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**A NOTE ON LATENT SEGMENTATION MODELS**

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# A Note On Latent Segmentation Models

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## Summary

Developing polytomous models of choice which could enable in identifying market segments has been the focus of attention in many papers in recent literature (Ramaswamy et al., 1996 and the references therein).

In this note, we show that latent segmentation models of the type developed by Ramaswamy et al. is *always* nonidentifiable, and hence the parameters of interest are nonestimable.

The joint segmentation model of Ramaswamy et al. (1996) considers data from consumers on two sets of categorical variables that serve as distinct bases for segmentation and are interdependent (specifically, benefits sought and services used). It is apparent that the data collected provides no information concerning the allocation of consumers to particular segments (indexed by  $(i, j)$ ). Therefore, it is readily apparent that it is impossible to make inference (i.e., estimate) the parameter  $\theta_{ij}$ , which is the joint probability of a consumer belonging to segment  $i$  and segment  $j$ . Consequently, the model they construct is nonidentifiable.

To illustrate this point we consider an analagous model: let observations from  $N$  individuals be categorised from the finite set  $\{1, \dots, m\}$ , and let  $x_{nk} = 1$  if the  $n$ th individual is in category  $k$  and zero otherwise, and let  $\pi_k$  be the probability that an individual is in category  $k$ . Now let  $\theta_i$  represent the probability of a consumer being in segment  $i$  and  $\pi_{ik}$  the probability of an individual in segment  $i$  being in category  $k$ . Obviously we have  $\sum_k \pi_{ik} = 1$ ,  $\sum_i \theta_i = 1$  and  $\sum_k \pi_k = 1$ . The likelihood is given by

$$\prod_n \left( \sum_i \theta_i \left[ \prod_{k=1}^m \pi_{ik}^{x_{nk}} \right] \right).$$

This likelihood expression is essentially a special version of the type of likelihood appearing in equation (2) of Ramaswamy et al. (1996). Clearly our likelihood can be written as

$$\prod_n \left( \sum_i \theta_i \pi_{ik_n} \right),$$

where  $k_n$  is the category of the  $n$ th individual. This becomes

$$\prod_{k=1}^m \left( \left[ \sum_i \theta_i \pi_{ik} \right]^{n_k} \right),$$

where  $n_k$  is the number of individuals located in the  $k$ th category. Finally, our likelihood becomes

$$\prod_{k=1}^m \pi_k^{n_k},$$

which is precisely what we should have started with in the first place. It is obvious the  $\theta_i$  and  $\pi_{ik}$  are nonestimable. An exact argument to the one

above can now be applied to the likelihood appearing in Ramaswamy et al. (1996, equation (2)). The details are now given.

Let  $L$  denote equation (2) from Ramaswamy et al. (1996), i.e. the likelihood function. We can write this likelihood in a different way:

$$L = \prod_{n=1}^N \sum_{ij} \theta_{ij} \left( \prod_{k=1}^K \gamma_{ikl_{kn}} \prod_{m=1}^M \delta_{jml_{mn}} \right),$$

where, for example,  $ikl_{kn}$  denotes that the  $n$ th individual is in the  $l$ th category, ( $l = 1, \dots, L_k$ ), for the  $k$ th variable, ( $k = 1, \dots, K$ ), in segment  $i$ ; and so,  $\gamma_{ikl_{kn}}$  denotes the probability of this event. Let  $\mathcal{C}$  denote the set of possible categories, and  $n_c$  denote the number of individuals who are in  $c \in \mathcal{C}$ . Then, we have

$$L = \prod_{c \in \mathcal{C}} \left( \sum_{ij} \theta_{ij} \pi_{ijc} \right)^{n_c},$$

where  $\pi_{ijc}$  is the probability of being in category  $c$  conditional on being in segments  $i$  and  $j$ . This last expression is identical to

$$\prod_{c \in \mathcal{C}} \pi_c^{n_c},$$

which, again, is the *correct* likelihood: so, clearly, this implies that  $\theta_{ij}$  is nonidentifiable.

The models discussed above should not be confused with *continuous* mixture models:

$$f(y; \theta, \phi) = \sum_s \theta_s f(y; \phi_s),$$

where  $\sum_s \theta_s = 1$ . Each  $(\theta, \phi)$  defines a *unique* density and hence such models are identifiable (in general). The *discrete* categorical model is given by:

$$p(y = k; \theta, \pi) = \sum_s \theta_s p(y = k|s).$$

Let  $\pi_{sk} = p(y = k|s)$ . Clearly  $\sum_s \theta_s \pi_{sk} = \tilde{\pi}_k$ , where  $\tilde{\pi}_k = p(y = k; \tilde{\pi})$ . Note, then, that for all  $(\theta, \pi)$ ,

$$p(y = k; \theta, \pi) = p(y = k; \tilde{\pi}).$$

Hence, the model is nonidentifiable for  $\theta$ ; i.e., many  $(\theta, \pi)$ , in fact an infinite number, lead to the same probability model for the data.

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### REFERENCE

Ramaswamy, V., Chatterjee, R. and Cohen, S.H. (1996). Joint segmentation on distinct interdependent bases with categorical data. *Journal of Marketing Research*, 33, 337-350.