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**Title:** Local climate determines vulnerability to camouflage mismatch in snowshoe hares

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**Biosketch:** Marketa Zimova is a postdoctoral researcher interested in the effects of global environmental change on wild populations. The majority of her work aims to understand species adaptive potential to climate change and to yield useful information that can guide conservation and management decisions necessary for maintaining biodiversity. The research team includes scientists from the United States and Canada who study climate change and the ecology and evolutionary biology of wild organisms.

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**Title:** Local climate determines vulnerability to camouflage mismatch in snowshoe hares

**Running title:** Local climate determines mismatch

**ABSTRACT:**

**Aim:** Phenological mismatches, when life-events become mistimed with optimal environmental conditions, have become increasingly common under climate change. Population-level susceptibility to mismatches depends on how phenology and phenotypic plasticity vary across a species distributional range. Here, we quantify the environmental drivers of color molt phenology, phenotypic plasticity, and the extent of phenological mismatch in seasonal camouflage to assess vulnerability to mismatch in a common North American mammal.

**Location:** North America.

**Time period:** 2010-2017.

**Major taxa studied:** Snowshoe hare (*Lepus americanus*).

**Methods:** We used >5,500 by-catch photographs of snowshoe hares from 448 remote camera trap sites at three independent study areas. To quantify molt phenology and phenotypic plasticity, we used multinomial logistic regression models that incorporated geospatial and high-resolution climate data. We estimated occurrence of camouflage mismatch between hares' coat color and the presence and absence of snow over seven years of monitoring.

**Results:** Spatial and temporal variation in molt phenology depended on local climate conditions more so than on latitude. First, hares in colder, snowier areas molted earlier in the fall and later in the spring. Next, hares exhibited phenotypic plasticity in molt phenology in response to annual variation in temperature and snow duration, especially in the spring. Finally, the occurrence of camouflage mismatch varied in space and time; white hares on dark, snowless background

31 occurred primarily during low-snow years in regions characterized by shallow, short-lasting  
32 snowpack.

33 **Main Conclusions:** Long-term climate and annual variation in snow and temperature determine  
34 coat color molt phenology in snowshoe hares. In most areas, climate change leads to shorter  
35 snow seasons, but the occurrence of camouflage mismatch varies across the species range. Our  
36 results underscore the population-specific susceptibility to climate change-induced stressors and  
37 the necessity to understand this variation to prioritize the populations most vulnerable under  
38 global environmental change.

39 **Keywords:** adaptation, camouflage mismatch, climate change, latitudinal gradient, phenological  
40 mismatch, phenotypic plasticity, range edge, snow, snowshoe hares.

41  
42 **Main text:**

## 43 INTRODUCTION

44 As a result of anthropogenic climate change, plant and animal populations are increasingly  
45 confronting environmental conditions different from the ones to which they are adapted.  
46 Organisms occupying seasonal environments have evolved mechanisms for the timing of their  
47 life cycles (i.e., phenology) to match with optimal environmental conditions and resources at  
48 their location (Bradshaw & Holzapfel, 2007; Williams et al., 2015). When phenology and  
49 favorable environmental conditions become asynchronized, organisms can suffer negative fitness  
50 costs (Both et al., 2006; Post & Forchhammer, 2008; Lane et al., 2012; Zimova et al., 2016;  
51 Senner et al., 2017). Such phenological mismatches are becoming increasingly common under  
52 climate change (Parmesan & Yohe, 2003; Thackeray et al., 2010; Cohen et al., 2018), which in  
53 the absence of adaptive responses could lead to population declines and local extinctions (Visser  
54 & Gienapp, 2019). Thus, there is a pressing need to understand the degree to which climate  
55 change leads to phenological mismatches and the capacity of wild populations to withstand  
56 attendant fitness costs.

57       Understanding the variation in phenology and its environmental drivers is fundamental  
58 for assessing current and future species' vulnerability to phenological mismatches. For the  
59 majority of traits in plants and animals in temperate regions, photoperiod serves as the principal  
60 cue for phenology, with temperature and other environmental factors exerting lesser influence

61 (Hofman, 2004; Bradshaw & Holzapfel, 2007). Because photoperiod, latitude, and climate  
62 covary across most species' ranges, variation in phenology is often distributed along latitudinal  
63 gradients. Northern populations typically experience harsher and colder climates that correspond  
64 with later initiation of spring and earlier initiation of fall events compared to southern  
65 populations (Bradshaw & Holzapfel, 2007; Hut et al., 2013). However, latitude may not be a  
66 reliable predictor of phenology for two reasons. First, the covariance between latitude and  
67 climate is imperfect when other geographic factors such as elevation also affect climatic  
68 conditions (Chaine, 2010; Visser et al., 2010). Secondly, species show year-to-year in situ  
69 variation in phenological traits in response to annual variation in climate. This temporal  
70 phenological variation is referred to as "population-level" phenotypic plasticity, and is different  
71 from between-individual level phenotypic plasticity (Phillimore et al., 2010; Gienapp &  
72 Brommer, 2014). Overall, both spatial and temporal variation in phenology may have  
73 consequences for susceptibility to phenological mismatch on both the local population and  
74 broader species levels.

75         Seasonal coat color molt, a key phenological trait, has received increased attention as a  
76 trait shaped by climate. Across the Northern Hemisphere, 21 species of mammals and birds  
77 change coat or plumage color from brown in the summer to white in the winter to match snow-  
78 covered landscapes (Mills et al., 2018; Zimova et al., 2018). As with other phenological traits,  
79 photoperiod serves as the principal cue for molt phenology, with some evidence that year-to-year  
80 variation in winter weather modulates the progression of molt (Hofman, 2004; Zimova et al.,  
81 2018). However, decreasing duration of snow cover due to climate change (Choi et al., 2010;  
82 Vaughan et al., 2013; Kunkel et al., 2016) may result in phenological mismatch, whereby winter  
83 white animals become color mismatched against dark, snowless backgrounds (Mills et al., 2013).  
84 Field studies indicate that mismatch in seasonal coat color and snow presence or absence leads to  
85 high fitness costs due to increased predator-induced mortality (Zimova et al., 2016; Atmeh et al.,  
86 2018) and may have already contributed to range contractions for several species including  
87 *Lagopus* and *Lepus* spp. (Imperio et al., 2013; Diefenbach et al., 2016; Sultaire et al., 2016;  
88 Pedersen et al., 2017).

89         The snowshoe hare (*Lepus americanus*), a key prey species of the boreal forest of North  
90 America (Krebs et al., 2001), exhibits seasonal color molts in the majority of its range  
91 (Nagorsen, 1983; Gigliotti et al., 2017; Mills et al., 2018). Because of their broad distribution

92 (Fig. 1a), hares inhabit a large range of environmental conditions, making them an ideal species  
93 for investigating variation in molt phenology and camouflage mismatch. In the only two areas  
94 where molts have been recently investigated in relation to climate change (i.e., Montana and  
95 Wisconsin, USA), phenotypic plasticity is not sufficient to prevent camouflage mismatch (e.g.,  
96 Mills et al., 2013; Wilson et al., 2018). For example, hares in Montana experience about a week  
97 of mismatch annually whereby hares are in the wrong coat color in relation to their background  
98 (i.e., white hares on snowless background or brown hares on snow; Mills et al., 2013; Zimova et  
99 al., 2014, 2016). Because hares rely heavily on their camouflage for survival, mismatch has  
100 strong negative fitness costs (i.e., 7-12% reduced weekly survival) that in the absence of  
101 evolutionary shifts and under projected snow declines would be sufficient to cause steep  
102 population declines and local extinction (Zimova et al. 2016, Wilson et al. 2018). To date, no  
103 study has evaluated the phenological drivers and camouflage mismatch across the heterogeneous  
104 snowshoe hare range in a unified analytical framework.

105 To understand snowshoe hare vulnerability to camouflage mismatch, we quantified the  
106 spatial and temporal variation in color molt phenology, phenotypic plasticity, and the occurrence  
107 of camouflage mismatch across three disjunct, climatically and geographically distinct  
108 populations. First, we hypothesized that the spatial variation in molt phenology would be  
109 determined by latitude and local climate. We tested two predictions: i) populations in more  
110 northern sites molt earlier in the fall and later in the spring, ii) populations in colder and snowier  
111 sites molt earlier in the fall and later in the spring. Second, we hypothesized that hares exhibit  
112 temperature- and snow-mediated phenotypic plasticity in molt phenology and we predicted that  
113 molts occur earlier in fall and later in spring during colder and or snowier years. Third, we  
114 quantified the occurrence of camouflage mismatch for each population in spring and fall and  
115 assessed which snowshoe hare populations may be the most vulnerable to camouflage mismatch  
116 now and under future climate change.

## 117 MATERIALS AND METHODS

### 118 **Study Areas**

119 Our analysis integrated three regional monitoring studies in North America: the Canadian  
120 Rockies, the San Juan Mountains in Colorado, and northern New England (Fig. 1a). Together,  
121 these areas encompass wide latitudinal, altitudinal and habitat variation across the distributional

122 range of snowshoe hares (Table 1). Our northernmost study area in the Canadian Rockies  
123 included three national parks (Banff, Yoho, and Kootenay National Parks) and was characterized  
124 by rugged, forested mountainous regions with a long snow season. This area is located within a  
125 homogeneous boreal habitat at the core of the snowshoe hare distribution range (Cheng et al.,  
126 2014). The southernmost study area was located in the San Juan Mountains of southwest  
127 Colorado; an isolated patch of high-elevation southern boreal forest near hares' southern range  
128 boundary. The northern New England study area, also near the southern edge of hare  
129 distribution, encompassed portions of the Green Mountains in Vermont and the White Mountains  
130 in New Hampshire. This area stretched across a transition zone between the northern hardwood  
131 and boreal forests and had, on average, the mildest climate of the three areas. Major hare  
132 predators at all areas included coyotes (*Canis latrans*), red fox (*Vulpes vulpes*), Canada lynx  
133 (*Lynx canadensis*), bobcats (*L. rufus*), martens (*Martes spp.*), weasels (*Mustela spp.*), northern  
134 goshawks (*Accipiter gentilis*), and great-horned owls (*Bubo virginianus*). The study area in  
135 Canada also had wolves (*C. lupus*) and grizzly bears (*Ursus arctos*).

### 136 **Camera trapping design**

137 We obtained our snowshoe hare phenology data as by-catch photos from three long-term  
138 independent studies of forest carnivores. A combination of motion-triggered camera models was  
139 used across the sites, but all produced comparable high-quality images during the day and night.  
140 The camera spacing differed between regions, but at any given time the minimum distance  
141 between cameras was >1 km; grid sizes varied by study area (10 × 10 km in Canada, 4 × 4 km in  
142 Colorado, and 2 × 2 km in New England). Each camera site was stationary and remained within  
143 a grid cell for a minimum duration of one year. We combined all observations for the cameras  
144 that were moved between years within <1 km distance and assigned them an average latitude,  
145 longitude and elevation. The 1-km threshold was chosen because it far exceeds average  
146 individual hare seasonal movement and home range size at all three areas (Hodges, 1999; Mills  
147 et al., 2005; Ivan, 2011).

### 148 **Coat color monitoring**

149 Coat color was visually estimated from the by-catch hare photographs by one observer,  
150 following a standardized protocol for color classification and image quality filtering. Briefly,  
151 hares were classified as 1) white when >90% of the body (excluding belly and feet) was white, 2)



152 brown when <10% of the body was white, and 3) molting for all other instances. As per  
153 previously developed classification methods (Zimova et al., 2014), we excluded the feet and  
154 belly color as these stay white all year in all three areas (Grange, 1932). Because hares cannot be  
155 individually identified from photos, we considered photos spaced by >24 hours as independent  
156 events, unless different individuals were distinguishable (e.g., two hares in one photo, a hare in a  
157 different molt stage recorded within 24 hrs).

158 In Canada, 121 cameras operated continuously from September 2011 to November 2016,  
159 yielding 1,888 independent coat color observations. In Colorado, 206 cameras operated from  
160 September 2010 to June 2011 and continuously from January 2014 to August 2017, yielding  
161 2,027 independent observations. In New England, 121 cameras operated continuously from  
162 January 2014 to January 2018, yielding 1,608 independent observations.

### 163 **Statistical analysis**

164 We used R (version 3.4.3) (R Core Team, 2016) for all statistical analyses.

#### 165 **Climate variables**

166 To characterize annual and long-term climate conditions at each camera site, we prepared a set of  
167 temperature and snow cover variables relevant to molt phenology (Mills et al., 2018; Zimova et  
168 al., 2018). Annual minimum (tmin) and maximum (tmax) temperature during spring and fall  
169 were calculated for each year of monitoring (2010-2017) at each camera site using daily gridded  
170 1 x 1 km resolution data (Daymet; Thornton et al., 2018). We used the same dataset to calculate  
171 mean seasonal tmin and tmax during the 30-year period 1980-2009 to describe long-term climate  
172 at each camera site. The seasons were defined as spring (March 1- May 31) and fall (September  
173 1- November 30), because at all areas the majority of molting occurs during those months.

174 We used modeled snow water equivalent (SWE) to quantify the duration of the  
175 continuous snow season and the total number of snow days for the spring and fall. For the years  
176 of monitoring (2010-2017), we used daily gridded 1 x 1 km resolution data by Snow Data  
177 Assimilation (hereafter SNODAS: Barrett, 2003). Because SNODAS is unavailable prior to  
178 2000, we used daily gridded 6 x 6 km resolution data (Livneh et al., 2015) to describe the mean  
179 long-term snow conditions during the 30-year period (1980-2009). The duration of spring snow  
180 season each year was calculated as the number of days between January 1st and snowmelt date,  
181 i.e. the first day when snow is absent (SWE= 0) at a camera site for a minimum of 7 days. Fall

182 snow season was estimated as the number of days following the summer (i.e., the longest  
183 snowless period at each camera site each year), between snow onset date (i.e., first day when  
184 SWE >0 for a minimum of 7 days) and December 31st. We used the 7-day buffers to discard  
185 spurious early spring snowmelts followed by further extended snow season and to account for  
186 spurious snow flurries in the fall. The total number of snow days was calculated as the sum of  
187 days with SWE >0 from January 1st to August 31st for the spring and from September 1st to  
188 December 31st for the fall.

### 189 **Molt phenology**

190 We used a hierarchical multinomial logistic regression analysis within a Bayesian framework to  
191 describe molt phenology and its phenotypic plasticity. For all models, we estimated the  
192 probability of a hare being in color category  $i$  at a camera site  $j$  on a Julian day  $d$  as:

$$193 \quad \Pr(y = i) = \frac{e^{\alpha_i + \beta 1_i * d + s_{i,j}}}{1 + \sum_{k=1}^{i-1} e^{\alpha_k + \beta 1_k * d + s_{k,j}}}$$

194 Coat color was treated as a categorical variable, such that a hare on day  $d$  was either brown  
195 ( $p_{\text{brown}}$ ), white ( $p_{\text{white}}$ ), or molting ( $p_{\text{molt}}$ ) and  $\Sigma(p_{1:3, j, d}) = 1$ . Camera site was coded as a random  
196 covariate  $s_{i,j}$  to reflect the hierarchical structure of the dataset and allow for repeated  
197 measurements.  $\alpha_i$  was the intercept and  $\beta 1_i$  was the effect of Julian day on the probability of  
198 being either brown, white, or molting. Fall and spring molts were modeled separately.

### 199 **Spatial variation in molt phenology**

200 First, to quantify average molt phenology of each population (Canada, Colorado, New England),  
201 we combined color observations from all years at that area and ran the model separately for each.  
202 We used the estimated probabilities to derive approximate dates when hares initiated and  
203 completed their molts as ‘initiation’ and ‘completion’ dates at each area. Fall start was specified  
204 as the first Julian day when mean  $p_{\text{brown}} < 0.9$  and end date when mean  $p_{\text{white}} > 0.9$ ; the opposite  
205 condition was used to estimate the spring dates (i.e., start when  $p_{\text{white}} < 0.9$  and end when  
206  $p_{\text{brown}} > 0.9$ ).

207 Next, to test the effect of local environmental covariates on phenology, we combined  
208 color observations from all years and populations in one dataset and constructed a set of  
209 univariate models. Each model included a single fixed effect of an environmental covariate  $\beta 2_i$   
210 on the probability of being brown, white, or molting. The environmental covariates were latitude,

211 elevation, and the 30-year long-term temperature and snow conditions (i.e., tmin and tmax,  
212 duration of snow season) in spring and fall during each season at each camera site. We used  
213 univariate models to avoid problems associated with multicollinearity as most environmental  
214 covariates were highly correlated (Pearson correlation coefficients  $> |0.60|$ ; Table S1). To  
215 facilitate comparisons between models, all covariates were standardized to a mean of 0 and SD  
216 of 1.

217 The resulting  $\beta$  coefficients represented the increase in the probability of being in a  
218 certain color category on the multinomial-logit scale for every one-unit (SD) change in the  
219 covariate. We considered a covariate to have a significant effect on molt phenology when the  
220 resulting  $\beta$  coefficients' 95% credible intervals (CRIs) did not overlap zero. Because we were not  
221 interested in quantification of the effect size per se, but rather on the direction and the relative  
222 sizes of the different covariates, we did not convert the coefficients to normal scale. We  
223 primarily focused on the covariate effects on probability of the season's final color (i.e.,  $\beta_{2_{\text{white}}}$  in  
224 the fall,  $\beta_{2_{\text{brown}}}$  in the spring) not the initial color or molting color category, in part to simplify  
225 the reporting of results. Furthermore, because photoperiod is known to strongly control the  
226 hormonal cascade that triggers the molt, we expected the effects of climate to be more apparent  
227 in the final rather than the initial stage of the molt (i.e., follicle stimulation and hair growth  
228 initiation versus the appearance of the newly grown hairs and shedding of the old hairs; Zimova  
229 et al., 2018). Finally, in most cases, the significance or absolute effect size of initial and final  
230 color probabilities were similar (all  $\beta$  coefficients shown in Tables S2 and S3).

### 231 Temporal variation in molt phenology

232 Next, to test the effect of annual variation in temperature and snow season duration, we  
233 constructed an alternative set of univariate models. Each model included a single fixed effect of  
234 climate covariate  $\beta_{2_i}$  on the probability of being in a certain color category to avoid  
235 multicollinearity issues (Table S1). The covariates included mean annual tmin, tmax, and  
236 duration of snow season in spring and fall at each camera site. The resulting  $\beta_{2_i}$  coefficients were  
237 the slopes of reaction norms of the climate covariates on probabilities of being brown ( $\beta_{2_{\text{brown}}}$ ) or  
238 white ( $\beta_{2_{\text{white}}}$ ).

239 For all models, we obtained posterior distributions of all parameters along with their 95%  
240 CRIs using Markov chain Monte Carlo (MCMC) implemented in JAGS (version 4.0.1), which

241 we called using the R2jags package (Su & Yajima, 2012). Model convergence was assessed  
242 using the Gelman–Rubin statistic, where values  $<1.1$  indicated convergence (Gelman & Rubin,  
243 1992). We generated three chains of 300,000 iterations after a burn-in of 150,000 iterations and  
244 thinned by three. Parameters  $\alpha_i$ ,  $\beta_{1_i}$ , and  $\beta_{2_i}$  received a vague prior of  $N(0, 0.001)$ , and the  
245 standard deviation of random effect  $s_{i,j}$  received uniform prior of  $U(0, 100)$ .

#### 246 Camouflage mismatch

247 Camouflage mismatch was calculated based on the daily presence or absence of snow and the  
248 modeled coat color at each camera site. Snow was present at a camera site when  $SWE > 0$ , and  
249 absent when  $SWE = 0$ , based on daily gridded 1 x 1 km resolution data (SNODAS, for validation  
250 of dataset see Siren et al., 2018). Next, we defined white hares as when mean  $p_{white} \geq 60\%$  and  
251 brown hares as  $p_{brown} > 60\%$  as these thresholds included mostly white or brown hares,  
252 respectively, when compared to observations. The camera days with brown and white hares were  
253 calculated using color probabilities from the models that included the best annual climate  
254 predictor (effect sizes in Table 3) in order to account for inter-annual variation in phenology.

255 To quantify the annual frequency of camouflage mismatch within each population, we  
256 calculated the number of days at all camera sites (camera days) when the color of hares would  
257 either match or mismatch against the background color. ‘White mismatch’ occurred on days  
258 when hares were white and snow was absent at the site. ‘Brown mismatch’ occurred on days  
259 when hares were brown and snow was present (Mills et al., 2013). ‘Match’ occurred on days  
260 when hares were white (brown) and snow was present (absent). The proportion of white  
261 mismatch occurrence, for example, was calculated as the count of all camera days when hares  
262 would experience white mismatch, divided by the total number of camera days (i.e., number of  
263 camera sites at each area multiplied by the total number of days in a season). We calculated the  
264 mismatch occurrence for two four-month periods when mismatch might occur at all three study  
265 areas; February 1- May 31 (spring) and September 1- December 31 (fall). All proportions were  
266 multiplied by 100 for interpretation in %.

267 RESULTS

268 Spatial variation in molt phenology

269 Snowshoe hare molt phenology varied across the three study areas (Fig. 1). In the fall,  
270 populations in Colorado and Canada initiated fall molts in early October, with some evidence  
271 that hares in Canada molted faster (32 days total) than hares in Colorado (42 days; Fig. 1b). The  
272 hare population in New England initiated fall molts almost three weeks later and took the longest  
273 to complete the molts (46 days). In the spring, hares in New England initiated the white-to-  
274 brown molts first, in late March, followed almost two weeks later by populations in Canada and  
275 four weeks later by the southernmost population in Colorado. The Colorado population took the  
276 shortest to complete the transitions (44 days vs. 48 and 51 days in Canada and New England,  
277 respectively) and became brown only 2 and 3 weeks later than the populations in Canada and  
278 New England, respectively (Fig. 1b).

279 Variation in molt phenology between populations did not follow the north-south  
280 latitudinal gradient as predicted. Among populations, latitude had a significant effect on the  
281 spring molt phenology, but the effect was negative; hares at higher (i.e., northern) latitudes  
282 became brown earlier than hares in lower latitudes (Table 2). In the fall, latitude had no effect on  
283 molt phenology (Table 2). Local climate and elevation were strong predictors of molt phenology  
284 and always in the predicted direction in both seasons; earlier fall and later spring molts were  
285 associated with sites that are generally colder, snowier and located in higher elevation. Elevation  
286 and long-term minimum temperature had the strongest effect on molt phenology in both seasons  
287 (Table 2).

288 Temporal variation in molt phenology

289 We found evidence of population-level temperature- and snow-mediated plasticity in molt  
290 phenology. All annual temperature and snow covariates affected molt phenology in the predicted  
291 direction; molts occurred later in the spring and potentially earlier in the fall during colder and or  
292 snowier years (Table 3). In the spring, this annual variation in temperature and snow duration  
293 resulted in two- to three-week differences in mean population initiation and completion dates  
294 between some years in Canada and New England (Fig. 2). In contrast, we found no significant  
295 differences between spring initiation and completion dates in the Colorado population.

296 Furthermore, we did not detect any differences in the fall molt phenology start or end dates in  
297 any population (Fig. S3).

### 298 Camouflage mismatch

299 The occurrence of camouflage mismatch varied between study areas and years (Fig. 3, Table 4,  
300 Fig. S1). White mismatch (white hare against snowless background) was relatively infrequent at  
301 all sites and occurred only during low snow years in New England and Colorado. Hares in New  
302 England experienced the highest frequency of white mismatch during both seasons (Table 4),  
303 with the highest proportions of camera days with white mismatch in fall 2015 (15%) and in  
304 spring 2016 (9%); seasons with 37-40 fewer snow days than observed during 1980-2009 (Fig. 3).  
305 Brown mismatch was rare in the New England population, with the exception of fall 2016, with  
306 6% brown mismatch (Fig. S1). Snowshoe hares in Colorado had lower proportions of white  
307 mismatch days than in New England but reached 7% in the fall of 2016 with very short snow  
308 season (Fig. 3). In the springs, brown mismatch was more common than white mismatch in  
309 Colorado (Table 4). In Canada, white mismatch was very rare, exceeding 1% in only 3 out of the  
310 9 seasons of monitoring (Fig. 3), but brown mismatch was frequent (Table 4, Fig. S1).

## 311 DISCUSSION

312 Using by-catch photographs from remote camera traps, we quantified the spatial and temporal  
313 variation in snowshoe hare molt phenology and camouflage mismatch across nearly the full  
314 range of environmental conditions experienced by the species. To our knowledge, this was the  
315 first study to evaluate these processes at such resolution in any seasonally molting species. First,  
316 local climate conditions were the main drivers of spatial and temporal variation in molt  
317 phenology; hares in colder and snowier sites and in higher elevations molt later in the spring and  
318 earlier in the fall. Second, hare populations exhibited temperature- and snow-mediated plasticity  
319 in molt phenology; we found strong evidence of later spring- and suggestive evidence of earlier  
320 fall-molts during colder and or snowier years. Third, the occurrence of camouflage mismatch  
321 varied in space and time, but white mismatch was more common in areas characterized by  
322 shallow, short-lasting snowpack. Finally, because areas with shallow snowpack are expected to  
323 face the steepest declines in snow cover duration (Lute et al., 2015; Ning & Bradley, 2015), we  
324 conclude that hares occupying those southern edge portions of their range are the most  
325 vulnerable to camouflage mismatch under current and future climate change.

326 Local climate drives variation in molt phenology

327 Our analysis based on 448 camera sites and spanning over 15 degrees of latitude showed that  
328 local climate influenced molt phenology in snowshoe hares (Fig. 1). Mean phenology in both  
329 seasons structured strongly along elevational, temperature, and snow cover gradients. Similar  
330 findings were described for other phenological traits across taxa (e.g., migration, reproduction,  
331 and hibernation) where phenology correlated with local variation in climate (e.g., Fielding et al.,  
332 1999; Sheriff et al., 2011; While & Uller, 2014; Duursma et al., 2018). As for seasonal color  
333 molt phenology, local climatic factors have been previously described to determine phenology in  
334 color molting species including snowshoe hares (Grange, 1932; Zimova et al., 2014), mountain  
335 hare (Watson, 1963), stoats (Feder, 1990), and rock ptarmigan (Salomonsen, 1939), although all  
336 were examined over relatively small spatial scales or conclusions were based on opportunistic  
337 observations and low sample sizes. Likewise, for multiple color molting species, the global  
338 distribution of genetically determined winter brown and winter white coat color morphs is driven  
339 by variation in snow cover duration (Mills et al., 2018).

340 We found no evidence that molt phenology variation was distributed along a latitudinal  
341 (i.e., north-south) gradient. While the effects of climate and elevation were strong in both  
342 seasons, the effects of latitude were not detected in the fall and in the opposite direction than  
343 expected in the spring (i.e., during the spring molt, the probability of being brown increased with  
344 latitude, Table 2). We think that the positive effect of latitude in the spring was driven by spatial  
345 variation in climate; the southernmost study area, Colorado, was as cold as the northernmost  
346 area. Although our study lacked the spatial coverage to give a definite conclusion on the role of  
347 latitude on molt phenology, we showed that latitude alone cannot predict molt phenology across  
348 a species range. Future studies that include more replicates across the latitudinal gradient will  
349 help elucidate the interacting effects of latitude and local environmental conditions on seasonal  
350 molt phenology.

351 Phenotypic plasticity in response to temperature and snow

352 We found that annual temperature and snow affect the molt phenology of snowshoe hares.  
353 However, this variation resulted in differences in population molt initiation and completion dates  
354 between years only in the spring (Fig. 2). For example, hares in Canada became brown 21 days  
355 earlier in spring 2016 which was on average 2.9°C warmer ( $t_{min}$ ) with a snow season 22 days

356 shorter than spring 2014 (Fig. 2, Fig. S2). Similar difference was observed between the same two  
357 years in New England, where the 2.4°C higher tmin and 21-day shorter snow season in spring  
358 2016 than in 2014 corresponded with an 11- and 14-day advances in molt initiation and  
359 completion dates, respectively (Fig. 2, Fig. S2). In contrast, we did not detect any differences  
360 between the spring molt phenologies in the Colorado population during all years of monitoring  
361 (Fig. 2). However, annual temperature and snow season length were less variable in Colorado  
362 than in the other two study areas (Fig. S2). For example, the 11-day difference in snow duration  
363 and 1.3°C difference in tmin between the two most extreme springs (2011 and 2015) might not  
364 have been sufficient to observe significant differences between phenology dates (Fig. 2). To  
365 determine whether hares in Colorado have lower phenotypic plasticity or whether this finding  
366 was caused by lower inter-annual variation will require analyses of additional years across a  
367 wider range of climatic conditions.

368 We found some evidence for population-level phenotypic plasticity in the fall molt  
369 phenology (Table 3), but we did not detect significant differences between initiation or  
370 completion dates in any study area (Fig. S3). Lower phenological plasticity in the fall than in the  
371 spring was previously described in least weasels and connected to lower variation in the fall than  
372 in the spring temperatures (Atmeh et al., 2018). In this study, the lower inter-annual variation in  
373 climatic conditions might have similarly contributed to the seasonal variation in plasticity in  
374 some cases (e.g., low variation in tmax in fall in Canada; Fig S2). However, in most cases, the  
375 inter-annual variation in temperature and snow at our study areas was comparable between  
376 spring and fall seasons (Fig. S2). For example, snow season duration differed by up to 44 days  
377 and tmax by 2.3°C between the most extreme falls in New England (2015 and 2016; Fig. S2), yet  
378 we observed no differences between molt phenology dates (Fig. S3). Furthermore, low plasticity  
379 in the fall and similar level of phenotypic plasticity in the spring as observed here were  
380 previously described for individual snowshoe hares in the Rocky Mountains in Montana (i.e., 2-3  
381 weeks in the spring) (Mills et al., 2013; Zimova et al., 2014). However, in this study, the number  
382 of independent snowshoe hare color observations was lower during the falls than in the spring at  
383 most of the sites (Table 1), which may have resulted in reduced statistical power to detect  
384 differences during this season.

385 Finally, we note that the methodology prevented phenology monitoring of the same  
386 individuals - and their fates - over multiple years. Therefore, we were unable to investigate to



387 what extent natural selection against camouflage mismatch might have contributed to the  
388 observed differences in temporal variation molt phenology. Specifically, during the few very  
389 short snow duration years, natural selection might have removed highly mismatched individuals  
390 from the population and contributed to the observed shift in the phenological distribution.  
391 Quantifying the relative importance of natural selection and individual plasticity on population-  
392 level responses in phenology will require intensive field studies with tagged individuals over  
393 multiple years.

#### 394 Spatial variation in camouflage mismatch

395 We observed high variation in the number of snow days each year but, overall, snow cover  
396 duration has decreased over the last 30 years in all three areas (Fig S4, Harpold et al., 2012; Ning  
397 & Bradley, 2015; Kunkel et al., 2016; Fassnacht et al., 2018; Mote et al., 2018). The declines in  
398 snow cover manifested differently in each population, however. In the Canadian population,  
399 white mismatch was very low each year, while the number of snow days ranged from 1 to 28  
400 fewer days than observed on average during 1980-2009 (Fig. 3). This is likely due to the very  
401 deep, long-lasting snowpack in the study area (Table 1, Fig. 4). For example, during the years of  
402 monitoring, spring snow season ended on average on June 9 (Fig. 4), more than two weeks after  
403 hares finished their white-to-brown molts (i.e., molts initiated on April 5 and completed on May  
404 23; Fig. 1b). Yet, annual snow season duration in the Canadian Rockies is predicted to decrease  
405 by about one month by the end of the century (Pomeroy et al., 2015). Those predictions would  
406 advance average snowmelt to about two weeks prior to when hares become fully brown.  
407 Although increase in white camouflage mismatch will likely occur in very short snow years,  
408 based on our results a mean decrease of snow season by one month will unlikely lead to dramatic  
409 increases in the frequency of white hares on snowless background (see Fig. 3).

410 In contrast, hares in Colorado and New England are much more vulnerable to white  
411 camouflage mismatch. For both populations, the proportion of white mismatch increased rapidly  
412 with fewer snow days, especially once the snow days anomaly exceeded 21 days (Fig. 3).  
413 Beyond this three-week threshold, hares began to experience elevated white mismatch (i.e., 15%  
414 in fall and 9% in spring in New England; 7% in fall in Colorado). This suggests that phenotypic  
415 plasticity in molt phenology is insufficient to buffer against the snow declines in those marginal  
416 areas where snow season is already short (Table 1, Fig. 4). To contrast the New England study

417 area with the previous example from Canada, the continuous snow season ended on average on  
418 April 18 (Fig. 4) but hares underwent spring molts from March 23 to May 13 (Fig. 1b).  
419 Therefore, with mean snowmelt occurring before hares are halfway through the molt, early snow  
420 melt years (e.g., 2016 seasonal snow melted on April 1; Fig. S2) result in steep increases in white  
421 mismatch.

422 Based on our results, in the absence of evolutionary shift in phenology, snowshoe hares  
423 in New England and Colorado will suffer drastic increases in white mismatch over the next  
424 century. In the Colorado study area, the frequency of low snow years is expected to increase and  
425 the annual snow season to decrease by 20-50 days (for comparison see mismatch in fall 2016,  
426 Fig. 3) by mid-century (Lute et al., 2015). Similarly, in many parts of New England, snow cover  
427 duration is predicted to shorten from the current five months to three months by 2100 (Ning &  
428 Bradley, 2015). Under such projections, hare in both study areas will by the end of the century  
429 routinely experience the same amount of mismatch as they experienced during the very low  
430 snow years of fall 2016 in Colorado and 2015 in New England.

431 The pattern observed with white mismatch was somewhat mirrored by that of brown  
432 mismatch. In both seasons in New England and during the fall in Colorado, white mismatch was  
433 more frequent than brown mismatch as snow cover was more likely to be absent when hares  
434 were white than vice versa (Fig. 4). In contrast, during springs in Canada and Colorado, hares  
435 experienced long periods when they had already molted to summer brown pelage, but snow was  
436 still present on the ground (Fig. 4). Furthermore, the onset of snow during fall often occurred  
437 prior to hares completing their brown-to-white molts in Canada. Overall, the relatively high  
438 occurrence of brown camouflage mismatch, despite the recorded snow declines, was unexpected.  
439 Eventually, however, as climate change will continue to lead to shorter snow duration across  
440 most of the snowshoe hare range (Danco et al., 2016; Easterling et al., 2017; Fyfe et al., 2017),  
441 the occurrence of brown mismatch will decrease.

442 The relative fitness costs of white versus brown mismatch are unknown, but we suspect  
443 that white mismatch has a higher survival cost than brown mismatch based on our experience in  
444 the field while locating radio-collared hares. First, brown animals and objects (e.g., branches,  
445 tree trunks, brown animals) are relatively common year-round, but white animals and objects are  
446 rare outside of winter. Perhaps due in part to this frequency difference in the two mismatch  
447 types, a white hare against a snowless background appears far more conspicuous than a brown

448 hare resting on snow. Previous quantifications of survival costs were carried out for ‘absolute  
449 mismatch’, that is both white and brown mismatch combined (Zimova et al., 2016; Wilson et al.,  
450 2018). Nonetheless, as documented here and elsewhere (Zimova et al., 2016; Wilson et al.,  
451 2018), white mismatch is already high in some populations and will increase under climate  
452 change (Mills et al., 2013). Therefore, fitness costs of white and brown mismatch should be  
453 quantified to inform conservation efforts, notably in situ management actions that foster  
454 evolutionary rescue, or genetic rescue by assisted gene flow of individuals with preadapted molt  
455 phenologies or winter coat color (Mills et al. 2018).

456 Understanding the spatial and temporal variation in phenological traits is critical for  
457 understanding the impact of climate change and species’ adaptive potential to environmental  
458 stressors. Here, we showed that snowshoe hare molt phenology is determined by local climate,  
459 but populations vary in their susceptibility to camouflage mismatch. Snowshoe hares responded  
460 to annual variation in temperature and snow via some adjustments in molt phenology, but the  
461 buffering effects of phenotypic plasticity were diminished in populations distributed along the  
462 southern edge of their range. In those areas, characterized by mild climate and shallow, short-  
463 lived snowpack, climate change mediated snow declines led to higher phenological mismatch.  
464 Therefore, hares occupying southern, marginal areas will in the absence of rapid evolution  
465 experience steep increases in camouflage mismatch, as those areas are expected to experience the  
466 largest declines in snow cover duration (Ning & Bradley, 2015; Fyfe et al., 2017), consistent  
467 with theoretical expectations of range contraction (Sirén & Morelli, 2019). More generally, our  
468 results underscore that populations vary in their susceptibility to environmental stressors and  
469 management efforts should consider this intra-specific variation to identify populations most  
470 vulnerable under global environmental change (Hampe & Petit, 2005; Nadeau et al., 2017).

471

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642 DATA ACCESSIBILITY STATEMENT

643 The data used in this study are available from the corresponding author upon request.

644

645 SUPPORTING INFORMATION

646 Additional supporting information may be found online in the Supporting Information section at  
647 the end of the article.

648 TABLES

649 **Table 1.** Geospatial and long-term climate details regarding the camera trap networks in the study areas. Number of camera sites and  
650 independent coat color observations are given for each season and area. Other metrics include mean values across all camera sites  
651 within a study area with standard deviation in parentheses. Long-term mean minimum (tmin) and maximum (tmax) temperature and  
652 snow season duration are based on 1980-2009 period.

Study Area	Latitude (degrees)	Elevation (m)	Camera sites (n)	Obs. (n)	Fall			Spring				
					tmin (°C)	tmax (°C)	Snow season (days)	Camera sites (n)	Obs. (n)	tmin (°C)	tmax (°C)	Snow season (days)
Canada	51.38 (0.33)	1850.98 (266.20)	98	967	-5.41 (1.33)	5.22 (1.86)	93.08 (10.93)	91	921	-6.62 (1.68)	5.17 (2.33)	168.69 (18.18)
Colorado	37.63 (0.30)	3200.79 (215.46)	110	322	-4.68 (1.09)	9.70 (1.78)	72.83 (6.91)	183	1,705	-7.24 (1.28)	8.03 (1.71)	155.70 (11.07)
New England	44.54 (0.51)	627.18 (284.59)	65	486	0.25 (0.89)	10.61 (1.63)	52.07 (7.30)	105	1,122	-3.44 (0.93)	8.35 (1.84)	125.01 (12.49)

653

654 **Table 2.** Effect of latitude, elevation and long-term climate covariates on snowshoe hare molt  
 655 phenology. Mean effect sizes and 95% credible interval (CRI) estimates for slopes for univariate  
 656 models including data from all years and populations combined. Betas indicate effects of  
 657 covariates on the probability of the molt's final color category ( $\beta_{2_{\text{brown}}}$  in the spring,  $\beta_{2_{\text{white}}}$  in the  
 658 fall). Snow is the duration of continuous snow season (days), tmax and tmin are the mean  
 659 minimum and maximum temperature (°C) in springs and falls during 1980-2009. Asterisks  
 660 indicate CRIs not overlapping 0. Values reflect standardized data.

Covariate	Fall $\beta_{2_{\text{white}}}$	Spring $\beta_{2_{\text{brown}}}$
Latitude	0.566 (-0.136, 1.293)	0.689* (0.376, 1.012)
Elevation	2.165* (1.450, 3.033)	-1.325* (-1.631, -1.039)
Snow	0.446 (-0.214, 1.123)	-0.809* (-1.143, -0.492)
tmax	-1.855* (-2.479, -1.288)	0.776* (0.440, 1.123)
tmin	-2.370* (-2.894, -1.909)	1.280* (0.998, 1.579)

661  
662

663 **Table 3:** Effect of annual temperature and snow season duration on molt phenology in snowshoe  
 664 hares. Betas are the slopes of reaction norms  $\beta_2$  (=mean effect size of annual climate covariate)  
 665 and their 95% credible intervals (CRI) on the probability of the molt's final color category.  
 666 Asterisks indicate CRIs not overlapping zero. Values reflect standardized data.

Covariate	Fall $\beta_{2\text{white}}$	Spring $\beta_{2\text{brown}}$
Snow annual	1.466* (1.009, 1.929)	-1.627* (-1.969, -1.303)
tmax annual	-2.070* (-2.850, -1.432)	1.587* (1.208, 2.003)
tmin annual	-2.344* (-2.943 -1.845)	1.273* (0.921, 1.655)

667

668 **Table 4.** Modeled mean percent of camera days with white and brown mismatch at each study  
 669 area. Mean percent were calculated based on annual estimates (Canada: 6 falls and 5 springs,  
 670 Colorado and New England: 4 springs and 4 falls). Standard deviations are given in parentheses.

	<b>Canada</b>		<b>Colorado</b>		<b>New England</b>	
	white	brown	white	brown	white	brown
spring	0.56 (0.55)	11.96 (4.31)	1.74 (1.34)	6.39 (0.68)	2.41 (4.36)	0.98 (0.73)
fall	0.55 (0.47)	9.54 (3.60)	2.69 (2.80)	1.11 (0.66)	4.72 (7.00)	2.42 (2.68)

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671 FIGURE LEGENDS

672

673 **Figure 1.** Camera site locations and snowshoe hare molt phenologies and molt dates in the  
674 Canada, Colorado and New England study areas. (a) snowshoe hare range (IUCN) is colored and  
675 shaded by the mean annual number of snow days (Dietz et al. 2015). Colored points represent the  
676 448 remote-camera trap sites. (b) bold lines depict predicted probabilities of being in the final  
677 color category (white in the fall, brown in the spring) over time. The dashed lines show 95%  
678 credible intervals. The horizontal dashed lines at 0.90 intersect with population means to identify  
679 molt completion dates. Population mean molt initiation and completion dates are depicted as a  
680 date range in the bottom right corners, with the completion dates in bold. The predicted  
681 probabilities and dates were estimated for each season and population based on the model  
682 without covariates.

683

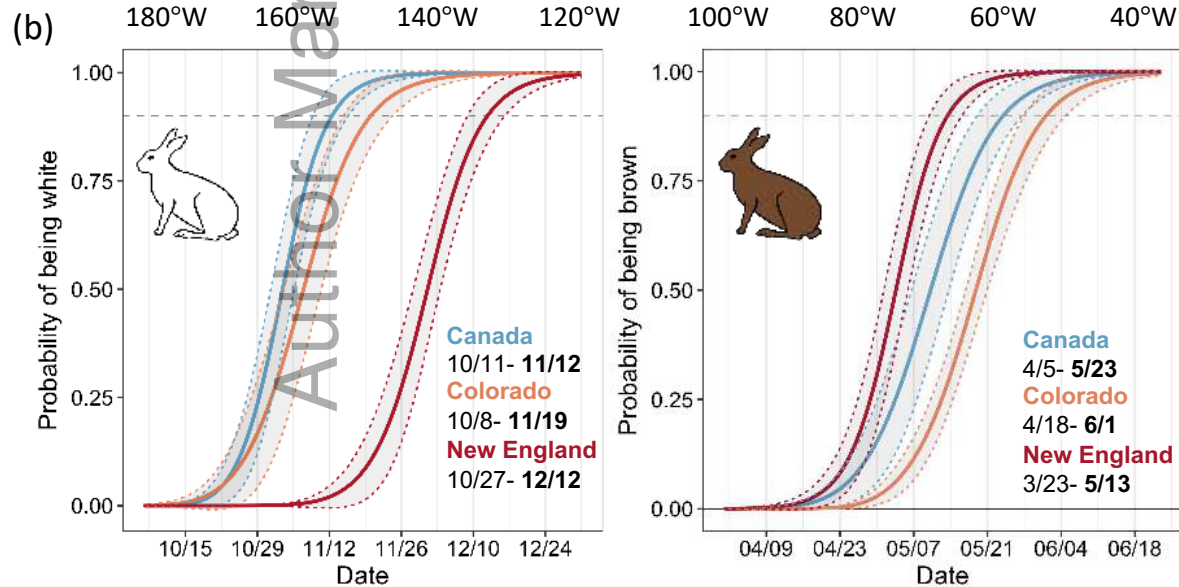
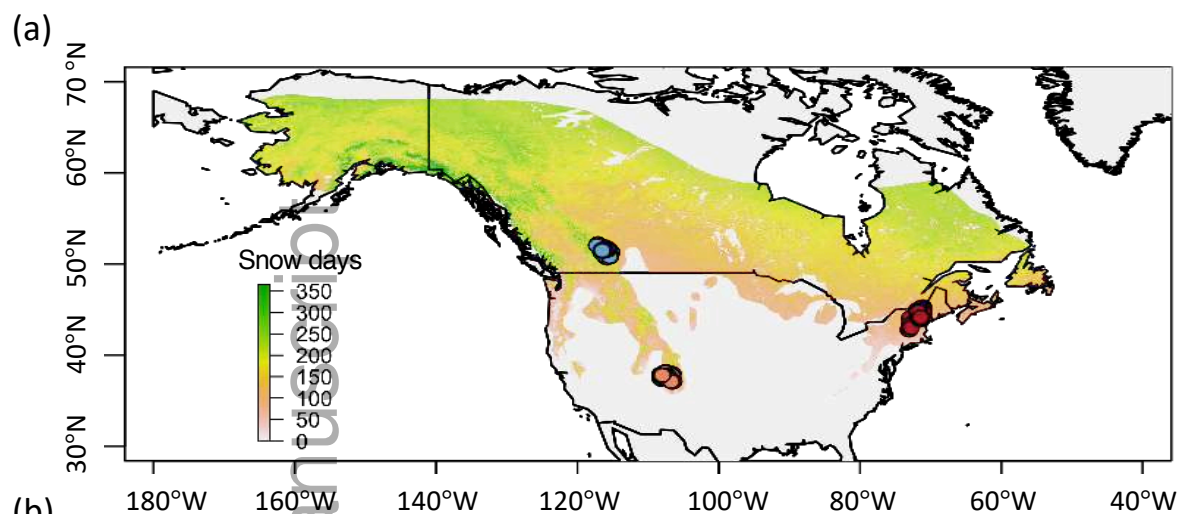
684 **Figure 2.** Estimated annual spring molt initiation ('i') and completion ('c') dates in the studied  
685 hare populations in Canada, Colorado and New England. Points show mean date estimates and  
686 are colored by the annual duration of spring snow season (in days). Horizontal lines show 95%  
687 credible intervals (overlapping CRIs identify same dates).

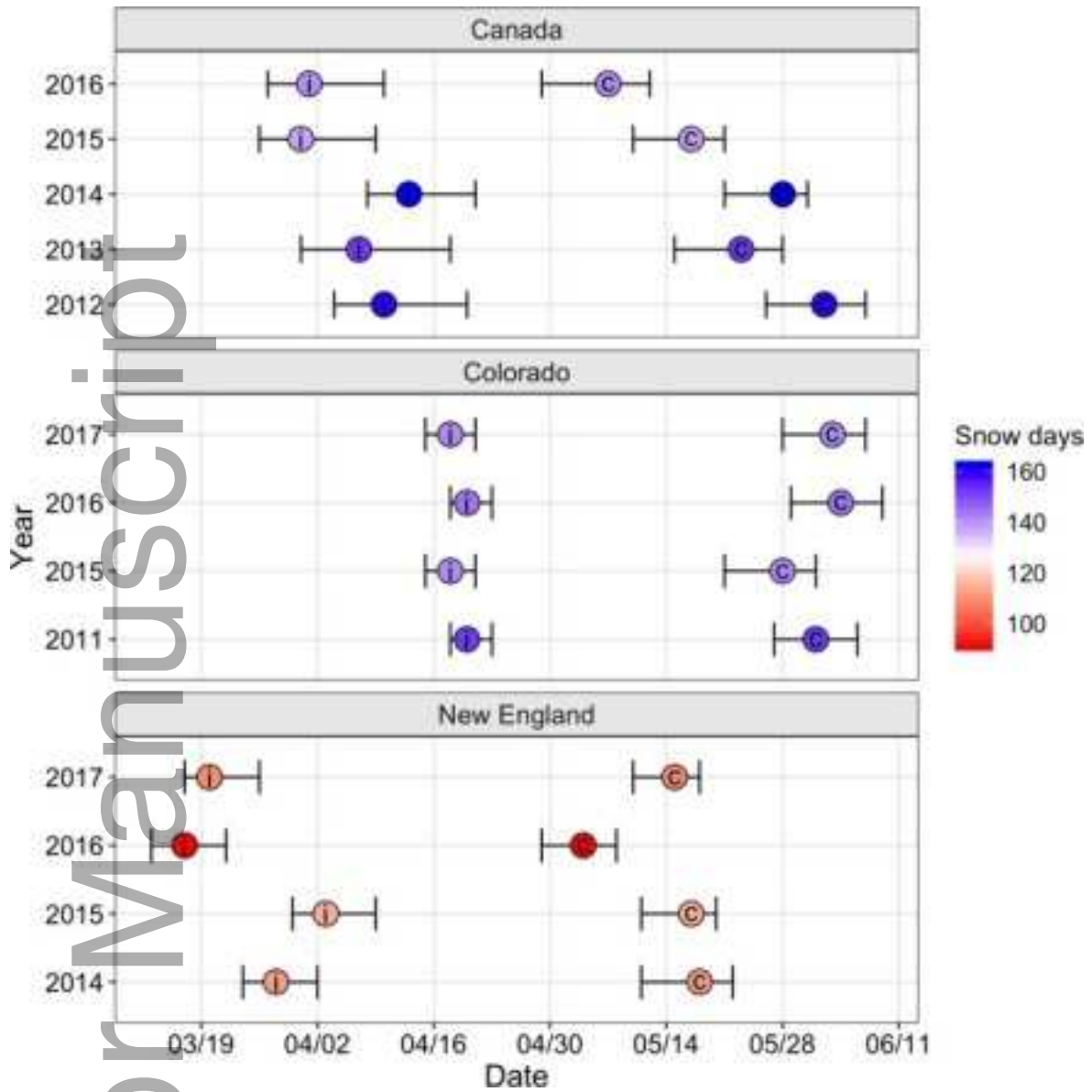
688

689 **Figure 3.** Annual proportions of camera days with white mismatch occurrences plotted against  
690 anomalies in the number of snow days each season in Canada, Colorado and New England.  
691 Study area-specific anomaly in the number of snow days was calculated for each year as the  
692 difference between the mean number of snow days during each season and the mean number of  
693 snow days during 1980-2009 at all camera sites. Photo depicts white mismatched hare.

694

695 **Figure 4.** Mean daily snow water equivalent (SWE; mm) at the remote camera sites for the years  
696 of molt phenology monitoring. Colored circles along the x axes indicate the population mean  
697 molt completion dates for each year, with spring molts on the left and fall molts on the right.  
698 Mean number of snow days for years of monitoring is given in facet titles for spring and fall.





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