Alcohol Use Initiation is Associated with Changes in Personality Trait Trajectories from Early Adolescence to Young Adulthood

Daniel M. Blonigen, C. Emily Durbin, Brian M. Hicks, Wendy Johnson, Matt McGue, and William G. Iacono

Background: Recent work has demonstrated the codevelopment of personality traits and alcohol use characteristics from early adolescence to young adulthood. Few studies, however, have tested whether alcohol use initiation impacts trajectories of personality over this time period. We examined the effect of alcohol use initiation on personality development from early adolescence to young adulthood.

Methods: Participants were male ($n_{men} = 2,350$) and female ($n_{women} = 2,618$) twins and adoptees from 3 community-based longitudinal studies conducted at the Minnesota Center for Twin and Family Research. Data on personality traits of Positive Emotionality (PEM—Well-being), Negative Emotionality (NEM—Stress Reaction, Alienation, and Aggression), and Constraint (CON—Control and Harm Avoidance)—assessed via the Multidimensional Personality Questionnaire (MPQ)—and age of first drink were collected for up to 4 waves spanning ages 10 to 32.

Results: Alcohol use initiation was associated with significant decreases in levels of Well-being and CON traits, most notably Control; and significant increases in levels of all NEM traits, particularly Aggression. In general, the effects of alcohol use initiation on personality traits were moderated by gender and enhanced among those with earlier age of first drink.

Conclusions: From early adolescence to young adulthood, alcohol use initiation predicts deviations from normative patterns of personality maturation. Such findings offer a potential mechanism underlying the codevelopment of personality traits and alcohol use characteristics during this formative period of development.

Key Words: Personality Traits, Alcohol Use Initiation, Adolescence, Young Adulthood.

Recent studies highlight the codevelopment of personality traits and alcohol use characteristics from adolescence to young adulthood (e.g., Littlefield et al., 2009, 2010; Quinn and Harden, 2013) and suggest that changes in one should be evaluated in the context of changes in the other. Two developmental models are relevant to this issue. A precursor model posits that traits increase risk for alcohol use. In support of this, prospective studies have found that higher scores on NEM traits and lower scores on CON and PEM traits predict subsequent increases in drinking during adolescence and young adulthood (MacPherson et al., 2010; Sher et al., 2000) and are linked to an earlier age of first drink (Malmberg et al., 2012; McGue et al., 2001).

An alternative model, which has received less attention in the literature, posits that alcohol use may affect the normative maturation of personality from adolescence to adulthood (Chassin and Haller, 2010; White et al., 2011). Specifically, most individuals decrease on NEM traits and increase on CON and PEM traits as they enter young adulthood (Donnellan et al., 2007; Roberts et al., 2006). However, there is significant individual-level variability in trait trajectories across this period (Vaidya et al., 2008), particularly during adolescence (Soto et al., 2011). Alcohol use initiation, which typically occurs during adolescence, is a common life event that may partly account for these individual differences.
patterns. There are relatively few tests of whether the timing of alcohol use initiation accounts for variation in the development of personality traits; however, some studies have found earlier age of first drink to predict decreases in traits related to CON (White et al., 2011; Zernicke et al., 2010).

**The Present Study**

Using a large, mixed-gender sample spanning preadolescence (age 10) through young adulthood (age 32), we examined the extent to which alcohol use initiation had a discontinuous effect on the trajectories of personality traits of PEM (i.e., Well-being), NEM (i.e., Stress Reaction, Alienation, and Aggression), and CON (i.e., Control and Harm Avoidance). Specifically, we tested whether there was a change in the levels of these traits following alcohol initiation — over and above that predicted by aging alone. Further, we also explored whether effects of alcohol use initiation on trait levels were moderated by gender or timing of drinking onset. The later provides a test of the sensitive period hypothesis, which posits that the effect of drinking onset on risk for later alcohol problems may be magnified during early adolescence (DeWit et al., 2000).

We predicted alcohol use initiation would be associated with decreases in levels of Well-being and CON traits and increases in levels of NEM traits. We also predicted that traits previously shown to exhibit gender × time interactions during the transition from adolescence and young adulthood would also exhibit significant gender moderation in terms of the effect of alcohol initiation on trait levels. For example, prior research has observed larger increases in neuroticism and CON traits among girls than boys (Blonigen et al., 2008; Soto et al., 2011); thus, we predicted that the effects of alcohol use initiation on levels of Stress Reaction and CON traits to be larger among girls than boys. Finally, given that prior work has found an early versus later onset of drinking is associated with elevated lifetime rates of alcohol use disorders (AUD; Agrawal et al., 2009; Dawson et al., 2008), and lower levels of PEM and CON traits, and higher levels of NEM traits (Hill et al., 2000; Malmberg et al., 2012), we predicted that the effects alcohol initiation on trait levels would be greater among those with earlier age of first drink.

**MATERIALS AND METHODS**

**Sample**

Participants were members of 3 ongoing, longitudinal epidemiological studies conducted at the Minnesota Center for Twin and Family Research (MCTFR; Iacono et al., 2006): Minnesota Twin Family Study (MTFS; Iacono et al., 1999); Enrichment Study (MTFS-ES; Keyes et al., 2009); and Sibling Interaction and Behavior Study (SIBS; McGue et al., 2007)—an adoption study. For all studies, families were eligible if they met the familial relationship conditions of the relevant study, lived within a day’s drive of the MCTFR laboratories, and the targeted children did not have cognitive or physical disabilities that would preclude participation in the day-long assessment.

For the MTFS and MTFS-ES, same-sex twins born in Minnesota from 1972 to 1984 and 1988 to 1994, respectively, were identified and located via public birth records. For the MTFS, families were recruited to participate the year the twins turned either 11 ($n = 1,512$) or 17 years old ($n = 1,252$). For the MTFS-ES, families were recruited the year the twins turned 11 years old ($n = 998$), and additional screening procedures were employed to ensure that one-half of participating families had at least 1 twin with conduct problems. For any given birth year, over 90% of twin families were located and over 80% of those who were eligible agreed to participate. Participating families were representative of the demographic profile of Minnesotans during the target birth years (Iacono et al., 1999; Keyes et al., 2009). Participants across the 2 studies were predominantly of European American ancestry (96%) and slightly more women (52%) than men.

The SIBS consists of 409 and 208 adoptive and nonadoptive families, respectively ($n = 1,232$), each with 2 siblings within 5 years of age of one another and between ages of 11 and 21 years old at intake. Adoptive families were identified from private adoption agencies in Minnesota and included 2 unrelated siblings (63.2% participation rate). Approximately two-thirds of adoptions were international (primarily East Asian ancestry) with a mean age of placement of 4.7 months (SD = 3.4). Nonadoption families were ascertained via public birth records and selected to include a pair of biological siblings comparable in age and gender to adoptive sibling pairs (57.3% participation rate). Across all sibling pairs, small majorities were same-sex (60.8%), female (54.9%), and of European American ancestry (55.8%); the other major racial/ethnic groups were composed of Korean adoptees.

Participants from each study were invited to return for follow-up assessments every 3 to 5 years. Despite ongoing and/or unequal number of assessments across participants, aggregation of the data across all samples and follow-up assessments allowed for modeling of the data across ages 10 to 32. The age distribution and number of participants across each wave of assessment are shown in Table 1. Wave 1 denotes the intake assessment and thus includes the total number of participants across all studies. The first follow-up assessment (Wave 2) has been completed for all studies (92.1% retention rate). Across the remaining waves, the decreasing $N$ is a function of ongoing assessments and study design. Ongoing assessments included the second follow-up for the MTFS-ES and SIBS (Wave 3) and Wave 4 for the MTFS-ES. Participants from the MTFS age 17 cohort had completed all scheduled assessments (Waves 1 to 4). MTFS assessments are scheduled to terminate when participants reach about age 30. SIBS participants are only scheduled to complete 3 waves of assessment. The mean true retention rate was approximately 90% across all completed follow-up assessments for each of the 3 studies and across different age and gender cohorts in the MTFS and MTFS-ES.

**Assessment**

**Personality.** Personality was assessed with the Multidimensional Personality Questionnaire (MPQ; Tellegen and Waller, 2008). This self-report questionnaire measures individuals’ typical affective and behavioral styles using a higher-order 3-factor structure of PEM, NEM, and CON. For the primary scales that were included in this study, the range of Cronbach’s alpha across the most common assessment ages (14, 17, 20, 24, and 29) is presented below. PEM captures a propensity to experience positive emotions and was measured in the current study by the primary scale of Well-being ($\alpha = 0.88$ to 0.92). NEM captures a propensity to experience negative emotions and comprises scales of Stress Reaction ($\alpha = 0.87$ to 0.90), Alienation ($\alpha = 0.87$ to 0.92), and Aggression ($\alpha = 0.87$ to 0.91). CON taps a tendency to be planful and cautious and averse to risk, and was measured in the current study by the primary scales
of Control ($\alpha = 0.84$ to $0.88$) and Harm Avoidance ($\alpha = 0.84$ to $0.87$).

These 6 primary scales were selected for analysis because they (i) have each been linked to drinking problems during adolescence and/or young adulthood (Dom et al., 2006; Elkins et al., 2004; Jackson and Sher, 2003) and (ii) were administered at each assessment wave in MCTFR studies. Specifically, the 198-item version of the MPQ was administered at the ages 17, 24, and 29 assessments of the MTFS and MTFS-ES. An abbreviated version of the MPQ that included only scales of Well-being, Stress Reaction, Alienation, Aggression, Control, and Harm Avoidance was administered to twins at the age 14 assessment and to female twins from the age 17 cohort at their age 20 assessment. SIBS participants 16 years of age and older completed the 198-item version of the MPQ, while those younger than 16 years old completed the abbreviated MPQ. All MPQ data prior to age 13 were collected from SIBS participants. For any given assessment, 83 to 93% of participants had MPQ data. Across all studies, 1,340 and 632 participants had personality data at 3 and 4 waves of assessment, respectively. Per Little’s MCAR test, these data for the MPQ scales were missing completely at random ($\chi^2 (82) = 101.94$, $p = 0.067$).

### Alcohol Use Initiation and Other Drinking Characteristics.

At each assessment, participants were asked information on their history and pattern of drinking, which included the following question —“How old were you the first time you drank alcohol (on your own; more than your parents allowed you to)?”—as part of the Substance Abuse Module of the Composite International Diagnostic Interview (Robins et al., 1988) ($M = 15.9$ years; $SD = 2.2$; range = 5.0 to 24.5). Among those who had ever consumed alcohol, the average frequency of drinking in the past 12 months was less than once per month at age 14, once per month at age 17, and 2 to 3 times per month at ages 20 and 24. The average number of drinks on a typical drinking day in the past 12 months was 2.1 ($SD = 2.6$) at age 14, 3.3 ($SD = 3.5$) at age 17, 4.3 ($SD = 3.5$) at age 20, and 3.5 ($SD = 2.7$) at age 24.

### Statistical Analyses

For each trait, multilevel models were fit in HLM 7.0 (Scientific Software International Inc., Skokie, IL) to test (i) for effects of alcohol use initiation on the levels of MPQ traits and (ii) whether effects of initiation were moderated by gender or timing of drinking onset. Each model included 3 levels to account for nesting of participants within families and assessments within participants: Level 3 (nesting of participants within families); Level 2 (individual participants); and Level 1 (repeated measures across time within participants).

The following choices were made when modeling MPQ trait trajectories. First, as indicated by Durbin and colleagues’ (in press) analysis of this data set, models that include linear and nonlinear (i.e., quadratic and cubic) effects of age provide the best fit to the MPQ trait trajectories from early adolescence to young adulthood (see also Harden and Tucker-Drob, 2011). Thus, both linear and nonlinear effects of age were included in the models for this study. Data were included for up to 4 waves for each participant, and because participants varied in chronological age at the time of each assessment, change was modeled as a function of participants’ actual chronological ages at each assessment. We centered age at the first assessment wave (age 11). Although the number of participants was smaller for the earliest ages (11 to 14) than for ages 17 and older, we wanted to generate the most robust estimates for trait levels across the full developmental span covered in the sample so as to more accurately estimate the effect of alcohol initiation on traits, beyond the effects of age-related change. Second, we modeled age-related change in traits as fixed effects estimated for the sample as a whole, while the effect of alcohol initiation was allowed to vary across participants. Per Durbin and colleagues (in press), models in which linear and nonlinear effects are freely estimated did not provide a better fit than models in which they were fixed (i.e., models without random effects for the change parameters did not fit worse than models with the additional random effects). Third, effects of alcohol initiation on trait levels were modeled as within-subjects effects (time-varying coefficients at Level 1); thus, only participants who had MPQ data for at least 1 assessment prior and 1 following their age of drinking onset contributed to the estimation of parameters for the effects of alcohol initiation on trait levels.

To examine initiation effects, we entered a time-varying predictor at Level 1 (ALC NO/YES) to account for participant-specific timing of this event and its effect on trait elevation across the sample. A value of 0 was entered for this variable at Level 1 for each assessment prior to participants’ age of first drink, and a value of 1 was entered for each assessment wave after age of first drink. This allowed for tests of within-subject differences in trait levels before versus after drinking onset, that is, the extent to which alcohol use initiation was associated with subsequent differences in trait levels above and beyond the effects of age-related change, with age centered at 11:

\[
\text{Stress Reaction}_{tij} = \gamma_{000} + \gamma_{100} \times (\text{AGE} - 11)_{tij} + \gamma_{200} \\
\times (\text{AGE} - 11)_{tij}^2 + \gamma_{300} \times (\text{AGE} - 11)_{tij}^3 \\
+ \gamma_{400} \times (\text{ALC NO/YES})_{tij} + \phi_{ij} + u_{tij} + e_{tij}
\]

Due to the number of data points before and after the typical age of initiation, we were unable to contrast slopes and examine effects of alcohol use initiation on rate of linear or nonlinear change in traits over time. Consequently, the initiation effects (and the graphs to be presented) only demonstrate, in the context of the average trait trajectory, the magnitude of the change in trait levels immediately after the average age of alcohol initiation, not the impact of initiation effects in the rate of change in traits over time. Next, for traits in which alcohol use initiation was associated with subsequent...
differences in trait levels before versus after drinking onset, we tested separate models in which gender and age of first drink moderated the magnitude of the initiation effects on trait levels. These moderators were entered at Level 2 as predictors of variation in the Level 1 effect of alcohol use initiation:

$$\text{Stress Reaction}_{ij} = \gamma_{000} + \gamma_{100} \cdot (\text{AGE} - 11_{ij}) + \gamma_{200} \cdot (\text{AGE} - 11_{ij})^2 + \gamma_{300} \cdot (\text{AGE} - 11_{ij})^3 + \gamma_{400} \cdot (\text{ALC NO/YES}_{ij}) + \gamma_{410} \cdot (\text{ALC NO/YES}_{ij}) \cdot (\text{GENDER}_{ij}) + \tau_{0ij} + \epsilon_{ij}.$$

For the final moderation models, we accounted for the main effects of gender and age of first drink on the basic change parameters (i.e., intercept, age, age$^2$, and age$^3$) of each trait; that is, the moderator analyses included either gender or age of first drink as a significant Level 2 predictor of the Level 1 effects of these change parameters.

RESULTS

Model Fitting Results for the Effects of Alcohol Use Initiation

The fit of models that included only age-related (linear and nonlinear) change versus models that also included effects of alcohol initiation was compared using Akaike's information criterion (AIC; lower values indicate a better fit to the data)—see Table 2. For each trait, a model that included the effects of initiation provided a better fit than a model that included only age-related change.

Effects of Alcohol Use Initiation on the Level of Personality Traits

Table 3 shows the results of the best-fitting models, which tested whether there were significant differences in levels of MPQ traits before versus after drinking onset (accounting for effects of age-related change on traits, i.e., initial intercept; and age, age$^2$, and age$^3$ slopes). Effects of alcohol use initiation on trait level were significant for all traits. For Well-being and both CON traits, there were significant decreases in the level of these traits after drinking onset. Conversely, for NEM scales, there were significant increases in the level of all traits after drinking onset. A graphical representation of the age-related effects and initiation effects for each trait at the mean age of drinking onset (i.e., age 16) is provided in Fig. 1. To facilitate interpretation of initiation effects, we used the grand SD for each trait (Well-being = 7.98; Stress Reaction = 9.37; Alienation = 9.01; Aggression = 9.66; Control = 8.17; Harm Avoidance = 10.86) to express the initiation effects in terms of their size relative to the scale SD. For example, after the onset of drinking, the level of Well-being decreased by 0.16 SD (i.e., 1.24/7.98), and levels of Control and Harm Avoidance decreased by 0.33 and 0.15 SD, respectively, whereas levels of Stress Reaction, Alienation, and Aggression increased by 0.13, 0.18, and 0.32 SD, respectively.

Per Table 4, gender was a significant moderator of the initiation effects for Aggression, Control, and Harm Avoidance, that is, the magnitudes of the effect of alcohol use initiation on these trait levels were significantly different between men and women. For Aggression, there were increases for both men and women; however, the effect was larger in men. For Control, there were decreases for both men and women after drinking onset, although the effect was larger in men. For Harm Avoidance, there was a significant decrease for men, but minimal change for women.

Per Table 5, age of first drink was a significant moderator of the initiation effects for most traits, with the exception of Aggression and Harm Avoidance, that is, the magnitude of the effect of alcohol use initiation on a trait level varied significantly as a function of the timing of alcohol use initiation. Based on the age-specific initiation effects (1 SD below [≤13.7 years] and 1 SD above [≥18.1 years] the mean age of initiation), the initiation effect for each trait was enhanced among those with an earlier age of first drink. Specifically, for Well-being and Control, decreases in the levels of these traits as a function of drinking onset were larger for those with an earlier (vs. later) age of onset, whereas for Stress

<table>
<thead>
<tr>
<th>Trait</th>
<th>Linear + nonlinear age-related change</th>
<th>Age-related change + effects of initiation</th>
</tr>
</thead>
<tbody>
<tr>
<td>Well-being</td>
<td>77340.57</td>
<td>67211.96</td>
</tr>
<tr>
<td>Stress Reaction</td>
<td>80307.65</td>
<td>69811.06</td>
</tr>
<tr>
<td>Alienation</td>
<td>79092.42</td>
<td>68934.68</td>
</tr>
<tr>
<td>Aggression</td>
<td>79598.00</td>
<td>68598.05</td>
</tr>
<tr>
<td>Control</td>
<td>77182.75</td>
<td>66875.31</td>
</tr>
<tr>
<td>Harm Avoidance</td>
<td>81904.56</td>
<td>77104.79</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Trait</th>
<th>Linear + nonlinear age-related change</th>
<th>Age-related change + effects of initiation</th>
</tr>
</thead>
<tbody>
<tr>
<td>Well-being</td>
<td>77340.57</td>
<td>67211.96</td>
</tr>
<tr>
<td>Stress Reaction</td>
<td>80307.65</td>
<td>69811.06</td>
</tr>
<tr>
<td>Alienation</td>
<td>79092.42</td>
<td>68934.68</td>
</tr>
<tr>
<td>Aggression</td>
<td>79598.00</td>
<td>68598.05</td>
</tr>
<tr>
<td>Control</td>
<td>77182.75</td>
<td>66875.31</td>
</tr>
<tr>
<td>Harm Avoidance</td>
<td>81904.56</td>
<td>77104.79</td>
</tr>
</tbody>
</table>

N = 3,211. AIC = Akaike's information criterion. MPQ = Multidimensional Personality Questionnaire. Linear + nonlinear age-related change = models that include effects for linear, quadratic, and cubic age-related change in traits over time. Age-related change + effects of initiation = models that include effects for linear, quadratic, and cubic age-related change in traits over time and an effect of alcohol use initiation on level of the MPQ traits. Lower AIC values denote better fit.
Reaction and Alienation, increases in the levels of these traits as a function of drinking onset were larger for those with an earlier (vs. later) age of onset.

**DISCUSSION**

This study tested whether alcohol use initiation impacted trajectories of personality development from early adolescence to young adulthood. Alcohol use initiation was associated with changes in levels of all trait trajectories examined; that is, levels of Well-being and CON traits (Control and Harm Avoidance) decreased after the onset of drinking, whereas levels of NEM traits (Stress Reaction, Alienation, and Aggression) increased after drinking onset. Notably, the magnitudes of these effects were greatest for Control and Aggression—2 of the strongest trait-based predictors of
problematic drinking during the transition to adulthood (Elkins et al., 2006; Krueger, 1999).

The current findings help to refine models regarding the association between personality and alcohol use characteristics during adolescence and young adulthood (Littlefield et al., 2012; Quinn and Harden, 2013). Specifically, prior prospective studies have tended to focus on a precursor model, observing that higher NEM traits and lower CON and PEM traits are linked to an earlier age of first drink (Malmberg et al., 2012; McGue et al., 2001) and predict subsequent increases in drinking during adolescence and young adulthood (MacPherson et al., 2010; Sher et al., 2000). The current findings support an alternative model in which alcohol use affects the developmental trajectories of personality traits during this developmental timeframe. Indeed, these 2 models are not mutually exclusive as the effects between personality and alcohol use may operate in a reciprocal fashion. For example, Malmberg and colleagues (2013) reported reciprocal effects between onset of substance use and traits related to CON (i.e., sensation-seeking and impulsivity).

The present study did not test for reciprocal effects per se; however, the findings complement prior tests of the precursor model by suggesting that alcohol use initiation is not simply predicted by deviations from normative patterns of personality maturation, but may also predict such patterns of development. Collectively, support for the 2 models aligns with the corresponsive principle of personality development, which posits that life experiences enhance the traits that lead

<table>
<thead>
<tr>
<th>Table 4. Moderation of Effects of Alcohol Use Initiation on the Level of MPQ Personality Traits by Gender</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Level 1 predictors</strong></td>
</tr>
<tr>
<td>Trait</td>
</tr>
<tr>
<td>Well-being</td>
</tr>
<tr>
<td>Stress</td>
</tr>
<tr>
<td>Alienation</td>
</tr>
<tr>
<td>Aggression</td>
</tr>
<tr>
<td>Control</td>
</tr>
<tr>
<td>Harm</td>
</tr>
</tbody>
</table>

*p < 0.05, **p < 0.01, ***p < 0.001, with no adjustment for multiple testing.

N = 3,211. Gender (1 = male). “Intercept” values reflect model-derived scores at age 11. Values for “age,” “age²,” and “age³” reflect the linear, quadratic, and cubic change, respectively, per year in that scale after the age of the intercept. “Gender on initiation effects” reflects the effect of including gender at Level 2 as a predictor of individual differences in the initiation effects at Level 1. This coefficient denotes whether the magnitude of the effects of alcohol use initiation on a trait level was significantly different between men and women.

<table>
<thead>
<tr>
<th>Table 5. Moderation of Effects of Alcohol Use Initiation on the Level of Multidimensional Personality Questionnaire Personality Traits by Age of First Drink</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Level 1 predictors</strong></td>
</tr>
<tr>
<td>Trait</td>
</tr>
<tr>
<td>Well-being</td>
</tr>
<tr>
<td>Stress</td>
</tr>
<tr>
<td>Alienation</td>
</tr>
<tr>
<td>Aggression</td>
</tr>
<tr>
<td>Control</td>
</tr>
<tr>
<td>Harm</td>
</tr>
</tbody>
</table>

*p < 0.05, **p < 0.01, ***p < 0.001, with no adjustment for multiple testing.

N = 3,211. “Intercept” values reflect model-derived scores at age 11. Values for “age,” “age²,” and “age³” reflect the linear, quadratic, and cubic change, respectively, per year in that scale after the age of the intercept. “Age of first drink on initiation effects” reflects the effect of including age of first drink (centered) at Level 2 as a predictor of individual differences in the initiation effects at Level 1. This coefficient denotes whether the magnitude of the effect of alcohol use initiation on a trait level varied significantly as a function of the timing of alcohol use initiation. The “Age-specific initiation effects” gives the respective effects of alcohol use initiation on trait levels at 1 SD below and 1 SD above the mean age of first drink.
people into those experiences in the first place (Roberts et al., 2003). This principle has been shown to be applicable to the association between personality development and work experiences and marital quality—for example, Le and colleagues (2013)—but to our knowledge has not previously been applied to alcohol use characteristics. The present findings suggest that the corresponsive principle may serve as a useful theoretical framework to describe the codevelopment of personality and drinking during adolescence and young adulthood.

Another notable finding from our analyses was that the strength and directions of the initiation effects were moderated by gender and timing of drinking onset. Regarding gender, changes in the levels of most traits following initiation were in the same directions for men and women, but served to magnify gender differences in these traits during adolescence and young adulthood; that is, prior research has found Stress Reaction, Control, and Harm Avoidance to be higher among women than men during these periods, and traits of Alienation and Aggression, to be higher among men than women (Blonigen et al., 2008; Soto et al., 2011). The effects of alcohol use initiation enhanced these gender differences. Notably, changes in levels of Harm Avoidance diverged somewhat for men and women such that the onset of drinking was followed by increases among men in risk-taking tendencies, but no change in such tendencies for women. These findings suggest that transactions between personality development and alcohol use may follow distinct pathways to problematic drinking for men and women (Hicks et al., 2007).

The findings for moderation by age of first drink highlight the salience of this variable within the addiction literature. Earlier age of first drink may have detrimental consequences on brain development (Clark et al., 2008) and has been consistently linked to increased risk for AUD (Agrawal et al., 2009; Dawson et al., 2008). Given that high NEM and low CON trait levels are robust predictors of AUD, the current findings in which earlier age of initiation was associated with greater increases in NEM and decreases in CON may account for why an early onset of drinking increases risk for developing AUD. Further, this finding suggests that the relevance of the corresponsive principle may not be uniform across development, with personality development more sensitive to effects of life events at some ages than others (DeWit et al., 2000). Specifically, effects of alcohol use initiation on levels of trait trajectories may be stronger when the onset of drinking occurs early in adolescence.

Although this study had several strengths including a large community-based, mixed-gender sample and prospective design over 4 time points, several limitations should be acknowledged. First, personality was assessed only via self-report questionnaires. Use of informant reports and behavioral tasks in future work will help corroborate the present findings and identify which effects are robust across assessment methods. Second, although the racial and ethnic profile of participants was consistent with the demographics of the particular geographic region and cohort, it was primarily of European American descent. Thus, generalizability of the findings may be limited. Third, for the sake of simplicity, our analysis of alcohol use characteristics was limited to age of first drink rather than regular or hazardous drinking patterns, for example binge drinking. However, age of first drink is a strong predictor of AUD and is therefore a variable of tremendous significance within the broader addiction literature. Finally, although our models of within-subject change provide a strong test of the “effect” of alcohol use initiation on personality trait development, true causal influence cannot be assumed.

In summary, alcohol use initiation predicted deviations from normative patterns of personality maturation during early adolescence to young adulthood. Future studies should examine these effects in concert with tests of the precursor model (cf., Malmberg et al., 2013), as well as examine whether reciprocal effects between personality traits and alcohol use characteristics are relevant for other drinking milestones (e.g., onset of AUD symptoms). Such studies would greatly expand our understanding of the interplay between personality and alcohol use characteristics during the formative years of adolescence and young adulthood.

ACKNOWLEDGMENTS

This research was supported in part by USPHS grants DA 05147, AA 09367, DA 13240, and DA 034606. DMB was supported by a Career Development Award from VA Office of Research and Development (Clinical Science Research & Development). BMH was supported by K01 DA 025868. None of the authors have a conflict of interest to report. The views expressed in this article are those of the authors and do not necessarily reflect the position or policy of the Department of Veteran Affairs.

REFERENCES


