

***Labor Hoarding In Russia:  
Where Does It Come From?***

*By: Rouslan Koumakhov and Boris Najman*

William Davidson Working Paper Number 394  
October 2001

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Rouslan Koumakhov, Boris Najman<sup>1</sup>

Rouslan Koumakhov  
Forum, Université Paris X-Nanterre, bat. K  
200, avenue de la République, 92001 Nanterre Cedex  
Tel.: 331 53 61 21 82  
fax: 331 40 97 71 83

Boris Najman Delta-ENS  
48, boul. Jourdan 75014 Paris  
Tel.: 331 42 19 07 91  
fax: 331 43 13 63 10

**June, 2000**

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<sup>1</sup> Suggestions and comments are welcome: [najman@delta.ens.fr](mailto:najman@delta.ens.fr), [rouslan.koumakhov@u-paris10.fr](mailto:rouslan.koumakhov@u-paris10.fr)  
We want to thank Mathilde Maurel, Georges de Ménénil, Jean-François Nivet and Dominique Redor for their extensive help and devotion in this research. A preliminary version of the paper was presented at the Econometric Society world congress in Seattle in August 2000.  
University of Nanterre Paris X; Delta, Ecoles des Hautes Etudes en Sciences Sociales. This paper is financed through CNRS (ROSES University of Paris I, MSE, 106-112, bd. de l'Hôpital 75647 Paris cedex 13).

## Abstract

The paper focuses on the labor "hoarding" problem in Russian. We studied two forms of "hoarding": unpaid leaves and short-time work. Our research is based on the *Russian Longitudinal Monitoring Survey* (RLMS) database. The paper exploits individual panel data between 1994 and 1996.

We show that unpaid leaves and short-time work do not represent a form of hidden unemployment. Both types of labor "hoarding" reflect the nature of employees' professional competencies. First, unpaid leaves concern primarily the employees with firm-specific knowledge, while short-time work affects strongly unskilled workers. Second, external mobility is mostly related to young people and unskilled blue-collar workers while employees with specific competencies do not change jobs so much.

The paper insists on significant internal adjustments which are taking place through unpaid leaves and short-time work. This explains why there has been no massive unemployment in Russia until now.

In conclusion, Russian labor market is characterized rather by internal flexibility than by labor "hoarding".

Key words: labor market, internal adjustments, flexibility, Russia, skills.

JEL classification: J2, C23, P23

## Introduction

One of the surprises of the Post-Soviet transformation is the relationship between output and labor: in the context of dramatic output fall in 1991-1997 there was still no open mass unemployment and no large decrease in employment<sup>1</sup> in Russia. To explain this puzzling result, different Russian and Western economists and sociologists draw our attention to the enterprise level, pointing out, in particular, "labor surplus", or "labor hoarding" as "guilty" practices which prevent reallocation and restructuring process in Post-Soviet economy. Two forms of labor "retaining", i.e. "administrative" (forced unpaid or partially paid) leave and short-time (reduced hours) work, are at the center of much of economic studies.

Two main problems arise from this reasoning.

First, the labor hoarding hypothesis seems to contradict the evidence of the high labor mobility in the enterprise data (Gimpelson et Lippoldt 1997) which is also stated in various survey studies (Smirnov 1998). The intensive labor rotation, which in itself points at the reallocation process, rejects the overhang idea: not only managers "hoard" labor by limiting employment reductions, but they try to replace separated workers by "outsiders".

Second, and more generally, we do not know why Russian enterprises should hoard labor.

Commander et al. (1995) insist on the social functions of the Russian firms: "And the decision to enforce severance may further be complicated because in a large number of cases firms are the source of a broad range of benefits and income supports for workers. Loss of employment likely implies loss of access to such benefits, frequently with no clear alternative suppliers" (p. 157-158). Layard and Richter (1994) evoke "paternalistic" behavior as one of the reasons for the job keeping.

This argument does not seem to be satisfactory.

First, social support cannot be considered as a reason of labor retaining, since we still do not understand the reasons of "paternalistic" commitment. Notice that the social support argument was advanced from the viewpoint of employees, but not from the viewpoint of firms. Why might managers keep social functions in the face of the drastic and continuous shocks to output? In other terms, explaining the patterns of "abnormal" behavior by another "abnormal" pattern which in turn has to be explained does not add a lot to our understanding.

Second, since hiring rates remain rather high managers cannot be committed only to "their" employees but to all workers they engage regularly from external labor market which does not seem very realistic.

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<sup>1</sup> Between 1994 and 1996, labor macro data shows a decline of employment by 1.9 millions (3.1 % of the total employment), and similarly the unemployment (ILO definition) increased by about 1 million people (1.8 % of the labor force), (Russian Economic Trends, 1999).

Notice, that the "technological" solution proposed by Commander et al. (1995) can neither be considered as satisfactory: if technological rigidities restrain the amount of labor adjustment, employee-retaining is no more justified by "overhang" or "hoarding" <sup>2</sup>.

In this paper, we analyze two forms of labor "retaining": "administrative" leave and reduced hours. Focusing our attention on these issues is justified by two main reasons.

First, both forms are often treated, though not empirically analyzed, as an evidence of hidden unemployment. For instance, Commander et al. (1995) characterize Goskomstat figures and the World Bank survey data on involuntary leave or short-time work as "a crude measure of the current employment overhang": "we can see that releasing the marginally employed into unemployment would have raised the unemployment rate by at least 300 percent" (p. 157). In the same vein, Standing (1994, 1996, 1998) who considers reduced hours and unpaid leaves as hiding the "real" unemployment state, argues that employees put on "administrative" leave have "minimal or no prospect of being recalled" (Standing 1994, p. 37)<sup>3</sup>.

Second, the choice of issues as unpaid leaves and short-time work is linked to the data we have utilized.

The data source in this paper is a survey of individuals and households (Russia Longitudinal Monitoring Survey, RLMS)<sup>4</sup>. We used and constructed variables on individuals and households (see appendix 4 and 5 for the definitions and measurement of variables) and enterprise characteristics, while focusing on the 1994-1996 sample in order to follow the same individuals all over the period. We studied people between 15 and 72 years old, the sample is composed 4607 persons.

One may say that the dataset which refers to individuals is not relevant to consider the issues of labor hoarding as, contrary to firm surveys, it does not contain information on enterprises.

We can counter-argue that the RLMS survey allows us to identify clearly a large set of variables related to the individual behavior within the enterprise and reflecting the firm policies towards their employees. This is, in particular, the case of unpaid leave and short-time work<sup>5</sup> that can be directly measured from the RLMS database. After all, we have no evidence that information reported by employees on this issues is less valuable than data given by managers in the firm surveys.

Moreover, data provided by individuals could be more precise in some cases. For example, a lot of employment practices are not formally documented within a firm and/or applied on the workshop or team level. Therefore, managers (and especially top-managers,

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<sup>2</sup> See Fay and Medoff (1985) for the general argument and Clarke (1998) for the Russian case.

<sup>3</sup> Linz (1988) suggests also that unpaid leaves could be considered as a "signal of management's commitment to maintaining employment levels" (p. 15) without giving any explanation for the "job rights" keeping.

<sup>4</sup> This survey stands as part of an international research project conducted jointly by the University of Chapel Hill (USA), the Institute of Sociology, Russian Academy of Sciences and the Russian Institute of Nutrition.

<sup>5</sup> We have measured short-time work as the time spent on primary work under 120 hours per month. As for unpaid leave we used the results of explicit question on this subject.

i.e. those who respond usually to the firm survey) may count only "official" leaves or working-hours' reductions, thus seriously underestimating the extent of a real labor "retaining". As for an individual, he faces directly these practices, which allows measuring them more accurately in the corresponding questions.

Finally, a firm survey usually do not contain any information on the individuals after they quitted enterprise, it can only report data on hirings and separations. Thus, such issues as unemployment or job-to-job movements, crucial for our paper are ignored. On the contrary, the RLMS database allows us to highlight these issues, for it follows the same individuals inside and outside the firms.

In fact, from 1994 to 1996 Russian labor market face a large reallocation of workers through short-time work and unpaid leaves (see table 1). The survey data shows an important reduction in working time. For instance the number of days –about 50 days– on unpaid leaves correspond to a 20 % reduction in the total working days over a year<sup>6</sup>.

Table 1: Employment enterprise policies (in % of employed population)

	% of people on short-time work	% of people on unpaid leaves this year	Average number of days on unpaid leaves*
1994	20.1	12.3	39
1995	18.9	6.6	58
1996	18.1	8.1	49

*Sources: RLMS, rounds V, VI, VII*

\* On average per person on "Unpaid leaves".

The structure of this paper is as follows.

We start by testing the "hidden unemployment" hypothesis: do "administrative" leave and short-time work represent preliminary states of unemployment? (Section I). We analyze then the determinants of these two practices (Section II) and the relationship between job-to-job mobility and labor-retaining (Section III). Finally we discuss our findings and propose corresponding conclusions (Sections IV, V).

<sup>6</sup> The total of working days in 1996 was about 250.

## I. Do "administrative" leave and reduced hours represent preliminary states of unemployment?

There can be two-fold relationship between the forms of labor "hoarding" and unemployment.

First, if workers put on unpaid leave or short-time work have "minimal or no project of being recalled", one should find a higher probability of becoming unemployed for these workers. In order to measure unemployment, we choose the data on self-reported unemployment. This definition is voluntary large in order into account all forms of unemployment. We did not want to underestimate the unemployment phenomenon. Second, the high risk of unemployment that both forms of labor "retaining" are supposed to "hide" should be reflected in corresponding subjective perception of the current job position. Such perception is an important characteristic of individual behavior with respect to unemployment. For instance, people that have few job opportunities are obviously more likely to accept various employment policies, in terms of work disciplining or labor adjustment. We measure this "subjective" characteristic using the RLMS data which contained two following questions<sup>7</sup>: "Are you afraid of losing your job?" and "How certain are you that you will be able to find a work, no worse than your present job?" We start by a general picture of the unemployment threat without any relation with the problem of labor hoarding. Table 2 gives corresponding percentages – of subjective perceptions and unemployment – for the total sample of individuals between 1994 and 1996 (on the basis of the RLMS data ). One may see that while the rate of unemployment (ILO definition) is rather low, especially in 1994-1995, the "subjective" perception of unemployment risk is high. People in general are very concerned about job losing and of not finding another one. This result suggests that at least a large part of employees could be subject to a permanent and strong pressure on behalf of managers who are able to impose them different "precarious" employment practices.

**Table 2:** employment expectations of the employees (in % of employed people) and unemployment (in % of the labor force, ILO definition)

	% of people very concerned about job losing	% of people absolutely uncertain to find a new job	% of people unemployed (ILO definition)
1994	41.0	37.8	6.4
1995	39.9	37.2	6.1
1996	41.2	38.9	7.9

Sources: RLMS, Rounds V, VI, VII

We study now actual and perceived risks of unemployment in relation with unpaid leaves and reduced hours. We examine the ordered *probit* (Greene, 1993, p.672-676) of "subjective" variables and the simple *probit* (Greene, 1993, p. 643-655) for unemployment controlling for all our individual, households and enterprise variables.

There are four main results emerging from regressions in Tables 3 and 4:

- (i) Employees who had experience with unpaid leaves are more afraid of losing their job and not to find a new job corresponding to their specialization, in case of eventual job loss.
- (ii) For the short-time experience, the "subjective" perception is quite on the opposite.
- (iii) There is no evidence that inflows in unemployment in 1996 are related to having been put on unpaid leaves in 1995.

<sup>7</sup> In the questionnaire, the evaluation of these "subjective" perceptions is marked from 1 to 5. In table 2, we present only the worse perceptions i.e. when the variables are equal to 5.

(iv) By contrast, there is a strong positive relation between short-time-working in 1995 and becoming unemployed in 1996.

Given these results, we can try to answer the main question of this section: can we state that both forms of employment practices, i.e. "administrative" leave and short-time work, represent preliminary states of unemployment? As for unpaid leaves, the answer is No: even if unpaid-leave-workers feel more vulnerable regarding their job opportunities, there is no proof that in reality these workers are facing higher risk of unemployment than other groups of workers. As for the practice of reduced hours, contrary to "administrative" leaves, the preliminary answer is Yes: short-time work may represent a preliminary stage of the unemployment.

We will now shadow these conclusions since the results of the regressions in Tables 3 and 4 provide us one more important finding. Both forms of employment practices studied here imply a kind of internal paradox. Workers who experience first form, unpaid leaves, are rather afraid of losing their job, but in reality few of them are likely to become unemployed. Workers who experience the second form, short-time work, are more likely to become unemployed but they are far more optimistic regarding their future job opportunities.

In the following considerations we try to propose a comprehensive understanding of this double paradox.

(i) Consider workers who experienced unpaid leaves. Their fear of unemployment provides a part of explanation why a large part of workers accept this "precarious" state of employment. Unpaid leaves could be, therefore, used as a means of pressure by managers. However, this instrument of disciplining is not likely to be utilized in actual managerial strategies. Otherwise, the outflows from "administrative" leaves into unemployment would be much more intensive. Hence, this form of employment policies, and *only* this form, i.e. unpaid leaves, is consistent with the generalization made by Commander et al (1995) : "In short, managers appear to have discretionary powers with respect to wages and employment, but they evidently choose not to exert them to enforce large employment separations or restructuring" (p. 178).

We suggest that the workers' "attachment" to the firm could be largely due to their firm-specific skills. On one hand, those who accept unpaid or partially paid leaves are afraid of not finding a workplace that conforms to their professional competencies outside the borders of their "enterprise of origin"<sup>8</sup>. On the other hand, firms in general are neither interested in releasing those employees whom they select for this practice and who accept it. It seems that managers are not sure, that they could easily replace specialized workers through the market, therefore, these workers are been kept within the firms. Thus, unpaid leaves translate *via* the role of firm-specific skills the process of mutual "attachment" between corresponding workers and enterprises.

We can, therefore, formulate the following hypothesis:

H1: the practice of putting employees on unpaid leaves is closely related to the firm-specific skills of these workers.

(ii) Consider short-time workers now. If unemployment risk appear to be a real threat for this group of workers in general, it is somewhat shadowed by the "subjective" perception of this risk: short-time workers are not afraid of losing job or of not finding another job. Notice, that in general, the determinants of "subjective" variables seem to correspond rather accurately to the reality<sup>9</sup>. The relative "optimism" of short-time workers suggests that at least a part of them may experience good opportunities on the labor market. This in itself does not

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<sup>8</sup> This seems to be a more coherent explanation of why employees stay within firms than a "personal identity" argument suggested by Layard and Richter (1994).

<sup>9</sup> Thus, the perception concerning losing or finding a job is fare better for people involved in informal activities. Is also better for male, for inhabitants urban settlements, and especially for the one from Moscow and really worse for those with low level of expenditures per capita.



seem surprisingly given the above-mentioned high inter-firm mobility in Russia. As Clarke (1998) pointed out, "the result of the very high rate of hiring is that redundancy and unemployment are not the threat that they would otherwise be" (p 45). This seems to indicate that short-time-workers behavior is strongly heterogeneous and that the heterogeneity is in accordance with employment opportunities.

We can, therefore, propose a second hypothesis:

H2: Short-time workers differ by professional segments which translate different outside job options.

In order to corroborate these hypotheses, we address two issues.

First, we need to study more deeply the characteristics of workers who were faced with administrative leaves or short-time work. In Section II we examine the determinants of both practices, controlling for professional qualifications which we consider as an indicator of skill specificity.

Second, if unemployment risk differs by professional segments of employees, it is necessary to analyze the likely relationship between job opportunities and labor segmentation. Section III focuses on the characteristics of the job-to-job mobility with respect to unpaid leaves and reduced hours. As before, the role of firm-specific skills is directly addressed.

**Table 3: Probability to be certain of losing job or not to be able to find one in 1995 (Ordered probit regression)**

Variables in 1995	How concerned are you that you might lose your job?		How certain are you that you will be able to find work?	
	Coef.	Standard error	Coef.	Standard error
Gender	-0.265**	0.060	-0.522**	0.060
Age	0.002	0.002	0.012**	0.002
Education (year)	-0.019	0.014	-0.015	0.014
Couple	0.037	0.063	0.013	0.064
Plots	-0.039	0.058	0.108*	0.059
Informal Activity	-0.205*	0.106	-0.260**	0.105
Manager	-0.406**	0.133	-0.517**	0.138
Intelligentsia	-0.260**	0.109	-0.471**	0.111
Engineer	0.016	0.134	-0.041	0.138
State servant	-0.156	0.156	-0.302*	0.157
Trader-financier	0.068	0.135	-0.292*	0.137
Technician	0.063	0.106	-0.305**	0.107
Workman	0.020	0.095	-0.207*	0.097
Farmer	0.051	0.133	-0.107	0.137
Elementary job	-0.066	0.108	-0.258**	0.109
In the forces	-0.079	0.251	0.163	0.256
Driver	-0.108	0.128	-0.155	0.130
Foreign	-0.108	0.147	-0.240	0.149
Public	-0.124*	0.070	-0.051	0.070
Private	-0.162*	0.082	-0.102	0.083
# Workers<100	-0.062	0.073	-0.184**	0.074
# Workers>500	-0.003	0.071	-0.174**	0.072
Duration. arrears	-0.006	0.010	0.023*	0.011
Am. Arrears	-0.001	0.003	-0.003	0.003
Unpaid leaves	0.361**	0.104	0.367**	0.104
Short-time work	-0.168**	0.063	-0.152**	0.064
In kind remuneration	-0.041	0.090	0.082	0.091
# of child in the household.	0.033	0.028	-0.024	0.028
# of elderly in the household	-0.016	0.044	0.044	0.044
Low hh. expenditure per capita	0.227**	0.055	0.241**	0.055
Medium hh. expenditure per capita	-0.039	0.069	0.096	0.069
Moscow-St Petersburg	-0.378**	0.116	-0.415**	0.118
Northwest	-0.134	0.120	-0.111	0.122
Center	-0.096	0.095	-0.043	0.096
Volga	0.082	0.095	-0.028	0.096
Caucasus	-0.110	0.106	-0.220*	0.108
Urals	-0.005	0.098	-0.061	0.099
Or. Siberia	0.255*	0.115	0.039	0.115
Semi-urban	0.458**	0.103	0.474**	0.102
Rural	0.214**	0.071	0.350**	0.072
$\mu(1)$	-1.430**	0.250	-1.452**	0.253
$\mu(2)$	-0.860**	0.249	-0.914**	0.251
$\mu(3)$	-0.583**	0.249	-0.515*	0.251
$\mu(4)$	0.009	0.249	0.223	0.251
Log Likelihood		-3151		-3059
Number of Observations		2190		2167
Chi <sup>2</sup> (39)		200.9		361.6
Prob. >chi <sup>2</sup>		0.0000		0.0000
Pseudo R <sup>2</sup>		0.03		0.056

**Notes:** *Dependent variable**Sources:* Panel data from RLMS, Rounds V, VI, VII

y=1 if the person is not concerned at all by losing job in 1995  
y=2 if the person is not very concerned by losing job in 1995  
y=3 if the person is both yes and no  
y=4 if the person is a little concerned by losing job in 1995  
y=5 if the person is very concerned by losing job in 1995

y=1 if the person is absolutely certain to find in 1995  
y=2 if the person is fairly certain to find in 1995  
y=3 if the person is both yes and no to find in 1995  
y=4 if the person is fairly uncertain to find in 1995  
y=5 if the person is absolutely uncertain to find in 1995

*Legend:* \* = statistically significant at the 5% level

\*\* = statistically significant at the 1% level

*Reference variables:* Clerk; Semi-public, # of employees between 100 and 500, High household expenditures per capita, Occidental Siberia and Urban

**Table 4: Probability to be unemployed (Self reported) in 1996 (probit regression)**

<b>Variables in 1996/1995</b>	<b>Coef.</b>	<b>Standard error</b>
Gender	0.133	0.138
Age	-0.029**	0.006
Education (year)	-0.079**	0.034
Couple	0.027	0.148
Plots	-0.021	0.126
Informal Activity	0.765**	0.147
Manager in 95	0.416	0.310
Intelligentsia in 95	-0.810*	0.448
Engineer in 95	0.439	0.326
State servant in 95	-0.767	0.708
Trader-financier in 95	0.480*	0.271
Technician in 95	-0.010	0.268
Workman in 95	0.250	0.228
Farmer in 95	0.579*	0.270
Elementary job in 95	-0.089	0.263
In the forces in 95	0.529	0.490
Driver in 95	0.368	0.269
Foreign in 95	0.309	0.351
Public in 95	0.317*	0.180
Private in 95	0.511**	0.193
# Workers<100 in 95	0.099	0.177
# Workers>500 in 95	0.004	0.175
Duration. arrears in 95	0.018	0.018
Am. arrears in 95	0.009	0.005
Unpaid leaves in 95	-0.337	0.259
Short-time work in 95	0.413**	0.134
In kind remuneration	-0.240	0.197
# of child in the household	0.036	0.060
# of elderly in the household.	-0.057	0.099
Low hh expenditure per capita	0.263*	0.141
Medium hh expenditure per capita	0.166	0.163
Moscow-St Petersburg	-0.269	0.312
Northwest	-0.355	0.317
Center	0.215	0.214
Volga	0.005	0.216
Caucasus	0.270	0.230
Urals	0.095	0.222
Or. Siberia	-0.019	0.253
Semi-urban	-0.186	0.254
Rural	0.114	0.149
Constant	-0.784	0.599
Log Likelihood	-338.7	
Number of Observations	2206	
Chi <sup>2</sup> (39)	160.9	
Prob. >chi <sup>2</sup>	0.0000	
Pseudo R <sup>2</sup>	0.19	

**Notes:****Dependent variable:**

y=1 if the person is unemployed in 1996 (self reported)

y=0 if not

Sources: Panel data from RLMS, Rounds V, VI, VII

Legend:

\* = statistically significant at the 5% level

\*\* = statistically significant at the 1% level

**Reference variables:** Clerk; Semi-public, # of employees between 100 and 500, High household expenditures per capita, Occidental Siberia and Urban Settlement.

## II. Determinants of "administrative" leave and short-time work.

### II.1. Detail analysis of the results.

We start this Section by analyzing the probability to be put on unpaid leaves and on short-time work controlling for all individual, household and enterprise variables. For this purpose we use simple *probit* model (Greene, 1993, p.646-655).

Table 5 provides regression results. First of all, we found out some common determinants of both employment policies: the probability to be put on short-time work or unpaid leaves is higher for female worker and for those involved in informal activities.

Other basic results are the following.

(i) Duration of wage arrears has a positive impact on the probability of being concerned by the "administrative" leave during the current year and no impact on the probability of being on short-time work.

(ii) Education influence is negative for "administrative" leave and has no impact on short-time work.

(iii) Young age increases the probability to have a reduced-hours work and does not play any role for the "administrative" leaves.

(iv) Particular professions increase or decrease both probabilities. For instance, elementary job – without any qualification – and belonging to intelligentsia increase the probability of short-time work while being a farmer plays in the opposite sense. Other professions, workmen and engineers, increase the probability to be on "administrative" leaves.

(v) Employment in small firms is negatively correlated with "administrative" leaves.

**Table 5: Probability to be on Short-time work job or on Unpaid leaves in 1996 (probit regression)**

Variables in 1996	Short-time work Job		Unpaid leaves	
	Coef.	Standard error	Coef.	Standard error
Gender	-0.212**	0.085	-0.355**	0.105
Age	-0.007*	0.003	0.001	0.004
Education (year)	-0.011	0.020	-0.058*	0.026
Couple	0.036	0.088	0.120	0.113
Plots	0.187*	0.087	0.201*	0.103
Informal Activity	0.249*	0.138	0.205	0.161
Manager	-0.382	0.336	-0.079	0.460
Intelligentsia	0.418**	0.150	-0.039	0.212
Engineer	0.100	0.200	0.840**	0.219
State servant	0.226	0.215	0.232	0.269
Trader-financier	0.116	0.177	-0.206	0.249
Technician	0.089	0.153	0.038	0.200
Workman	0.043	0.146	0.438**	0.177
Farmer	-0.721**	0.275	-0.443	0.334
Elementary job	0.468**	0.156	-0.172	0.219
In the forces	-0.448	0.504	0.237	0.520
Driver	-0.218	0.244	0.094	0.276
Foreign	-0.449*	0.259	-0.062	0.248
Public	-0.005	0.112	-0.302**	0.119
Private	0.071	0.126	0.066	0.133
# Workers<100	0.036	0.104	-0.267*	0.134
# Workers>500	-0.013	0.104	0.042	0.121
Duration. arrears	0.018	0.013	0.040**	0.013
Am. arrears	0.000	0.003	0.000	0.003
In kind remuneration	-0.092	0.123	-0.020	0.140
# of child in the household	-0.064	0.044	-0.010	0.054
# of elderly in the household	-0.036	0.066	0.075	0.078
Low hh expenditure per capita	0.041	0.087	0.282**	0.111
Medium hh expenditure per cap.	0.075	0.100	0.272*	0.126
Moscow-St Petersburg	0.143	0.165	0.296	0.210
Northwest	0.228	0.177	-0.575*	0.302
Center	0.101	0.145	-0.014	0.185
Volga	-0.069	0.147	0.038	0.187
Caucasus	-0.237	0.163	0.477**	0.193
Urals	-0.006	0.147	0.163	0.183
Or. Siberia	0.169	0.170	0.053	0.229
Semi-urban	0.008	0.150	0.276	0.168
Rural	-0.009	0.111	-0.272*	0.145
Constant	-0.688*	0.366	-1.138**	0.466
Log Likelihood		-842.9		-524.0
Number of Observations		1922		2205
Chi <sup>2</sup> (37)		95.5		148.7
Prob. >chi <sup>2</sup>		0.0000		0.0000
Pseudo R <sup>2</sup>		0.054		0.125

Sources: Panel data from RLMS, Rounds VII

**Notes:***Dependent variable:*

y=1 if the person is on short-time work or unpaid leaves in 1996

y=0 if not

Legend: \*= statistically significant at the 5% level

\*\*= statistically significant at the 1% level

Reference variables: Clerk; Semi-public, # of employees between 100 and 500, High household expenditures per capita, Occidental Siberia and Urban Settlement

*II.2. Preliminary conclusions.*

(i) We can observe some distinct signs of the segmentation of employees. However, this segmentation is not reduced to conventional division blue collars vs. white collars or low skills vs. high skills. The probability of short-time work is important for elementary professions, but for the intellectuals as well<sup>10</sup>. Both workmen and engineers are likely to be concerned by the practice of unpaid leaves. In other words, if one considers both policies – unpaid leaves and short-time work – as "precarious" states for the employees, there is no evidence that the blue-collar workers are a particularly disgraced group.

Not only these findings corroborate heavily the hypothesis (H2) formulated in the previous section but they even extend it: labor segmentation according to professional qualification is not limited to the division within the pool of short-time workers but characterizes those employees who faced unpaid leaves as well.

(ii) The relation between elementary qualification and short-time work suggests that both forms of employment practices are rather different and even opposite regarding the reasons of labor "retaining". Elementary workers are less qualified, most of them do not have any specific skills while it is just the opposite for the basic characteristics of the employees who had been put on "administrative" leave. The fact that the probability to be on unpaid leaves is higher for engineers and workmen seems to be related to their specific competencies within the firm, while short-time workers do not have any firm-attachment in terms of skills.

This confirms our first hypothesis (H1) related to the relationship between the forms of labor "retaining" and the nature of employees' competencies.

(iii) Thus, the forms of labor "retaining" appear as forms of internal adjustment, which affect differently workers according to their profession. If this is the case, labor demand of Russian firms differs by professional competencies, which gives us a new argument to suggest that the external the mobility of employees will probably be linked to their profession.

We need now to investigate the mobility itself in order to find the impact of our two proxies of labor "hoarding" on job-to-job mobility.

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<sup>10</sup> However intellectual professions are usually working outside the enterprise, they are probably on short time not because of the enterprise employment policy but rather because of their own activity, as teacher for instance.

### III. Labor "retaining" policies and *job-to-job* mobility.

#### III. 1. Three measures of external mobility

Until now, we have considered exits from enterprises as being mainly limited to inflows in unemployment, which is only one component of external labor mobility. We shall consider now another component of this mobility: movement across the jobs.

Measuring this component using the RLMS database is not evident. Thus, labor turnover defined as hiring plus separations could not be calculated using households' or individual data. Earle and Sabrianova (1999) argue that the RLMS data does not allow us to estimate (external) labor mobility because of "ambiguity in the explicit question on job-changing" (p. 11). In fact, with this question<sup>11</sup> we actually could not distinguish individuals staying within the enterprise from individuals moving outside.

Nevertheless, measuring external mobility on the basis of the RLMS survey is possible by using three types of variables. We have searched to measure more precisely *job to job* mobility. In our paper, it differs from the *activity to activity* mobility<sup>12</sup>; and is restricted to primary wage-earning job.

First, we "filtered" the information provided by the "ambivalent" question mentioned above with the data on professional mobility and we constructed on this basis a dummy variable "New job". This approach appeared as justified, as we have found out in the regression that "New job" depended positively on some variables that usually characterize the labor turnover in Russia (see III.2).

Second, using the RLMS data we can define the date of entry in the firm as an information of mobility. We constructed dummy variables "Change job" for people changing firm between 1994 and 1995, between 1995 and 1996 as well as over the all period (1994-1996)<sup>13</sup>. We checked that these people did really work the year before they changed job<sup>14</sup>.

Third, we defined another variable measuring the tenure of a person within a firm. We constructed two dummies of less than one year of tenure and less than three years (see table 6). The interest of the "Tenure" variable is double. First, it allows measuring external mobility. Second, as a measure of firm-specific competencies, it allows us to directly test the stated above conclusions on the importance of the firm-specific skills.

The "New job" and the "Change job" dummies are explained by lagged independent variables, while the "Tenure" dummies are explained by contemporary independent variables. "Backward" and "forward" analysis could be conducted using these dummies in a simple

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<sup>11</sup> The question is "tell me, please, did you change your place of work or profession by comparison with December of last year?"

<sup>12</sup> See B.Najman, A. Pailhé (2000) where the external labor mobility in Russia is defined as "activity to activity" movement which includes formal employment, informal and individual activities, plots, and multi-activity.

<sup>13</sup> Workers changing enterprise either in 1995 or in 1996.

<sup>14</sup> We also checked that in every year chronology of entrance in the enterprise is coherent. For example in 1996 round the date of entry could not be reported before the date of entry reported in 1994 round.

*probit* model (Greene, 1993, p.646-655). All dummies reflect employees' behavior in relation with their skills. The more workers stay within "their" firm the more they get firm-specific competencies.

The following table (6) presents the extent of the job-to-job mobility (in terms of percentage and number of people) according to RLMS data. We observe a high percentage of people with low tenure (12-15 % with less than one year and 34-35% with less than three years). This suggests a high job-to-job mobility. The variables "Change job" and "New job" reflect somehow these results. We should stress that RLMS data probably largely underestimate the labor mobility because in the panel data we are following people which are not moving geographically.

**Table 6: Job to Job Mobility in Russian Firms (in number of people and %)**

	YES	Yes (as a % of total)	NO
Changed Job Between 1994 and 1995	262	11.5%	2013
Changed Job Between 1995 and 1996	284	12.3%	2014
Changed Job Between 1994 and 1996 <sup>15</sup>	500	20.1%	1984
New job in 1996	342	12.9%	2315
Tenure less than 1 year in 1994	420	15.0%	2376
Tenure less than 3 years in 1994	981	35.1%	1884
Tenure less than 1 year in 1996	322	12.3%	2289
Tenure less than 3 years in 1996	883	33.8%	1536

*Sources: RLMS, rounds V, VI, VII*

### III. 2. Results

The regression results are presented in Tables 8-11 and summarized in Table 7.

(i) The determinants of "New job" are closely related to the main characteristics of the labor mobility rotation in Russia. For instance, it follows from the Goskomstat data and the case studies (Gimpelson and Lippoldt 1997, Smirnov 1998) that young men with no qualification are the bulk of Russian high mobility. It is the same for our study: having a new job in 1996 is negatively correlated with age (young people have more chances), and positively with one profession, elementary workers. Moreover, the "New job" movement is

<sup>15</sup> The person changed job either between 1994 and 1995 or between 1995 and 1996



linked to the other signs of extreme labor mobility in Russia, such as working in the small, private or foreign firms and living in particular regions like Moscow and St-Petersburg.

Notice also the positive correlation of the "New job" variable with the substitution of remuneration "in kind" (goods or services) for the pay in money<sup>16</sup>.

(ii) The key determinants of the "Change job" variable are close to those of "New job", thus confirming most of characteristics of external mobility revealed by analyzing the latter variable. The probability of moving across jobs is higher for young, male, and elementary workers. In 1995-1996, changing job is highly probable for the inhabitants of Moscow and St-Petersburg for employees in foreign firms and those accepting In kind remuneration in lieu of wage pay.

(iii) Both variables of tenure in 1996, less than one year and less than three year, are positively correlated with "traditional" determinants of mobility (being young, male, and elementary worker) and being employed in private firm. Tenure is negatively correlated with the amount of wage arrears and experience of unpaid leaves and positively to the "demonetisations" of remuneration. Tenure of less than three years is negatively correlated with intelligentsia, and positively with employment in a small or private firm. Tenure of less than one year is positively correlated with living in Moscow-St Petersburg.

For 1994, the determinants of both tenure variables are largely similar. The main differences are related to the appearance of negative correlation between being engineer and having a tenure of less than three years.

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<sup>16</sup> Russian firms generalize more and more the practice of offering goods or services instead of regular wage pays. The RLMS survey contained an explicit question on this subject.

**Table 7. The main determinants of the job-to-job mobility (see tables 8, 9, 10, and 11)**

New job	Change job	Short Tenure (in 1994)	Short Tenure (in 1996)
+young and male	+young and male	+young and male -unpaid leaves	+young and male -unpaid leaves
+elementary job	+elementary job	+elementary job -intelligentsia <sup>3</sup> -engineer <sup>3</sup>	+elementary job -intelligentsia <sup>3</sup> -driver
+small firm		+small firm <sup>3</sup>	+small firm <sup>3</sup>
+private firm		+private firm	+private firm <sup>3</sup>
+foreign firm	+foreign firm***	+public firm	
		-duration of arrears <sup>1</sup>	-amount of arrears
+In kind remuneration	+In kind remuneration***		+In kind remuneration
+Moscow-St-Petersburg	+Moscow-St-Petersburg***		+Moscow-St-Petersburg <sup>1</sup>

Here "+" and "-" indicate the sign of the correlation

<sup>1</sup>less than 1 year

<sup>3</sup>less than 3 years

\*\*\*Only between 1995 and 1996

### III.3. Conclusions.

(i) The positive relation between moving from job to job and elementary workers provides an explanation to why at least one group of short-time workers is not afraid of losing their workplace (see Section I). Being always prepared for job-changing, elementary workers feel secure, since they do not actually face a direct threat of unemployment.

Notice that not all short-timers are prepared to quit the firm as elementary workers. For instance, being female increases the probability of reduced-hours experience, but it is not consistent with the status of a male worker mostly related to external mobility.

Thus, the hypothesis (H2) related to the segmentation of the short-time workers is confirmed regarding outside job options.

(ii) If H1 is correct, that is if it is true that unpaid leaves are related to specific skills, one should expect a positive correlation between unpaid leaves and longer tenure. Our regression results show that this is actually the case for all four tenure variables (in 1994 and in 1996).

(iii) In the same vein, professional structure of external mobility largely confirms this relationship between the nature of employees' skills and the forms of labor "retaining". Our findings show that quits are strongly related to the absence of firm-specific competencies (elementary professions). On the contrary, despite the fact that we have already established the link between staying within a firm and having specific skills (see previous point (ii) on longer tenure and unpaid leaves), the negative correlation between some of the short-tenure variables and employees with specific knowledge (engineers, drivers) still holds. The Hypothesis 1 is again confirmed.

(iv) The wage arrears appear to produce a coherent effect. The duration and the amount of arrears are likely to increase the tenure. We propose a following explanation of this result. Accepting wage arrears is primarily related to the limited job opportunities. Specific skills could play an important role in employees' "tolerant" attitude. Thus, staying in firms with wage debts or accepting to be put on unpaid leaves have the same reasons: this type of employees wish to retain their jobs since they are not sure to utilize their knowledge elsewhere. The positive correlation between the duration of arrears and probability of being on unpaid leaves (see Table 5) seems to support this suggestion.

There is no evidence in the survey that social benefits play a substantial role in the workers' attachment to the firms with "wage problems". On the other hand, "demonetisation" of the employment relationship (goods and services instead of monetary wages) provokes immediately separations that, therefore, seem to be voluntary quits.

**Table 8: Probability to Have a New Job  
in 1996 (probit regression)**

Variables in 1996 except for labor "retaining" wage arrears and In kind remuneration in 1995	New Job	
	Coef.	Standard error
Gender	0.270**	0.102
Age	-0.015**	0.004
Education (year)	0.020	0.026
Couple	-0.086	0.107
Plots	-0.284**	0.110
Informal Activity	0.253	0.168
Manager	-0.119	0.350
Intelligentsia	0.110	0.195
Engineer	-0.252	0.280
State servant	-0.010	0.288
Trader-financier	0.348*	0.215
Technician	0.099	0.194
Workman	0.166	0.180
Farmer	-0.230	0.261
Elementary job	0.534**	0.192
In the forces	-0.101	0.553
Driver	0.028	0.263
Foreign	0.416*	0.237
Public	0.002	0.130
Private	0.327*	0.145
# Workers<100	0.281*	0.129
# Workers>500	0.043	0.129
Duration. Arrears in 1995	0.005	0.019
Am. arrears in 1995	0.004	0.007
Unpaid leaves in 1995	-0.041	0.177
Short-time work in 1995	0.087	0.110
In kind remuneration in 1995	0.372**	0.144
# of child in the household	-0.067	0.050
# of elderly in the household	0.000	0.077
Low hh expenditure per capita in 1995	0.194*	0.095
Medium hh expenditure per capita in 1995	-0.134	0.132
Moscow-St Petersburg	0.603**	0.198
Northwest	0.495**	0.204
Center	0.177	0.181
Volga	0.069	0.184
Caucasus	-0.037	0.214
Urals	0.287	0.181
Or. Siberia	0.296	0.205
Semi-urban	0.011	0.183
Rural	0.075	0.130
Constant	-1.597**	0.451
<hr/>		
Log Likelihood		-568.5
Number of Observations		1975
Chi <sup>2</sup> (39)		127.1
Prob. >chi <sup>2</sup>		0.0000
Pseudo R <sup>2</sup>		0.10

**Notes: Dependent variable**

y=1 if the person has a new job in 1996

y=0 if the person had no new job in 1996

*Legend:*

\* = statistically significant at the 5% level

\*\* = statistically significant at the 1% level

*Sources:* Panel data from RLMS, Rounds V, VI, VII*Reference variables:* Clerk; Semi-public, # of employees between 100 and 500, High household expenditures per capita, Occidental Siberia and Urban Settlement

**Table 9: Probability to Have Tenure Under 1 or 3 years in 1994 (probit regression)**

Variables in 1994	Tenure Less Than One Year		Tenure Less Than Three Years	
	Coef.	Standard error	Coef.	Standard error
Gender	<b>0.380**</b>	0.090	<b>0.500**</b>	0.075
Age	<b>-0.027**</b>	0.004	<b>-0.031</b>	0.003
Education (year)	<b>0.002</b>	0.022	<b>0.020</b>	0.018
Couple	<b>-0.059</b>	0.096	<b>-0.082</b>	0.080
Plots	<b>0.035</b>	0.128	<b>-0.165</b>	0.103
Informal Activity	<b>-0.115</b>	0.141	<b>0.115</b>	0.114
Manager	<b>0.046</b>	0.270	<b>0.008</b>	0.221
Intelligentsia	<b>-0.232</b>	0.172	<b>-0.284*</b>	0.137
Engineer	<b>-0.103</b>	0.214	<b>-0.384*</b>	0.171
State servant	<b>-0.251</b>	0.237	<b>0.044</b>	0.187
Trader-financier	<b>0.211</b>	0.186	<b>0.141</b>	0.158
Technician	<b>0.027</b>	0.155	<b>-0.091</b>	0.126
Workman	<b>0.051</b>	0.149	<b>-0.178</b>	0.122
Farmer	<b>-0.046</b>	0.210	<b>-0.115</b>	0.168
Elementary job	<b>0.563**</b>	0.163	<b>0.592**</b>	0.138
In the forces	<b>-0.821</b>	0.542	<b>-0.638*</b>	0.377
Driver	<b>-0.237</b>	0.213	<b>-0.259</b>	0.170
Foreign	<b>0.257</b>	0.234	<b>0.269</b>	0.192
Public	<b>0.267*</b>	0.113	<b>0.174*</b>	0.087
Private	<b>0.450**</b>	0.127	<b>0.320**</b>	0.101
# Workers<100	<b>0.114</b>	0.099	<b>0.241**</b>	0.082
# Workers>500	<b>-0.127</b>	0.102	<b>-0.024</b>	0.083
Duration. arrears	<b>-0.088**</b>	0.030	<b>-0.031*</b>	0.019
Am. Arrears	<b>-0.003</b>	0.011	<b>-0.011</b>	0.008
Unpaid leaves	<b>-0.804**</b>	0.176	<b>-0.549**</b>	0.108
Short-time work	<b>-0.126</b>	0.099	<b>0.012</b>	0.078
In kind remuneration	<b>0.127</b>	0.139	<b>0.081</b>	0.114
# of child in the household	<b>-0.021</b>	0.042	<b>-0.083**</b>	0.035
# of elderly in the household	<b>-0.052</b>	0.070	<b>-0.047</b>	0.057
Low hh. expenditure per capita	<b>0.009</b>	0.092	<b>-0.153*</b>	0.075
Medium hh. expenditure per capita	<b>0.011</b>	0.103	<b>-0.063</b>	0.084
Constant	<b>-0.170</b>	0.369	<b>0.607*</b>	0.308
Log Likelihood	-786.0		-1234.0	
Number of Observations	2190		2190	
Chi <sup>2</sup> (41)	217.6		351.9	
Prob. >chi <sup>2</sup>	0.0000		0.0000	
Pseudo R <sup>2</sup>	0.12		0.125	

**Notes:** *Dependent variable:**Sources:* Panel data from RLMS, Rounds VI, VII

y=1 if the person had less than (1 or 3) year of tenure in 1994  
y=0 if the person had more than (1 or 3) year of tenure in 1994

Regression is run with regional dummies and settlement type dummies.

*Legend:*

\* = statistically significant at the 5% level

\*\* = statistically significant at the 1% level

*Reference variables:* Clerk; Semi-public, # of employees between 100 and 500, Occidental Siberia and Urban Settlement, high level of household expenditures per capita.

**Table 10: Probability to Have Tenure Under 1 or 3 years in 1996 (probit regression)**

Variables in 1996	Tenure Less Than One Year		Tenure Less Than Three Years	
	Coef.	Standard error	Coef.	Standard error
Gender	<b>0.255**</b>	0.097	<b>0.334**</b>	0.078
Age	<b>-0.022**</b>	0.004	<b>-0.027**</b>	0.003
Education (year)	<b>0.001</b>	0.024	<b>0.014</b>	0.019
Couple	<b>-0.173*</b>	0.099	<b>-0.149*</b>	0.081
Plots	<b>-0.171</b>	0.109	<b>-0.144*</b>	0.081
Informal Activity	<b>0.391**</b>	0.151	<b>0.242*</b>	0.131
Manager	<b>-0.213</b>	0.325	<b>-0.245</b>	0.255
Intelligentsia	<b>-0.158</b>	0.174	<b>-0.294/</b>	0.141
Engineer	<b>-0.383</b>	0.253	<b>-0.134</b>	0.180
State servant	<b>-0.401</b>	0.277	<b>0.020</b>	0.195
Trader-financier	<b>0.017</b>	0.188	<b>0.010</b>	0.161
Technician	<b>-0.253</b>	0.177	<b>-0.125</b>	0.139
Workman	<b>-0.187</b>	0.160	<b>-0.105</b>	0.128
Farmer	<b>-0.193</b>	0.237	<b>-0.039</b>	0.185
Elementary job	<b>0.319*</b>	0.169	<b>0.527**</b>	0.144
In the forces	<b>-0.359</b>	0.570	<b>-0.020</b>	0.346
Driver	<b>-0.702**</b>	0.291	<b>-0.371*</b>	0.203
Foreign	<b>0.215</b>	0.240	<b>0.184</b>	0.193
Public	<b>0.080</b>	0.136	<b>0.022</b>	0.100
Private	<b>0.401**</b>	0.147	<b>0.387**</b>	0.112
# Workers<100	<b>0.011</b>	0.125	<b>0.147</b>	0.095
# Workers>500	<b>0.000</b>	0.122	<b>-0.024</b>	0.094
Duration. Arrears	<b>0.001</b>	0.018	<b>-0.020</b>	0.014
Am. Arrears	<b>-0.017**</b>	0.005	<b>-0.008**</b>	0.003
Unpaid leaves	<b>-0.408*</b>	0.200	<b>-0.356**</b>	0.137
Short-time work	<b>0.071</b>	0.105	<b>-0.030</b>	0.086
In kind remuneration	<b>0.346**</b>	0.131	<b>0.199*</b>	0.107
# of child in the household	<b>0.041</b>	0.049	<b>-0.047</b>	0.039
# of elderly in the household	<b>0.022</b>	0.075	<b>0.018</b>	0.060
Low hh expenditure per capita	<b>0.174*</b>	0.102	<b>0.002</b>	0.079
Medium hh expenditure per capita	<b>0.098</b>	0.119	<b>0.030</b>	0.092
Semi-urban	<b>-0.116</b>	0.196	<b>-0.265*</b>	0.144
Rural	<b>-0.053</b>	0.134	<b>-0.106</b>	0.102
Constant	<b>-0.468</b>	0.415	<b>0.590*</b>	0.334
Log Likelihood	-616.7		-1067.0	
Number of Observations	1886		1886	
Chi <sup>2</sup> (41)	177.0		276.1	
Prob. >chi <sup>2</sup>	0.0000		0.0000	
Pseudo R <sup>2</sup>	0.126		0.11	

**Notes:** *Dependent variable:*

y=1 if the person had less than (1 or 3) year of tenure in 1996

y=0 if the person had more than (1 or 3) year of tenure in 1996

*Sources:* Panel data from RLMS, Rounds V, VI, VII

Regression is run with regional dummies

*Legend:*

\* = statistically significant at the 5% level

\*\* = statistically significant at the 1% level

*Reference variables:* Clerk; Semi-public, # of employees between 100 and 500, Occidental Siberia and Urban Settlement, high level of household expenditures per capita

**Table 11: Probability to change enterprise, between 1994 and 1996 (probit regression)**

Variables (are lagged, t-1)	Changed Job Between 1994-1995		Changed Job Between 1995-1996		Changed Job Between 1994-1996	
	Coef.	Standard error	Coef.	Standard error	Coef.	Standard error
Gender	0.227*	0.104	0.261**	0.101	0.300**	0.087
Age	-0.020**	0.004	-0.011**	0.004	-0.019**	0.004
Education (year)	0.077**	0.027	0.022	0.026	0.050*	0.022
Couple	-0.054	0.112	-0.185*	0.105	-0.148	0.094
Plots	0.045	0.151	-0.124	0.104	-0.071	0.128
Informal Activity	0.168	0.148	0.334*	0.155	0.239*	0.127
Manager	-0.140	0.326	0.176	0.224	-0.254	0.279
Intelligentsia	-0.240	0.198	0.061	0.193	-0.201	0.163
Engineer	-0.048	0.232	-0.291	0.255	-0.128	0.195
State servant	-0.280	0.277	-0.218	0.298	-0.325	0.229
Trader-financier	-0.017	0.239	0.208	0.227	0.049	0.194
Technician	-0.132	0.191	-0.004	0.188	-0.106	0.155
Workman	0.054	0.177	0.075	0.168	-0.001	0.145
Farmer	0.324	0.234	-0.158	0.243	0.101	0.198
Elementary job	0.359*	0.201	0.372*	0.184	0.369*	0.166
In the forces	-	-	-0.283	0.527	-0.473	0.467
Driver	0.264	0.233	0.034	0.223	0.124	0.197
Foreign	0.218	0.253	0.419*	0.219	0.137	0.216
Public	0.121	0.122	-0.064	0.120	0.071	0.101
Private	0.242*	0.140	0.129	0.137	0.167	0.117
# Workers<100	0.010	0.113	0.208	0.128	0.145	0.097
# Workers>500	-0.169	0.114	-0.003	0.127	-0.056	0.098
Duration. arrears	0.025	0.026	0.023	0.015	0.028	0.022
Am. arrears	-0.008	0.013	-0.001	0.007	-0.015	0.012
Unpaid leaves	-0.023	0.136	0.109	0.158	0.021	0.112
Short-time work	-0.036	0.109	0.151	0.105	-0.075	0.093
In kind remuneration	0.007	0.157	0.300*	0.138	-0.051	0.139
# of child in the household	-0.016	0.050	0.013	0.048	0.015	0.042
# of elderly in the household	0.001	0.081	0.109	0.074	0.012	0.066
Low hh. expenditure per capita	0.088	0.106	0.076	0.094	0.060	0.089
Medium hh. expenditure per capita	-0.040	0.122	-0.121	0.124	-0.024	0.101
Moscow-St Petersburg	0.066	0.195	0.467**	0.185	0.213	0.162
Northwest	-0.250	0.208	0.257	0.196	-0.123	0.171
Center	-0.002	0.156	0.020	0.167	-0.080	0.135
Volga	-0.255	0.166	0.019	0.167	-0.168	0.139
Caucasus	0.026	0.176	-0.247	0.203	-0.115	0.152
Urals	-0.300*	0.170	0.067	0.170	-0.291*	0.142
Or. Siberia	0.177	0.191	0.106	0.195	0.017	0.166
Semi-urban	0.082	0.158	-0.406*	0.205	-0.128	0.144
Rural	-0.362**	0.140	-0.055	0.126	-0.204*	0.113
Constant	-1.453**	0.439	-1.352**	0.432	-0.784*	0.368
Log Likelihood	-574.2		-601.5		-854.0	
Number of Observations	1825		1853		1902	
Chi <sup>2</sup> (39)	83.6		102.5		107	
Prob. >chi <sup>2</sup>	0.0000		0.0000		0.0000	
Pseudo R <sup>2</sup>	0.068		0.078		0.059	

Sources: Panel data from RLMS, Rounds V, VI, VII

**Notes: Dependent variable**

y=1 if the person changed enterprise during the period

y=0 if the person did not change enterprise during the period

**Legend:**

\* = statistically significant at the 5% level \*\* = statistically significant at the 1% level

**Reference variables:** Clerk; Semi-public, # of employees between 100 and 500, Occidental Siberia and Urban Settlement, high level of household expenditures per capita.

#### IV. Discussion.

We can conclude that labor "retaining" mechanism is not homogeneous. Firm's behavior in this domain is linked to the labor segmentation. Each segment is composed by a large group of employees whose characteristics determine the specificity of employment policies. We can thus distinguish two major policies corresponding to different groups of workers (see scheme 1).

(i) "Administrative"-leave-policies reflect firms' efforts to keep employees with specific skills. Most of these workers have no choice since firm-specific competencies are not likely to be employed elsewhere. "Administrative" leaves represent a form of internal flexibility and they have only a secondary effect on external mobility: keeping firm-specific workers favors pushing elementary workers outside the borders of the firm.

(ii) Short-time work reflects another form of internal flexibility, which is especially, related to continuous, though not always regular demand for elementary professions. Contrary to the employees endowed with specific competencies, most of elementary workers are not integrated into the enterprise, they can easily quit it, thus playing an important role in the connection between internal mobility (participating in the reduced hours practices) and external mobility (job-to-job changing across the firms). Finally, the role of wage arrears in the internal labor adjustment is similar to unpaid leaves practices. Wage arrears for both amount and duration of arrears, are likely to be related to long tenure. For employees who have few external opportunities and, therefore, are "locked" within their firm there is practically no other solution than to accept wage arrears. In this case, managers utilize wage arrears as an additional means of internal adjustment. The fact that the probability to be on unpaid leaves is positively correlated with the duration of arrears supports this interpretation. However, there could be also a problem of wage arrears definition, since employees sent on "administrative" leaves often consider themselves as suffering from wage debts even if formally firms do not owe them any wages.

These results are consistent with the evidence provided by much of case studies (Koshman, 1995; Samara Research Group, 1996): enterprises continue to experience labor shortages when they have to meet orders. Facing extreme uncertainty on the product market managers' reaction consists in using various forms of internal flexibility. Therefore the rational behavior is to create and keep a permanent reserve of human resources every time a firm (or one of its departments or workshops) has to meet a sudden increase in demand.

The question is: in what kind of human resources are Russian enterprises interested?

The outlined results allow us to test the "technological" argument (Commander et al. 1995). This argument points out firm-specific nature of human resources and hence the extremely complementary characteristics of labor-capital relation (Grosfeld et al., 1999). Our results show that keeping firm-specific competencies as a part of the firms' strategies represents a form of internal flexibility, i.e. labor adjustment, which explains relative employment stability. The distinction between blue collar and white collar employees seems to be undetermined in this context since the firm-specific competencies concern more diversified groups of employees, including in particular engineers and most of workers.

Nevertheless, the firm-specific explanation provides only a part, though important, of understanding relative employment stability.

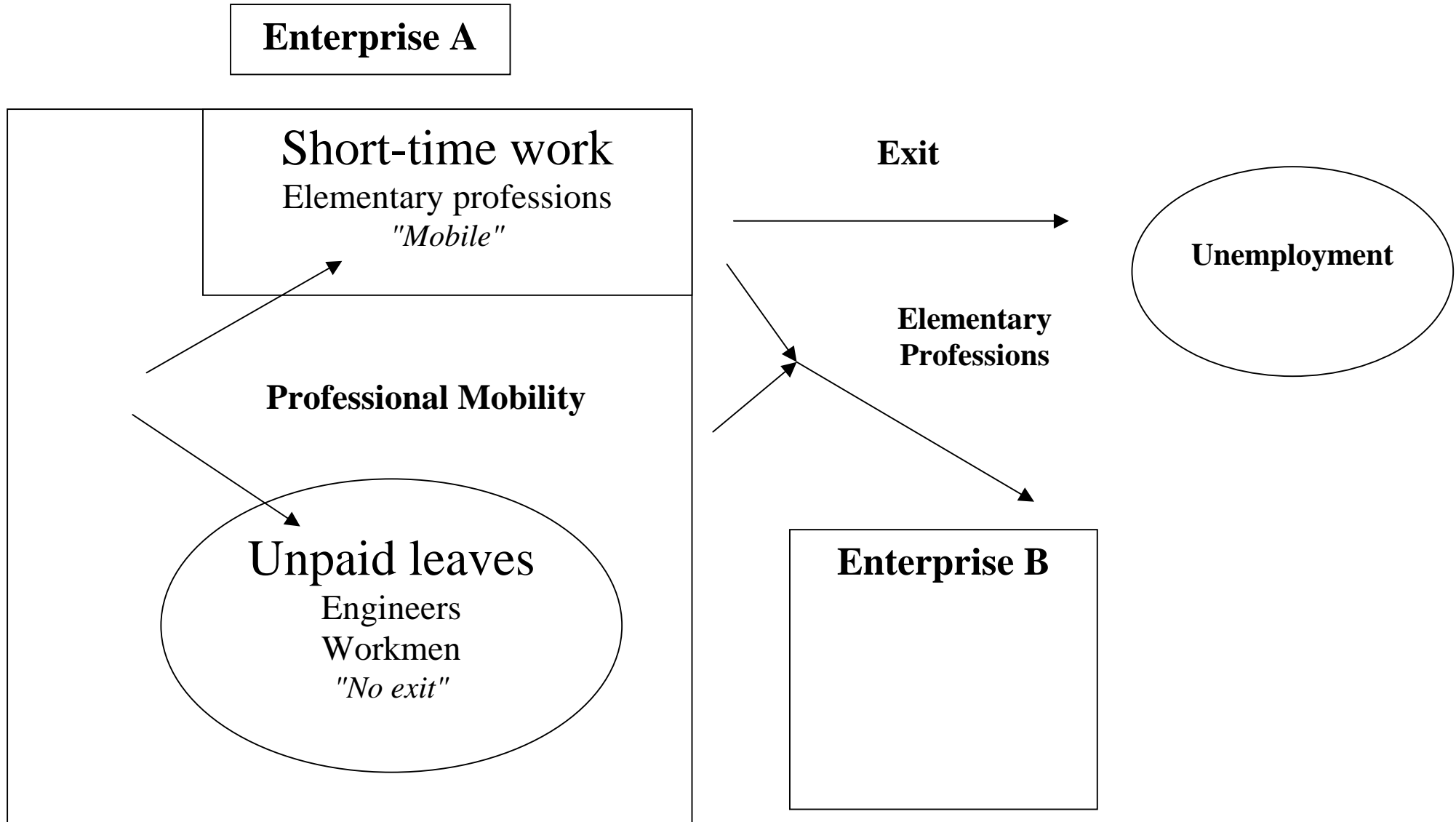
The results show also that the second form of labor adjustment is related to the evident need to hire "labor in general". In other words, in the context of continuing degradation of physical capital stock, labor tends to substitute capital which is necessary to face the extreme product market uncertainty. This evidence is also in accordance with the results of different case studies (Donova, 1996). Investment fall in Russian economy creates permanent needs in unskilled workers (or former skilled employees who accept to work as unskilled workers) which firms try to meet by the recourse to the internal mobility (reduced-hours policy) or to the external labor market.



## V. Conclusion

- First, what is usually qualified, as labor hoarding seems to represent particular forms of labor adjustment and internal flexibility. Two forms have been investigated in this paper: unpaid leaves and short-time work work.
- Second, unpaid leaves and short-time work correspond to different segments of labor. However, this segmentation is not reduced to a simplified opposition blue collar versus white collar.
- Third, at least these two forms of labor adjustment provide us a robust explanation of labor-output Russian paradox. Both of them stimulate internal labor reallocation.

# Scheme 1: Internal Flexibility and External Mobility



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## Appendix

### APPENDIX 1: Job Tenure average by professions in 1996

	Tenure less than 1 year (In % of the total)	Tenure less than 3 years (In % of the total)	Tenure		N With Tenure
			Average years	Standard error	
<b>Total</b>	<b>24 %</b>	<b>41%</b>	<b>8.7</b>	<b>9.3</b>	<b>2711</b>
<i>engineer</i>	18%	31%	12.0	10.6	130
<i>intelligentsia</i>	21%	34%	10.4	10.1	372
<i>workman</i>	23%	39%	9.8	10.0	742
<i>farmer</i>	23%	41%	9.7	10.1	198
<i>technician</i>	19%	37%	9.3	9.2	292
<i>driver</i>	19%	32%	8.8	8.3	125
<i>clerk</i>	26%	40%	8.1	8.6	242
<i>manager</i>	24%	51%	7.0	8.9	46
<i>state servant</i>	21%	48%	6.6	7.1	100
<i>in the forces</i>	12 %	40 %	5.8	4.2	26
<i>trader-financier</i>	34%	56%	5.3	6.5	170
<i>elementary job</i>	40%	60 %	4.6	5.9	268

Sources: RLMS, round VII

**APPENDIX 2: Tenure within individual and household characteristics in 1996**

	Tenure less than 1 year (In % of the total)	Tenure less than 3 years (In % of the total)	Tenure		N With Tenure
			Average years	Standard error	
<b>Gender</b>					
<i>Male</i>	26%	45%	8.5	9.5	1209
<i>Female</i>	23%	38%	8.9	9.1	1404
<b>Age</b>					
<i>Less than 25</i>	56%	80%	1.9	2.3	207
<i>25-40</i>	26%	44%	5.9	5.4	1075
<i>40-55/60</i>	19%	33%	11.5	10.2	1093
<i>Over 55/60</i>	16%	29%	15.0	13.6	238
<b>Education</b>					
<i>Primary</i>	17%	34%	11.1	11.1	102
<i>Secondary</i>	25%	43%	8.5	9.1	1261
<i>University</i>	24%	40%	8.8	9.3	1250
<b>Type of Settlement</b>					
<i>Rural</i>	20%	37%	9.0	9.4	616
<i>Semi-urban</i>	17%	36%	8.5	8.8	179
<i>Urban</i>	27%	43%	8.8	9.3	1818
<b>Household Expenditures<sup>1</sup></b>					
<i>High</i>	23%	42%	8.8	9.4	781
<i>Medium</i>	26%	44%	8.9	9.8	539
<i>Low</i>	24%	40%	8.6	9.1	1293
<b>Work on Plots</b>					
<i>No</i>	27%	45%	7.9	8.6	1851
<i>Yes</i>	18%	33%	10.7	10.5	762
<b>Informal Activities</b>					
<i>No</i>	24%	41%	8.7	9.3	2447
<i>Yes</i>	30%	43%	8.2	9.2	166

Sources: RLMS, round VII

<sup>1</sup> The households expenditures are per capita, and adjusted by the oxford scale (here 0.5 for child, and 0.75 for elderly people)

**APPENDIX 3: Tenure within enterprise characteristics and employment policy in 1996**

	Tenure less than 1 year (In % of the total)	Tenure less than 3 years (In % of the total)	Tenure		N With Tenure
			Average years	Standard error	
<b>Firm size</b>					
<i>&lt;20</i>	<b>30%</b>	<b>50%</b>	<b>7.2</b>	8.8	341
<i>21-100</i>	<b>26%</b>	<b>41%</b>	<b>8.4</b>	8.9	594
<i>100-200</i>	<b>17%</b>	<b>31%</b>	<b>10.1</b>	9.6	211
<i>201-500</i>	<b>24%</b>	<b>43%</b>	<b>9</b>	9.5	236
<i>Over 500</i>	<b>13%</b>	<b>26%</b>	<b>12.1</b>	<i>10.1</i>	341
<b>Ownership</b>					
<i>State</i>	<b>22%</b>	<b>39%</b>	<b>8.9</b>	9.3	1580
<i>Semi-public or Partnership</i>	<b>19%</b>	<b>37%</b>	<b>10.1</b>	9.8	458
<i>Private</i>	<b>32%</b>	<b>48%</b>	<b>7.6</b>	9.0	557
<b>On Short-time work</b>					
<i>No</i>	<b>23%</b>	<b>40%</b>	<b>8.7</b>	9.2	2207
<i>Yes</i>	<b>30%</b>	<b>44%</b>	<b>8.7</b>	9.7	406
<b>On unpaid leaves</b>					
<i>No</i>	<b>25%</b>	<b>43%</b>	<b>8.5</b>	9.2	2373
<i>Yes</i>	<b>16%</b>	<b>27%</b>	<b>11.8</b>	<i>10.2</i>	215

Sources: RLMS, round VII

**APPENDIX 4: Definition of variables**

<i>Variable</i>	<i>Definition</i>
Gender	Equals 1 if the person is a man, else 0
Age	Age of the person
Education	Education number of years
Couple	Equals 1 if the person lives as a couple, else 0
Plots	Equals 1 if the person worked on a <i>plots</i> , else 0
Informal Activity	Equals 1 if the person worked in an <i>Informal Activity</i> , else 0
Unemployment self-rep.	Equals 1 if the person reported it self as unemployed , else 0
Tenure	Number of years of tenure
Tenure less then 1 year	Equals 1 if the tenure of the person is 1 year or less, else 0
Tenure less than 3 years	Equals 1 if the tenure of the person is 3 year or less, else 0
Change job	Equals 1 if the person changed job during the period, else 0
# children	Number of children in the family (respondent excluded)
# elderly	Number of elderly in the family (respondent excluded)
Low hh expenditure 94	Equals 1 if the person bellows to low per capita expenditure (real) category in 1994, else 0
Medium expenditure 94	Equals 1 if the person bellows to medium per capita expenditure (real) category in 1994, else 0
Low expenditure 95	Equals 1 if the person bellows to low per capita expenditure (real) category in 1995, else 0
Medium expenditure 95	Equals 1 if the person bellows to medium per capita expenditure (real) category in 1995, else 0
Center	Equals 1 if the person lives in Central Russia, else 0
Northwest	Equals 1 if the person lives in Northwest Russia, else 0
Volga	Equals 1 if the person lives in Volga, else 0
Caucasus	Equals 1 if the person lives in Caucasus, else 0
Urals	Equals 1 if the person lives in Urals, else 0
Occ. Siberia	Equals 1 if the person lives in Occidental Siberia, else 0
Or. Siberia	Equals 1 if the person lives in Oriental Siberia, else 0
Manager	Equals 1 if the person is a manager, else 0
Intelligentsia	Equals 1 if the person belongs to intelligentsia, else 0
Engineer	Equals 1 if the person is an engineer, else 0
State servant	Equals 1 if the person is a state servant, else 0
Trader-financier	Equals 1 if the person works in trade or finance, else 0
Technician	Equals 1 if the person is a technician, else 0
Clerk	Equals 1 if the person is a clerk, else 0
Workman	Equals 1 if the person is a workman, else 0
Farmer	Equals 1 if the person is a farmer, else 0
Elementary job	Equals 1 if the person has an elementary job, else 0
In the forces	Equals 1 if the person is in the army, else 0
Driver	Equals 1 if the person is a driver, else 0
# workers<100	Equals 1 if the person works in an enterprise of less than 100 employees, else 0
# workers>500	Equals 1 if the person works in an enterprise of more than 500 employees, else 0
Private	Equals 1 if the person works in a private establishment, else 0
Public	Equals 1 if the person works in a state-run establishment, else 0
Foreign	Equals 1 if the person works in a firm with foreign share, else 0
Newjob 96	Equals 1 if the person found a new job since December 1995, else 0
Short-time work 94	Equals 1 if the person works 120 hours by month and less in 1994, else 0
Short-time work 95	Equals 1 if the person works 120 hours by month and less in 1995, else 0
Short-time work 96	Equals 1 if the person works 120 hours by month and less in 1996, else 0
Unpaid leaves 94	Equals 1 if the person has been sent on compulsory unpaid leaves in 1994
Unpaid leaves 95	Equals 1 if the person has been sent on compulsory unpaid leaves in 1995
Unpaid leaves 96	Equals 1 if the person has been sent on compulsory unpaid leaves in 1996
Am. arrears 94	Total amount of money not been paid to worker in 1994 (*100000)
Am. arrears 95	Total amount of money not been paid to worker in 1995 (*100000)
Duration arrears 94	Wage arrears duration in months in 1994
Duration arrears 95	Wage arrears duration in months in 1995
In kind remuneration 96	Equals 1 if the person has been paid in goods or services in 1996, else 0



**APPENDIX 5: Number of observations, mean, standard deviation, minimum and maximum of all variables in 1996, 1995 and 1994**

<i>Variable</i>	<i>Nb. of Obs.</i>	<i>Mean</i>	<i>Stand. Dev.</i>	<i>Min</i>	<i>Max</i>
Gender	4607	0.43	0.50	0	1
Age in 1996	4607	44.67	15.46	16	72
Education in 1996	4607	10.94	2.76	0	15
Couple in 1996	4607	0.70	0.46	0	1
Plots in 1996	4607	0.34	0.47	0	1
Informal Activity in 1996	4607	0.07	0.25	0	1
Unemployment self reported 96	4607	0.08	0.27	0	1
Tenure less then 1 year in 1996	2611	0.12	0.33	0	1
Tenure less than 3 years in 1996	2611	0.34	0.47	0	1
Tenure in 1996	2611	9.67	9.20	1	55
Tenure in 1995	2747	9.57	9.21	1	54
Change job between 96 & 95	2298	0.12	0.33	0	1
Change job between 95 & 94	2275	0.12	0.32	0	1
Change job in 1996 or 1995	2484	0.20	0.40	0	1
Urban	4607	0.65	0.48	0	1
semi-urban	4607	0.06	0.24	0	1
Rural	4607	0.28	0.45	0	1
# Children	4607	0.84	0.98	0	8
# Elderly	4607	0.38	0.59	0	3
Low expenditure 96	4605	0.56	0.50	0	1
Medium expenditure 96	4605	0.18	0.39	0	1
High expenditure 96	4605	0.26	0.44	0	1
Center	4607	0.21	0.40	0	1
Volga	4607	0.19	0.40	0	1
Caucasus	4607	0.13	0.34	0	1
Urals	4607	0.16	0.36	0	1
Or. Siberia	4607	0.09	0.28	0	1
Occ. Siberia	4607	0.10	0.30	0	1
Northwest	4607	0.06	0.24	0	1
Moscow-St Pe	4607	0.06	0.24	0	1
Manager	4607	0.01	0.10	0	1
Intelligentsia	4607	0.08	0.27	0	1
Engineer	4607	0.03	0.17	0	1
State servant	4607	0.02	0.14	0	1
Clerk	4607	0.05	0.22	0	1
Workman	4607	0.04	0.19	0	1
Trader-financier	4607	0.06	0.24	0	1
Technician	4607	0.16	0.37	0	1
Farmer	4607	0.04	0.20	0	1
In the forces	4607	0.06	0.23	0	1
Driver	4607	0.01	0.07	0	1
Elementary job	4607	0.03	0.16	0	1
# workers<100	2668	0.38	0.48	0	1
# workers>500	2668	0.47	0.50	0	1
Private in 1996	2668	0.20	0.40	0	1
Foreign in 1996	2668	0.03	0.17	0	1
Public in 1996	2668	0.60	0.49	0	1
Am. arrears 96/100000	2287	8.87	15.12	0	240
Duration. arrears 96	2501	2.32	3.81	0	36
Short-time work 94	2782	0.12	0.33	0	1
Short-time work 95	2769	0.07	0.25	0	1
Short-time work 96	2643	0.08	0.27	0	1
Unpaid leaves 94	2461	0.20	0.40	0	1
Unpaid leaves 95	2425	0.19	0.39	0	1
Unpaid leaves 96	2283	0.18	0.39	0	1
In kind remuneration in 94	2781	0.09	0.29	0	1
In kind remuneration in 95	2764	0.09	0.28	0	1
In kind remuneration in 96	2641	0.13	0.33	0	1

*Sources: RLMS, rounds V, VI, VII*

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