Title: Socioeconomic disparities in survival after high-risk neuroblastoma treatment with modern therapy



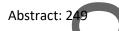
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Figures: 3

Supporting Files: 1

Short Running Title: Socioeconomic disparities in high-risk neuroblastoma

Keywords: Neuroblastoma, healthcare disparities, pediatric oncology, health services research, poverty, insurance



Abbreviation	Full Term
OS	Overall survival
SES	Socioeconomic status
HR	Hazard ratio
ASCT	Autologous stem cell transplant
SEER	Surveillance, Epidemiology, and End Results
ACS	American Community Survey
FPL	Federal poverty level
Cox-PH	Cox proportional hazards
Abstract:	

Background: Modern therapeutic advances in high-risk neuroblastoma have improved overall survival (O5), but it is unclear whether these survival gains have been equitable. This study examined the relationship between socioeconomic status (SES) and OS in children with high-risk neuroblastoma, and whether SES-associated disparities have changed over time.



Procedure: In this population-based cohort study, children <18 years diagnosed with high-risk neuroblastoma (diagnosis at age \geq 12 months with metastatic disease) from 1991-2015 were identified through the National Cancer Institute's Surveillance, Epidemiology, and End Results database. Associations of county-level SES variables and OS were tested with univariate Cox proportional hazards regression. For a sub-cohort diagnosed after 2007, insurance status was examined as an individual-level SES variable. Multivariable regression analyses with treatment era and interaction terms were performed when SES variables reached near-significance (p \leq 0.1) in univariate and bivariate modeling with treatment era.

Results: Among 1,217 children, 2-year OS improved from $53.0\pm3.4\%$ in 1991-1998 to $76.9\pm2.9\%$ in 2011-2015 (p<0.001). In univariate analyses, children in high-poverty counties (hazard ratio [HR]=1.74, 95% confidence interval [CI]=1.17-2.60, p=0.007), and those with Medicaid (HR=1.40, 95% CI=1.05-1.86, p=0.02) experienced an increased hazard of death. No interactions between treatment era and SES variables were statistically significant in multivariable analyses, indicating that differences in the OS between SES groups did not change over time.

Conclusions: Survival disparities among children with high-risk neuroblastoma have not widened over time, suggesting equitable access to and benefit from therapeutic advances. However, children of low SES experience persistently inferior survival. Interventions to narrow this disparity are paramount.

Author

1 Introduction

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Neuroblastoma is the most common solid extracranial tumor in childhood, with over 600 cases
diagnosed per year in the United States.¹ High-risk neuroblastoma is associated with significant risk
of relapse and death. However, advances in treatment for children with high-risk disease have led to
impressive increases in survival over recent decades. Patients who receive the full complement of
standard of care therapy (chemotherapy, radiation, surgery, autologous stem cell transplant (ASCT),
and cytokine/immunotherapy) now experience a two-year overall survival (OS) as high as 86%, a
striking survival gain over two decades.²⁻⁵

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While these therapeutic advances hold incredible potential for improved patient outcomes, they 11 require **complex and** highly-intensive treatment, which may not be equally accessible to all patients. 12 Specifically, modern high-risk neuroblastoma treatment is typically delivered at large referral centers 13 14 that have the capacity to provide ASCT, cytokine/immunotherapy, and advanced supportive care measures. This treatment lasts for around 18-months and includes multiple inpatient admissions. 15 16 The adoption of centralized, intensive, and prolonged treatment as a modern standard of care may exacerbate existing disparities based on geographic distance from referral centers, differential or 17 biased clinical trial enrollment, and/or family ability to adhere to treatment demands; all of which 18

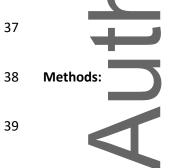
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19 may cumulatively result in unequal survival gains benefitting children of higher socioeconomic

20 status.⁶⁻⁸

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22 A recent analysis of a heterogeneous population of patients with childhood cancer identified through the National Cancer Institute's Surveillance, Epidemiology, and End Results (SEER) database 23 demonstrated that SES significantly mediated racial and ethnic survival disparities for a number of 24 cancer diagnoses including neuroblastoma.⁹ This high-level analysis identified the applicability of SES 25 to survival outcome disparities in pediatric cancer but did not examine disease-specific risk groups to 26 27 differentiate outcome disparities due to stage versus SES-related access disparities. Other specific 28 analyses of neuroblastoma cohorts have identified poverty as an independent risk factor for relapse and death and have found that minority patients have higher prevalence of high-risk disease.^{10,11} 29 30 However, these analyses were restricted to patients enrolled on clinical trials. This poses limitations, given the possibility that families of lower SES and without private insurance may have more 31 difficulty accessing and participating in trials and experimental therapies. The SEER database, as a 32 33 national population-based registry, mitigates this limitation. Our specific aims for this study were to: (1) describe differences in OS among patients with high-risk neuroblastoma by individual- and 34 35 county-level SES characteristics; and (2) investigate whether changes in OS over time in patients with high-risk neuroblastoma differ by these characteristics. 36



41 The analytic cohort was selected using individual- and population-based cancer registry data from 42 the National Cancer Institute's SEER database using SEER*STAT 8.3.4 (Washington, D.C.). Pediatric patients (age <18 years) diagnosed with neuroblastoma from January 1, 1991 through December 31, 43 2015 were selected to allow for a minimum of 2-year follow-up at the time of SEER 8.3.4 release.¹² 44 To approximate characteristics of high-risk (stage M) neuroblastoma according to the International 45 Neuroblastoma Risk Group staging system used in most neuroblastoma clinical trials,¹³ we restricted 46 the analyses to patients with distant metastases and age ≥ 12 months at time of diagnosis. Based on 47 48 data availability, for patients diagnosed from 1991-1992, we used the SEER-9 Registries; for patients diagnosed from 1992-2000, we used the SEER-13 Registries; for patients diagnosed from 2000 49 onwards, we used the SEER-18 Registries.¹⁴ Geographic distribution varies depending on era, with 50 the most recent/expansive registries (SEER-18) covering approximately 28% of the U.S. population 51 and including 18 geographic registries (Native Alaska; Connecticut; Detroit, Michigan; Atlanta, 52 53 Georgia; rural Georgia; greater Georgia excluding the two previous regions; San Francisco-Oakland, 54 California; San Jose-Monterey, California; Los Angeles, California; California excluding the previous 55 metropolitan areas; Hawaii; Iowa; Kentucky; Louisiana; New Mexico; New Jersey; Seattle-Puget Sound, Washington; Utah). Representative of the demographics of the entire U.S. population, this 56 57 broad coverage allows SEER to account for diverse populations throughout the U.S., including 66.5% of Native Hawaiian/Pacific Islanders, 50.4% of Asians, 38.4% of Hispanics, 30.6% of American 58 Indian/Alaska Natives, 24.9% of whites, and 25.6% of Black residents.¹⁵ The study was deemed 59 60 exempt from review by the Dana-Farber Cancer Institute's Institutional Review Board (DFCI protocol 61 18-409).

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63 Measures of Socioeconomic Status

We examined measures of SES at the county level for the entire cohort, and at the individual level 64 65 for a sub-cohort of patients. SEER includes county-level variables from the American Community Survey (ACS) County Attributes data.¹⁶ County-level variables were determined based on the 66 patient's residency county code at diagnosis. SEER utilizes ACS data based on the cancer case/year of 67 diagnosis. County-level measures of SES were chosen a priori based on previous SEER analyses and 68 disparities literature.¹⁷⁻¹⁹ We included county-based poverty (proportion of households living below 69 the federal poverty level (FPL)), educational attainment (proportion of individuals in county >25 70 years of age with less than high school education), unemployment (proportion of individuals in 71 72 county >16 years of age unemployed), language isolation (proportion of households in county with no household member age ≥14 years who speaks English), and urban-rural status (population >1 73 million vs. population of 250,000 to 1 million vs. population <250,000). These variables were defined 74 per SEER and ACS.²⁰ To maximally highlight disparities should they exist, we dichotomized each 75 county-level variable at the 90th percentile cut-point (language isolation, education, unemployment) 76 to define low-SES and high-SES cohorts. High-poverty counties were defined as those with greater 77 than or equal to 20% of households living below 100% FPL.²¹ As a sensitivity analysis, we also 78 analyzed county-based poverty using a cut-point at the 90th percentile. 79

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We examined individual-level insurance data for the sub-cohort of patients for whom it was
available (diagnosed from 2007 onward) in addition to county-level. Insurance status was
dichotomized as any public insurance coverage (i.e. Medicaid) versus non-Medicaid insurance (those

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with Medicaid as a second insurer were coded as Medicaid, per SEER database convention). Given its
rarity in pediatrics, patients without documented insurance (N=14) were excluded. Patients with
Medicaid insurance were a priori considered low-SES. Supplementary Table 1 details exact SEER
variable names and descriptions.

90 The primary outcome was OS, derived from SEER's "Survival Months" attribute, defined as months 91 from date of cancer diagnosis to date of death from any cause, censored at date of last contact. We 92 used OS rather than cancer survival given its lack of ambiguity and frequency of use in oncology 93 clinical trials.²² and the rarity of death from non-cancer causes among children with cancer. We 94 reported 2 year OS based on available follow-up data for our cohort at the time of analysis. Three-95 year OS, given its consistency with recent publications,^{3,23} was included as a sensitivity analysis 96 (Supplementary Table 2) for those patients with sufficient follow-up time.

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Outcome

98 Covariates

99 Covariates included sex, race (white, black, or other), ethnicity (Spanish/Hispanic/Latino vs. non100 Spanish/non-Hispanic/non-Latino) and diagnostic treatment era. We examined race and ethnicity as
101 distinct constructs, given prior reports of differential health outcomes according to race and/or
102 ethnicity.²⁴ Treatment eras were defined based on major advances in the standard-of-care for
103 children with high-risk neuroblastoma: 1991-1998 (early treatment era) vs. 1999-2004 (multimodal
104 treatment including ASCT)⁵ vs. 2005-2010 (improved supportive care) vs. 2011-2015

105 (immunotherapy)² to allow for exploration of the potential interaction between treatment era and

- 106 SES.
- 107
- 108 Statistical analysis
- Descriptive statistics were used to summarize baseline cohort characteristics. Kaplan-Meier curves of
 OS were generated for the overall cohort and stratified by county- and individual-level variables, as
 well as insurance for the post-2007 sub-cohort. OS was compared between groups using the logrank test. Univariate Cox proportional hazards (Cox-PH) regression was used to test the association
 of each SES variable with OS. The proportional hazards assumption was tested by visually examining
 log-log plots and by testing the interaction of selected covariates with time. Our results indicated no
 violation of the proportional hazards assumption.
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We a priori defined a stepwise procedure to evaluate the relationship between SES variables and 117 118 treatment era. First, for SES variables with p≤0.1 in univariate analyses, bivariable Cox-PH regressions were performed to test the effect of each SES variable and (continuous) treatment era 119 120 on OS. Second, the variables with p≤0.1 in these bivariable analyses were then included in separate 121 multivariable models for each SES variable. Lastly, the multivariable Cox-PH regressions tested each 122 SES variable, treatment era, and the interaction of treatment era with the SES variable. If the interaction term was statistically significant, this would provide evidence that there was a significant 123 difference in the change in OS over time between examined SES groups. 124

126	Subjects with missing data (<2% in all measured variables) were excluded from analysis. Notably,
127	insurance status was only available in SEER for those diagnosed after 2007. Thus, analyses of
128	insurance were performed solely in the sub-cohort of patients diagnosed after 2007.
129	
130	We performed additional multivariable regression analyses with county-level poverty, race,
131	ethnicity, and treatment era to explore the relative contributions of these variables to survival. For
132	the sub-cohort of patients with insurance data, we built a second multivariable model including
133	insurance.
134	
135	Statistical analyses were performed using SAS version 9.4 (SAS Institute, Cary, NC). Two-sided p-
136	values ≤ 0.05 were considered statistically significant. The data that support the findings of this study
137	are publicly available through SEER. ²⁵
138	
139	Results:
140	$\overline{\mathbf{O}}$
141	Study population
142	The analytic cohort included 1,217 patients (Figure 1). 550 patients were diagnosed after 2007 and
143	formed the sub-cohort available for analyses of insurance status. Twenty-six percent of patients
144	were non-White race and 20% were of Spanish/Hispanic/Latino ethnicity (Table 1). In the post-2007

- sub-cohort (those for whom insurance data were available), 61% (338/550) were insured by
- 146 Medicaid.
- 147
- 148 Univariate analysis of SES factors and treatment era with OS
- 149 Median follow-up was 6.08 years [range=0-24.92 years]. OS improved by treatment era (p<0.001) 150 with 2-year OS (±standard error) increasing from 53.0±3.4% to 76.9±2.9% between 1991-1998 and 2011-2015 (Figure 2). In univariate analysis of OS (Table 2), increased hazard of death was seen in 151 152 patients in high-poverty counties (≥20% of households below 100% FPL) compared to those in lowpoverty counties (hazard ratio [HR]=1.74, 95% confidence interval [CI]=1.17-2.60, p=0.007). No other 153 county-level SES factors were found to be statistically significant. In the post-2007 sub-cohort, 154 individuals with any Medicaid experienced increased hazard of death compared to those with other 155 insurance (HR=1.40, 95% CI=1.05-1.86, p=0.02). 156
- 157
- 158 Multivariable analysis of change in OS by SES factors over time

In multivariable analysis, we included the SES variables that were near significant in bivariable analysis ($p \le 0.1$), treatment era, and the specific SES*treatment era interaction to assess whether SES-associated hazard ratio for OS changed over time. A separate model was created for each SES variable. As displayed in **Figure 3**, the SES*treatment era interaction is not statistically significant for either Spanish/Hispanic/Latino ethnicity (interaction *p*=0.800) or higher county-based poverty (interaction *p*=0.923), indicating that the SES-associated survival hazard did not change significantly over time.

166

167 Sensitivity analyses

A sensitivity analysis using the 90th percentile as a cut-point to define a high-poverty county, while 168 169 not statistically significant, did not differ greatly from the primary analysis (HR=1.15, 95% CI=0.89-170 1.49, p=0.27). In multivariable modeling, the interaction between treatment era and county-level poverty using this cut-point was similarly not statistically significant (p = 0.45). Additional sensitivity 171 analyses considering 3-year OS for the patients with sufficient follow-up time (N=1156) are included 172 173 in **Supplementary Table 2** and demonstrate similar findings to our primary analyses. 174 175 Multivariable analyses including race and ethnicity in the model 176 In a multivariable regression model for the entire cohort considering county-based poverty, race, ethnicity, and treatment era, there was an increased hazard of death associated with higher county-177 based poverty (HR=2.08, p<0.001, Table 3) and a lower hazard of death associated with later 178 179 treatment era (HR=0.78, p<0.001). Race and ethnicity were not significantly associated with survival. 180 In the post-2007 sub-cohort, higher county-based poverty (HR=2.38, p=0.001) and any Medicaid 181 insurance (HR=1.38, p=0.04) were statistically significantly associated with increased hazard of 182 death. Race, ethnicity, and treatment era were not statistically significant. Treatment era was 183 184 included in this second model for consistency although this sub-cohort notably only includes 185 patients diagnosed after 2007, limiting power of this analysis.

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Discussion:

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189	In a representative population of U.S. children with high-risk neuroblastoma, children living in high-
190	poverty counties experienced a 74% increased hazard of death compared to those living in low-
191	poverty counties. In a sub-cohort of children with available insurance data, those with any Medicaid
192	insurance experienced a 40% increased hazard of death compared to those with other insurance. As
193	reported elsewhere ⁴ , OS for the entire cohort improved steadily over time (from 53% to 77%).
194	Notably, while SES-related survival disparities persisted over time, they did not widen, despite a shift
195	to a modern standard of care which includes highly-centralized and intensive therapy.
196	G
197	Our data build on prior work identifying SES-associated survival disparities in children with high-risk
198	neuroblastoma treated on clinical trials ¹¹ by investigating whether outcome disparities have
199	increased with advances in care. This question is highly pertinent in the current era when the dual

potential of precision medicine to improve outcomes while simultaneously worsening health
disparities if these advances are not delivered equitably must be considered.²⁶⁻²⁸ Our data
demonstrate that despite a shift to a standard of care requiring resource-intense and highly
centralized treatment, existing SES-associated disparities have not worsened. Although analyses of
this cohort are limited by absence of patient-level treatment data in the SEER database, the
equitable survival gains across all groups over time are encouraging and stand in contrast to
widening disparities observed in other populations (e.g. asthma, adult cancers).^{29,30} These data

- 207 suggest that access to treatment advances with known survival benefits has been generally
- equitable among children with high-risk neuroblastoma, perhaps due to the high reliance on clinical
 trial enrollment and delivery of protocolized care in pediatric oncology.
- We did, however, observe persistent and clinically meaningful survival disparities associated with
 SES. That U.S. children of lower SES with high-risk neuroblastoma continue to die at higher rates
 than their higher SES counterparts must also be highlighted.¹¹ Despite focused policy statements and
 advocacy efforts identifying poverty as a major determinant and predictor of adverse health
 outcomes in children, these data demonstrate no improvements in this area over time. Recent

studies in other cancer populations have similarly found insurance and neighborhood (i.e. countybased) poverty to be predictors of inferior outcomes.^{31,32}

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Of note, children with high-risk neuroblastoma receive intensive (largely inpatient) treatment for 18-219 220 months at specialized, tertiary care centers. Therefore, this population is in many ways optimally 221 positioned to minimize disparities in care. Our finding that survival disparities persist even in this 222 population suggests that there may be other fundamental mechanisms driving SES-related gaps, warranting exploration of mechanisms beyond access to care. Non-adherence to chemotherapy is a 223 224 mechanism leading to disparate outcomes that has been well-described in pediatric acute lymphoblastic leukemia (ALL).^{33,34} Future work could investigate if there are similar findings in high-225 risk neuroblastoma with outpatient GM-CSF/IL-2 and oral cis-retinoic acid. However, it is important 226 227 to note that ALL treatment (which includes two years of outpatient-based therapy) contrasts sharply 228 with the predominantly inpatient treatment of high-risk neuroblastoma. The intensity of high-risk

229 neuroblastoma treatment including multiple prolonged admissions may place excess and 230 disproportionate burden on lower SES families who may have fewer resources for childcare, missed work/compensation, and household material needs. Whether access to relapse therapies, which are 231 similarly intensive, is inferior in families of lower SES, leading to earlier death following relapse, 232 should be investigated. Promising SES-focused interventions are being developed across various 233 234 healthcare settings to address potential mechanisms underlying observed disparities. Examples 235 include systematic provision of grocery and transportation services, clinic-based screening and referrals for community-based resources for a range of basic material needs, and free clinic-236 embedded tax services to assist families in obtaining earned income tax credits.³⁵⁻³⁷ Future 237 238 evaluation of such interventions to improve health equity in pediatric oncology is paramount.

239

Our study benefits from a robust, population-based sample across more than two decades of high-240 risk neuroblastoma treatment, allowing for analysis of changes in survival over time. The SEER 241 database is uniquely positioned as a population-based registry to provide disparities data as it is 242 243 intentionally biased to oversample minority populations and those that have traditionally been underrepresented in clinical trials.³⁸ Interestingly, we did not identify racial and ethnic disparities 244 observed in other neuroblastoma cohorts.^{9,10} Importantly, our cohort differs from these previous 245 246 studies in that we restricted analyses to those with high-risk disease, which is more prevalent in 247 minority populations. As such, our data may thus reflect prior findings that inferior survival observed for Black patients compared to their white counterparts was attributable to their higher prevalence 248 of high-risk disease at diagnosis.¹⁰ Our data build on recent publications demonstrating that SES 249 mediates racial and ethnic survival disparities across pediatric cancer,⁹ and that low-SES is associated 250

251 with inferior survival in the context of modern-era clinical trials.¹¹ These disparities

disproportionately impact children of racial and ethnic minority status who disproportionately live in
low-SES households due to structural disadvantages and biases. We importantly find that these
disparities are not, however, worsening in the modern era of complex treatment delivery. Future
attention to characterizing the relationships between SES and outcomes is essential to begin to
narrow the survival gap.

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There are important limitations to our data. Inherent to any large registry, SEER data are limited by 258 missing data/unrecorded variables, coding reliability, and selection bias.³⁹ We utilized a proxy for 259 260 high-risk (stage M) neuroblastoma in the absence of histological, genetic, and staging variables in SEER. While this proxy definition approximates elements of modern staging criteria, we may have 261 misclassified children with lower risk disease. We similarly lacked access to patient-level treatment 262 data, though our findings are consistent with SES-associated disparities in the clinical trial setting. 263 264 Finally, we had access to individual-level (insurance) SES data for only the sub-cohort of patients 265 diagnosed after 2007, limiting our ability to consider the impact of insurance across all treatment eras. Given the magnitude of the effect of insurance status on survival, however, similar findings 266 267 would be expected prior to 2007. SEER also codes patients with Medicaid-only (e.g. based on income 268 eligibility) and Medicaid as a second-insurer identically. Consequently, some patients with Medicaid 269 as a second insurer may have been misclassified as low-SES, an error which would bias toward the null, lending additional weight to our finding of survival disparities according to insurance status. 270 271

272 Our findings are important for two reasons. First, our data demonstrate that a steady shift toward 273 highly-centralized and intensive therapy has not resulted in worsened survival disparities for children of lower-SES despite the very reasonable concern that these children may have inferior access to 274 modern therapeutic advances. These findings suggest that the highly-centralized and structured care 275 delivery model of pediatric oncology allows for equitable integration of novel therapies into the 276 277 standard of care They suggest a model of care that could be applied to other patient populations— 278 in oncology and more generally—for whom such therapies and other resource-intensive treatment modalities are entering the clinical space. Second, our data demonstrate that in the 21st century, 279 280 children of low-SES continue to die at higher rates than their higher-SES counterparts. In other words, disparities have not worsened, but they have certainly not improved. These data should be a 281 call to action in an era purportedly focused on issues of equity. Future research, while continuing to 282 283 aim to improve survival and decrease toxicity for all children with cancer, must also work to 284 elucidate mechanisms underlying SES-related disparities and incorporate interventions to address social determinants of health to ensure that these gains are experienced equitably. 285

Author

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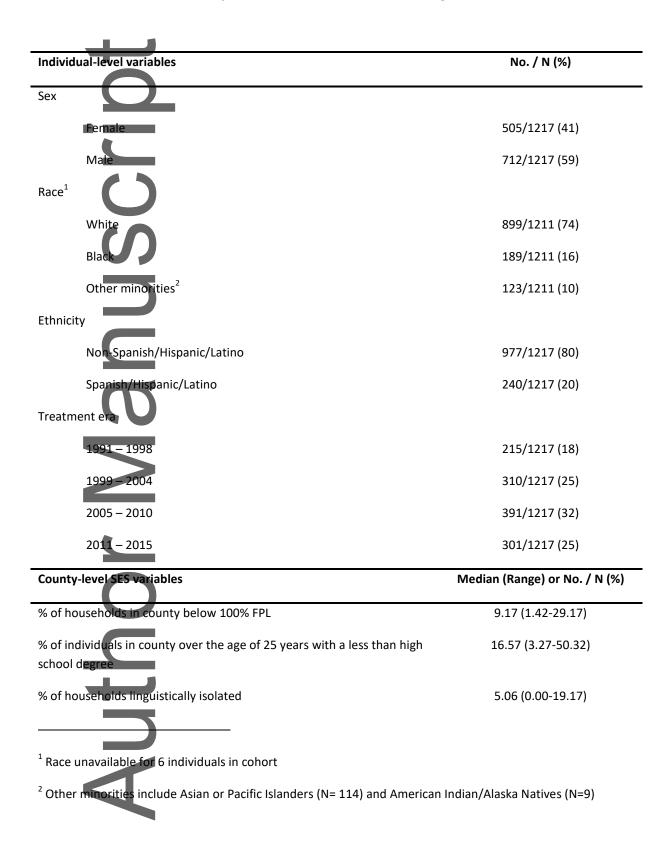


 Table 1. Individual- and County-level Patient Characteristics at Diagnosis (N = 1,217)

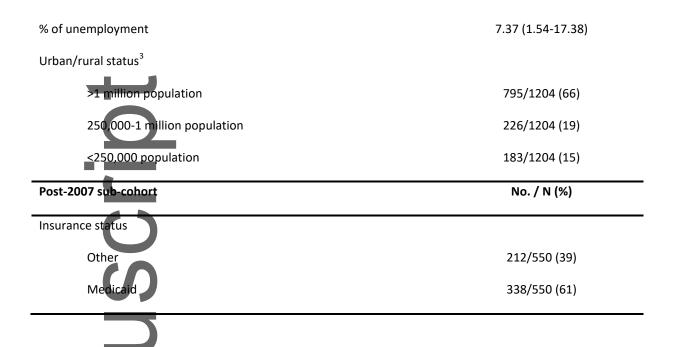


Table 2. Univariate Analysis for Individual- and County-level Factors on Overall Survival (OS)

ognostic Factors	No.	2-Year OS ± SE (%)	HR (95% CI)	P value
Female	505	68.6 ± 2.1	Reference	.66
Male	712	69.1 ± 1.8	1.04 (0.88, 1.22)	
ce				
White	899	69.3 ± 1.6	Reference	.46
Black	189	70.6 ± 3.4	0.94 (0.75, 1.18)	
Other minorities	123	62.7 ± 4.6	1.15 (0.88, 1.49)	
nnicity				
Non-Spanish/Hispanic/Latino	977	69.0 ± 1.5	Reference	.10
Spanish/Hispanic/Latino	240	68.4 ± 3.2	1.18 (0.97, 1.43)	

Treatment Era (trend)⁴

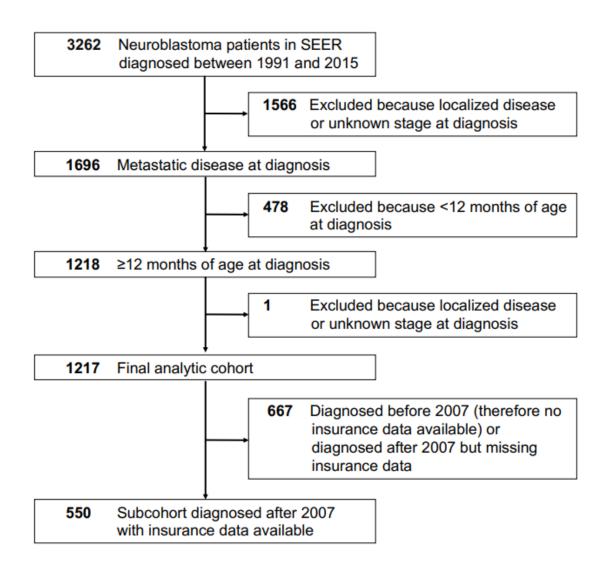
1991 – 1998	215	53.0 ± 3.4	0.79 (0.73, 0.86)	<.001
1999 – 2004	310	69.0 ± 2.6		
2005 – 2010	391	72.9 ± 2.3		
2011 – 2015	301	76.9 ± 2.9		
% of households in county below 1009	6			
FPL				
Low poverty (<20%)	1181	69.3 ± 1.4	Reference	.007
High poverty (≥20%)	36	55.4 ± 8.6	1.74 (1.17, 2.60)	
% of individuals in county over the age		55.4 2 0.0	1.74 (1.17, 2.00)	
of 25 years with a less than high schoo				
degree				
<90 th percentile (30%)				
≥90 th percentile	1115	69.0 ± 1.4	Reference	.48
(U	102	68.6 ± 4.6	1.10 (0.85, 1.42)	
% of households linguistically isolated				
<90 th percentile (13%)	1103	68.5 ± 1.5	Reference	.85
$\geq 90^{\text{th}}$ percentile	114	73.4 ± 4.4	0.97 (0.74, 1.29)	
% of unemployment				
<90 th percentile (12%)	1102	68.3 ± 1.5	Reference	.54
≥90 th percentile	115	74.9 ± 4.4	0.91 (0.68, 1.22)	
Urban/rural status by population				
>1 million	795	69.5 ± 1.7	Reference	.37
				

⁴ Continuous treatment year was used in univariate analyses. HR indicates survival gain over time. For example, there is a 21% lower hazard of survival for patients diagnosed in 1999-2004 compared to those diagnosed in 1991-1998.

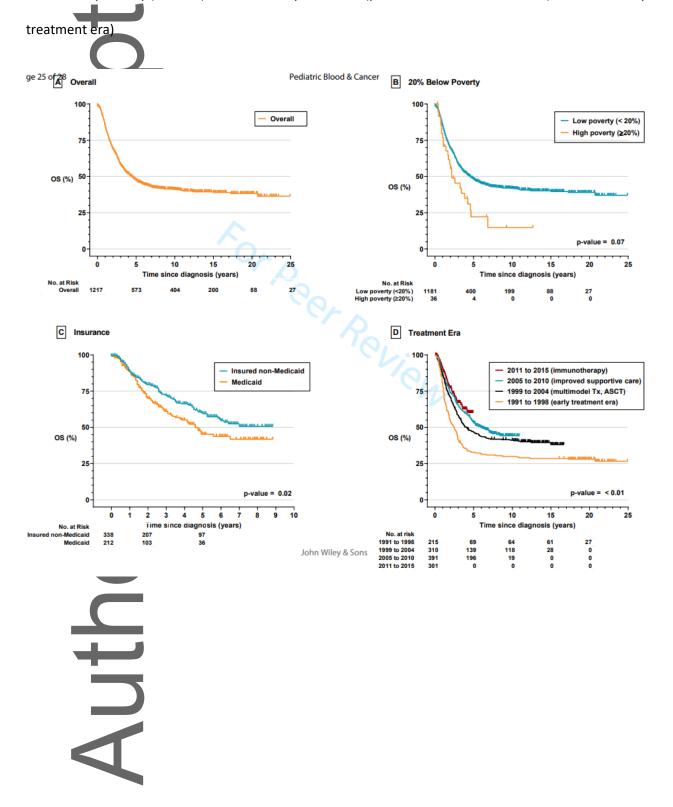
226	73 2 + 3 1	0 94 (0 76	1 16)
-			
183	63.6 ± 3.7	1.14 (0.91,	1.41)
338	79.6 ± 2.3	Referen	ce .0
212	69.9 ± 3.5	1.40 (1.05,	1.86)
Models Includi	ng Race, Ethnicity,	, and Treatment E	ra
		Hazard Patio	Adjusted <i>P</i> -value
>			Adjusted P-value
% of households	below poverty	2.08	<0.001
		Global p: 0.430	
		Reference	Reference
		0.98	0.85
		1.18	0.22
		1.22	0.05
		0.78	<.001
% of households	below poverty	2.38	0.001
		Global p: 0.542	
		Globa	l p: 0.542
		Globa Reference	l p: 0.542 Reference
n cohort			
	212 Models Includi	183 63.6 ± 3.7 338 79.6 ± 2.3 212 69.9 ± 3.5	183 63.6 ± 3.7 1.14 (0.91, 338 79.6 ± 2.3 Referen 212 69.9 ± 3.5 1.40 (1.05, Models Including Race, Ethnicity, and Treatment E s Hazard Ratio % of households below poverty 2.08 Globa Reference 0.98 1.18 1.22 0.78



Figure Legends J. anu



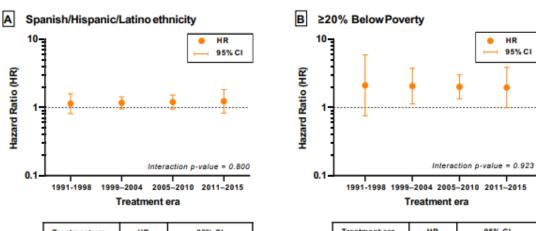
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Low-Poverty County (N=1217); C. Stratified by Insurance (post-2007 Sub-Cohort, N=550); D. Stratified by

Figure 2. Kaplan Meier curves of Overall Survival (A. Overall Cohort (N=1217); B. Stratified by High- and

Figure 3. Hazard ratio plot with 95% CI of (A) Spanish/Hispanic/Latino ethnicity and (B) High-poverty county in multivariate analyses controlling for treatment era and SES*treatment era interaction on overall survival. Interaction p-value demonstrates the interaction effect of SES*treatment era. CI, confidence interva



Treatment era	HR	95% CI
1991-1998	1.140	0.814-1.596
1999-2004	1.171	0.953-1.438
2005-2010	1.203	0.948-1.526
2011-2015	1.235	0.832-1.834

HR 95% CI Treatment era 1991-1998 2.118 0.756-5.932 1999-2004 2.067 1.130-3.780 2005-2010 2.016 1.342-3.029 2011-2015 0.996-3.886 1.967

HR •

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95% CI

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