

Supporting Information for “Cluster Non-Gaussian Functional Data” by

Qingzhi Zhong¹, Huazhen Lin^{1*}, Yi Li²

1. Center of Statistical Research and School of Statistics,

Southwestern University of Finance and Economics, Chengdu, China 611130

2. Department of Biostatistics,

University of Michigan, Ann Arbor, MI48109, USA

**email*: linhz@swufe.edu.cn

1 Web Appendix A: Notations

$$U_n = \{\boldsymbol{\alpha}'\mathbf{B}_n(t) : \boldsymbol{\alpha} = (\alpha_1, \dots, \alpha_{q_n})' \in \mathcal{R}^{q_n}, \max_{1 \leq i \leq q_n} |\alpha_i| \leq c_0, t \in [0, 1]\},$$

$$\Phi_n = \{\boldsymbol{\beta}'\mathbf{B}_n(t) : \boldsymbol{\beta} = (\beta_1, \dots, \beta_{q_n})' \in \mathcal{R}^{q_n}, \max_{1 \leq i \leq q_n} |\beta_i| \leq c_0, t \in [0, 1]\},$$

$$\boldsymbol{\Omega}_n^* = \{\boldsymbol{\Omega}_n = (\boldsymbol{\Lambda}', \boldsymbol{\sigma}^{2'}, \boldsymbol{\pi}', \boldsymbol{\mu}', \boldsymbol{\phi}')' \in R_+^{\sum_{g=1}^C K_g} \otimes R_+^C \otimes [0, 1]^C \otimes U_n^C \otimes \Phi_n^{\sum_{g=1}^C K_g}\},$$

$$d(\boldsymbol{\Omega}_1, \boldsymbol{\Omega}_2) = (\|\boldsymbol{\Lambda}_1 - \boldsymbol{\Lambda}_2\|^2 + \|\boldsymbol{\sigma}_1^2 - \boldsymbol{\sigma}_2^2\|^2 + \|\boldsymbol{\pi}_1 - \boldsymbol{\pi}_2\|^2 + \|\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2\|_2^2 + \|\boldsymbol{\phi}_1 - \boldsymbol{\phi}_2\|_2^2)^{1/2},$$

$$\sigma_{ij,g}(\boldsymbol{\Omega}_n) = \boldsymbol{\phi}_g(t_{ij})'\boldsymbol{\Lambda}_g\boldsymbol{\phi}_g(t_{ij}) + \sigma_g^2,$$

$$G(\omega; y, \boldsymbol{\Omega}_n) = E \left[\sum_{g=1}^C \pi_{g0} \Phi \left\{ \frac{H_0(y) - \mu_{g0}(t_{ij})}{\sqrt{\sigma_{ij,g}(\boldsymbol{\Omega}_0)}} \right\} - \pi_g \Phi \left\{ \frac{\omega - \mathbf{B}_n(t_{ij})'\boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\boldsymbol{\Omega}_n)}} \right\} \right],$$

$$G_n(\omega; y, \boldsymbol{\Omega}_n) = \frac{1}{\sum_{i=1}^C n_i} \sum_{i=1}^n \sum_{j=1}^{n_i} \left[I(Y_{ij} \leq y) - \sum_{g=1}^C \pi_g \Phi \left\{ \frac{\omega - \mathbf{B}_n(t_{ij})'\boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\boldsymbol{\Omega}_n)}} \right\} \right],$$

$$l_i(\boldsymbol{\Omega}_n; H) = \log \left\{ \sum_{g=1}^C \pi_g f_{gi}(\boldsymbol{\Omega}_n; H) \right\} - \lambda \sum_{g=1}^C \log \left\{ \frac{\epsilon + \pi_g}{\epsilon} \right\}$$

$$Pl(\boldsymbol{\Omega}_n; H) = El_i(\boldsymbol{\Omega}_n; H), P_n l(\boldsymbol{\Omega}_n; H) = \frac{1}{n} \sum_{i=1}^n l_i(\boldsymbol{\Omega}_n; H).$$

Writing $\sqrt{n}(P_n - P)l(\boldsymbol{\Omega}_n; H)$ for the empirical process indexed by $l_i(\boldsymbol{\Omega}_n; H)$. $\widehat{H}_n(y; \boldsymbol{\Omega}_n)$ is the estimator of $H(y)$ given $\boldsymbol{\Omega}_n$ and is the solution of (3.17) with respect to $H(y)$.

2 Web Appendix B: Lemmas

In this section, we sketch the proofs of Theorems 1-3 of the paper. To prove Theorems 1-3, we will employ the theory of empirical processes, and some techniques commonly used in semiparametric literature. Define the class of functions $\mathcal{L}_n = \{l_i(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) : \boldsymbol{\Omega}_n \in \boldsymbol{\Omega}_n^*\}$. For any $\varepsilon > 0$, the $L_1(P_n)$ covering number $N(\varepsilon, \mathcal{L}_n, L_1(P_n))$ of \mathcal{L}_n is the smallest value κ for which there exist $\{\boldsymbol{\Omega}_{n,j} \in \boldsymbol{\Omega}_n^*, j = 1, \dots, \kappa\}$, such

that for any $\Omega_n \in \Omega_n^*$,

$$\min_{j \in \{1, \dots, \kappa\}} \frac{1}{n} \sum_{i=1}^n |l_i(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - l_i(\Omega_{n,j}; \widehat{H}_n(\cdot; \Omega_{n,j}))| \leq \varepsilon.$$

We will define $N(\varepsilon, \mathcal{L}_n, L_1(P_n)) = \infty$ if no such κ exists. To obtain the proofs, we first give some lemmas.

Lemma 1 *By Lemma 2.5 and Corollary 2.6 in Van de Geer (2000), we can get the covering number of Ω_n^* satisfying*

$$N(\varepsilon, \Omega_n^*, L_2) \leq c_1 c_0^K M_n^d \cdot \varepsilon^{-(d+K+2C)},$$

where c_0, c_1 are finite constants independent of n, K , and $K = \sum_{g=1}^C K_g$, $d = (K + C)q_n$.

Lemma 2 *Assume that Conditions (C1)-(C5) hold, then*

$$\sup_{\Omega_n \in \Omega_n^*, y \in [y_1, y_2]} |\widehat{H}_n(y; \Omega_n) - H(y; \Omega_n)| \rightarrow 0,$$

where $H(y; \Omega_n)$ is the solution to

$$G(H(y; \Omega_n); y, \Omega_n) = 0, \tag{2.1}$$

for any given $\Omega_n \in \Omega_n^*, y \in [y_1, y_2]$.

Proof: We first prove the existence of $\widehat{H}_n(y; \Omega_n)$, then uniform convergence of $\widehat{H}_n(y; \Omega_n)$ to $H(y; \Omega_n)$ for $y \in [y_1, y_2]$ and $\Omega_n \in \Omega_n^*$. By the Lemma 1 and Theorem 19.4 of Van der Vaart (1998), Ω_n^* is P-Glivenko-Cantelli class. Since $\pi_g \Phi \left\{ \frac{H_0(y) - \mathbf{B}_n(t_{ij})' \boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\Omega_n)}} - \vartheta \right\}$ is a continuous function on Ω_n^* and the indicator function class is VC, and both are bounded by 1. Then by Theorem 3.7 of Van de Geer (2000) and the monotonicity of $H_0(y)$, we have

$$\begin{aligned} & \frac{1}{\sum_{i=1}^C n_i} \sum_{i=1}^n \sum_{j=1}^{n_i} \left[I(Y_{ij} \leq y) - \sum_{g=1}^C \pi_g \Phi \left\{ \frac{H_0(y) - \mathbf{B}_n(t_{ij})' \boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\Omega_n)}} - \vartheta \right\} \right] \\ & \rightarrow \sum_{g=1}^C E \left[\pi_{g0} \Phi \left\{ \frac{H_0(y) - \mu_{g0}(t_{ij})}{\sqrt{\sigma_{ij,g}(\Omega_0)}} \right\} - \pi_g \Phi \left\{ \frac{H_0(y) - \mathbf{B}_n(t_{ij})' \boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\Omega_n)}} - \vartheta \right\} \right], \end{aligned} \tag{2.2}$$

almost surely as $n \rightarrow \infty$, for any $\vartheta \geq 0$, and uniformly for $\boldsymbol{\Omega}_n \in \boldsymbol{\Omega}_n^*$ and $y \in [y_1, y_2]$.

By (2.2), for large n and sufficiently large ϑ ,

$$\frac{1}{\sum_{i=1}^C n_i} \sum_{i=1}^n \sum_{j=1}^{n_i} \left[I(Y_{ij} \leq y) - \sum_{g=1}^C \pi_g \Phi \left\{ \frac{H_0(y) - \mathbf{B}_n(t_{ij})' \boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\boldsymbol{\Omega}_n)}} - \vartheta \right\} \right] > 0, \quad (2.3)$$

$$\frac{1}{\sum_{i=1}^C n_i} \sum_{i=1}^n \sum_{j=1}^{n_i} \left[I(Y_{ij} \leq y) - \sum_{g=1}^C \pi_g \Phi \left\{ \frac{H_0(y) - \mathbf{B}_n(t_{ij})' \boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\boldsymbol{\Omega}_n)}} + \vartheta \right\} \right] < 0, \quad (2.4)$$

This together with the monotonicity and continuity of Φ implies that there exists a unique $\widehat{H}_n(y; \boldsymbol{\Omega}_n)$ such that

$$\frac{1}{\sum_{i=1}^C n_i} \sum_{i=1}^n \sum_{j=1}^{n_i} \left[I(Y_{ij} \leq y) - \sum_{g=1}^C \pi_g \Phi \left\{ \frac{\widehat{H}_n(y; \boldsymbol{\Omega}_n) - \mathbf{B}_n(t_{ij})' \boldsymbol{\alpha}_g}{\sqrt{\sigma_{ij,g}(\boldsymbol{\Omega}_n)}} \right\} \right] = 0. \quad (2.5)$$

Furthermore, by Lemma 1 and the uniform strong law of large numbers, we have

$$G_n(H(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n) \rightarrow G(H(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n),$$

uniformly for $y \in [y_1, y_2]$ and $\boldsymbol{\Omega}_n \in \boldsymbol{\Omega}_n^*$.

Denote $\zeta_n = \sup_{\boldsymbol{\Omega}_n \in \boldsymbol{\Omega}_n^*, y \in [y_1, y_2]} \|G_n(H(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n)\|$, by the definition of $G(H(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n)$ in Supplementary 1 and the definition of $H(y; \boldsymbol{\Omega}_n)$ in (2.1), we have $G(H(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n) = 0$, hence $\zeta_n \rightarrow 0$. Note that

$$G_n(\widehat{H}_n(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n) = G_n(\widehat{H}_n(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n) - G_n(H(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n) + G_n(H(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n),$$

we have

$$0 = \|G_n(\widehat{H}_n(y; \boldsymbol{\Omega}_n); y, \boldsymbol{\Omega}_n)\| \geq M \|\widehat{H}_n(y; \boldsymbol{\Omega}_n) - H(y; \boldsymbol{\Omega}_n)\| - \zeta_n,$$

and hence $\widehat{H}_n(y; \boldsymbol{\Omega}_n) \rightarrow H(y; \boldsymbol{\Omega}_n)$ uniformly in $y \in [y_1, y_2]$ and $\boldsymbol{\Omega}_n \in \boldsymbol{\Omega}_n^*$.

Lemma 3 *Under Conditions (C1)-(C5), the covering number of \mathcal{L}_n satisfies*

$$N(\varepsilon, \mathcal{L}_n, L_1(P_n)) \leq c_1 c_0^K M_n^d \cdot \varepsilon^{-(d+K+2C)}.$$

Proof: For any $\Omega^{(1)} = (\Lambda^{(1)'}, \sigma^{(1)2'}, \pi^{(1)'}, \mu^{(1)'}, \phi^{(1)'})'$, $\Omega^{(2)} = (\Lambda^{(2)'}, \sigma^{(2)2'}, \pi^{(2)'}, \mu^{(2)'}, \phi^{(2)'})' \in \Omega_n^*$. By using the Taylor expansion and $\lambda\sqrt{n} \rightarrow 0$, we have

$$\begin{aligned} |l_i(\Omega^{(1)}; H(\cdot; \Omega^{(1)})) - l_i(\Omega^{(2)}; H(\cdot; \Omega^{(2)}))| &\leq M(\|\Lambda^{(1)} - \Lambda^{(2)}\| + \|\sigma^{(1)2} - \sigma^{(2)2}\| \\ &+ \|\pi^{(1)} - \pi^{(2)}\| + \sum_{g=1}^C \|\mu_g^{(1)} - \mu_g^{(2)}\|_\infty + \sum_{g=1}^C \sum_{k=1}^{K_g} \|\phi_{gk}^{(1)} - \phi_{gk}^{(2)}\|_\infty). \end{aligned} \quad (2.6)$$

Note that

$$\begin{aligned} \|\mu_g^{(1)} - \mu_g^{(2)}\|_\infty &= \sup_t \left| \sum_{j=1}^{q_n} \alpha_{gj}^{(1)} b_j(t) - \sum_{j=1}^{q_n} \alpha_{gj}^{(2)} b_j(t) \right| \\ &\leq M \max_{1 \leq j \leq q_n} |\alpha_{gj}^{(1)} - \alpha_{gj}^{(2)}| = M \|\alpha_g^{(1)} - \alpha_g^{(2)}\|_\infty. \end{aligned}$$

Similarly, we can have $\|\phi_{gk}^{(1)} - \phi_{gk}^{(2)}\|_\infty \leq M \|\beta_{gk}^{(1)} - \beta_{gk}^{(2)}\|_\infty$. Then by (2.6), we have that

$$\begin{aligned} P_n |l(\Omega^{(1)}; H(\cdot; \Omega^{(1)})) - l(\Omega^{(2)}; H(\cdot; \Omega^{(2)}))| &\leq M(\|\Lambda^{(1)} - \Lambda^{(2)}\| + \|\sigma^{(1)2} - \sigma^{(2)2}\| \\ &+ \|\pi^{(1)} - \pi^{(2)}\| + \sum_{g=1}^C \|\alpha_g^{(1)} - \alpha_g^{(2)}\|_\infty + \sum_{g=1}^C \sum_{k=1}^{K_g} \|\beta_{gk}^{(1)} - \beta_{gk}^{(2)}\|_\infty). \end{aligned} \quad (2.7)$$

By the uniform convergence of $\widehat{H}_n(y; \Omega_n)$ to $H(y; \Omega_n)$ in Lemma 2, we have

$$P_n l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) = P_n l(\Omega_n; H(\cdot; \Omega_n)) + o_p(1).$$

This combining with (2.7) implies that, given $\Omega \in \Omega_n^*$, there exists $\Omega^{(j)} = (\Lambda^{(j)'}, \sigma^{(j)2'}, \pi^{(j)'}, \mu^{(j)'}, \phi^{(j)'})' \in \Omega_n^*$ such that

$$\begin{aligned} P_n |l(\Omega; \widehat{H}_n(\cdot; \Omega)) - l(\Omega^{(j)}; \widehat{H}_n(\cdot; \Omega^{(j)}))| &\leq M(\|\Lambda - \Lambda^{(j)}\| + \|\sigma^2 - \sigma^{(j)2}\| + \|\pi - \pi^{(j)}\| \\ &+ \sum_{g=1}^C \|\alpha_g - \alpha_g^{(j)}\|_\infty + \sum_{g=1}^C \sum_{k=1}^{K_g} \|\beta_{gk} - \beta_{gk}^{(j)}\|_\infty). \end{aligned} \quad (2.8)$$

By (2.8) and Lemma 1 and following the calculation on page 94 of Van der Vaart & Weller(1996), we have

$$N(\varepsilon, \mathcal{L}_n, L_1(P_n)) \leq c_1 c_0^K M_n^d \cdot \varepsilon^{-(d+K+2C)}.$$

This completes the proof.

Lemma 4 *Suppose that Conditions (C1)-(C5) hold. Then*

$$\sup_{\Omega_n \in \Omega_n^*} |P_n l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n))| \rightarrow 0 \text{ almost surely.}$$

Proof: Note that

$$\begin{aligned} P_n l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n)) &= P_n l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - P_n l(\Omega_n; H(\cdot; \Omega_n)) \\ &+ P_n l(\Omega_n; H(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n)). \end{aligned}$$

According to uniform convergence of $\widehat{H}_n(y; \Omega_n)$ in Lemma 2, we have

$$\sup_{\Omega_n \in \Omega_n^*} |P_n l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - P_n l(\Omega_n; H(\cdot; \Omega_n))| \leq \sup_{\Omega_n \in \Omega_n^*} |P_n l(\Omega_n; H(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n))| + o_p(1).$$

Hence, we only need to prove $\sup_{\Omega_n \in \Omega_n^*} |P_n l(\Omega_n; H(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n))| \rightarrow 0$.

By Condition (C5), let $(v + e)/2 < \phi < 1/2$ and $\alpha_n = n^{-1/2+\phi}(\log n)^{1/2}$. Here, $\{\alpha_n\}$

is a non-increasing positive numbers sequence, and for the given $\varepsilon > 0$ in Lemma 1,

let $\varepsilon_n = \varepsilon \alpha_n$. Under Condition (C2), $Pl^2(\Omega_n; H(\cdot; \Omega_n))$ is bounded. Then for any

$\Omega_n \in \Omega_n^*$ and sufficiently large n , we have

$$\frac{\text{var}[P_n l(\Omega_n; H(\cdot; \Omega_n))]}{(4\varepsilon_n)^2} \leq \frac{(1/n)Pl^2(\Omega_n; H(\cdot; \Omega_n))}{16\varepsilon^2\alpha_n^2} \leq \frac{c_0}{16n\varepsilon^2\alpha_n^2} \leq \frac{1}{16\varepsilon^2 n^{2\phi} \log n} \leq \frac{1}{2},$$

where c_0 is finite constant independent of n .

Furthermore, by the inequality (31) and Lemma 33 of Pollard (1984) and Lemma

3, we have

$$\begin{aligned}
& P\left[\sup_{\Omega_n \in \Omega_n^*} |P_n l(\Omega_n; H(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n))| > 8\varepsilon_n\right] \\
& \leq 8N(\varepsilon_n, \mathcal{L}_n, L_1(P_n)) \exp(-n\varepsilon_n^2/128) P\left[\sup_{\Omega_n \in \Omega_n^*} |P_n l^2(\Omega_n; H(\cdot; \Omega_n))| \leq 64\right] \\
& \quad + P\left[\sup_{\Omega_n \in \Omega_n^*} |P_n l^2(\Omega_n; H(\cdot; \Omega_n))| > 64\right] \\
& \leq M c_0^K M_n^d \cdot \varepsilon_n^{-(d+K+2C)} \exp(-n\varepsilon_n^2/128) \\
& \leq M \exp[(d+K+2C)v \log n - (d+K+2C) \log\{\varepsilon n^{-1/2+\phi}(\log n)^{1/2}\} - n\varepsilon^2 n^{-1+2\phi} \log n/128] \\
& = M \exp[(d+K+2C)\{(v-1/2+\phi) \log n - \log \log n/2 - \log \varepsilon\} - \varepsilon^2 n^{2\phi} \log n/128] \\
& \leq M \exp(-Mn^{2\phi} \log n),
\end{aligned}$$

where M is a constant. Hence, $\sum_{n=1}^{\infty} P[\sup_{\Omega_n \in \Omega_n^*} |P_n l(\Omega_n; H(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n))| > 8\varepsilon_n] < \infty$. By the Borel-Cantelli lemma (Chandra, 2012, pp. 15), we have $\sup_{\Omega_n \in \Omega_n^*} |P_n l(\Omega_n; H(\cdot; \Omega_n)) - Pl(\Omega_n; H(\cdot; \Omega_n))| \rightarrow 0$ almost sure. This completes the proof.

3 Web Appendix C: Proof of Theorems 1-3

Proof of Theorem 1

The proof of Theorem 1 is split into three steps. The first step proves the consistency of $\widehat{\Omega}_n$ and $\widehat{H}_n(\cdot; \widehat{\Omega}_n)$. The second step consists of the convergence rate of $\widehat{\Omega}_n$. Finally, we obtain the selection consistency for the cluster number.

Step 1: consistency

Under Conditions (C1) and (C2), by Corollary 6.21 of Schumaker (1981), there

exist $\mu_{n,g} = \boldsymbol{\alpha}'_{g0} \mathbf{B}_n(t)$ and $\phi_{n,gk} = \boldsymbol{\beta}'_{gk0} \mathbf{B}_n(t)$ such that

$$\begin{aligned} \sup_{t \in [0,1]} |\mu_{n,g}(t) - \mu_{g0}(t)| &= O(q_n^{-r}), \text{ and} \\ \sup_{t \in [0,1]} |\phi_{n,gk}(t) - \phi_{gk0}(t)| &= O(q_n^{-r}), \end{aligned}$$

where $\mu_{g0}(\cdot)$ and $\phi_{gk0}(\cdot)$ denote the true functions of $\mu_g(\cdot)$ and $\phi_{gk}(\cdot)$, respectively, for $g = 1, \dots, C, k = 1, \dots, K_g$. Let $\boldsymbol{\Omega}_{n0} = (\boldsymbol{\Lambda}'_0, \boldsymbol{\sigma}'_0, \boldsymbol{\pi}'_0, \boldsymbol{\mu}'_{n0}, \boldsymbol{\phi}'_{n0})'$, where $\boldsymbol{\mu}_{n0} = (\mu_{n,1}, \dots, \mu_{n,C})'$, $\boldsymbol{\phi}_{n0} = (\phi_{n,gk}, g = 1, \dots, C, k = 1, \dots, K_g)'$. Then we have

$$d(\boldsymbol{\Omega}_0, \boldsymbol{\Omega}_{n0}) = O(n^{-rv+e/2}). \quad (3.1)$$

Let $\boldsymbol{\Theta}_0 = (\boldsymbol{\pi}_0', \boldsymbol{\mu}_0', \boldsymbol{\Sigma}_0)'$ denote the true value of $\boldsymbol{\Theta}$, $\boldsymbol{\Theta}_{n0} = (\boldsymbol{\pi}_0', \boldsymbol{\mu}_{n0}', \boldsymbol{\Sigma}_{n0})'$, and

$$\begin{aligned} \Sigma_{g0}(s, t) &= \sum_{k=1}^{K_g} \phi_{gk0}(t) \lambda_{gk0} \phi_{gk0}(s) + \sigma_{g0}^2 I(s = t), \\ \Sigma_{n,g0}(s, t) &= \sum_{k=1}^{K_g} \phi_{n,gk}(t) \lambda_{gk0} \phi_{n,gk}(s) + \sigma_{g0}^2 I(s = t). \end{aligned}$$

From (3.1), we have that

$$\begin{aligned} &\sup_{(t,s) \in [0,1]^2} |\Sigma_{g0}(s, t) - \Sigma_{n,g0}(s, t)| \\ &\leq \sum_{k=1}^{K_g} \lambda_{gk0} \sup_{(t,s) \in [0,1]^2} |\phi_{gk0}(t) \phi_{gk0}(s) - \phi_{n,gk}(t) \phi_{n,gk}(s)| \\ &\leq \sum_{k=1}^{K_g} \lambda_{gk0} \left\{ \sup_{(t,s) \in [0,1]^2} |\phi_{gk0}(s)| |\phi_{gk0}(t) - \phi_{n,gk}(t)| + \sup_{(t,s) \in [0,1]^2} |\phi_{n,gk}(t)| |\phi_{gk0}(s) - \phi_{n,gk}(s)| \right\} \\ &= O(K^{1/2} q_n^{-r}). \end{aligned}$$

Then

$$d(\boldsymbol{\Theta}_{n0}, \boldsymbol{\Theta}_0) = O(n^{-rv+e/2}). \quad (3.2)$$

Let $M_i(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) = -l_i(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n))$, and $PM(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) = E\{M_i(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n))\}$, where ‘‘E’’ is expectation over i . $P_n M(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) = \frac{1}{n} \sum_{i=1}^n M_i(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n))$,

$K_\varsigma = \{\boldsymbol{\Omega}_n : d(\boldsymbol{\Omega}_n, \boldsymbol{\Omega}_{n0}) \geq \varsigma, \boldsymbol{\Omega}_n \in \boldsymbol{\Omega}_n^*\}$ for $\varsigma > 0$. Then, we have

$$\begin{aligned} \inf_{K_\varsigma} PM(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) &= \inf_{K_\varsigma} \left[PM(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) - P_n M(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) + P_n M(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) \right] \\ &\leq \zeta_{1n} + \inf_{K_\varsigma} P_n M(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)), \end{aligned} \quad (3.3)$$

where $\zeta_{1n} = \sup_{\boldsymbol{\Omega}_n \in \boldsymbol{\Omega}_n^*} |P_n M(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) - PM(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n))|$.

If $\widehat{\boldsymbol{\Omega}}_n \in K_\varsigma$, one can show that

$$\begin{aligned} \inf_{K_\varsigma} P_n M(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) &= P_n M(\widehat{\boldsymbol{\Omega}}_n; \widehat{H}_n(\cdot; \widehat{\boldsymbol{\Omega}}_n)) \\ &\leq P_n M(\boldsymbol{\Omega}_{n0}; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_{n0})) = \zeta_{2n} + PM(\boldsymbol{\Omega}_{n0}; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_{n0})), \end{aligned} \quad (3.4)$$

with $\zeta_{2n} = P_n M(\boldsymbol{\Omega}_{n0}; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_{n0})) - PM(\boldsymbol{\Omega}_{n0}; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_{n0}))$. By (3.3) and (3.4), we have

$$\inf_{K_\varsigma} PM(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) \leq \zeta_{1n} + \zeta_{2n} + PM(\boldsymbol{\Omega}_{n0}; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_{n0})) = \zeta_n + PM(\boldsymbol{\Omega}_{n0}; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_{n0})),$$

with $\zeta_n = \zeta_{1n} + \zeta_{2n}$. It is obvious that $\zeta_n \geq \delta_\varsigma \hat{=} \inf_{K_\varsigma} PM(\boldsymbol{\Omega}_n; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_n)) - PM(\boldsymbol{\Omega}_{n0}; \widehat{H}_n(\cdot; \boldsymbol{\Omega}_{n0}))$ which is larger than zero under condition $\lambda\sqrt{n} \rightarrow 0$ when n is large enough. Hence

$$\{\widehat{\boldsymbol{\Omega}}_n \in K_\varsigma\} \subseteq \{\zeta_n \geq \delta_\varsigma\}. \quad (3.5)$$

By Lemma 2 and Law of Large Numbers, we have $\zeta_{1n} \rightarrow 0$ and $\zeta_{2n} \rightarrow 0$ then $\zeta_n \rightarrow 0$ almost surely. Hence, when n is large enough $\bigcup_{i=1}^{\infty} \bigcap_{n=i}^{\infty} \{\zeta_n \geq \delta_\varsigma\}$ is null set. By (3.5), we have $\bigcup_{i=1}^{\infty} \bigcap_{n=i}^{\infty} \{\widehat{\boldsymbol{\Omega}}_n \in K_\varsigma\}$ is null set. Coupling with the definition of K_ς , we have

$$d(\widehat{\boldsymbol{\Omega}}_n, \boldsymbol{\Omega}_{n0}) \rightarrow 0, \quad (3.6)$$

almost surely as $n \rightarrow \infty$. This together with (3.1), we have $d(\widehat{\boldsymbol{\Omega}}_n, \boldsymbol{\Omega}_0) \rightarrow 0$. Further-

more, since

$$\begin{aligned}
\widehat{\Sigma}_g(s, t) - \Sigma_{n, g0}(s, t) &= \sum_{k=1}^{K_g} \left\{ \widehat{\lambda}_{gk} \widehat{\phi}_{gk}(t) \widehat{\phi}_{gk}(s) - \lambda_{gk0} \phi_{n, gk}(t) \phi_{n, gk}(s) \right\} + (\widehat{\sigma}_g^2 - \sigma_{g0}^2) I(s = t) \\
&= \sum_{k=1}^{K_g} \widehat{\lambda}_{gk} \left[\widehat{\phi}_{gk}(t) \{ \widehat{\phi}_{gk}(s) - \phi_{n, gk}(s) \} + \phi_{n, gk}(s) \{ \widehat{\phi}_{gk}(t) - \phi_{n, gk}(t) \} \right] \\
&\quad + \{ \widehat{\lambda}_{gk} - \lambda_{gk0} \} \phi_{n, gk}(t) \phi_{n, gk}(s) + (\widehat{\sigma}_g^2 - \sigma_{g0}^2) I(s = t), \tag{3.7}
\end{aligned}$$

therefore $\widehat{\Sigma}_g(s, t) \rightarrow \Sigma_{n, g0}(s, t)$ almost surely for $g = 1, \dots, C$ by (3.6). Then $d(\widehat{\Theta}_n, \Theta_{n0}) \rightarrow 0$, together with (3.2), we prove $d(\widehat{\Theta}_n, \Theta_0) \rightarrow 0$. Since $\widehat{H}_n(y) \equiv \widehat{H}_n(y; \widehat{\Omega}_n)$, coupling with Lemma 2, we have $\widehat{H}_n(y) \rightarrow H_0(y)$ uniformly in $y \in [y_1, y_2]$.

Step 2: convergence rate

To show the convergence rate, for any $\eta > 0$, define class $\mathcal{L}_\eta = \{l_i(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - l_i(\Omega_{n0}; \widehat{H}_n(\cdot; \Omega_{n0})) : \Omega_n \in \Omega_n^*, \eta/2 \leq d(\Omega_n, \Omega_{n0}) \leq \eta\}$, and $P[l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - l(\Omega_{n0}; \widehat{H}_n(\cdot; \Omega_{n0}))] = E[l_i(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - l_i(\Omega_{n0}; \widehat{H}_n(\cdot; \Omega_{n0}))]$. Following the calculation of Shen and Wong (1994, p.597), for $0 < \varepsilon < \eta$, we can establish that $\log N_{[]}(\varepsilon, \mathcal{L}_\eta, L_2(P)) \leq MKq_n \log(\eta/\varepsilon)$. Moreover, for large n and $\lambda\sqrt{n} \rightarrow 0$, we have $P[l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - l(\Omega_{n0}; \widehat{H}_n(\cdot; \Omega_{n0}))] \leq M\eta^2$, for any $l(\Omega_n; \widehat{H}_n(\cdot; \Omega_n)) - l(\Omega_{n0}; \widehat{H}_n(\cdot; \Omega_{n0})) \in \mathcal{L}_\eta$. Therefore, by Lemma 3.4.2 of Van der Vaart and Wellner (1996) and the condition $\lambda\sqrt{n} \rightarrow 0$, we have

$$E_P \|\sqrt{n}(P_n - P)\|_{\mathcal{L}_\eta} \leq MJ_{[]}(\eta, \mathcal{L}_\eta, L_2(P)) \left[1 + \frac{J_{[]}(\eta, \mathcal{L}_\eta, L_2(P))}{n^{1/2}\eta^2} \right], \tag{3.8}$$

where $J_{[]}(\eta, \mathcal{L}_\eta, L_2(P)) = \int_0^\eta \sqrt{1 + \log N_{[]}(\varepsilon, \mathcal{L}_\eta, L_2(P))} d\varepsilon \leq M(Kq_n)^{1/2}\eta$. The right hand of (3.8) yields that the key function $\phi_n(\eta) = \sqrt{Kq_n}\eta + Kq_n/n^{1/2}$. Note that, $\phi_n(\eta)/\eta$ is decreasing in η , and $r_n^2\phi_n(1/r_n) = r_n\sqrt{Kq_n} + r_n^2Kq_n/n^{1/2} \leq Mn^{1/2}$, where $r_n = (Kq_n)^{-1/2}n^{1/2} = n^{(1-v-e)/2}$. Hence $n^{(1-v-e)/2}d(\widehat{\Omega}_n, \Omega_{n0}) = O_p(1)$ by Theorem 3.4.1 of Van der Vaart and Wellner (1996). This with (3.1), yields that $d(\widehat{\Omega}_n, \Omega_0) = O_p\{n^{-\min(rv-e/2, (1-v-e)/2)}\}$. Furthermore, by $d(\widehat{\Omega}_n, \Omega_{n0}) = O_p\{n^{-(1-v-e)/2}\}$ and (3.7),

we have $d(\widehat{\Theta}_n, \Theta_{n0}) = O_p\{n^{-(1-v-e)/2}\}$, combining with (3.2), yields that

$$d(\widehat{\Theta}_n, \Theta_0) = O_p\{n^{-(1-v-e)/2} + n^{-rv+e/2}\} = O_p\{n^{-\min(rv-e/2, (1-v-e)/2)}\}. \quad (3.9)$$

Step 3: selection consistency

Denote $\dot{Q}_{ng}(\Omega_n; H(\cdot)) = \frac{\partial Q_n(\Omega_n; H(\cdot))}{\partial \pi_g}$. Following the lines of the proof for Theorem 3, for any given number of clusters C , we have that the estimates of parametric part have an $n^{1/2}$ rate of convergence, therefore, $\pi_g = \pi_{g0} + O_p(n^{-1/2})$ for $g = 1, \dots, C$ with the condition $\lambda\sqrt{n} \rightarrow 0$. According to the proof of Huang et al. (2017), we only need to consider the solution of $\dot{Q}_{ng}(\Omega_n; \widehat{H}(\cdot; \Omega_n)) = 0$ with $d(\Omega_n, \Omega_0) = O_p\{n^{-\min(rv-e/2, (1-v-e)/2)}\}$, $\pi_g < \frac{1}{\sqrt{n} \log(n)}$ and $g > C_0$. According to Lemma 2, it is sufficient to show that

$$\dot{Q}_{ng}(\Omega_n; H(\cdot; \Omega_n)) < 0 \quad \text{for} \quad 0 < \pi_g < \frac{1}{\sqrt{n} \log(n)} \text{ and } g > C_0, \quad (3.10)$$

with probability tending to 1 as $n \rightarrow \infty$. With the constraint $\sum_{g=1}^C \pi_g = 1$, we can write

$$\begin{aligned} \dot{Q}_{ng}(\Omega_n; H(\cdot; \Omega_n)) &= \sum_{i=1}^n \frac{f_{gi}(\Omega_n; H(\cdot; \Omega_n))}{\sum_{g=1}^C \pi_g f_{gi}(\Omega_n; H(\cdot; \Omega_n))} - \sum_{i=1}^n \frac{f_{1i}(\Omega_n; H(\cdot; \Omega_n))}{\sum_{g=1}^C \pi_g f_{gi}(\Omega_n; H(\cdot; \Omega_n))} \\ &\quad - n\lambda \frac{1}{\epsilon + \pi_g} + n\lambda \frac{1}{\epsilon + \pi_1} \hat{=} R_1 - R_2 - R_3 + R_4. \end{aligned} \quad (3.11)$$

By the law of large numbers, it is obvious that R_1 and R_2 are of order $O_p(n)$. Furthermore, we have $\pi_g = \pi_{g0} + O_p(n^{-1/2}) > \frac{1}{2} \cdot \min\{\pi_{10}, \dots, \pi_{C_0,0}\}$ when $g = 1$. Hence, R_4 should be $O_p(n\lambda) = o_p(n)$ by the condition $\lambda = o(1)$.

Because $\epsilon = o\{\frac{1}{\sqrt{n} \log(n)}\}$ in Theorem 1 when $g > C_0$ and $0 < \pi_g < \frac{1}{\sqrt{n} \log(n)}$, we have $R_3 = O_p(n\lambda\sqrt{n} \log n)$, with probability tending to one. Since R_1 and R_2 are of order $O_p(n)$ and R_4 is of order $o_p(n)$, hence, R_3 dominates R_1, R_2 and R_4 with the condition $\lambda\sqrt{n} \log n \rightarrow \infty$. Therefore, we prove (3.10), or equivalently $\pi_g = 0$ for $g > C_0$ with probability tending to one when $n \rightarrow \infty$.

Proof of Theorem 2

The proof of Theorem 2 follows by Steps 1 and 2 in the proof of Theorem 1.

Proof of Theorem 3

Let $\Omega_{\setminus 0}^*$ to be Ω^* excluding the true value Ω_0 , where Ω^* denotes the parameter space. Denote $\delta_n = n^{-\min(rv-e/2, (1-v-e)/2)}$. Let V denote the linear span of $\Omega_{\setminus 0}^*$ and define the Fisher inner product on the space V as $\langle \mathbf{v}, \tilde{\mathbf{v}} \rangle = P \left\{ \dot{l}_i(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}] \dot{l}_i(\Omega_0; H(\cdot; \Omega_0))[\tilde{\mathbf{v}}] \right\}$ for $\mathbf{v}, \tilde{\mathbf{v}} \in V$, the Fisher norm $\|\mathbf{v}\|^2 = \langle \mathbf{v}, \mathbf{v} \rangle$, where $\dot{l}_i(\Omega; H(\cdot; \Omega))[\mathbf{v}] = \left. \frac{dl_i(\Omega + s\mathbf{v}; H(\cdot; \Omega))}{ds} \right|_{s=0}$ is the first order directional derivative of $l_i(\Omega; H(\cdot; \Omega))$ at the direction $\mathbf{v} \in V$, and $\Omega \in \{\Omega \in \Omega^* : d(\Omega, \Omega_0) = O(\delta_n)\}$. Let \bar{V} be the closed linear span of V under the Fisher norm, then $(\bar{V}, \|\cdot\|)$ is a Hilbert space. Let $\dot{l}_{i,\Omega}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}] = \dot{l}_i(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}] + \dot{l}_{i,H}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}]$, where $\dot{l}_{i,H}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}] = \left. \frac{dl_i(\Omega_0; H(\cdot; \Omega_0 + s\mathbf{v}))}{ds} \right|_{s=0}$. Denote $\langle \mathbf{v}, \tilde{\mathbf{v}} \rangle_{\Omega} = P \left\{ \dot{l}_i(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}] \dot{l}_{i,\Omega}(\Omega_0; H(\cdot; \Omega_0))[\tilde{\mathbf{v}}] \right\}$. For a $2C_0$ -dimensional vector $\Gamma = (\Gamma'_1, \Gamma'_2)'$ with $\|\Gamma\| \leq 1$, let $\rho(\Omega; H(\cdot; \Omega)) = \Gamma' \boldsymbol{\theta} = \boldsymbol{\sigma}^{2'} \Gamma_1 + \boldsymbol{\pi}' \Gamma_2$ denote a smooth functional of Ω and $H(\cdot; \Omega)$, and $\dot{\rho}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}] = \left. \frac{d\rho(\Omega_0 + s\mathbf{v}; H(\cdot; \Omega_0 + s\mathbf{v}))}{ds} \right|_{s=0}$. Note that $\rho(\Omega; H(\cdot; \Omega)) - \rho(\Omega_0; H(\cdot; \Omega_0)) = \dot{\rho}(\Omega_0; H(\cdot; \Omega_0))[\Omega - \Omega_0]$. According to the Riesz representation theorem, for any given $\mathbf{v} \in V$, there exists $\mathbf{v}^* \in \bar{V}$ such that $\dot{\rho}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}] = \langle \mathbf{v}^*, \mathbf{v} \rangle_{\Omega}$. Thus, according to the Cramér-Wold device, to prove Theorem 3, it suffices to show that

$$\sqrt{n} \langle \mathbf{v}^*, \hat{\Omega}_n - \Omega_0 \rangle_{\Omega} \xrightarrow{d} N\{0, \Gamma' \mathbf{I}^{-1}(\boldsymbol{\theta}_0) \Gamma\}, \quad (3.12)$$

due to $\Gamma'(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0) = \rho(\hat{\Omega}_n; H(\cdot; \hat{\Omega}_n)) - \rho(\Omega_0; H(\cdot; \Omega_0)) = \dot{\rho}(\Omega_0; H(\cdot; \Omega_0))[\hat{\Omega}_n - \Omega_0] = \langle \mathbf{v}^*, \hat{\Omega}_n - \Omega_0 \rangle_{\Omega}$. In fact, (3.12) holds when $\sqrt{n} \langle \mathbf{v}^*, \hat{\Omega}_n - \Omega_0 \rangle_{\Omega} \xrightarrow{d} N(0, \|\mathbf{v}^*\|^2)$ and $\|\mathbf{v}^*\|^2 = \Gamma' \mathbf{I}^{-1}(\boldsymbol{\theta}_0) \Gamma$.

We will take two steps to prove (3.12). First, we prove $\sqrt{n} \langle \mathbf{v}^*, \hat{\Omega}_n - \Omega_0 \rangle_{\Omega} \xrightarrow{d} N(0, \|\mathbf{v}^*\|^2)$. By Corollary 6.21 of Schumaker (1981), there exists $\Pi_n \mathbf{v}^* \in \Omega_n^*$ such that $\|\Pi_n \mathbf{v}^* - \mathbf{v}^*\| = o(1)$ and $\delta_n \|\Pi_n \mathbf{v}^* - \mathbf{v}^*\| = o(n^{-1/2})$. For any $\Omega \in \{\Omega \in$

$\Omega^* : d(\Omega, \Omega_0) = O(\delta_n)$, define $\ddot{l}_i(\Omega; H(\cdot; \Omega))[\mathbf{v}, \tilde{\mathbf{v}}] = \left. \frac{d\dot{l}_i(\Omega + s\tilde{\mathbf{v}}; H(\cdot; \Omega + s\tilde{\mathbf{v}}))[\mathbf{v}]}{ds} \right|_{\tilde{s}=0}$, and $r_i[\Pi_n \mathbf{v}^*, \Omega - \Omega_0] = \dot{l}_i(\Omega; H(\cdot; \Omega))[\Pi_n \mathbf{v}^*] - \dot{l}_i(\Omega_0; H(\cdot; \Omega_0))[\Pi_n \mathbf{v}^*]$.

Since $P_n \dot{l}(\widehat{\Omega}_n; \widehat{H}(\cdot; \widehat{\Omega}_n))[\Pi_n \mathbf{v}^*] = 0$ by the definition of $\widehat{\Omega}_n$. Then by Lemma 2, we have

$$\begin{aligned}
0 &= P_n \dot{l}(\widehat{\Omega}_n; H(\cdot; \widehat{\Omega}_n))[\Pi_n \mathbf{v}^*] = P_n \dot{l}(\Omega_0; H(\cdot; \Omega_0))[\Pi_n \mathbf{v}^*] + P_n r[\Pi_n \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0] \\
&= P_n \dot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*] + P_n \dot{l}(\Omega_0; H(\cdot; \Omega_0))[\Pi_n \mathbf{v}^* - \mathbf{v}^*] + (P_n - P)r[\Pi_n \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0] \\
&\quad + Pr[\Pi_n \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0] \\
&= P_n \dot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*] + I_1 + I_2 + I_3.
\end{aligned} \tag{3.13}$$

We will investigate the asymptotic behavior of I_1, I_2 and I_3 . For I_1 , by Chebyshev inequality, $\lambda\sqrt{n} \rightarrow 0$, $P\dot{l}(\Omega_0; H(\cdot; \Omega_0))[\Pi_n \mathbf{v}^* - \mathbf{v}^*] = 0$ and $\|\Pi_n \mathbf{v}^* - \mathbf{v}^*\| = o(1)$, we have that

$$I_1 = o_p(n^{-1/2}). \tag{3.14}$$

For I_2 , by the definition of $r[\Pi_n \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0]$, we can get

$$I_2 = (P_n - P) \left\{ \dot{l}(\widehat{\Omega}_n; H(\cdot; \widehat{\Omega}_n)) - \dot{l}(\Omega_0; H(\cdot; \Omega_0)) \right\} [\Pi_n \mathbf{v}^*].$$

By Theorem 2.8.3 of Van der Varrrt and Wellner (1996), $\{\dot{l}(\Omega; H(\cdot; \Omega))[\Pi_n \mathbf{v}^*] : \|\Omega - \Omega_0\| = O_p(\delta_n)\}$ is a Donsker class. Hence by Theorem 2.11.23 of Van der Varrrt and Wellner (1996), we have

$$I_2 = o_p(n^{-1/2}). \tag{3.15}$$

For I_3 , by the definition of $r[\Pi_n \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0]$ and the Mean Value Theorem (Flett, 1958), we have

$$\begin{aligned}
I_3 &= P \left\{ \ddot{l}(\tilde{\Omega}; H(\cdot; \tilde{\Omega})) - \ddot{l}(\Omega_0; H(\cdot; \Omega_0)) \right\} [\Pi_n \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0] + P\ddot{l}(\Omega_0; H(\cdot; \Omega_0))[\Pi_n \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0] \\
&= o_p(n^{-1/2}) + P\ddot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*, \widehat{\Omega}_n - \Omega_0] + P\ddot{l}(\Omega_0; H(\cdot; \Omega_0))[\Pi_n \mathbf{v}^* - \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0] \\
&= o_p(n^{-1/2}) + P\ddot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*, \widehat{\Omega}_n - \Omega_0],
\end{aligned} \tag{3.16}$$

where $\tilde{\Omega}$ is between $\widehat{\Omega}_n$ and Ω_0 , the second equation follows by using a Taylor expansion, Condition (C2) and $\|\Pi_n \mathbf{v}^*\|^2 \rightarrow \|\mathbf{v}^*\|^2$, the last equation holds since $\delta_n \|\Pi_n \mathbf{v}^* - \mathbf{v}^*\| = o(n^{-1/2})$. Hence, by (3.14), (3.13), (3.15), (3.16), together with $\lambda\sqrt{n} \rightarrow 0$, $P\dot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*] = 0$ and the definition of $\langle \cdot, \cdot \rangle_{\Omega}$, we have that

$$\begin{aligned} 0 &= P_n \dot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*] - \langle \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0 \rangle_{\Omega} + o_p(n^{-1/2}) \\ &= (P_n - P) \dot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*] - \langle \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0 \rangle_{\Omega} + o_p(n^{-1/2}). \end{aligned}$$

Hence

$$\sqrt{n} \langle \mathbf{v}^*, \widehat{\Omega}_n - \Omega_0 \rangle_{\Omega} = \sqrt{n} (P_n - P) \dot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*] + o_p(1) \rightarrow N(0, \|\mathbf{v}^*\|^2),$$

with $\|\mathbf{v}^*\|^2 = \|\dot{l}(\Omega_0; H(\cdot; \Omega_0))[\mathbf{v}^*]\|^2$.

In the second step, we calculate $\|\mathbf{v}^*\|^2$. Rewrite $\boldsymbol{\theta} = (\boldsymbol{\sigma}^{2'}, \boldsymbol{\pi}')' = (\theta_1, \dots, \theta_{2C_0})'$, let $K_0 = \sum_{g=1}^{C_0} K_g$ and $l_{i,b_j} = \frac{\partial l_i(\Omega_0; H(\cdot; \Omega_0))}{\partial b_j}$. For each $\theta_q, q = 1, 2, \dots, 2C_0$, denote $\psi_q^* = \{b_{1q}^*, b_{2q}^*, \dots, b_{(C_0+K_0)q+K_0}^*, b_{H,q}^*\}$ be the minimizer

$$E \left\{ l_{i,\boldsymbol{\theta}} \cdot e_q - l_{i,b_1}[b_{1q}] - l_{i,b_2}[b_{2q}] - \dots - l_{i,b_{C_0+2K_0}}[b_{(C_0+K_0)q+K_0}] - l_{i,b_H}[b_{H,q}] \right\}^2,$$

with respect to $\psi_q = \{b_{1q}, b_{2q}, \dots, b_{(C_0+K_0)q+K_0}, b_{H,q}\}$, where e_q is a $2C_0$ -dimensional vector of zeros except the q th element equal to 1, $l_{i,\boldsymbol{\theta}} = (l'_{i,\boldsymbol{\sigma}^2}, l'_{i,\boldsymbol{\pi}})'$, $l_{i,\boldsymbol{\sigma}^2} = (l_{i,\sigma_1^2}, \dots, l_{i,\sigma_{C_0}^2})'$, $l_{i,\boldsymbol{\pi}} = (l_{i,\pi_1}, \dots, l_{i,\pi_{C_0}})'$, $l_{i,\sigma_g^2} = \frac{\partial l_i(\Omega_0; H(\cdot; \Omega_0))}{\partial \sigma_g^2}$, $l_{i,\pi_g} = \frac{\partial l_i(\Omega_0; H(\cdot; \Omega_0))}{\partial \pi_g}$, $g = 1, \dots, C_0$, and $(l_{i,b_j}[b_{jq}], j = 1, \dots, C_0)$ is the directional derivative of $\boldsymbol{\mu}$, $(l_{i,b_j}[b_{jq}], j = C_0 + 1, \dots, C_0 + K_0)$ is the directional derivative of $\boldsymbol{\phi}$, $(l_{i,b_j}[b_j], j = C_0 + K_0 + 1, \dots, C_0 + 2K_0)$ is the directional derivative of $\boldsymbol{\Lambda}$, and $l_{i,b_H}[b_{H,q}]$ is the directional derivative of H . By the similar calculation of Chen et al. (2006), we can obtain $\|\mathbf{v}^*\|^2 = \boldsymbol{\Gamma}' [E\{S(\boldsymbol{\theta}_0)S(\boldsymbol{\theta}_0)'\}]^{-1} \boldsymbol{\Gamma} = \boldsymbol{\Gamma}' \boldsymbol{\Gamma}^{-1}(\boldsymbol{\theta}_0) \boldsymbol{\Gamma}$, where $S(\boldsymbol{\theta}_0)$ is a $2C_0$ -dimensional vector, with the q th element as $l_{i,\boldsymbol{\theta}} \cdot e_q - l_{i,b_1}[b_{1q}^*] - l_{i,b_2}[b_{2q}^*] - \dots - l_{i,b_{C_0+2K_0}}[b_{(C_0+K_0)q+K_0}^*] - l_{i,b_H}[b_{H,q}^*]$. Then, we complete the proof of Theorem 3.

4 Web Appendix D: Tables 6-8 and Figures 3-9

Table 6: Results for Case 2 of Simulation 1.

	Proposed(C=7)		CT(C=3)		CT(C=7)		WoT(C=3)	
	bias(sd)	RMSE	bias(sd)	RMSE	bias(sd)	RMSE	bias(sd)	RMSE
π_1	0.005(0.043)	0.043	0.002(0.031)	0.031	0.003(0.041)	0.041	0.058(0.054)	0.080
π_2	0.009(0.052)	0.053	0.000(0.032)	0.032	0.008(0.046)	0.047	0.053(0.085)	0.100
π_3	0.003(0.052)	0.052	0.002(0.030)	0.030	0.003(0.038)	0.038	0.111(0.082)	0.138
σ_1^2	0.002(0.016)	0.016	0.000(0.011)	0.011	0.001(0.011)	0.011	0.094(0.002)	0.094
σ_2^2	0.003(0.018)	0.018	0.001(0.014)	0.014	0.001(0.014)	0.014	0.141(0.002)	0.141
σ_3^2	0.002(0.020)	0.021	0.001(0.015)	0.015	0.001(0.019)	0.019	0.188(0.001)	0.188
λ_{11}	0.022(0.152)	0.153	0.017(0.132)	0.133	0.022(0.140)	0.141	0.960(0.008)	0.960
λ_{12}	0.009(0.037)	0.038	0.003(0.036)	0.036	0.003(0.036)	0.036	0.236(0.005)	0.236
λ_{21}	0.031(0.152)	0.155	0.006(0.130)	0.130	0.021(0.131)	0.133	1.044(0.014)	1.044
λ_{22}	0.001(0.033)	0.033	0.001(0.010)	0.010	0.001(0.016)	0.016	0.003(0.009)	0.010
λ_{31}	0.040(0.151)	0.156	0.031(0.105)	0.110	0.033(0.135)	0.139	0.937(0.009)	0.937
λ_{32}	0.030(0.086)	0.091	0.015(0.058)	0.060	0.015(0.061)	0.063	0.132(0.005)	0.132
$\mu_1(\cdot)$	0.007(0.131)	0.132	0.006(0.091)	0.091	0.007(0.093)	0.093	0.362(0.074)	0.369
$\mu_2(\cdot)$	0.013(0.135)	0.135	0.007(0.105)	0.105	0.012(0.109)	0.110	0.458(0.099)	0.468
$\mu_3(\cdot)$	0.012(0.133)	0.134	0.010(0.089)	0.090	0.011(0.115)	0.115	1.252(0.073)	1.255
$\phi_{11}(\cdot)$	0.007(0.069)	0.069	0.005(0.063)	0.063	0.007(0.065)	0.065	0.187(0.475)	0.511
$\phi_{12}(\cdot)$	0.008(0.081)	0.082	0.007(0.075)	0.075	0.008(0.077)	0.077	0.184(0.495)	0.528
$\phi_{21}(\cdot)$	0.016(0.061)	0.063	0.012(0.059)	0.060	0.016(0.061)	0.063	0.127(0.316)	0.341
$\phi_{22}(\cdot)$	0.008(0.090)	0.090	0.007(0.087)	0.087	0.008(0.089)	0.090	0.456(0.819)	0.937
$\phi_{31}(\cdot)$	0.024(0.093)	0.096	0.018(0.055)	0.058	0.018(0.060)	0.063	0.068(0.197)	0.208
$\phi_{32}(\cdot)$	0.007(0.046)	0.046	0.004(0.011)	0.011	0.006(0.041)	0.041	0.105(0.113)	0.154
$\Sigma_1(\cdot, \cdot)$	0.027(0.174)	0.176	0.022(0.154)	0.156	0.024(0.161)	0.163	0.996(0.092)	1.000
$\Sigma_2(\cdot, \cdot)$	0.043(0.172)	0.178	0.037(0.156)	0.160	0.041(0.159)	0.164	1.048(0.145)	1.058
$\Sigma_3(\cdot, \cdot)$	0.048(0.177)	0.184	0.031(0.135)	0.139	0.046(0.151)	0.158	1.582(0.067)	1.583
#cluster	0.050(0.219)	0.224	-	-	0.003(0.208)	0.209	-	-

“-” not available

Table 7: Results for Simulation 2.

	Proposed(C=7)		CT(C=3)		CT(C=7)		WoT(C=3)	
	bias(sd)	RMSE	bias(sd)	RMSE	bias(sd)	RMSE	bias(sd)	RMSE
π_1	0.024(0.091)	0.094	0.017(0.080)	0.081	0.022(0.084)	0.087	0.024(0.400)	0.401
π_2	0.032(0.090)	0.096	0.013(0.064)	0.066	0.013(0.068)	0.069	0.077(0.391)	0.399
π_3	0.027(0.116)	0.119	0.017(0.109)	0.111	0.026(0.114)	0.117	0.053(0.410)	0.413
σ_1^2	0.004(0.017)	0.018	0.003(0.013)	0.013	0.004(0.018)	0.018	0.020(0.003)	0.020
σ_2^2	0.016(0.090)	0.092	0.014(0.085)	0.086	0.014(0.085)	0.086	0.150(0.003)	0.150
σ_3^2	0.002(0.085)	0.085	0.001(0.078)	0.078	0.001(0.082)	0.082	0.237(0.003)	0.237
λ_{11}	0.017(0.087)	0.089	0.009(0.074)	0.074	0.015(0.083)	0.085	0.198(0.011)	0.198
λ_{31}	0.006(0.041)	0.041	0.003(0.029)	0.029	0.005(0.029)	0.030	0.137(0.013)	0.138
λ_{32}	0.008(0.038)	0.038	0.004(0.021)	0.021	0.007(0.032)	0.033	0.050(0.012)	0.052
$\mu_1(\cdot)$	0.003(0.033)	0.033	0.003(0.014)	0.014	0.003(0.014)	0.014	1.024(0.002)	1.024
$\mu_2(\cdot)$	0.004(0.023)	0.024	0.002(0.018)	0.018	0.002(0.019)	0.019	0.920(0.002)	0.920
$\mu_3(\cdot)$	0.002(0.042)	0.042	0.001(0.038)	0.038	0.001(0.042)	0.042	1.027(0.002)	1.027
$\phi_{11}(\cdot)$	0.006(0.038)	0.039	0.006(0.030)	0.031	0.006(0.031)	0.032	0.705(0.657)	0.964
$\phi_{31}(\cdot)$	0.017(0.069)	0.071	0.013(0.061)	0.062	0.016(0.062)	0.064	0.181(0.550)	0.579
$\phi_{32}(\cdot)$	0.005(0.037)	0.037	0.004(0.032)	0.032	0.004(0.032)	0.032	0.171(0.031)	0.174
$\Sigma_1(\cdot, \cdot)$	0.019(0.125)	0.127	0.012(0.095)	0.096	0.017(0.118)	0.119	0.929(0.819)	1.239
$\Sigma_2(\cdot, \cdot)$	0.016(0.090)	0.092	0.014(0.085)	0.086	0.014(0.085)	0.086	0.150(0.003)	0.150
$\Sigma_3(\cdot, \cdot)$	0.032(0.130)	0.134	0.029(0.088)	0.093	0.032(0.128)	0.132	0.982(0.707)	1.210
#cluster	0.021(0.185)	0.186	-	-	0.007(0.144)	0.144	-	-

“-” not available

Table 8: The proposed estimators under gamma distributions for Simulation 3

τ	normal		100		10		5		1	
Skewness	0		0.2		0.63		0.89		2	
Excess kurtosis	0		0.06		0.6		1.2		6	
	bias(sd)	RMSE	bias(sd)	RMSE	bias(sd)	RMSE	bias(sd)	RMSE	bias(sd)	RMSE
π_1	0.003(0.034)	0.034	0.004(0.034)	0.035	0.006(0.034)	0.035	0.007(0.034)	0.035	0.009(0.034)	0.035
π_2	0.009(0.039)	0.039	0.009(0.039)	0.039	0.009(0.040)	0.041	0.014(0.039)	0.041	0.030(0.053)	0.061
π_3	0.004(0.038)	0.038	0.006(0.038)	0.038	0.011(0.039)	0.041	0.017(0.040)	0.043	0.039(0.054)	0.067
σ_1^2	0.001(0.009)	0.009	0.001(0.009)	0.009	0.002(0.010)	0.010	0.001(0.010)	0.010	0.006(0.012)	0.013
σ_2^2	0.001(0.014)	0.014	0.004(0.019)	0.019	0.006(0.023)	0.024	0.008(0.024)	0.025	0.018(0.034)	0.038
σ_3^2	0.006(0.030)	0.031	0.007(0.032)	0.032	0.009(0.033)	0.034	0.009(0.035)	0.036	0.021(0.039)	0.044
λ_{11}	0.030(0.145)	0.148	0.040(0.164)	0.169	0.050(0.189)	0.195	0.063(0.204)	0.213	0.066(0.234)	0.243
λ_{12}	0.011(0.036)	0.038	0.012(0.041)	0.043	0.014(0.054)	0.055	0.016(0.056)	0.058	0.016(0.058)	0.061
λ_{21}	0.037(0.140)	0.145	0.043(0.177)	0.182	0.051(0.176)	0.183	0.073(0.187)	0.200	0.069(0.225)	0.235
λ_{22}	0.001(0.011)	0.011	0.006(0.035)	0.036	0.012(0.037)	0.038	0.015(0.039)	0.042	0.039(0.047)	0.062
λ_{31}	0.038(0.149)	0.154	0.041(0.170)	0.175	0.050(0.173)	0.180	0.067(0.178)	0.190	0.069(0.177)	0.190
λ_{32}	0.017(0.056)	0.058	0.021(0.057)	0.061	0.024(0.058)	0.063	0.035(0.061)	0.071	0.036(0.073)	0.081
$\mu_1(\cdot)$	0.020(0.130)	0.132	0.026(0.159)	0.161	0.029(0.165)	0.168	0.031(0.186)	0.189	0.035(0.222)	0.225
$\mu_2(\cdot)$	0.033(0.180)	0.183	0.040(0.230)	0.233	0.042(0.256)	0.259	0.047(0.272)	0.276	0.053(0.306)	0.310
$\mu_3(\cdot)$	0.049(0.184)	0.191	0.057(0.257)	0.263	0.059(0.251)	0.258	0.061(0.279)	0.285	0.071(0.293)	0.302
$\phi_{11}(\cdot)$	0.020(0.070)	0.073	0.022(0.088)	0.090	0.021(0.114)	0.116	0.020(0.138)	0.140	0.021(0.158)	0.159
$\phi_{12}(\cdot)$	0.013(0.083)	0.084	0.015(0.100)	0.101	0.018(0.135)	0.136	0.023(0.144)	0.146	0.023(0.165)	0.167
$\phi_{21}(\cdot)$	0.037(0.109)	0.115	0.040(0.107)	0.114	0.042(0.120)	0.127	0.045(0.160)	0.166	0.044(0.164)	0.170
$\phi_{22}(\cdot)$	0.014(0.115)	0.116	0.021(0.141)	0.142	0.029(0.152)	0.155	0.034(0.192)	0.195	0.082(0.308)	0.319
$\phi_{31}(\cdot)$	0.035(0.106)	0.112	0.042(0.118)	0.125	0.081(0.155)	0.175	0.085(0.173)	0.193	0.095(0.177)	0.201
$\phi_{32}(\cdot)$	0.023(0.107)	0.110	0.030(0.108)	0.112	0.034(0.117)	0.122	0.081(0.118)	0.144	0.091(0.134)	0.162
$\Sigma_1(\cdot, \cdot)$	0.061(0.267)	0.274	0.070(0.268)	0.277	0.082(0.276)	0.288	0.089(0.285)	0.298	0.095(0.461)	0.471
$\Sigma_2(\cdot, \cdot)$	0.072(0.271)	0.280	0.081(0.304)	0.314	0.092(0.297)	0.311	0.098(0.313)	0.328	0.101(0.398)	0.410
$\Sigma_3(\cdot, \cdot)$	0.073(0.293)	0.302	0.101(0.311)	0.327	0.114(0.353)	0.371	0.133(0.366)	0.389	0.142(0.372)	0.398

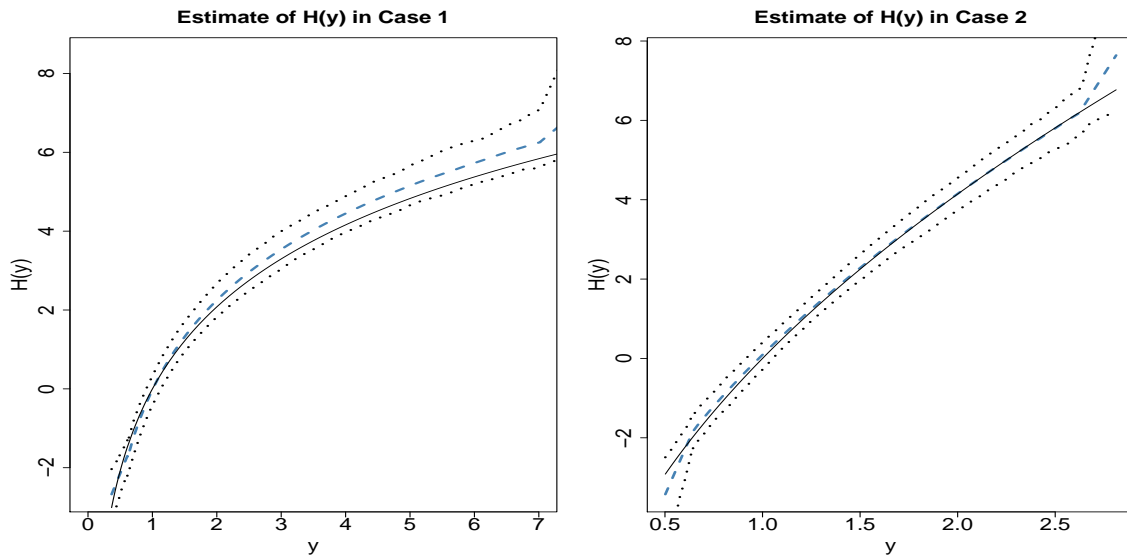


Figure 3: Estimated transformation functions for Cases 1 and 2 of Simulation 1 (solid—true function; dotted—95% confidence limits; dashed—average of the estimated function). This figure appears in color in the electronic version of this article, and any mention of color refers to that version.

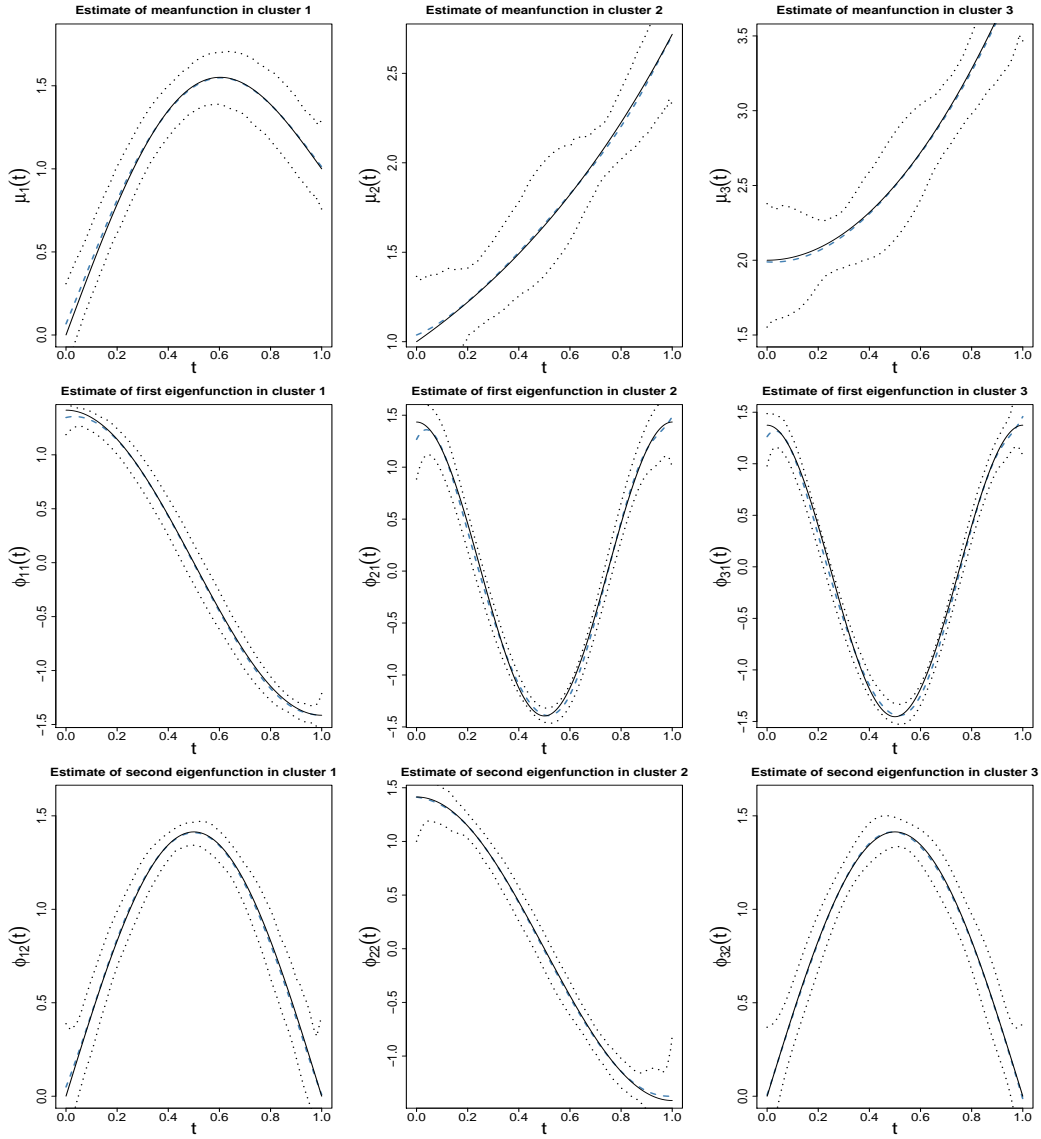


Figure 4: Estimated mean functions (top) and eigenfunctions (median and bottom) for Case 1 of Simulation 1 (solid–true function; dotted–95% confidence limits; dashed–average of the estimated function). This figure appears in color in the electronic version of this article, and any mention of color refers to that version.

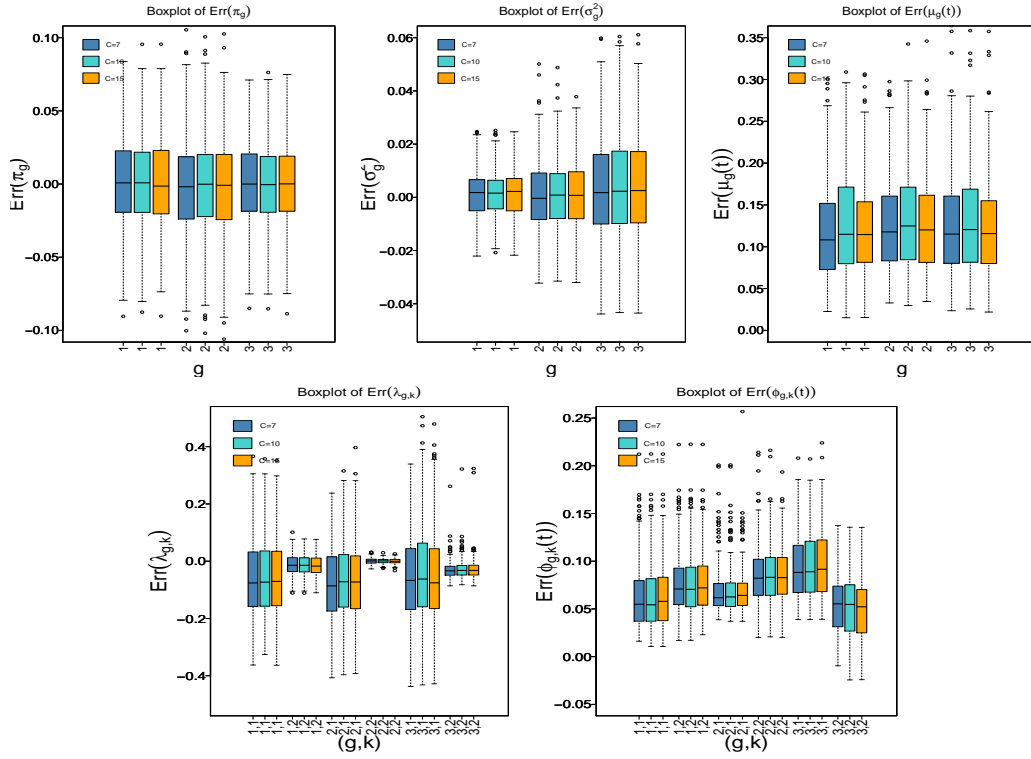


Figure 5: Comparison of the estimation error with different initial numbers of clusters, $C = 7, 10, 15$, from left to right for Case 1 of Simulation 1. This figure appears in color in the electronic version of this article, and any mention of color refers to that version.

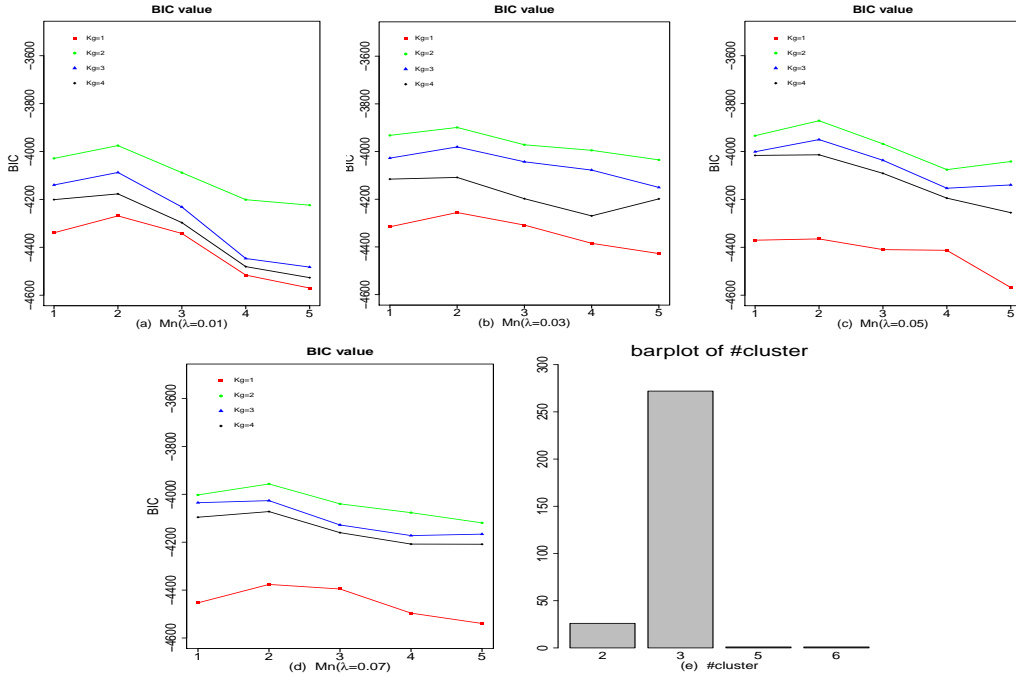


Figure 6: (a)- (d) The BIC values under various K_g , M_n and λ ; (e) The barplot of estimated number of clusters based on the BIC criterion for Case 1 of Simulation 1. This figure appears in color in the electronic version of this article, and any mention of color refers to that version.

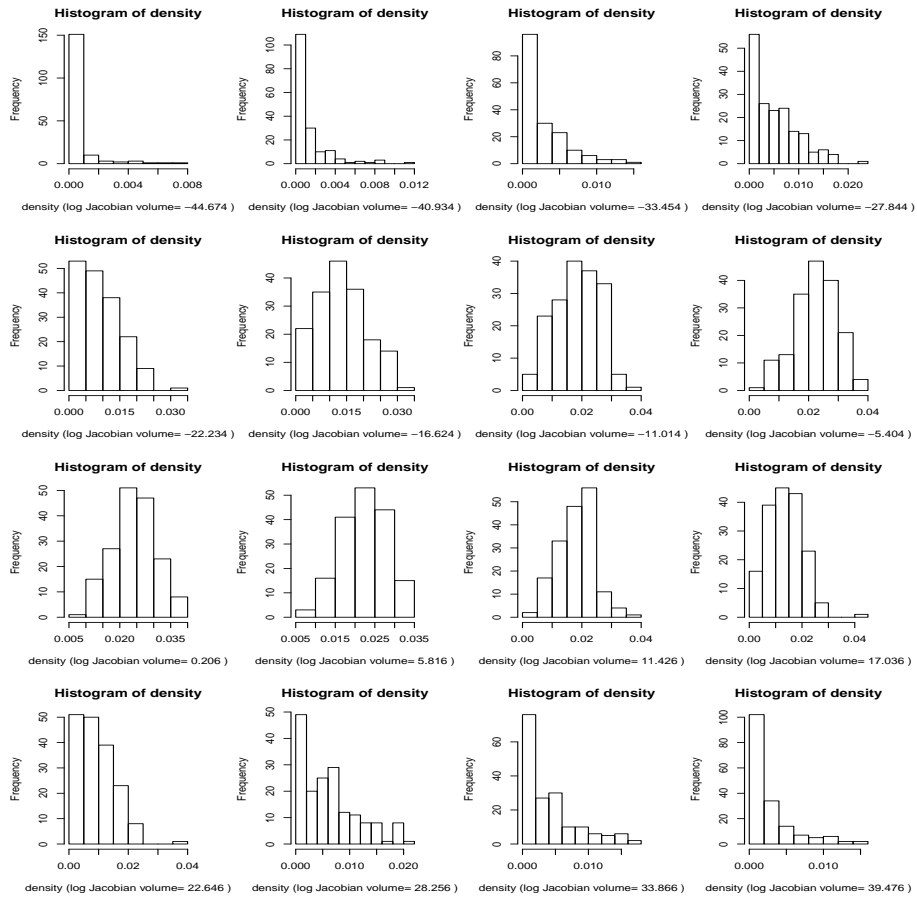


Figure 7: Histogram of density for cluster 1.

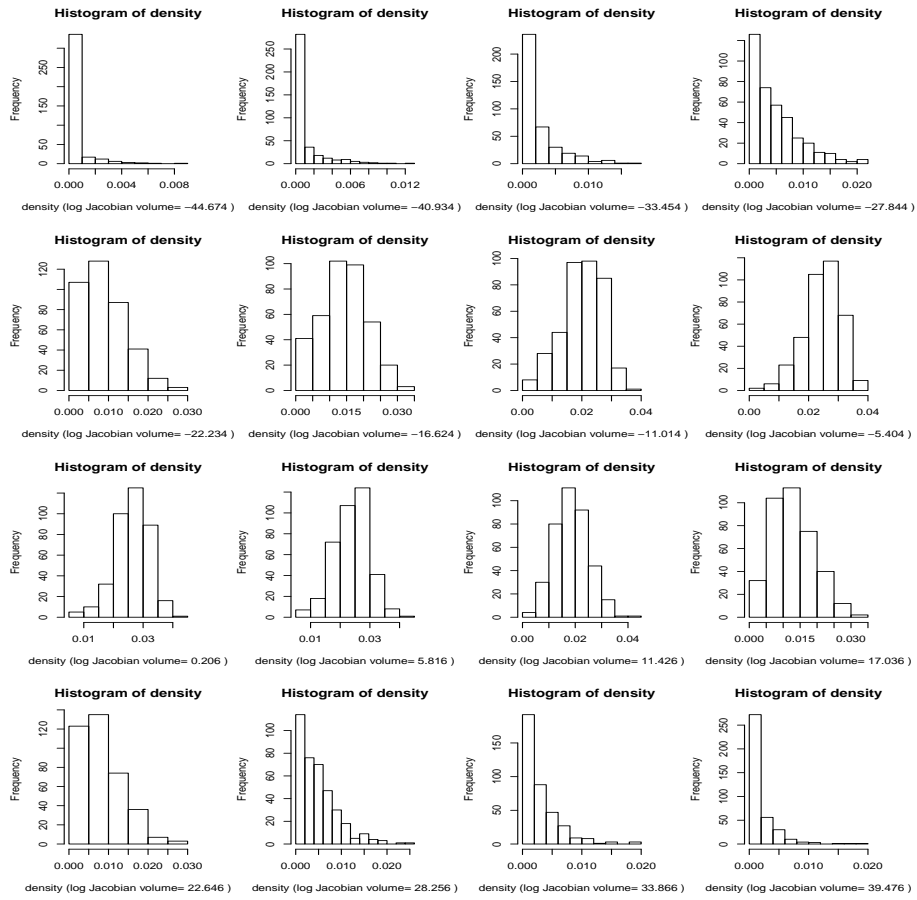


Figure 8: Histogram of density for cluster 2.

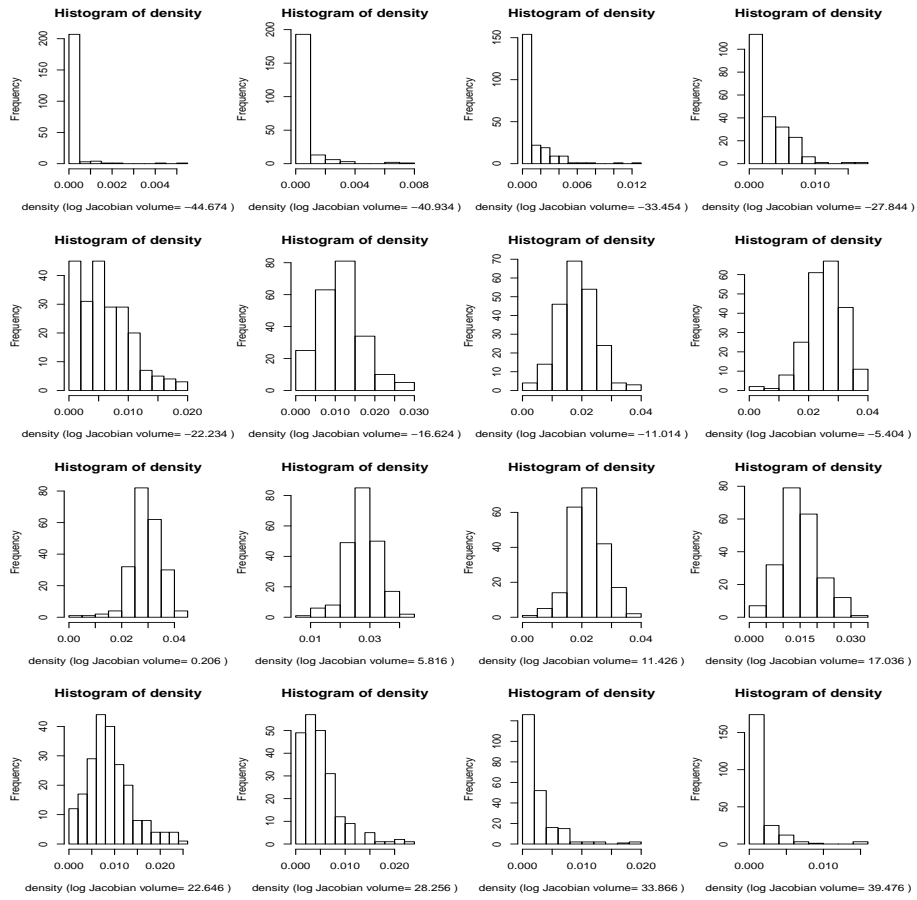


Figure 9: Histogram of density for cluster 3.

References

- Chen, X., Fan, Y., and Tsyrennikov, V. (2006). Efficient estimation of semiparametric multivariate copula models. *Journal of the American Statistical Association*, **101**, 1228-1240.
- Chandra, T. K. (2012). The Borel-Cantelli Lemma. Springer Science & Business Media.
- Flett, T. M. (1958). A mean value theorem. *The Mathematical Gazette*, **42**, 38-39.
- Huang, T., Peng, H., and Zhang, K.(2017). Model selection for gaussian mixture models. *Statistica Sinica*, **27**, 147-169.
- Pollard, D. (1984). Convergence of Stochastic Processes. Springer, New York.
- Schumaker, L. L. Spline functions: basic theory. 1981. John Wiley & Sons, New York.
- Shen, X., and Wong, W. H. (1994). Convergence rate of sieve estimates. *The Annals of Statistics*, **22**, 580-615.
- Van de Geer, S. (2000). Empirical Processes in M-estimation. Cambridge University Press.
- Van der Vaart, A. W. (1998). Asymptotic statistics (Vol. 3). Cambridge university press.
- Van der Varrrt, A. W. and Wellner, J. A. (1996). Weak Convergence and Empirical Processes. Springer, New York.