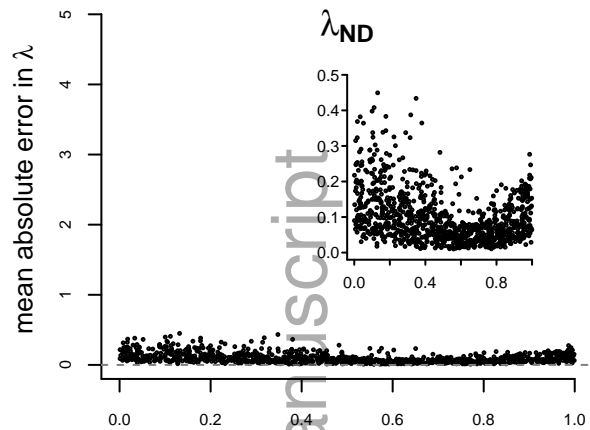


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744 **Figure 6.** Relationship between λ_{TRUE} , λ_{BAMM} , and λ_{DR} for a simulated phylogeny containing a
 745 single rate shift (orange circle). Subplots to the right of the tree illustrate true and estimated rates
 746 for each tip (left) and corresponding absolute error (right). Asterisks at the bottom denote mean
 747 per-tip error in tip rate metrics. Mean per-tip error is relatively low and similar between λ_{DR} and
 748 λ_{BAMM} , but the sample variance in λ_{DR} tip rates is high. In this example, the variance in absolute
 749 per-tip error in λ_{DR} is 0.002 versus 0.0003 for λ_{BAMM} . On average, λ_{DR} tends to either
 750 overestimate or underestimate rates relative to λ_{BAMM} , even if the mean per-tip error is relatively
 751 low for both metrics.
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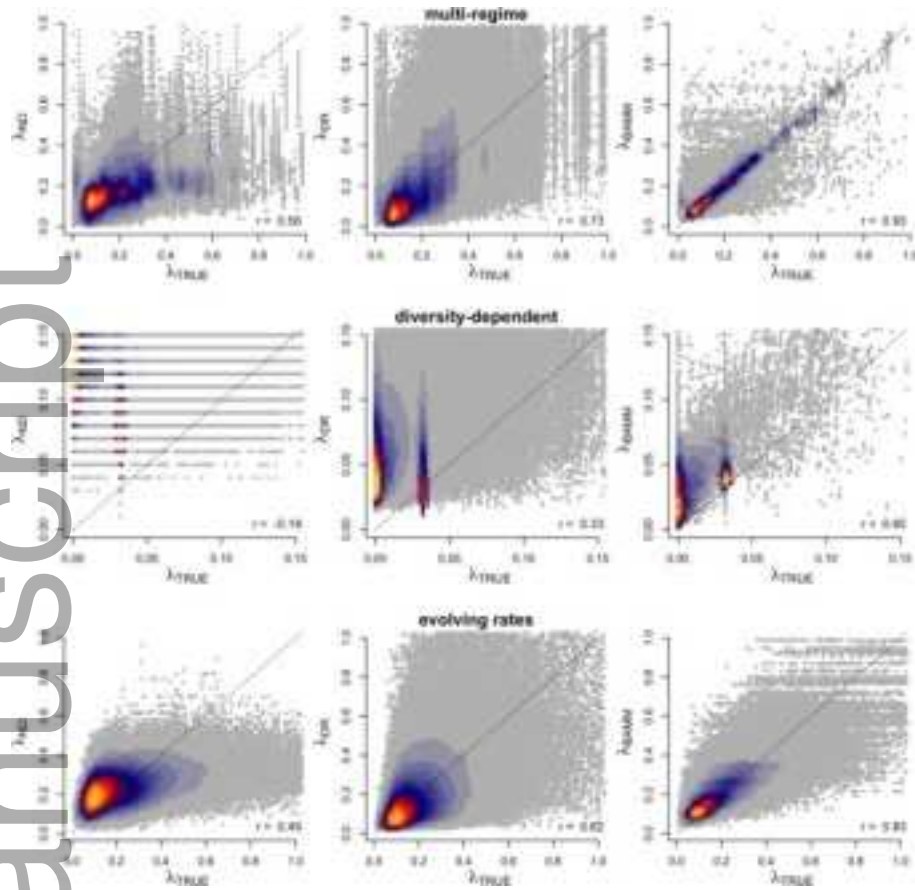


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 754 **Figure 7.** Relationship between λ_{DR} , λ_{BAMM} , and λ_{DR} for a phylogeny simulated under an
 755 “evolving rates” model, such that the speciation rate itself varies under a diffusion model. See
 756 Figure 6 for additional details. Neither metric is particularly well equipped to infer the true rate
 757 variation in this case. However λ_{BAMM} ’s conservative estimates are still more accurate relative to
 758 λ_{DR} , which is negatively impacted by high variance in tip rates. Here, variance in absolute per-
 759 tip error in λ_{DR} is 0.012 versus 0.003 for λ_{BAMM} .



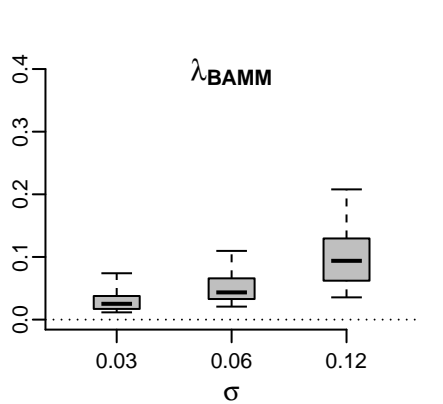
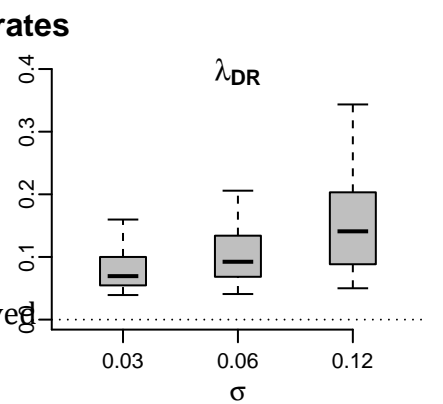
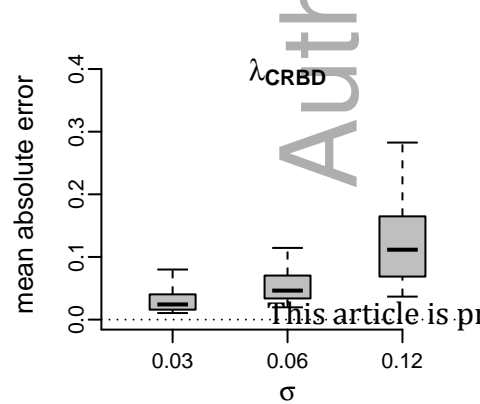
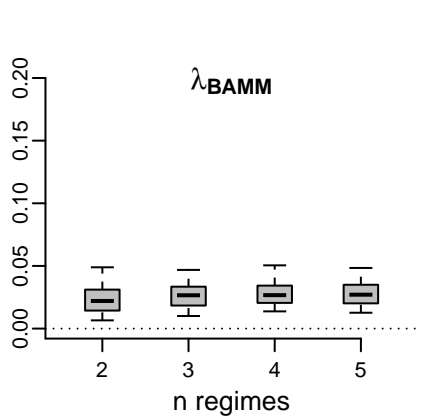
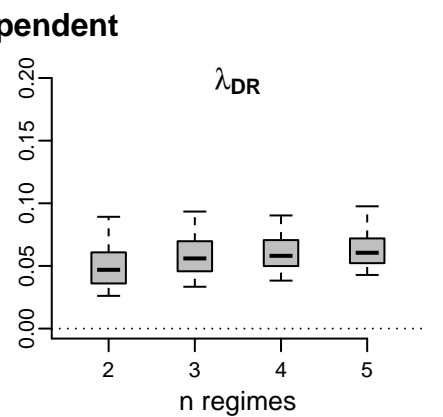
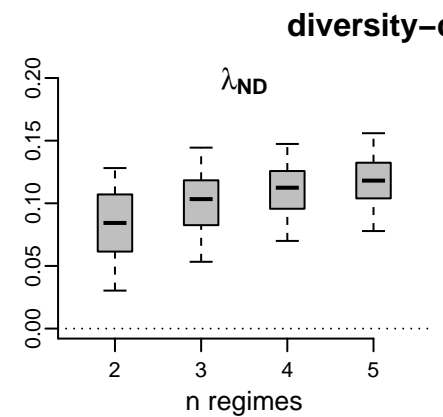
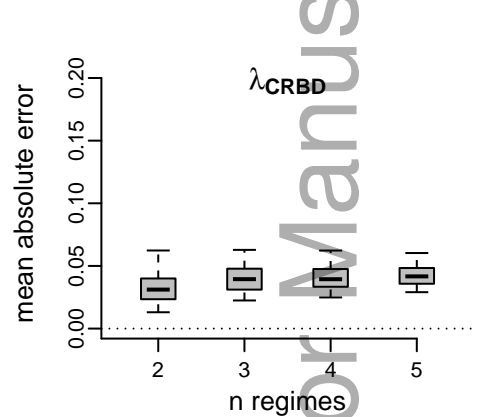
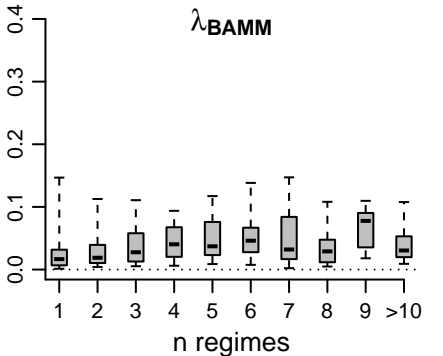
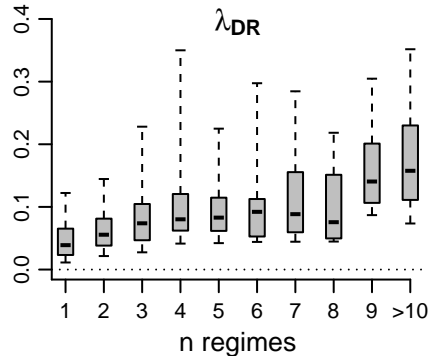
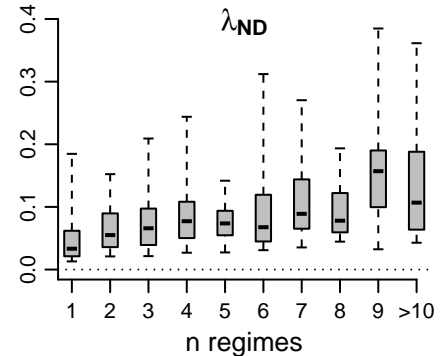
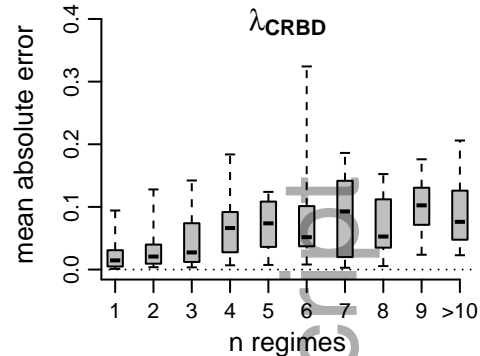
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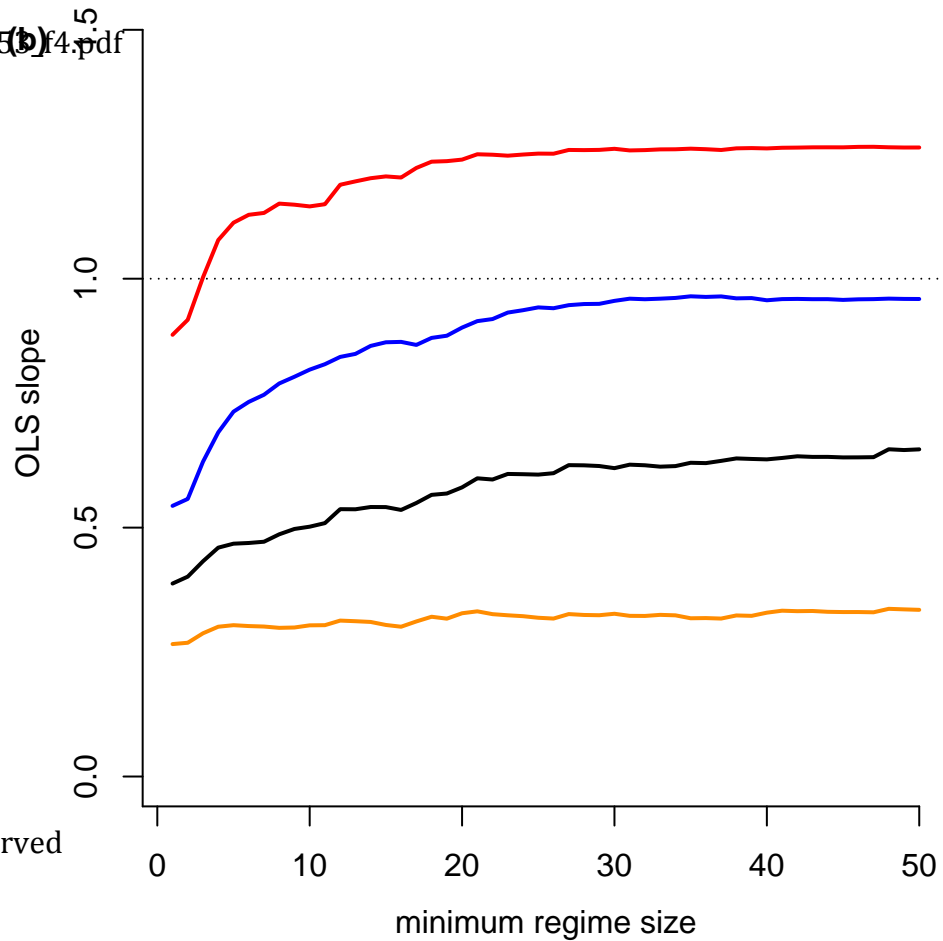
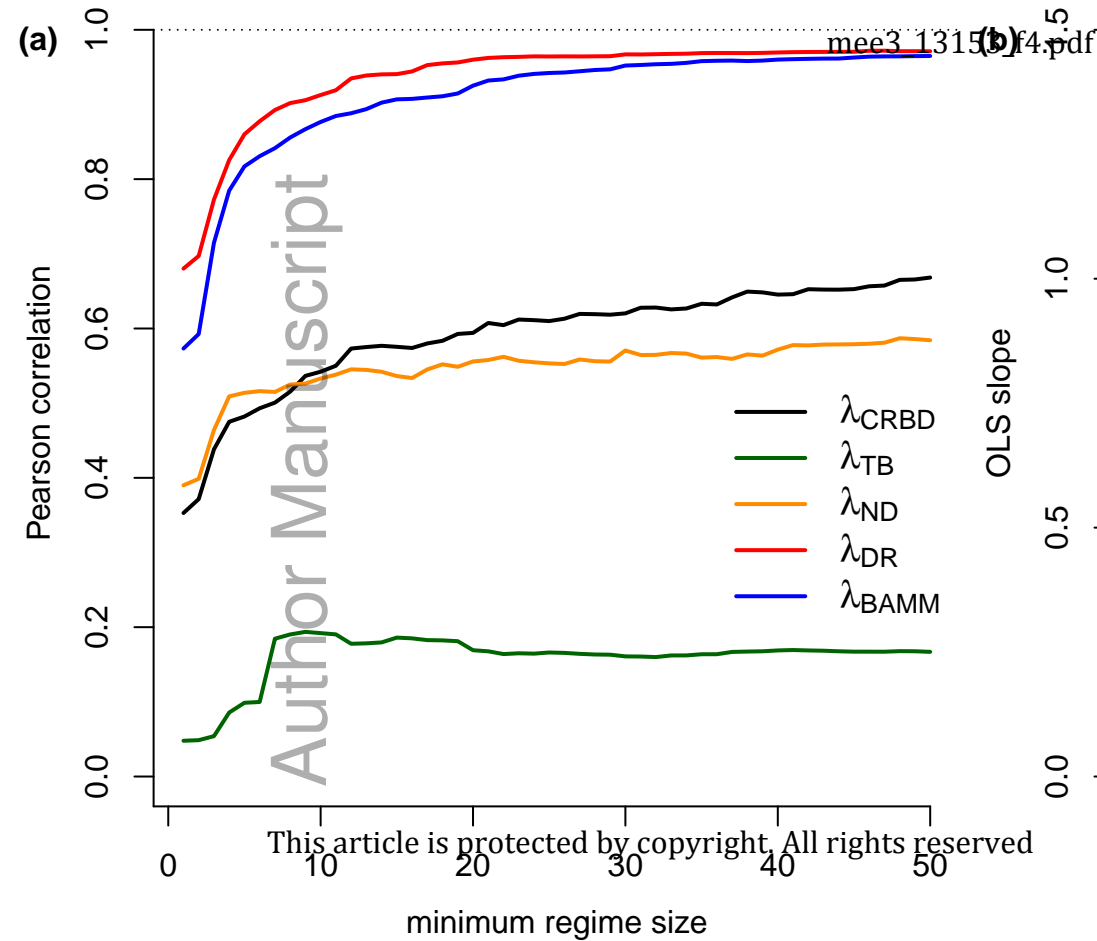


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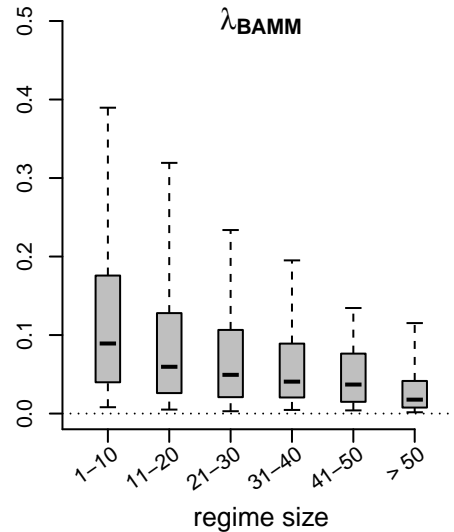
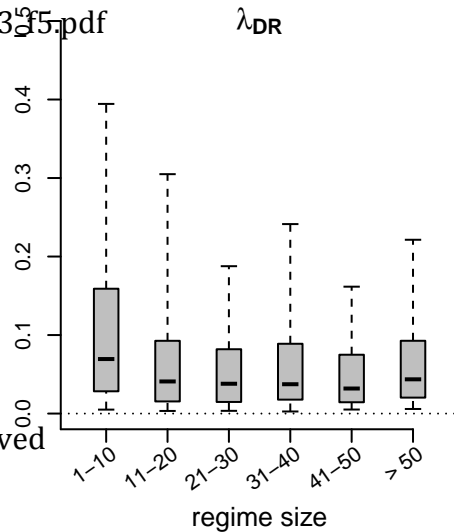
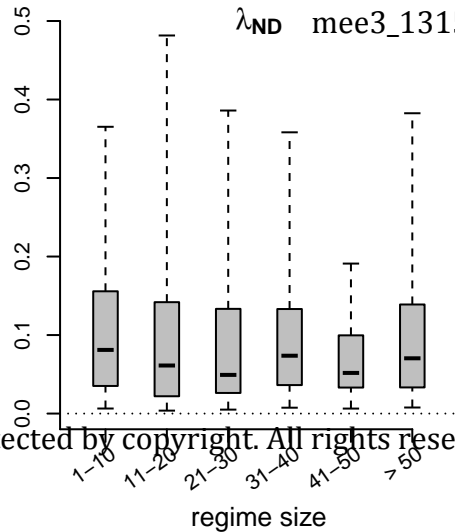
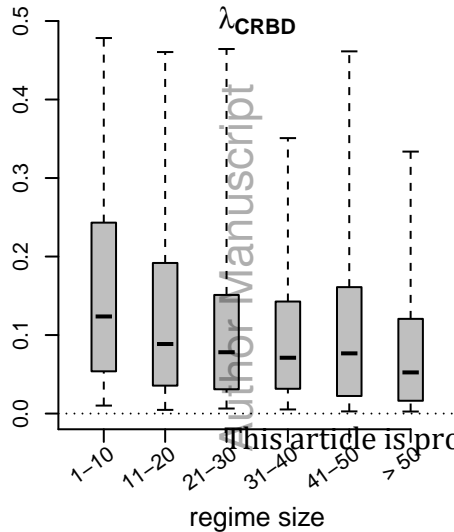
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