

# Gene Expression Associations with the Growth Inhibitory Effects of Small Molecules on Live Cells: Specificity of Effects and Uniformity of Mechanisms

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**Abstract:** The NCI60 human tumor cell line screen is a public resource for studying selective and nonselective growth inhibition of small molecules against cancer cells. By coupling growth inhibition screening data with biological characterizations of the different cell lines, it becomes possible to infer mechanisms of action underlying some of the observable patterns of selective activity. Using these data, mechanistic relationships have been identified including specific associations between single genes and small families of closely related compounds, and less specific relationships between biological processes involving several cooperating genes and broader families of compounds. Here, we aim to characterize the degree to which such specific and general relationships are present in these data. A related question is whether genes tend to act with a uniform mechanism for all associated compounds, or whether multiple mechanisms are commonly involved. We address these two issues in a statistical framework placing special emphasis on the effects of measurement error in the gene expression and chemical screening data. We find that as measurement accuracy increases, the pattern of apparent associations shifts from one dominated by isolated gene/compound pairs, to one in which families consisting of an average of 25 compounds are associated to the same gene. At the same time, the number of genes that appear to play a role in influencing compound activities decreases. For less than half of the genes, the presence of both positive and negative correlations indicates pleiotropic associations with molecules via different mechanisms of action. © 2009 Wiley Periodicals, Inc. *Statistical Analysis and Data Mining* 2: 175–185, 2009

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## 1. INTRODUCTION

The Developmental Therapeutics Program (DTP) of the National Cancer Institute has utilized a panel of 60 human tumor-derived cell lines to assess the cytotoxic activity of more than 75 000 compounds, in what is commonly known as the “NCI60 cell line screen” [1,2]. These cell lines are derived from various tissue types, and represent nine different cancers: leukemia, melanoma, and solid tumors of the lung, colon, brain, ovary, breast, prostate and kidney. The expression of a large number of genes has been measured in these cell lines using various transcriptional profiling techniques. As the gene expression is measured before exposure to any of the chemical agents, it can be used to identify biological factors predisposing a cell line to sensitivity or resistance to a particular compound or class of compounds. In particular, the NCI60 cell lines have been used as a platform for analyzing relationships between the selective

activity of potential anticancer agents, and the expression of genes encoding drug targets, drug transporters, and proteins involved in drug activation, cellular stress response, and detoxification [3–9].

A natural first step toward understanding this dataset is to identify pairwise associations between the expression of individual genes and the activity of individual compounds. Subsequently, higher order relationships involving multiple genes and/or compounds can be pursued [10–12]. In this paper, we focus on the identification of pairwise associations. On the basis of our statistical findings, we can consider a sequence of important scientific questions:

- *Do numerous distinct genes show associations with the activity of at least one compound, and do numerous distinct compounds show associations with the expression of at least one gene?* The answer to this question tells us in broad terms whether gene expression is an important determinant of

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chemosensitivity, and whether the influence of gene expression on chemosensitivity is or is not restricted to a few key genes.

- *Are associations between gene expression and compound activity highly specific, in that small groups of genes are uniquely associated with small groups of compounds; or are they more general, in that a single compound is often associated with many genes and a single gene is often associated with many compounds?* The answer to this question is informative about whether the mechanisms by which genes influence compound activity are highly specific, for example, through binding or molecular recognition events, or alternatively are broad, as in a particular gene marking a cell state that results in sensitivity or resistance to a diverse class of compounds.
- *Do genes act in a uniform manner, by always increasing or always decreasing chemosensitivity, or do they act differently for different compounds?* The answer to this question tells us whether the consequences of an individual gene's expression are mechanistically simple, in that the effect of the gene is the same for all associated compounds, or are mechanistically complex, in that different compounds are affected in qualitatively different ways.

Our ability to accurately identify associations between gene expression and compound activity, and consequently to address the three questions posed above, is largely limited by the diversity of the cell lines in the screen, and by the measurement accuracy of the experiments. Cell line diversity is important; because we will be unable to detect associations with biological states that are not represented in the assayed panel. Thus, the associations we identify here represent a subset of all true associations, and in particular some classes of compounds that show no associations in these data may show strong and interesting associations in a more biologically diverse collection of cell lines. Measurement error limits the identification of associations as it diminishes the apparent strength of true associations. In addition, as our goal will be to focus on associations that occur when gene expression and toxicity vary beyond some fixed level, the presence of measurement error makes it more difficult to accurately identify which genes and compounds have the required level of variability in their true, as opposed to measured, expression and activity levels. In this paper, we aim to identify associations between gene expression and compound activity, and subsequently to answer as fully as possible the three questions posed above, while accounting for and where possible compensating for the limited measurement accuracy in the data.

## 2. METHODS

### 2.1. Experimental Data

Compound activity data are available from the National Cancer Institute (NCI) Developmental Therapeutics Program (DTP) Human Tumor Cell Line Screen ([http://dtp.nci.nih.gov/docs/cancer/cancer\\_data.html](http://dtp.nci.nih.gov/docs/cancer/cancer_data.html)) [1]. We used the July 2007 data release. We used data for the 59 cell lines that had more than 1000 GI50 values and that also had microarray gene expression data (the MDA-N cell line was excluded). Compound activity results are reported as the negative of the  $\log_{10}$  GI50, where GI50 is the compound concentration required to inhibit cell growth by 50%. Compounds are tested in a series of dilutions with maximum concentration typically at  $10^{-4}$  M. Microarray gene expression data on the NCI cell line panel from multiple sources are available through the DTP web site (<http://dtp.nci.nih.gov/mtargets/download.html>). We used three gene expression datasets provided by GeneLogic and Novartis. The two GeneLogic datasets are single replications per cell line, one using the Affymetrix U133 microarrays and one using the Affymetrix U95A/B microarray. The Novartis data is the average of three replicate assays per cell line using the Affymetrix U95A microarray. The probe sets in the Novartis data are a subset of those in the GeneLogic U95A/B data, with only the A chip being used. A fraction of the GI50 data points are missing for quality control reasons. There are no missing values in the microarray data.

### 2.2. Filtering Genes and Compounds

All the analyses were performed on log-scale data ( $\log_{10}$  for GI50,  $\log_2$  for gene expression). Compounds were excluded if they had fewer than 50 experimental GI50 measurements for the 59 cell lines (the threshold was changed from 50 to 45 when the leukemia cell lines were omitted, see below for details). Compounds were also excluded if their  $\log_{10}$  GI50 standard deviation was below 0.2. When the standard deviation is 0.2, the probability that a particular GI50 value will deviate by more than twofold (higher or lower) from the mean value is approximately 1/8. Thus under this selection criterion, at least 1/8 of the (at most 59) GI50 values for a given compound will vary at least twofold from the mean. We chose the fraction 1/8 because it approximately represents the fraction of the total data from one tissue type. GI50 standard deviations were estimated using the robust biweight midvariance estimator [13]. Expression data for a given probe set was excluded if the standard deviation for  $\log_2$  data was below 0.65. The interpretation of this threshold is analogous to the threshold 0.2 used for the GI50 data; the numerical threshold differs only due to

the change in scale. To enable direct comparison, the standard deviation thresholds 0.2 and 0.65 were also used when the leukemia cell lines were omitted (see below).

### 2.3. Identifying Associations

All probe set/compound pairs that passed the filtering step were evaluated for association using the biweight mid-covariance [13] to robustly estimate correlation coefficients without excessive loss of statistical efficiency. All variances and covariances were calculated using a robust procedure as in earlier work with these data we found that many of the strongest Pearson correlations between gene expression and compound activity were due to a single cell line's data. While these "outlying" measurements may reflect true expression or activity values, we wanted to focus here on the associations that are reflected in multiple cell lines.

All available data for each compound/gene pair were used to calculate the correlation estimates (data for between 50 and 59 cell lines if all tissue types are included, or for between 45 and 51 cell lines if the leukemia panel is omitted). Estimated correlation values greater in magnitude than thresholds of either 0.5 or 0.55 were retained for subsequent analysis as described below. These thresholds were chosen as they gave false discovery rate (FDR) values around 0.05 and 0.01, respectively, in the A dataset. Owing to the coding of GI50 as  $-\log(\text{GI50})$ , a positive association means that higher gene expression is associated with chemosensitivity, and a negative association means that higher gene expression is associated with chemoresistance.

### 2.4. False Discovery Rate Analysis

The null expected number of associations and FDR were obtained by simulating independent standard normal value pairs (corresponding to the GI50 and expression data) and estimating their correlation coefficient using the biweight procedure discussed above. The sample sizes for the simulation analysis (the number of value pairs corresponding to cell lines with observed data) were frequency-matched to the distribution of sample sizes in the actual data. We simulated  $10^7$  sets of value pairs to estimate the null expected number of associations. False discovery rates were then estimated using the simple expression  $N_n/N_o$ , where  $N_o$  is the number of associations identified in the data, and the expected number of false positives  $N_n$  is estimated as  $N_t p_0$ , where  $N_t$  is the number of tests and  $p_0$  is the null probability of an association. Data from other distributions, and permuted experimental data were also used to explore the sensitivity of  $p_0$  to the data distribution. We elected to use this simple approach to FDR analysis, similar to the approach of Benjamini and Hochberg [14] rather than more sophisticated approaches [15,16] as the number of

tests is very large and the proportion of positives is very small ( $\ll 1\%$ ).

### 2.5. Tissue Type Specificity

Many genes show strong differential expression across the tissue types, and many compounds show strong patterns of tissue-specific activity. We aimed to identify associations between gene expression and GI50 that are primarily due to tissue type, as well as those that are not. Therefore, the filtering and association identification steps were carried out three times, as follows. One set of results, denoted A below, was obtained using all cell lines and compounds that pass the filtering step. A second set of results, denoted C below, was obtained after mean-centering gene expression and GI50 values within each tissue type. Because a major component of the tissue-specific effects are due to leukemia cells (and the underlying difference between cells grown on a solid surface and cells grown in suspension), a third set of results, denoted S below, was generated omitting the leukemia cell lines. For centering, the tissue types were assigned according to the updated designation of MDA-MB-435 and NCI/ADR-RES as melanoma and ovarian cancer in origin, respectively [17,18].

## 3. RESULTS

### 3.1. Overall Statistical Significance

We began by assessing the statistical strength of apparent associations in the data. Table 1 summarizes our association findings in each dataset, and for each of the inclusion and processing rules A, C, and S defined in the methods. For example, for the A data from the GeneLogic U133 platform with threshold 0.5 (first row of Table 1), the null probability of an association was  $p_0 = 4.4 \times 10^{-5}$ , so the expected number of false associations is  $26391 \times 11196 \times 4.4 \times 10^{-5} = 13001$ . As we observed 308 165 associations, the FDR is estimated as  $13001/308165 \approx 0.04$ .

To assess the sensitivity of these results to the choice of a normal distribution for determining the null probability of an association  $p_0$ , the analysis was repeated with other forms of null data. First, several parametric distributions were used, including Bernoulli trials with equal probabilities (0.5/0.5) and unequal probabilities (0.75/0.25), various finite mixtures of normals, a continuous uniform distribution, and standard exponential and Cauchy distributions. Most of these distributions gave fewer positives calls than in the normal case, suggesting that the use of a normal reference distribution is conservative if it is biased at all. Only a strongly skewed distribution (exponential) and a heavy-tailed distribution (Cauchy) gave more positive calls than

**Table 1.** Summary of association results compared to null expected findings. The primary datasets are coded in the first column as 1: GeneLogic U133, 2: Novartis U95A, 3: GeneLogic U95A/B. Three data filtering and processing schemes are indicated in the third column: all data were used without centering (A), all data were used following centering within tissue types (C), or only data from solid tumor-derived cell lines were used without tissue type centering (S).

Data set	Threshold	Processing	#Probe sets	#Compounds	#Assoc.	Null exp. #Assoc.	FDR(98%)
1	0.5	A	26 391	11 196	3 08 165	13 001	0.04(0.11)
2	0.5	A	5 144	11 196	1 45 365	2 534	0.02(0.08)
3	0.5	A	35 196	11 196	2 96 205	17 338	0.06(0.14)
1	0.55	A	26 391	11 196	1 02 098	1 861	0.02(0.09)
2	0.55	A	5 144	11 196	50 018	363	0.01(0.08)
3	0.55	A	35 196	11 196	92 515	2 483	0.03(0.12)
1	0.5	C	4 601	11 117	11 601	13 810	0.84
2	0.5	C	988	11 117	3 692	2 966	0.80
3	0.5	C	20 488	11 117	43 353	61 497	1.00
1	0.55	C	4 601	11 117	2 490	3 375	1.00
2	0.55	C	988	11 117	850	725	0.85
3	0.55	C	20 488	11 117	9 128	15 032	1.00
1	0.5	S	25 657	9 484	2 12 117	58 399	0.28(0.48)
2	0.5	S	4 833	9 484	82 899	11 001	0.13(0.36)
3	0.5	S	34 883	9 484	2 32 279	86 016	0.37(0.60)
1	0.55	S	25 657	9 484	66 538	7 300	0.11(0.34)
2	0.55	S	4 833	9 484	28 239	1 375	0.05(0.29)
3	0.55	S	34 883	9 484	68 754	9 924	0.14(0.41)

in the normal case, by around 25 and 50%, respectively. However, consideration of skew coefficients and tail indices in the data do not suggest that either the gene expression or GI50 data are often as skewed as an exponential distribution or as heavy-tailed as a Cauchy distribution. We also used a simple permutation approach, in which genes and compounds were randomly selected, and the GI50 values for the compound were randomly permuted prior to calculating the correlation coefficient. These results were within 10% of the value for the normal case (on the low side), which is not a statistically significant difference with respect to Monte-Carlo error, based on  $10^7$  permutations as we performed.

The number of positives for the A data tends to be an order of magnitude or more greater than the number for the C data, whereas the number of positives drops by around half when comparing the C to the S data. This suggests that a large fraction of associations are at least partially related to tissue-specific cell growth inhibition, and that not all of these effects are explained by the differences between solid tumor-derived and leukemic cell lines. For all three datasets, the number of positive calls as a function of the correlation threshold is steeply decreasing in the range of thresholds considered —there is around an order of magnitude drop in the number of positives as the threshold is raised from 0.5 to 0.55, even though scatterplots in this range of correlation values tend to be visually indistinguishable.

We performed a limited analysis of the structures of the compounds involved in associations to assess whether a restricted class of chemical structures gives rise to a

large fraction of the associations. Out of the 11 196 compounds passing the filtering step, 8605 had CACTVS fingerprints available from PubChem. Using these fingerprints, we calculated Tanimoto similarity measures between all compound pairs within the 5879 compounds involved in associations, and between all compound pairs within the 2726 compounds not involved in associations. The mean and standard deviation of Tanimoto coefficients in these two groups, respectively, are 0.56(0.24) and 0.42(0.15). Thus the compounds involved in associations are somewhat less diverse than the set of compounds that are not involved in associations, suggesting that some common structural families are contained within the set of compounds whose activity is influenced by gene expression. For example, P-glycoprotein substrates are a structurally restricted class of compounds that are selective against cell types lacking P-glycoprotein mediated drug efflux activity, which is reflected in expression of the ABCB1 gene.

The results in Table 1 indicate that for the A data, associations between GI50 and gene expression can be identified with high accuracy, whereas for the C data there is insufficient power to identify any true associations that might exist. For the A data, associations appear to be reliably identifiable at a correlation threshold of 0.5. In the S data, there is strong evidence that numerous associations exist (as the number of strong associations is far greater than expected), but at the 0.5 correlation threshold the number of false positives will be high. For the remainder of the analysis, we will use the 0.5 threshold for the A data and the 0.55 threshold for the S data. The Novartis data

has the lowest number of genes passing the filtering step among the three platforms, and also consistently has the lowest FDR value. As the Novartis data is the average of three experiments, it presumably has the least measurement error. This suggests that a number of spurious associations in the GeneLogic datasets are due to measurement error. Nevertheless, even the GeneLogic datasets have favorable FDR values, so we will continue to use all three datasets in the analysis.

To further assess the significance of these results, we hypothesized a mean/variance relationship for the null distribution of the number of gene/compound pairs associated at a given threshold level. As a working relationship, we took  $\text{variance} = f \times \text{mean}$  for some  $f > 1$ . If there were no dependence within the genes or within the compounds, and if each gene and each compound were tested only once (rather than in all gene/compound pairs), the null distribution would be Poisson and we would have  $f = 1$ . Because the C data showed no evidence of associations, we used these data to estimate  $f$ . Specifically, we used simple linear regression to regress  $(Y - X)^2$  against  $X$ , where  $Y$  is the observed number of associations and  $X$  is the null mean. The regression fit over the 6 points for the C data (as given in rows 7–12 of Table 1) had an adjusted  $R^2$  of 0.94, and an intercept that was not significantly different from zero ( $p > 0.1$ ), consistent with the hypothesized mean/variance relationship. The estimated slope was highly significant ( $p < 10^{-3}$ ), and gave an estimate  $\sqrt{f} = 90$  which we used as an inflation factor for standard deviations (i.e. estimating the standard deviation as  $90 \times \sqrt{\text{mean}}$ ). In this way, we calculated the 98<sup>th</sup> percentile of the FDR distribution for the A and S datasets with each threshold as the mean plus two standard deviations. For the A data with correlation threshold 0.5, these values are below 0.15, and for the S data with correlation threshold 0.55 these values are below 0.41. From this, we conclude that associations identified from the A and S data are unlikely to be dominated by false positives.

### 3.2. Positive Controls

As a positive control, we considered the relationship between triciribine (TCN, NSC154020) or triciribine phosphate (TCN-P, NSC280594) and adenosine kinase (ADK). Expression of ADK is necessary for the growth inhibitory activity of TCN and TCN-P, as it participates in the phosphorylation/ dephosphorylation of TCN that converts it between its toxic and cell-permeable forms [19,20]. The U95 array and U133 array each contain two probe sets for ADK. On the U133 array, one TCN/ADK pairing (NSC154020, 204119\_s\_at) had a correlation coefficient of 0.60, and another (NSC280594, 204119\_s\_at) had a correlation of 0.50. On the GeneLogic U95 array, the

**Table 2.** Percentages of genes associated with chemosensitivity or resistance for at least one compound. The datasets (1, 2, 3) and data filtering/processing schemes (A, C, S) are coded as in Table 1. Results are shown for correlation thresholds 0.5 and 0.55.

	0.5			0.55		
	1	2	3	1	2	3
A	64	79	60	33	53	28
C	66	69	63	26	30	24
S	80	89	79	46	62	43

pairing (NSC154020, 168\_at) had a correlation of 0.54. These signed correlations are in the expected direction. The correlations given are for the A data, and are equal or slightly stronger in the S data. In the Novartis U95 dataset, the two probe sets for ADK did not meet our filtering criteria.

### 3.3. Proportion of Measured Genes Associated with Chemosensitivity or Chemoresistance

The percentages of probe sets that passed the filtering steps and that are associated with at least one compound are shown in Table 2. Given the low false discovery rates for the A and S data, this suggests that a substantial fraction of all genes that vary in expression in the NCI cell line panel are truly associated with the activities of one or more compounds. Indeed, if the false positives are distributed independently of the genes, the proportion of genes associated with at least one compound drops by only 1–2% for the A data and by 4–5% for the S data. These percentages are obtained by considering the number of compounds associated to a particular gene, then calculating the probability that all of the associations are false positives based on the FDR values given in Table 1. Specifically, if the FDR is  $f$  and a given gene is associated to  $m$  compounds, the probability that the gene has at least one true association is  $1 - f^m$ . The values 1–2% and 4–5% cited above are expected values calculated by summing the probabilities  $1 - f^m$  over all genes. The reduction is smaller than the FDR as most genes are associated with multiple compounds.

For both the A and S data, the Novartis dataset with its presumed lower level of measurement error had the fewest probe sets passing the filtering, but the greatest proportion of nonfiltered probe sets involved in associations. The observation that fewer probe sets pass the filtering in the Novartis set is likely due to the unreplicated GeneLogic datasets overcalling genes with expression variation. If true associations are less common among the less variable genes, the proportion of genes with at least one association will be lower in the unreplicated datasets. The observation that the Novartis data has the greatest proportion of

nonfiltered probe sets involved in associations can also be explained in terms of measurement error. If the false positives are distributed randomly across the genes and compounds, and given that associations are rare given the number of tests performed, false positives will tend to introduce genes with associations to only one compound. This is discussed further below. Notably, when using the Novartis data, the fraction of substantially varying genes that are associated with GI50 for at least one compound exceeds  $1/2$ . This suggests that at a minimum, around half of measured genes are associated with compound activities. However, it is important to note that an unknown fraction of these associations reflect indirect rather than causal relationships.

### 3.4. Specificity of Associations

Figure 1(a) shows the distribution of the number of compounds associated with each nonfiltered probe set (restricted to probe sets associated with at least one compound). The results shown in the figure are for the A data, from each of the three array platforms. Probe sets associated with many compounds are likely to represent one or more mechanisms of selective activity or resistance. All three datasets have a small number of probe sets associated with more than 100 compounds, with a few genes being associated with more than 500 compounds. The Novartis dataset has the greatest fraction of probe sets associated with large numbers of compounds. Conversely, the two unreplicated GeneLogic sets have a greater fractions of genes associated with only one or two compounds.

As noted above, measurement error in the gene expression and GI50 measurements is likely to increase the fraction of isolated associations between genes and compounds

relative to situations where a gene is associated with numerous compounds. Consistent with this, the Novartis data has the lowest proportion of genes with small numbers of associated compounds.

Although the Novartis data presumably has the lowest measurement error among the expression datasets considered here (as it is an average of replicates), its measurement error is not zero, and there is also measurement error in the GI50 data. Therefore, we sought to understand what the distribution of the number of associated compounds per gene would look like in the complete absence of experimental measurement error in both the gene expression and GI50 data. Using the original data from which the Novartis averages were obtained, the measurement variance in the averaged Novartis data can be estimated as  $0.08 = 0.24/3$ , where 0.24 is the measurement variance within replicates averaged over the genes. Unlike the gene expression data, the GI50 data are not systematically replicated. However, there are around  $10^5$  compound/cell line pairs for which at least two experimental points are available. From these, we estimated the GI50 measurement error variance as 0.45, averaged over the compounds.

Given nominal levels for the gene expression and GI50 measurement variance, it becomes possible to employ a procedure analogous to the simulation-extrapolation (SIMEX) procedure [21], which is used to adjust coefficient estimates in regression models for the effects of measurement errors. To do this, we added simulated centered normal errors with variances 0.08 and 0.45 to the Novartis gene expression and the GI50 data, respectively (following the A processing). Then we recalculated all associations from these data. If  $f$  is the estimated density for the experimental data, and  $g$  is the estimated density after adding additional simulated errors of a similar magnitude to the

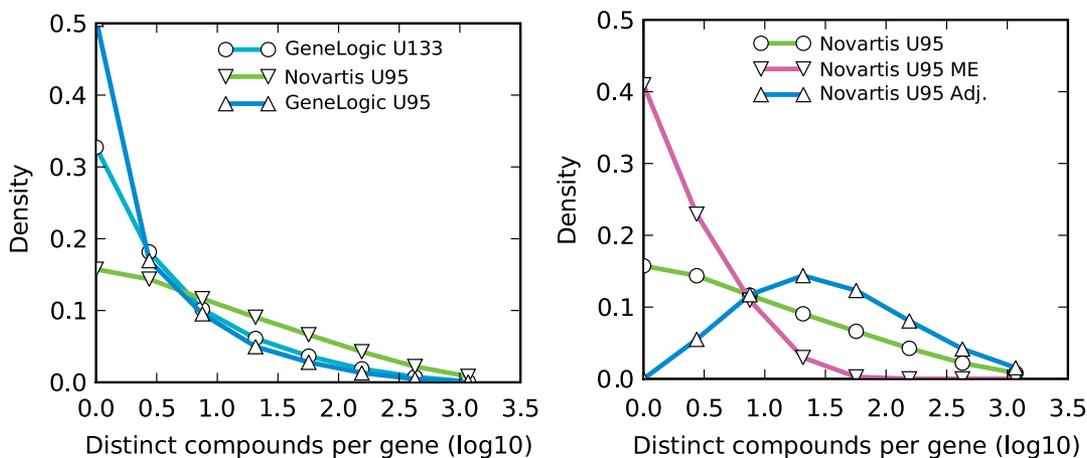


Fig. 1 (a) Estimated distribution of the number of associated compounds per gene for each of the three experimental datasets. (b) The experimental distribution of associated genes in the Novartis set, along with the number of associations observed when extra measurement error is added (Novartis U95 ME), and the adjusted distribution estimating what would be found in the absence of measurement error (Novartis U95 Adj.).

actual measurement errors, then  $g - f$  estimates the bias in the density estimate due to the measurement error. Thus  $f - (g - f) = 2f - g$  can be used as an approximately unbiased estimate of the density of pairwise associations. For interpretative purposes, we can inspect the graph of  $2f - g$ , but if it is desired to make this a proper density, we can truncate at zero and define  $f^* = (2f - g) \vee 0$ , then rescale so that  $f^*$  integrates to one. This estimate is justified under an assumption that for the magnitudes of measurement errors under consideration, the effect of measurement error on the density is approximately linear in a pointwise sense.

Figure 1(a) shows the empirical distributions of the number of compounds associated to a given gene for all three datasets, and Fig. 1(b) shows the results of the bias correction analysis for the Novartis dataset. The bias correction is only applied to the Novartis dataset as it is the only dataset where we have replication, allowing the measurement error variance to be directly estimated. The curve denoted with the  $\circ$  symbol in Fig. 1(b) is the unadjusted density estimate for the number of compounds associated to a given gene (identical to curve denoted with the  $\nabla$  symbol in Fig. 1(a)). The curve denoted with the  $\nabla$  symbol in Fig. 1(b) results from adding additional measurement error to the data, as dictated by the SIMEX procedure. Note that the density of pairwise associations becomes much higher at zero. Consequently, the bias-corrected distribution for the number of compounds associated to a gene when no measurement error is present (denoted by the  $\triangle$  symbol in the right panel in Fig. 1(b)) is markedly different from the direct estimate shown in Fig. 1(a). The direct estimate has a mode at zero, whereas the adjusted estimate has a mode at around 1.4 (on the log10 scale), and has very little mass at zero. Thus, while an analysis ignoring measurement error suggests that the associations are dominated by isolated gene/compound pairs, consideration of the measurement error suggests that very few such isolated pairs actually exist, and that a typical gene that is associated to at least one compound is associated to around  $10^{1.4} \approx 25$  compounds.

### 3.5. Balance Between Genes Conferring Chemosensitivity and Chemoresistance

Positive and negative directions of association between gene expression and  $-\log(\text{GI50})$  are indicative of different mechanisms of action. Genes associated with upregulation of detoxification mechanisms such as drug efflux and catalytic deactivation should lead to negative associations, but in general both directions are possible. For the A, C, and S datasets, the genes associated with  $-\log(\text{GI50})$  at either a 0.5 or 0.55 threshold are nearly perfectly balanced between positive and negative associations. For genes that

are associated with multiple compounds, the directions of associations can provide us with an indication of whether the gene acts with a uniform mechanism for all the compounds. If the gene acts via the same mechanism for all compounds, we expect the direction of association to be the same (either positive or negative) for all associated compounds.

We assessed whether there is evidence that individual genes exhibit both positive and negative associations with different sets of compounds. To examine this, we identified the probe sets that are associated with at least 20 compounds in a single direction in the A dataset at threshold 0.5. There are 2466, 1134, and 2397 probe sets satisfying this condition from the GeneLogic U133, Novartis U95A, and GeneLogic U95A/B arrays, respectively. The overlap between the Novartis and GeneLogic U95A arrays is 541 probe sets, or roughly half of the list from the Novartis data. The average proportion of associations within each of these probe sets that were oriented in the less frequent direction for that probe set was 0.12, 0.12, and 0.11 in the three datasets. In other words, almost 90% of the associations are in the same direction, on average.

We next aimed to determine whether the observed level of balance between positive associations and negative associations within a gene was consistent with the true associations for the gene being either entirely positive or entirely negative. We performed this by correlating the actual GI50 data to simulated independent and identically distributed (iid) standard normal data as a proxy for gene expression when no association is present. This provided us with estimates of the mean and variance of the number of false positives when testing a single gene against the 11 196 compounds in set A. The variance in the number of positives (based on threshold 0.5) was 22 times the mean number of positives. This substantial overdispersion relative to the Poisson distribution is due to the clustered nature of the compound set, which largely consists of small sets of related compounds with similar GI50 profiles (note that the overdispersion here is much less than found above when testing all genes against all compounds). From the simulation described above, the probability of a false association with either a positive or negative sign for a particular gene and compound is  $p_0 = 4.4 \times 10^{-5}$ , so the expected number of false associations with a particular sign for a given gene in the A data tested against all compounds is  $11196 \times 2.2 \times 10^{-5} = 0.25$ . Thus the upper 99.9th percentile for the number of false associations with a particular sign is approximately  $0.25 + 3\sqrt{0.25 \times 22} = 7.3$ .

On the basis of this analysis, genes with seven or fewer compounds associated in a negative direction can be presumed to have only positive true associations, and genes with seven or fewer compounds associated in a positive direction can be presumed to have only negative true

**Table 3.** Summary of results on uniformity of associations.  $N_+$ : genes with at least 20 positive associations;  $N_-$ : genes with at least 20 negative associations;  $N$ : genes with at least 20 positive or at least 20 negative associations;  $N_{++}$ : genes with at least 20 positive and at most 7 negative associations;  $N_{--}$ : genes with at least 20 negative and at most 7 positive associations.

Data	$N_+$	$N_-$	$N$	$N_{++}$	$N_{--}$
1	1 195	1 776	2 466	535	1 033
2	608	777	1 134	265	416
3	1 220	1 600	2 397	657	970

associations. Table 3 shows the results of this analysis. Among genes in the A data having 20 or more associations in at least one direction, the fraction having seven or fewer associations in the less common direction are 0.64, 0.60, 0.68, in the GeneLogic U133, Novartis U95A, and GeneLogic U95AB datasets, respectively (calculated as  $(N_{++} + N_{--})/N$  in Table 3). We conclude that the experimental data for these genes are consistent with a uniform mechanism of action for all associated compounds, although uniformity in the direction of association is not sufficient to conclude this. Conversely, we can infer that the 30–40% of genes that are associated with multiple compounds in differing directions are unlikely to do so with a uniform mechanism of action.

### 3.6. Genes Associated with Many Compounds

Some of the genes associated with GI50 for large number of compounds are involved in cellular processes identified as being related to small molecule toxicity in previous studies using the NCI60 cell line screening data. For example, in the S dataset, only three probe sets are associated with the GI50 of more than 300 compounds. Two of these three probe sets are for the cysteine/glutamate transporter SLC7A11, previously identified as being involved in chemoresistance [3,22]. In the A dataset, five separate probe sets for ATP-binding cassette C3 (ABCC3) are each positively associated with chemosensitivity for more than 250 compounds. ABCC3 is well known as playing a major role in multidrug resistance [23].

We analyzed the set of genes having 100 or more associations in at least one of the datasets using the DAVID/EASE functional classification tool [24]. For genes whose expression was associated with chemosensitivity, eight functional clusters were identified: these include clusters of various actins, integrins, and collagens. The clearest theme in the eight clusters is that they contain numerous genes related to maintenance of the extracellular membrane, and more generally to a differentiated cell state. For genes whose expression was associated with chemoresistance, six clusters were identified. These included clusters centered on

growth factors, transcription factors, and genes associated with cell signaling and proliferation.

For detailed manual analysis, we focused on the genes whose expression is associated with 200 or more compounds in the S data. A total of 146 probe sets on the three arrays satisfy this condition. Among the genes represented by these probe sets, there is a consistent trend of signal transduction genes being upregulated in cells with a resistant phenotype, and extracellular matrix and cytoskeletal genes being upregulated in cells with a sensitive phenotype. The subset of these genes associated with 200 or more compounds in at least two of the three S datasets is shown in Table 4.

Signal transduction genes whose expression is associated with chemoresistance include the epidermal growth factor receptor (EGFR), erythroblastic leukemia viral oncogene homolog 3 (ERBB3), cysteine-rich transmembrane regulator (CRIM1), fasciculation and elongation protein zeta 2 (FEZ2), tumor necrosis family receptor superfamily

**Table 4.** Genes associated with at least 200 compounds in at least two of the three datasets. Datasets 1, 2, 3 are denoted along the top margin as in Tables 1–3. For each dataset, three numbers are given: the number of distinct probe sets showing an association, the percentage of associations that are positive, and the total number of associations. The percentage and total are calculated over all compounds associated to any probe set for the gene.

	1			2			3		
CRIM1	2	20	474	1	4	212	1	19	284
MYLK	1	76	268	1	86	239	2	67	558
FHL1	1	65	232	1	74	200	1	61	224
MIPEP	1	73	276	1	67	248	1	37	281
LOXL2	1	59	240	1	76	285	1	58	208
TRPC1	1	86	260	2	80	472	1	82	232
CYR61	2	49	545	1	67	231	1	51	269
NMT2	1	31	220	1	11	202	1	29	217
ANLN	1	32	205				1	37	258
GLRB				1	85	259	1	81	265
PDGFC	2	78	482				1	69	251
CLIP4	1	1	208				1	0	219
COL4A1	2	74	497				1	63	252
C6ORF192	1	26	208				1	31	233
PEA15				1	63	233	1	30	210
THBS1	3	63	755				2	75	492
PALLD	2	82	457				1	87	230
BCAT1	1	60	281				2	51	479
JUB	1	17	232				1	20	290
MYH9	1	19	278				1	12	312
IL1R1				1	65	224	1	68	219
POLR1D	1	90	205				1	85	249
EGFR	2	5	445				2	8	461
TNFRSF12A	1	19	277				1	17	200
NUAK1	1	72	264				1	87	210

member 12A (TNFRSF12A), ubiquitin-associated SH3 domain containing protein A (CLIP4), and hypoxia inducible factor alpha (HIF1A). Interestingly, CLIP4 is a regulator of EGFR signaling, that acts by inhibiting internalization of EGFR into endocytic vesicles and therefore blocking its degradation. Also, the FEZ2 gene is a regulator of protein kinase C, which is downstream of the EGFR. CRIM1 has been implicated in growth factor secretion and HIF1A is an important signal transduction molecule that triggers an angiogenic response and induces drug resistance when it is upregulated in cancer cells in response to hypoxic conditions. Therefore, these genes are potentially linked to a common pathway involved in autocrine or paracrine growth factor signaling mechanisms associated with angiogenic phenotypes, which are upregulated in aggressive cancer cells.

As noted above, expression of extracellular matrix and cell adhesion genes tend to be associated with drug sensitivity. These include collagen 7A1 (COL7A1), one of the main extracellular matrix components of the basement membrane; dentin matrix protein (FAM20C) a component of the tooth extracellular matrix; versican (VCAN), which is chondroitin sulfate proteoglycan that binds hyaluronic acid and is a component of cartilage; fibrillin 1 (FBN1) which is the major component of extracellular matrix microfibrils; myosin light chain kinase (MYLK), which is associated with regulating contractility and cell shape; and palladin (PALLD), which is a component of actin stress fibers and has a role in cell adhesion. In addition to these genes, the integrins ITGAV and ITGA5 which are the vitronectin and fibronectin receptors are associated with drug sensitivity, although the overexpression of another integrin ITGA3 is more associated with resistance. Extracellular matrix and adhesion genes are likely to be associated with more differentiated cancer cells with anchorage-dependent growth regulation.

#### 4. DISCUSSION

We have identified several attributes of the NCI/DTP screening data that are not apparent from simple tabulation of pairwise associations. First, and least surprisingly, we found that if no attempt is made to remove associations that may be due to tissue type differences, a large number of associations can be identified with only a modest rate of false positives. However, as many associations that depend on tissue type variation may reflect confounding effects of tissue-specific gene expression, it is desirable to identify associations that are not completely due to the differences between the tissues. Our results suggest that attempts to minimize tissue-specific effects should be considered in the context of statistical power. Mean centering each cell

type in the panel is the most direct way to completely remove tissue-specific effects, but it appears to remove so much of the association signal that little power remains. It is likely that mean centering the tissues overcorrects for tissue-specific expression, and moreover, not all tissue-specific expression is uninformative about mechanisms of activity. As a less drastic strategy, focusing on solid tumors by removing leukemic cell lines enriches for relationships that are less tissue specific, at a modest cost in power.

The most important and original finding reported here is that accounting for the effects of measurement error in the gene expression and compound activity data dramatically changes the apparent distribution of the number of compounds associated to a given gene. To put this into context, a number of previous publications have focused on associations between single genes and single compounds, which are presumably indicative of chemically specific relationships, perhaps mediated through molecular recognition or binding. Other reported associations have been more general, including the identification of efflux and detoxification systems for which numerous compounds can serve as substrates. Our findings reported here suggest that genes truly associated with a single compound are rare.

An important issue that must be considered when analyzing the data considered here is the inhomogeneous and clustered variation among the compound structures and among the genes. Much of our analysis makes use of FDR. The expected value of the FDR estimator has been shown to be insensitive to the presence of many forms of dependencies among the test statistics [25]. The variability of the FDR estimator is inflated by the presence of clusters of related compounds and related genes. By calculating an overdispersion factor, we were able to account for this clustering when estimating upper confidence limits for the FDR.

While the FDR is mostly immune to correlations within the genes and compounds, the interpretations of some other quantities discussed here do depend on the composition of genes and/or compounds in the data set. For example, the distribution of the number of compounds associated to a gene shown in Fig. 1 should be interpreted relative to the genes that are represented on these particular microarray platforms, and that passed the filtering steps in our analysis. These genes may not be perfectly representative of the human genome. However, the microarrays are intended to be as comprehensive as is technically practical. Regardless of the composition of the set of genes that are analyzed, the trend in Fig. 1(b) toward a distribution with greatly reduced mass at zero should occur for any sampling of genes.

A similar issue arises with the compounds. The compounds in the DTP screening set are diverse, but the set does contain clusters of related structures, particularly around approved drugs. However, we estimated the mean number

of associated compounds per gene as 30, and clusters of 30 or more closely related compounds are uncommon. Therefore, it is unlikely that the gene/drug association clusters we have shown to exist are exclusively due to clustering of structural analogs in the compound set. Nevertheless, in future work it would be valuable to consider the diversity among chemical structures that are associated to a common gene.

A separate consequence of measurement error is that it degrades our ability to identify the most variable genes and compounds. Comparing our findings with replicated and unreplicated data, it appears that the unreplicated datasets substantially overcall the set of genes with a given threshold level of variability. The fact that our positive control gene ADK did not vary sufficiently in the Novartis data to be tested for associations suggests that the variance thresholds we chose were too high to capture all mechanistic relationships. However, the false positive rate presumably increases as the threshold is lowered. Thus, the situation with ADK may best be viewed as reflecting the limits of power of this dataset. Also of note is that previous work involving the association between ADK and TCN [7] began with a subset of the NCI screening compounds consisting of well-characterized compounds. Thus, the multiple testing issue in that context was of a much lower order, and the success of that analysis does not imply that such associations should be detectable in the entire dataset.

The balance between associations in the positive and negative directions within a gene can be informative about the uniformity of the mechanism underlying the associations. If substantial numbers of associations occur in both directions, multiple mechanisms are likely present, whereas predominance of one direction is consistent with a uniform mechanism. Our results suggest that more than half of genes may act with a uniform mechanism.

Finally, we performed systematic and manual analyses of genes associated with large numbers of compounds. These genes are presumed to be involved with general, less specific mechanisms of action. In contrast, genes associated with only a few compounds are more likely to engage in chemically specific interactions [7,9]. The genes associated with large numbers of compounds include transporters already identified as involved in drug efflux and cell detoxification [3,22]. In addition, we observed that the expression of a set of genes relating to EGFR signaling is related to chemoresistance, whereas genes relating to the extracellular membrane and cell adhesion are associated with chemosensitivity. The latter association may reflect a general tendency for the less differentiated cell lines, presumably derived from more advanced cancers, to be broadly less sensitive to cytotoxic agents.

In terms of general methodology for high throughput screening and biomolecular assay data, we have

demonstrated that the use of false discovery rates together with straightforward techniques to account for the effects of measurement error can yield insights about how associations among variables are distributed in a large and complex dataset. Our approach utilizes a simple formulation of FDR that was calibrated using simulations (due to the lack of an analytic variance formula for the robust association estimators we used). To accommodate the clustered nature of the gene expression and compound activity data, overdispersion factors were calculated from the data. We note that this part of the analysis could alternatively have been carried out with a permutation approach. For measurement error analysis, we used direct estimates of the measurement error variance based on replication in the gene expression and compound activity data sets. We then followed a simulation procedure much like the SIMEX approach to determine the distribution of associations that is expected in a dataset that is free of measurement error. We found this to be a straightforward and effective means to clarify the overall pattern of gene/compound associations, and propose that it could be meaningfully applied in a variety of other data analytic settings involving large sets of data measured with error.

Our findings indicate that the NCI60 cell line screen is most effective at identifying associations involving multiple genes or compounds as that provides a form of internal replication. Accordingly, our results point to the types of associations that are most likely to be detectable in this dataset. As it appears that genes associated with compound activity are often associated with multiple, distinct compounds, power for future studies will be maximized by focusing on that subset of the data.

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