

Web-based Supplementary Materials for  
“Semiparametric transformation models for  
semicompeting survival data” by Lin, Zhou, Li and Li

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## Supplementary Material A: Notations

The following notations are used in Theorem 2. Let

$$f^{(k_1, k_2, \dots)}(\mathbf{x}_1, \mathbf{x}_2, \dots) = \frac{d^{(k_1 + k_2 + \dots)} f(\mathbf{x}_1, \mathbf{x}_2, \dots)}{dx_1^{k_1} dx_2^{k_2} \dots}, \quad \gamma_{1i}(x, y) = -\frac{S_{\rho_0}^{(10)}(x - \mathbf{X}'_i \boldsymbol{\beta}_0, y - \mathbf{X}'_i \boldsymbol{\alpha}_0)}{S_{\rho_0}(x - \mathbf{X}'_i \boldsymbol{\beta}_0, y - \mathbf{X}'_i \boldsymbol{\alpha}_0)},$$

$$\gamma_{2i}(x, y) = -\frac{S_{\rho_0}^{(01)}(x - \mathbf{X}'_i \boldsymbol{\beta}_0, y - \mathbf{X}'_i \boldsymbol{\alpha}_0)}{S_{\rho_0}(x - \mathbf{X}'_i \boldsymbol{\beta}_0, y - \mathbf{X}'_i \boldsymbol{\alpha}_0)}, \quad \gamma_{3i}(x, y) = -\frac{\dot{S}_{\rho}(x - \mathbf{X}'_i \boldsymbol{\beta}, y - \mathbf{X}'_i \boldsymbol{\alpha})}{S_{\rho}(x - \mathbf{X}'_i \boldsymbol{\beta}, y - \mathbf{X}'_i \boldsymbol{\alpha})},$$

$$\dot{S}_{\rho}(x, y) = \frac{\partial S_{\rho}(x, y)}{\partial \rho}, \quad \gamma_1\{H_{10}(t)\} = \exp\left(\int_{t_0}^t \frac{E[Y_i(s) d\gamma_{1i}\{H_{10}(s), H_{20}(s)\}]}{E[Y_i(s) \gamma_{1i}\{H_{10}(s), H_{20}(s)\}]}\right),$$

$$\begin{aligned}
\gamma_2(H_{20}(t)) &= \exp \left( \int_{t_0}^t \frac{E[Y_i(s)d\gamma_{2i}\{H_{10}(s), H_{20}(s)\}]}{E[Y_i(s)\gamma_{2i}\{H_{10}(s), H_{20}(s)\}]} \right), \\
\lambda_{2i}(x) &= \frac{\phi(x - \mathbf{X}'_i \boldsymbol{\alpha}_0)}{1 - \Phi(x - \mathbf{X}'_i \boldsymbol{\alpha}_0)}, \quad \lambda_2(H_{20}(t)) = \exp \left( \int_{t_0}^t \frac{E[Y_{2i}(s)d\lambda_{2i}\{H_{20}(s)\}]}{E[Y_{2i}(s)\lambda_{2i}\{H_{20}(s)\}]} \right), \\
\mu(t) &= \int_{t_0}^t \frac{\lambda_2\{H_{20}(s)\} E \left[ Y_{2i}(s) \frac{\partial \mathbf{X}'_i \boldsymbol{\alpha}_0}{\partial \boldsymbol{\Theta}} d\lambda_{2i}\{H_{20}(s)\} \right]}{E[Y_{2i}(s)\lambda_{2i}\{H_{20}(s)\}]}}, \quad \mathcal{K}(s) = E[Y_{2i}(s)\lambda_{2i}\{H_{20}(s)\}], \\
B(s) &= E[Y_i(s)\gamma_{1i}\{H_{10}(s), H_{20}(s)\}], \quad \mathcal{A}(s) = E[Y_i(s)\gamma_{2i}\{H_{10}(s), H_{20}(s)\}], \\
\mathcal{D}_1(s) &= E \left[ \frac{\partial^2 \log \mathcal{L}_i(\boldsymbol{\Theta}_0; H_{10}, H_{20})}{\partial \boldsymbol{\Theta} \partial H_1(U_{1i})} \frac{Y_i(s)}{\gamma_1\{H_{10}(U_{1i})\}} \right] \frac{\gamma_1\{H_{10}(s)\}}{B(s)}, \\
\mathcal{D}_2(s) &= E \left\{ \frac{\partial^2 \log \mathcal{L}_i(\boldsymbol{\Theta}_0; H_{10}, H_{20})}{\partial \boldsymbol{\Theta} \partial H_2(U_{2i})} \frac{Y_{2i}(s)}{\lambda_2\{H_{20}(U_{2i})\}} \right\} \frac{\lambda_2\{H_{20}(s)\}}{\mathcal{K}(s)} - \frac{\mathcal{D}_1(s)\mathcal{A}(s)}{\mathcal{K}(s)} \\
&\quad - \int_s^\tau \frac{\mathcal{D}_1(v)\mathcal{A}(v)}{\mathcal{K}(s)} \frac{\lambda_2\{H_{20}(s)\}}{\gamma_2\{H_{20}(v)\}} d \frac{\gamma_2\{H_{20}(v)\}}{\lambda_2\{H_{20}(v)\}}, \\
C(t) &= \int_{t_0}^t E \left[ Y_i(s) \frac{\partial \mathbf{X}'_i \boldsymbol{\beta}_0}{\partial \boldsymbol{\Theta}} d\gamma_{1i}\{H_{10}(s), H_{20}(s)\} \right] + E \left[ Y_i(s) \frac{\partial \mathbf{X}'_i \boldsymbol{\alpha}_0}{\partial \boldsymbol{\Theta}} d\gamma_{2i}\{H_{10}(s), H_{20}(s)\} \right] \\
&\quad - \frac{\partial \rho_0}{\partial \boldsymbol{\Theta}} E[Y_i(s)d\gamma_{3i}\{H_{10}(s), H_{20}(s)\}], \\
\Delta &= E \left\{ \frac{\partial \log \mathcal{L}_i(\boldsymbol{\Theta}_0; H_{10}, H_{20})}{\partial \boldsymbol{\Theta}} + \int_{t_0}^\tau \mathcal{D}_1(s)dM_i(s) + \int_{t_0}^\tau \mathcal{D}_2(s)dM_{2i}(s) \right\}^2, \\
\Sigma &= E \left\{ \frac{\partial^2 \log \mathcal{L}_i(\boldsymbol{\Theta}_0; H_{10}, H_{20})}{\partial \boldsymbol{\Theta} \partial \boldsymbol{\Theta}'} \right\} + \int_{t_0}^\tau \mathcal{D}_1(s)dC'(s) \\
&\quad - \int_{t_0}^\tau \mathcal{D}_1(s)\mathcal{A}(s) \frac{d\mu'(s)}{\lambda_2\{H_{20}(s)\}} - \int_{t_0}^\tau \int_s^\tau \frac{\mathcal{D}_1(v)\mathcal{A}(v)}{\gamma_2\{H_{20}(v)\}} d \left[ \frac{\gamma_2\{H_{20}(v)\}}{\lambda_2\{H_{20}(v)\}} \right] d\mu'(s) \\
&\quad + \int_{t_0}^\tau E \left[ \frac{\partial^2 \log \mathcal{L}_i(\boldsymbol{\Theta}_0; H_{10}, H_{20})}{\partial \boldsymbol{\Theta} \partial H_2(U_{2i})} \frac{Y_{2i}(s)}{\lambda_2\{H_{20}(U_{2i})\}} \right] d\mu'(s).
\end{aligned}$$

The following notations are used in Theorem 3. Denote

$$\begin{aligned}
\zeta(t) &= \frac{1}{\gamma_1\{H_{10}(t)\}} \int_{t_0}^t \frac{\gamma_1\{H_{10}(s)\}dC(s)}{B(s)} - \frac{1}{\gamma_1\{H_{10}(t)\}} \int_{t_0}^t \left[ \frac{\mathcal{A}(s)\gamma_1\{H_{10}(s)\}}{B(s)\lambda_2\{H_{20}(s)\}} \right. \\
&\quad \left. + \int_s^t \frac{\gamma_1\{H_{10}(v)\}\mathcal{A}(v)}{\gamma_2\{H_{20}(v)\}B(v)} d \left( \frac{\gamma_2\{H_{20}(v)\}}{\lambda_2\{H_{20}(v)\}} \right) \right] d\mu(s).
\end{aligned}$$

$$\begin{aligned}
\omega_1(s, t) &= I(s \leq t) \frac{\gamma_1\{H_{10}(s)\}}{B(s)\gamma_1\{H_{10}(t)\}} - \zeta'(t)\Sigma^{-1}\mathcal{D}_1(s), \\
\omega_2(s, t) &= \zeta'(t)\Sigma^{-1}\mathcal{D}_2(s) + I(s \leq t) \left[ \frac{\mathcal{A}(s)\gamma_1\{H_{10}(s)\}}{B(s)\mathcal{K}(s)\gamma_1\{H_{10}(t)\}} \right. \\
&\quad \left. + \frac{\lambda_2\{H_{20}(s)\}}{\mathcal{K}(s)\gamma_1\{H_{10}(t)\}} \int_s^t \frac{\mathcal{A}(v)\gamma_1\{H_{10}(v)\}}{B(v)\gamma_2\{H_{20}(v)\}} d \frac{\gamma_2\{H_{20}(v)\}}{\lambda_2\{H_{20}(v)\}} \right], \\
\Sigma_1(t) &= E \left\{ -\zeta'(t)\Sigma^{-1} \frac{\partial \log \mathcal{L}_i(\Theta_0; H_{10}, H_{20})}{\partial \Theta} \right. \\
&\quad \left. + \int_{t_0}^{\tau} \omega_1(s, t) dM_i(s) - \int_{t_0}^{\tau} \omega_2(s, t) dM_{2i}(s) \right\}^2,
\end{aligned}$$

and

$$\begin{aligned}
\Sigma_2(t) &= \frac{1}{\lambda_2^2\{H_{20}(t)\}} E \left[ \int_{t_0}^t \frac{\lambda_2\{H_{20}(s)\}}{\mathcal{K}(s)} dM_{2i}(s) - \mu'(t)\Sigma^{-1} \frac{\partial \log \mathcal{L}_i(\Theta_0; H_{10}, H_{20})}{\partial \Theta} \right. \\
&\quad \left. - \mu'(t)\Sigma^{-1} \int_{t_0}^{\tau} \mathcal{D}_1(s) dM_i(s) - \mu'(t)\Sigma^{-1} \int_{t_0}^{\tau} \mathcal{D}_2(s) dM_{2i}(s) \right]^2.
\end{aligned}$$

## Supplementary Material B: Proofs

The consistency and asymptotic normality stated in Theorems 1 and 2 are proved using arguments similar to those of Chen *et al.* (2002), so we highlight only the steps that are different.

### Proof of Theorems 1 and 2.

*Step 1.* Using similar arguments to Step A1 of Chen *et al.* (2002), it can be shown that  $d\{\widehat{H}_2(\cdot, \Theta_0), H_{20}(\cdot)\} \rightarrow 0$  almost surely, where  $\widehat{H}_2(\cdot, \Theta)$  is the function implicitly defined as the unique solution of (2.2) for fixed  $\Theta$  and  $d(G_1, G_2) = \sup_{t \in [t_0, \tau]} \left| \exp\{G_1(t)\} - \exp\{G_2(t)\} \right|$ .

Now we show that  $D\{\widehat{H}_1(\cdot, \Theta_0), H_{10}(\cdot)\} \rightarrow 0$  almost surely, where  $\widehat{H}_1(\cdot, \Theta_0) \in \mathcal{H}_1$  is the function implicitly defined as the unique solution of (2.3) with  $\Theta = \Theta_0$ ,  $H_2(\cdot) = \widehat{H}_2(\cdot, \Theta_0)$  and  $D(G_1, G_2) = \sup_{t \in [t_0, \tau]} \left| E \left( \log [S_{\rho_0} \{G_1(t \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\beta}_0, H_{20}(t \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\alpha}_0\}] - \log [S_{\rho_0} \{G_2(t \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\beta}_0, H_{20}(t \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\alpha}_0\}] \right) \right|$ , for any two nondecreasing functions  $G_1$  and  $G_2$  on  $[t_0, \tau]$  such that  $G_1(t_0) = G_2(t_0) = -\infty$ . Denote  $\mathcal{H}_1 = \{H_1 : H_1 \text{ is nondecreasing step functions on } [t_0, \tau] \text{ with } H_1(t_0) = -\infty \text{ and with jumps only at the observed failure times } t_1, \dots, t_M\}$ , and  $A$  is a mapping defined by  $A(H_1)(t) = \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t \left( dN_i(t) + Y_i(t) d \log \left[ S_{\rho_0} \left\{ H_1(t) - \mathbf{X}'_i \boldsymbol{\beta}_0, \widehat{H}_2(t, \Theta_0) - \mathbf{X}'_i \boldsymbol{\alpha}_0 \right\} \right] \right)$ , where  $\rho_0$ ,  $\boldsymbol{\beta}_0$  and  $\boldsymbol{\alpha}_0$  are the true values of  $\rho$ ,  $\boldsymbol{\beta}$  and  $\boldsymbol{\alpha}$ . For an arbitrary but fixed  $\epsilon > 0$ , consider  $G_1$  and  $G_2$  such that  $D(G_1, G_2) > \epsilon$ , then there exists a  $t^* \in [t_0, \tau]$  such that

$$\left| E \left[ \log \{S_{\rho_0} (G_1(t^* \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\beta}_0, H_{20}(t^* \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\alpha}_0)\} - \log \{S_{\rho_0} (G_2(t^* \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\beta}_0, H_{20}(t^* \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\alpha}_0)\} \right] \right| \geq \epsilon/2.$$

Hence, coupling with  $d\{\widehat{H}_2(\cdot, \Theta_0), H_{20}(\cdot)\} \rightarrow 0$  almost surely, we have

$$\begin{aligned} & \sup_{t \in [t_0, \tau]} \left| A(G_1)(t) - A(G_2)(t) \right| \\ & \geq \frac{1}{n} \left| \sum_{i=1}^n \left( \log \left[ S_{\rho_0} \left\{ G_1(t^* \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\beta}_0, \widehat{H}_2(t^* \wedge U_{1i}, \Theta_0) - \mathbf{X}'_i \boldsymbol{\alpha}_0 \right\} \right] - \log \left[ S_{\rho_0} \left\{ G_2(t^* \wedge U_{1i}) - \mathbf{X}'_i \boldsymbol{\beta}_0, \widehat{H}_2(t^* \wedge U_{1i}, \Theta_0) - \mathbf{X}'_i \boldsymbol{\alpha}_0 \right\} \right] \right) \right| \geq \epsilon/2 \end{aligned}$$

when  $n$  is large enough. Choose  $H_{10}^* \in \mathcal{H}_1$  such that  $H_{10}^*(t_i) = H_{10}(t_i)$  for  $i = 1, \dots, M$ . By the law of large numbers, the continuity of  $H_{10}$  and  $d\{\widehat{H}_2(\cdot, \Theta_0), H_{20}(\cdot)\} \rightarrow 0$  implies that  $\sup\{A(H_{10}^*)(t) : t \in [t_0, \tau]\} \rightarrow 0$  almost surely. It follows that  $\sup_{t \in [t_0, \tau]} \left| A(H_{10}^*)(t) - A\{\widehat{H}_1(\cdot, \Theta_0)\}(t) \right| \rightarrow 0$  almost surely because, by definition of  $\widehat{H}_1(\cdot, \Theta_0)$ ,  $A\{\widehat{H}_1(\cdot, \Theta_0)\}(t) = 0$  for all  $t \in [t_0, \tau]$ . Then, with probability 1,  $\widehat{H}_1(\cdot, \Theta_0)$  is in the neighborhood of  $H_{10}^*$  of

radius  $\epsilon$  under the metric  $D(\cdot, \cdot)$ . Therefore,  $D(\widehat{H}_1(\cdot, \Theta_0), H_{10}) \rightarrow 0$  almost surely, since  $\epsilon > 0$  can be arbitrarily small and  $\widehat{H}_1(\cdot, \Theta_0)$  and  $H_{10}$  are monotone.

*Step 2.* Constructing the expressions of  $\widehat{H}_2(t; \Theta_0)$  and  $\widehat{H}_1(t; \Theta_0)$ . Let  $a > 0$ , let  $b$  be fixed finite numbers and define  $\mathcal{K}_1(t) = \int_a^t E[Y_{2i}(s)\lambda_{2i}^{(1)}\{H_{20}(s)\}]dH_{20}(s)$ ,  $\Gamma_1(x) = \int_b^x \gamma_1(s)ds$ ,  $\Lambda_2(x) = \int_b^x \lambda_2(s)ds$ , for  $t > t_0$  and  $x \in (-\infty, \infty)$ . We choose finite  $a > 0$  and  $b$  as the lower limits of the integration to ensure that the integrals are finite. Similar to Step A2 in Chen *et al.* (2002), we have, uniformly for  $t \in [t_0, \tau]$ ,

$$\Lambda_2\{\widehat{H}_2(t; \Theta_0)\} - \Lambda_2\{H_{20}(t)\} = \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t \frac{\lambda_2\{H_{20}(s)\}}{\mathcal{K}(s)} dM_{2i}(s) + o_p(n^{-1/2}). \quad (\text{S.1})$$

Now we consider the representative of  $\widehat{H}_1(t; \Theta_0)$ . Denote

$$B_1(t) = \int_a^t E[Y_i(s)d\gamma_{1i}\{H_{10}(s), H_{20}(s)\}], \quad \Gamma_2(x) = \int_b^x \gamma_2(s)ds,$$

$$\mathcal{A}_1(t) = \int_a^t E[Y_i(s)d\gamma_{2i}\{H_{10}(s), H_{20}(s)\}], \quad \mathcal{A}(t) = E[Y_i(t)\gamma_{2i}\{H_{10}(t), H_{20}(t)\}].$$

Thus, it is easy to see that  $d\gamma_1\{H_{10}(t)\} = [\gamma_1\{H_{10}(t)\}/B(t)]dB_1(t)$  and  $d\gamma_2\{H_{20}(t)\} = [\gamma_2\{H_{20}(t)\}/\mathcal{A}(t)]d\mathcal{A}_1(t)$  and therefore:

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n M_i(t) &= \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t Y_i(s) d \left( \frac{\gamma_{1i}\{H_{10}(s), H_{20}(s)\}}{\gamma_1\{H_{10}(s)\}} \left[ \Gamma_1\{\widehat{H}_1(s; \Theta_0)\} - \Gamma_1\{H_{10}(s)\} \right] \right) \\ &\quad + \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t Y_i(s) d \left( \frac{\gamma_{2i}\{H_{10}(s), H_{20}(s)\}}{\gamma_2\{H_{20}(s)\}} \left[ \Gamma_2\{\widehat{H}_2(s; \Theta_0)\} - \Gamma_2\{H_{20}(s)\} \right] \right) + o_p(n^{-1/2}) \\ &= \int_{t_0}^t \frac{B(s)}{\gamma_1\{H_{10}(s)\}} d \left[ \Gamma_1\{\widehat{H}_1(s; \Theta_0)\} - \Gamma_1\{H_{10}(s)\} \right] \\ &\quad + \int_{t_0}^t \frac{\mathcal{A}(s)}{\gamma_2\{H_{20}(s)\}} d \left[ \Gamma_2\{\widehat{H}_2(s; \Theta_0)\} - \Gamma_2\{H_{20}(s)\} \right] + o_p(n^{-1/2}). \end{aligned} \quad (\text{S.2})$$

Denote  $\Upsilon(t) = \Gamma_2\{\Lambda_2^{-1}(t)\}$ . By (S.1), uniformly for  $t \in [t_0, \tau]$ , we have

$$\begin{aligned} \Gamma_2\{\widehat{H}_2(t; \Theta_0)\} - \Gamma_2\{H_{20}(t)\} &= \Upsilon[\Lambda_2\{\widehat{H}_2(t; \Theta_0)\}] - \Upsilon[\Lambda_2\{H_{20}(t)\}] \\ &= \frac{\gamma_2\{H_{20}(t)\}}{n\lambda_2\{H_{20}(t)\}} \sum_{i=1}^n \int_{t_0}^t \frac{\lambda_2\{H_{20}(s)\}}{\mathcal{K}(s)} dM_{2i}(s) + o_p(n^{-1/2}). \end{aligned} \quad (\text{S.3})$$

Substituting it into (S.2), we get

$$\begin{aligned}
\Gamma_1\{\widehat{H}_1(t; \Theta_0)\} - \Gamma_1\{H_{10}(t)\} &= \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t \frac{\gamma_1\{H_{10}(s)\}}{B(s)} dM_i(s) - \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t \left[ \frac{\mathcal{A}(s)}{B(s)} \frac{\gamma_1\{H_{10}(s)\}}{\mathcal{K}(s)} \right. \\
&\quad \left. + \frac{\lambda_2\{H_{20}(s)\}}{\mathcal{K}(s)} \int_s^t \frac{\mathcal{A}(v)}{B(v)} \frac{\gamma_1\{H_{10}(v)\}}{\gamma_2\{H_{20}(v)\}} d \frac{\gamma_2\{H_{20}(v)\}}{\lambda_2\{H_{20}(v)\}} \right] dM_{2i}(s) + o_p(n^{-1/2}) \\
&\cong \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t \Psi_1(s) dM_i(s) - \frac{1}{n} \sum_{i=1}^n \int_{t_0}^t \Psi_2(s, t) dM_{2i}(s) + o_p(n^{-1/2}), \tag{S.4}
\end{aligned}$$

uniformly over  $t \in [t_0, \tau]$ .

*Step 3.* Denote  $\mathcal{L}(\Theta; H_1, H_2) = \prod_{i=1}^n \mathcal{L}_i(\Theta; H_1, H_2)$  and  $U(\Theta; H_1, H_2) = \frac{\partial \log \mathcal{L}(\Theta; H_1, H_2)}{\partial \Theta}$ . In the step, we compute  $V(\Theta) = \frac{\partial U(\Theta; \widehat{H}_1(\cdot; \Theta), \widehat{H}_2(\cdot; \Theta))}{n \partial \Theta}$  at  $\Theta = \Theta_0$ . By differentiating both side of (2.2) with  $H_2(t)$  replaced by  $\widehat{H}_2(t; \Theta)$ , respect to  $\Theta$ , we obtain the identity

$$\sum_{i=1}^n \int_{t_0}^t Y_{2i}(s) d \left[ \lambda_{2i}\{\widehat{H}_2(s; \Theta)\} \left\{ \frac{\partial \widehat{H}_2(s; \Theta)}{\partial \Theta} - \frac{\partial \mathbf{X}'_i \alpha}{\partial \Theta} \right\} \right] \Big|_{\Theta=\Theta_0} = 0. \tag{S.5}$$

Similar to Step 2, we see that

$$\frac{\partial \widehat{H}_2(t; \Theta_0)}{\partial \Theta} = \frac{\mu(t)}{\lambda_2\{H_{20}(t)\}} + o_p(1), \tag{S.6}$$

where  $\mu(t)$  is defined in Section 3.

However, by differentiating both side of (2.3) with  $H_1(t)$  and  $H_2(t)$  replaced by  $\widehat{H}_1(t; \Theta)$  and  $\widehat{H}_2(t; \Theta)$ , respect to  $\Theta$ , we obtain the identity

$$\begin{aligned}
\sum_{i=1}^n Y_i(t) d \left[ \gamma_{1i}\{\widehat{H}_1(t, \Theta), \widehat{H}_2(t, \Theta)\} \left\{ \frac{\partial \widehat{H}_1(t, \Theta)}{\partial \Theta} - \frac{\partial \mathbf{X}'_i \beta}{\partial \Theta} \right\} \right. \\
\left. + \gamma_{2i}\{\widehat{H}_1(t, \Theta), \widehat{H}_2(t, \Theta)\} \left\{ \frac{\partial \widehat{H}_2(t, \Theta)}{\partial \Theta} - \frac{\partial \mathbf{X}'_i \alpha}{\partial \Theta} \right\} + \gamma_{3i}\{\widehat{H}_1(t, \Theta), \widehat{H}_2(t, \Theta)\} \frac{\partial \rho}{\partial \Theta} \right] = 0,
\end{aligned}$$

where  $\gamma_{3i}(x, y) = -\frac{\dot{S}_\rho(x - \mathbf{X}'_i \beta, y - \mathbf{X}'_i \alpha)}{S_\rho(x - \mathbf{X}'_i \beta, y - \mathbf{X}'_i \alpha)}$  and  $\dot{S}_\rho(x, y) = \frac{\partial S_\rho(x, y)}{\partial \rho}$ . Similar to Step 2, we now

see that

$$\begin{aligned} \frac{\partial \widehat{H}_1(t, \Theta)}{\partial \Theta} \gamma_1\{H_{10}(t)\} + \int_{t_0}^t \frac{\mathcal{A}(s)\gamma_1\{H_{10}(s)\}}{B(s)\gamma_2\{H_{20}(s)\}} d \left[ \gamma_2\{H_{20}(s)\} \frac{\partial \widehat{H}_2(s, \Theta)}{\partial \Theta} \right] \\ = \int_{t_0}^t \gamma_1\{H_{10}(s)\} \frac{dC(s)}{B(s)} + o_p(1), \end{aligned}$$

for  $\Theta = \Theta_0$ , where  $C(s)$  is defined in Section 3. Then, using (S.6), we get for  $\Theta = \Theta_0$ ,

$$\begin{aligned} \frac{\partial \widehat{H}_1(t, \Theta)}{\partial \Theta} &= \frac{1}{\gamma_1\{H_{10}(t)\}} \int_{t_0}^t \frac{\gamma_1\{H_{10}(s)\}}{B(s)} dC(s) - \frac{1}{\gamma_1\{H_{10}(t)\}} \int_{t_0}^t \left( \frac{\mathcal{A}(s)\gamma_1\{H_{10}(s)\}}{B(s)\lambda_2\{H_{20}(s)\}} \right. \\ &\quad \left. + \int_s^t \frac{\gamma_1\{H_{10}(v)\}\mathcal{A}(v)}{\lambda_2\{H_{20}(v)\}B(v)} \left[ \frac{d\mathcal{A}_1(v)}{\mathcal{A}(v)} - \frac{d\lambda_2\{H_{20}(v)\}}{\lambda_2\{H_{20}(v)\}} \right] \right) d\mu(s) + o_p(1). \quad (\text{S.7}) \end{aligned}$$

It follows from the law of large numbers that

$$\begin{aligned} V(\Theta_0) &= \frac{1}{n} \frac{\partial^2 \log \mathcal{L}(\Theta_0; H_{10}, H_{20})}{\partial \Theta \partial \Theta'} + \frac{1}{n} \sum_{i=1}^n \frac{\partial^2 \log \mathcal{L}_i(\Theta_0; H_{10}, H_{20})}{\partial \Theta \partial H_1(U_{1i})} \frac{\partial \widehat{H}_1(U_{1i}; \Theta_0)}{\partial \Theta'} \\ &\quad + \frac{1}{n} \sum_{i=1}^n \frac{\partial^2 \log \mathcal{L}_i(\Theta_0; H_{10}, H_{20})}{\partial \Theta \partial H_2(U_{2i})} \frac{\partial \widehat{H}_2(U_{2i}; \Theta_0)}{\partial \Theta'} = \Sigma + o_p(1), \end{aligned}$$

where  $\mathcal{L}_i(\Theta; H_1, H_2)$  is the contribution of subject  $i$  to the likelihood function  $\mathcal{L}(\Theta; H_1, H_2)$ .

*Step 4.* In the step, we show the asymptotic normality of  $U(\Theta_0; \widehat{H}_1(\cdot, \Theta_0), \widehat{H}_2(\cdot, \Theta_0))$ .

Using the results of Steps 1 and 2 and some empirical process approximation techniques, we can write

$$\begin{aligned} \frac{1}{n} U\{\Theta_0; \widehat{H}_1(\cdot, \Theta_0), \widehat{H}_2(\cdot, \Theta_0)\} &= \frac{1}{n} \sum_{i=1}^n \frac{\partial \log \mathcal{L}_i(\Theta_0; H_{10}, H_{20})}{\partial \Theta} \\ &\quad + \frac{1}{n} \sum_{i=1}^n \frac{\partial^2 \log \mathcal{L}_i(\Theta_0; H_{10}, H_{20})}{\partial \Theta \partial H_1(U_{1i})} \frac{1}{\gamma_1\{H_{10}(U_{1i})\}} \left[ \Gamma_1\{\widehat{H}_1(U_{1i}; \Theta_0)\} - \Gamma_1\{H_{10}(U_{1i})\} \right] \\ &\quad + \frac{1}{n} \sum_{i=1}^n \frac{\partial^2 \log \mathcal{L}_i(\Theta_0; H_{10}, H_{20})}{\partial \Theta \partial H_2(U_{2i})} \frac{1}{\gamma_2\{H_{20}(U_{2i})\}} \left[ \Gamma_2\{\widehat{H}_2(U_{2i}; \Theta_0)\} - \Gamma_2\{H_{20}(U_{2i})\} \right] + o_p(n^{-1/2}) \end{aligned}$$

Substituting (S.4) and (S.3) into it, we get

$$\begin{aligned} & \frac{1}{n}U(\boldsymbol{\Theta}_0; \widehat{H}_1(\cdot, \boldsymbol{\Theta}_0), \widehat{H}_2(\cdot, \boldsymbol{\Theta}_0)) \\ &= \frac{1}{n} \sum_{i=1}^n \left\{ \frac{\partial \log \mathcal{L}_i(\boldsymbol{\Theta}_0; H_{10}, H_{20})}{\partial \boldsymbol{\Theta}} + \int_{t_0}^{\tau} \mathcal{D}_1(s) dM_i(s) + \int_{t_0}^{\tau} \mathcal{D}_2(s) dM_{2i}(s) \right\} + o_p(n^{-1/2}), \end{aligned}$$

where  $\mathcal{D}_1(s)$  and  $\mathcal{D}_2(s)$  are defined in Section 3. It then follows that

$$\frac{1}{\sqrt{n}}U\{\boldsymbol{\Theta}_0; \widehat{H}_1(\cdot, \boldsymbol{\Theta}_0), \widehat{H}_2(\cdot, \boldsymbol{\Theta}_0)\} \rightarrow N(0, \Delta),$$

where  $\Delta$  is defined in Section 3. The rest of the proof essentially proceeds along the lines of Chen *et al.* (2002) and is omitted here.

### Proof of Theorem 3.

By Taylor series expansions, (S.3), (S.6) and Theorem 2, we get,

$$\begin{aligned} & \Lambda_2\{\widehat{H}_2(t; \widehat{\boldsymbol{\Theta}})\} - \Lambda_2\{H_{20}(t)\} \\ &= \Lambda_2\{\widehat{H}_2(t; \widehat{\boldsymbol{\Theta}})\} - \Lambda_2\{\widehat{H}_2(t; \boldsymbol{\Theta}_0)\} + \Lambda_2\{\widehat{H}_2(t; \boldsymbol{\Theta}_0)\} - \Lambda_2\{H_{20}(t)\} \\ &= \lambda_2\{H_{20}(t)\} \frac{\partial \widehat{H}_2(t; \boldsymbol{\Theta}_0)}{\partial \boldsymbol{\Theta}} (\widehat{\boldsymbol{\Theta}} - \boldsymbol{\Theta}_0) + \Lambda_2\{\widehat{H}_2(t; \boldsymbol{\Theta}_0)\} - \Lambda_2\{H_{20}(t)\} + o_p(n^{-1/2}) \\ &= \frac{1}{n} \sum_{i=1}^n \left[ \int_{t_0}^t \frac{\lambda_2\{H_{20}(s)\}}{\mathcal{K}(s)} dM_{2i}(s) - \mu'(t)\Sigma^{-1} \frac{\partial \log \mathcal{L}_i(\boldsymbol{\Theta}_0; H_{10}, H_{20})}{\partial \boldsymbol{\Theta}} \right. \\ & \quad \left. - \mu'(t)\Sigma^{-1} \int_{t_0}^{\tau} \mathcal{D}_1(s) dM_i(s) - \mu'(t)\Sigma^{-1} \int_{t_0}^{\tau} \mathcal{D}_2(s) dM_{2i}(s) \right] + o_p(n^{-1/2}). \end{aligned}$$

Then the first part of Theorem 3 follows.

By Taylor series expansions, we get,

$$\begin{aligned} & \Gamma_1\{\widehat{H}_1(t)\} - \Gamma_1\{H_{10}(t)\} \\ &= \Gamma_1\{\widehat{H}_1(t; \widehat{\boldsymbol{\Theta}})\} - \Gamma_1\{\widehat{H}_1(t; \boldsymbol{\Theta}_0)\} + \Gamma_1\{\widehat{H}_1(t; \boldsymbol{\Theta}_0)\} - \Gamma_1\{H_{10}(t)\} \\ &= \gamma_1\{H_{10}(t)\} \frac{\partial \widehat{H}_1(t; \boldsymbol{\Theta}_0)}{\partial \boldsymbol{\Theta}} (\widehat{\boldsymbol{\Theta}} - \boldsymbol{\Theta}_0) + \Gamma_1\{\widehat{H}_1(t; \boldsymbol{\Theta}_0)\} - \Gamma_1\{H_{10}(t)\}. \end{aligned}$$



Then the second part of Theorem 3 follows from Theorem 2, (S.4) and (S.7).

### Proof of the Proposition.

Denote  $U^*(\Theta; H_1^*, H_2^*) = \sum_{i=1}^n \xi_i \frac{\partial \mathcal{L}_i(\Theta; H_1^*, H_2^*)}{\partial \Theta}$ . Then, for  $\|\Theta_1 - \Theta_2\| = o(1)$ , we have

$$\begin{aligned} \frac{1}{n} U^*(\Theta_1; H_1^*, H_2^*) - \frac{1}{n} U^*(\Theta_2; H_1^*, H_2^*) \\ = V^*(\Theta_2)(\Theta_1 - \Theta_2) + o(\|\Theta_1 - \Theta_2\|), \end{aligned} \quad (\text{S.8})$$

where  $V^*(\Theta) = -\frac{\partial U^*(\Theta; H_1^*, H_2^*)}{n \partial \Theta}$ . Similarly to Step 3 in the proof of Theorem 1 and Theorem 2, combining with  $E(\xi) = 1$  and  $Var(\xi) = 1$ , it follows from the strong law of large numbers that

$$\begin{aligned} d\{H_1^*(\cdot, \Theta), \widehat{H}_1(\cdot, \Theta)\} \rightarrow 0, \quad d\{H_2^*(\cdot, \Theta), \widehat{H}_2(\cdot, \Theta)\} \rightarrow 0, \\ \frac{\partial H_1^*(t; \widehat{\Theta})}{\partial \Theta} \rightarrow \zeta(t), \quad \frac{\partial H_2^*(t; \widehat{\Theta})}{\partial \Theta} \rightarrow \frac{\mu(t)}{\lambda_2 \{H_{20}(t)\}}, \end{aligned}$$

almost surely, and furthermore,

$$V^*(\widehat{\Theta}) \xrightarrow{a.s.} \Sigma. \quad (\text{S.9})$$

However,

$$\begin{aligned} \frac{1}{n} U^*(\widehat{\Theta}; H_1^*, H_2^*) \\ = \frac{1}{n} \sum_{i=1}^n (\xi_i - 1) \frac{\partial \mathcal{L}_i(\widehat{\Theta}; \widehat{H}_1, \widehat{H}_2)}{\partial \Theta} + \frac{1}{n} \sum_{i=1}^n \xi_i \left\{ \frac{\partial \mathcal{L}_i(\widehat{\Theta}; H_1^*, H_2^*)}{\partial \Theta} - \frac{\partial \mathcal{L}_i(\widehat{\Theta}; \widehat{H}_1, \widehat{H}_2)}{\partial \Theta} \right\}, \end{aligned}$$

where the second term is  $o_p(1)$ . Hence, following from the result of Step 4 in the proof of Theorem 1 and Theorem 2, we have

$$\frac{1}{\sqrt{n}} U^*\{\widehat{\Theta}; H_1^*, H_2^*\} \rightarrow N(0, \Delta). \quad (\text{S.10})$$

Finally, note that  $U^*(\Theta^*; H_1^*, H_2^*) = 0$ , and  $U^*(\Theta_0; H_1^*, H_2^*) \rightarrow U^*(\Theta_0; \widehat{H}_1, \widehat{H}_2) \rightarrow U(\Theta_0; \widehat{H}_1, \widehat{H}_2) \rightarrow 0$  almost surely; In addition, with the existence of  $\Sigma^{-1}$  and  $\widehat{\Theta} \xrightarrow{a.s.} \Theta_0$ , we have  $\|\Theta^* - \widehat{\Theta}\| = o(1)$ . Combining (S.8) with (S.10), the proof of the Proposition is completed.

## Supplementary Material C: Simulation 3

**Simulation 3.** For computational easy, we estimate the transformation function by estimating equations, which may be less efficient than the maximum likelihood estimator. To investigate the possible loss due to the using of the estimating equations instead of the maximum likelihood function, we compare the proposed method with the oracle parametric maximum likelihood (PML) method, where the transformation function is correctly specified and all of the parameters are estimated by the maximum likelihood function. We generate data from a simulation setting similar to Simulation 1, except that  $H_1(t) = \log(t)$  and  $H_2(t) = 10 \log(t) + t$  in Case 1;  $H_1(t) = t + 1.5$  and  $H_2(t) = 3t + 0.5$  in Case 2. To use the PML method,  $H_1(t)$  and  $H_2(t)$  are specified as  $H_1(t) = \theta_1 \log(t) + \theta_2$  and  $H_2(t) = \theta_3 \log(t) + \theta_4 t$  for Caes 1, and  $H_1(t) = \theta_1 t + \theta_2$  and  $H_2(t) = \theta_3 t + \theta_4$  for Case 2. The resulting estimators are displayed in Table 5. Table 5 shows that the performance of the proposed method is close to that of the PML although the transformation function of our method is unknown. The results in Table 5 confirm the proposed method is efficient.

**Table 5. The bias, SD and RMSE of estimators for Simulation 3.**

|        | $\hat{\alpha}$ |       | $\hat{\beta}$ |       | $\hat{\rho}$    |        |
|--------|----------------|-------|---------------|-------|-----------------|--------|
|        | Bias(SD)       | RMSE  | Bias(SD)      | RMSE  | Bias(SD)        | RMSE   |
| Case 1 |                |       |               |       |                 |        |
| Prop.  | 0.032(0.126)   | 0.130 | 0.021(0.121)  | 0.122 | 0.0064(0.0439)  | 0.0444 |
| PML    | 0.018(0.126)   | 0.127 | 0.019(0.119)  | 0.121 | -0.0001(0.0436) | 0.0436 |
| Case 2 |                |       |               |       |                 |        |
| Prop.  | 0.024(0.115)   | 0.117 | 0.021(0.116)  | 0.118 | 0.0082(0.0458)  | 0.0465 |
| PML    | 0.005(0.116)   | 0.116 | -0.001(0.116) | 0.116 | 0.0010(0.0450)  | 0.0450 |

## Supplementary Material D: Simulation 4

**Simulation 4.** In this simulation, we consider the data with two-dimensional covariate vector, which is a combination of continuous and discrete covariates. The setting in Simulation 4 is similar to that in Simulation 1, except that the covariate  $\mathbf{X} = (X_1, X_2)$ ,  $H_1(t) = t^3$  and  $H_2(t) = \Phi^{-1}(t/2)$ , where  $X_1$  is generated uniformly over  $[-2, 2]$ ,  $X_2$  is treatment indicator in which  $n/2$  subjects receive each of the two groups. The censoring random variable  $C$  is distributed uniformly on  $(0, 5)$ , so that about 20% of T is censored by  $C$  and about 24% of S is censored by  $C \wedge T$ . The resulting estimators are displayed in Table 6, suggesting the proposed method performs well in this setting.

**Table 6. The bias, empirical standard deviation and root of mean square error (RMSE) of estimators based on 500 simulations.**

| Method   | $\hat{\alpha}_1$ |        |        | $\hat{\alpha}_2$ |        |        |
|----------|------------------|--------|--------|------------------|--------|--------|
|          | Bias             | SD     | RMSE   | Bias             | SD     | RMSE   |
| Proposed | -0.0066          | 0.0625 | 0.0628 | 0.0438           | 0.1160 | 0.1240 |
| Proposed | $\hat{\beta}_1$  |        |        | $\hat{\beta}_2$  |        |        |
|          | -0.0070          | 0.0593 | 0.0597 | 0.0449           | 0.1167 | 0.1250 |
| Proposed | $\hat{\rho}$     |        |        |                  |        |        |
|          | 0.0046           | 0.0461 | 0.0463 |                  |        |        |

## Supplementary Material E: Figures 3 and 4

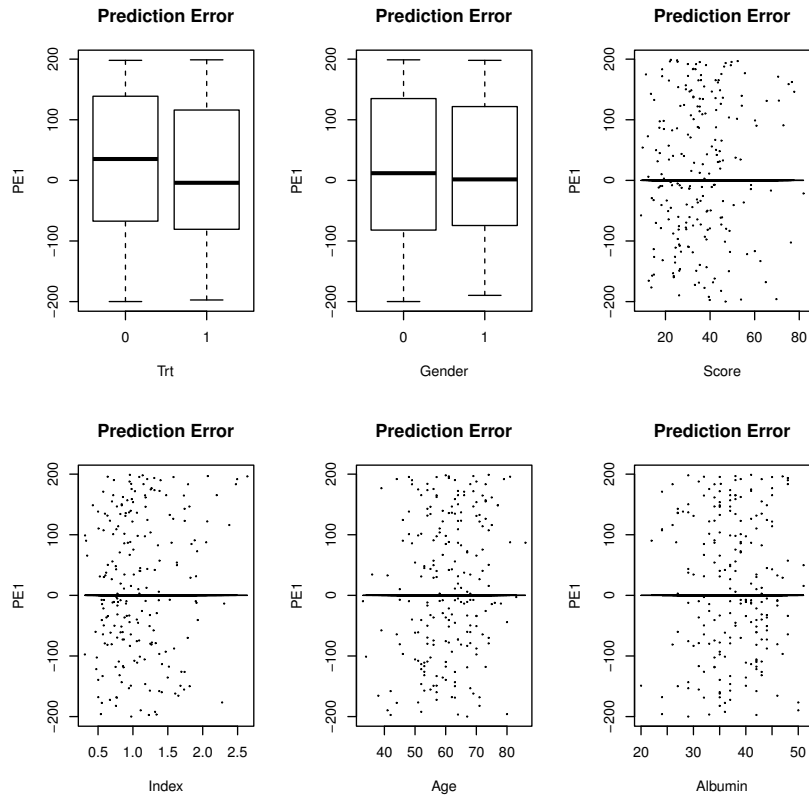


Figure 3: The plot of the prediction error  $PE_1$  versus the covariates using the proposed method for the myeloma data.

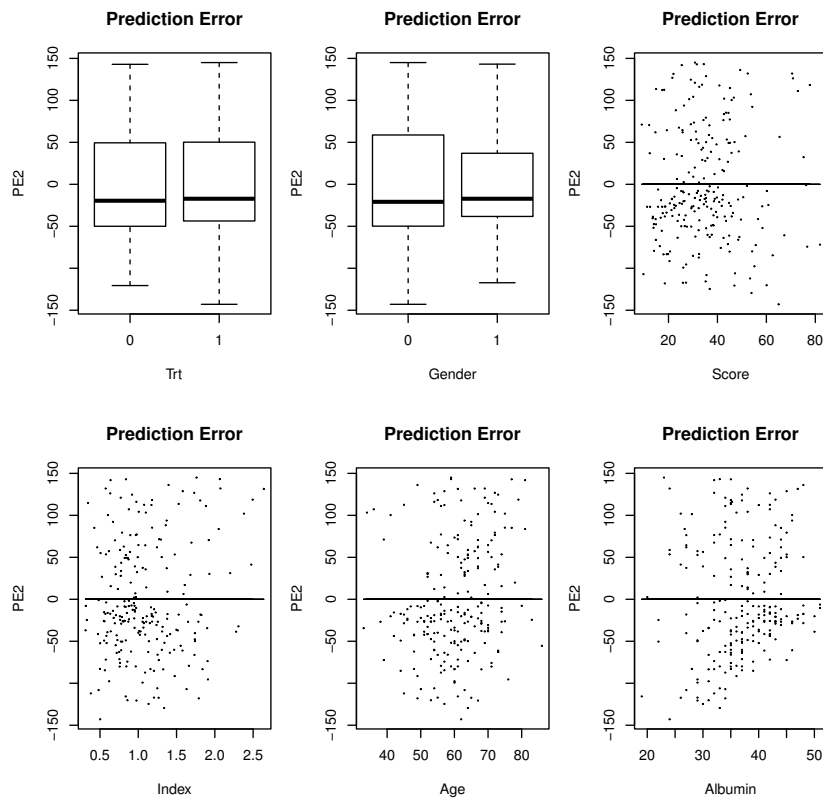


Figure 4: The plot of the prediction error  $PE_2$  versus the covariates using the proposed method for the myeloma data.