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Contrasting support for alternative models of genomic variation based on microhabitat preference: species-specific effects of climate change in alpine sedges

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ABSTRACT

30
31
32 Deterministic processes may uniquely affect co-distributed species' phylogeographic patterns
33 such that discordant genetic variation among taxa is predicted. Yet, explicitly testing
34 expectations of genomic discordance in a statistical framework remains challenging. Here, we
35 construct spatially and temporally dynamic models to investigate the hypothesized effect of
36 microhabitat preferences on the permeability of glaciated regions to gene flow in two closely
37 related montane species. Utilizing environmental niche models from the Last Glacial Maximum
38 and the present to inform demographic models of changes in habitat suitability over time, we
39 evaluate the relative probabilities of two alternative models using approximate Bayesian
40 computation (ABC) in which glaciated regions are either (i) permeable or (ii) a barrier to gene
41 flow. Results based on the fit of the empirical data to datasets simulated using a spatially explicit
42 coalescent under alternative models indicate that genomic data are consistent with predictions
43 about the hypothesized role of microhabitat in generating discordant patterns of genetic variation
44 among the taxa. Specifically, a model in which glaciated areas acted as a barrier was much more
45 probable based on patterns of genomic variation in *Carex nova*, a wet-adapted species. However,
46 in the dry-adapted *C. chalciolepis*, the permeable model was more probable, although the
47 difference in the support of the models was small. This work highlights how statistical inferences
48 can be used to distinguish deterministic processes that are expected to result in discordant
49 genomic patterns among species, including species-specific responses to climate change.

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INTRODUCTION

Understanding the contribution of species-specific attributes to observed patterns of genetic variation is critical for determining why taxa responded similarly (or dissimilarly) to historical climate changes. Fortunately, newly developed, spatially explicit methodologies that generate genetic expectations under alternative scenarios (e.g., Neuenschwander *et al.* 2008; Knowles & Alvarado-Serrano 2010; Brown & Knowles 2012) offer an opportunity to evaluate the relative support for competing hypotheses related to species' life history characteristics. For example, comparing the fit of empirical data to models that capture differences in species' habitat requirements or in the habitat stability across a landscape (e.g., He *et al.* 2013) may provides insight into the role of biotic factors in structuring genetic variation. However, given that the number of biologically informed models that could be tested for any given study system is limitless, the challenge is how to decide what models to test (Knowles 2009).

Even though it might be possible to compare a hundred different models (e.g., Pelletier & Carstens 2014), such an approach may not be desirable. For example, if the differences among the models are trivial (e.g., the models differ in what may be considered nuisance parameters that do not impact their interpretation), the biological insights provided by selecting one model over another will be limited. Such inherent constraints of model-based approaches reinforce the importance of developing models that illuminate processes of biological interest (Papadopoulou & Knowles 2015a, 2016; but see O'Meara *et al.* 2015 for a dissenting point of view). In other words, it is not only the analytical approach per se, but the creativity and intimate knowledge of a study system that a researcher brings to such tests that ultimately determines how much insight a model-based phylogeographic analysis might provide. Rather than building narratives centered upon how well empirical data should fit generic scenarios (e.g., testing for a correlation under an isolation by distance model), insights can be gained by developing narratives derived from organisms' natural histories, as well as historical ecological and climatic factors, to generate testable hypotheses (Papadopoulou & Knowles 2016). Given that not all biological characteristics (e.g., dispersal capability or habitat specificities) of a species may be important, or that their effects may vary depending upon the geographic or temporal scale of study (Papadopoulou & Knowles 2016), the key is to identify those characteristics that provide insights about the processes structuring genetic variation (e.g., the role of geographic barriers or changing

81 climatic conditions). When coupled with data simulated under biologically informed models, the
82 relative fit of empirical genetic data to alternative models forms the basis for rigorous statistical
83 tests of competing hypotheses (see also Bruggeman *et al.* 2010; Epperson *et al.* 2010; Landguth
84 *et al.* 2010; Morgan *et al.* 2011; Shirk *et al.* 2012; Papadopoulou & Knowles 2015b).

85 Here we test the role of species' microhabitat preferences and their potential interactions
86 with shifting distributions associated with climatic changes in structuring the geographic
87 distribution of genetic variation in two species of montane sedges (*Carex* L., Cyperaceae).
88 Briefly, although the species co-occur within montane habitat across the southern Rocky
89 Mountains, *C. nova*, an inhabitant of wet microhabitats (i.e., occurring predominantly within
90 drainages) may have been disproportionately displaced to lower elevations (relative to dry-
91 adapted species) because of the large accumulation of snow and ice in this microhabitat during
92 glacial periods (Fig. 1). As a consequence of population persistence being limited to lower
93 elevations around the margins of glaciated habitat, populations of *C. nova* may have been
94 relatively isolated during the glacial periods that predominated throughout the Pleistocene (i.e.,
95 most of the time during the Pleistocene is represented by glacial, not interglacial, periods). This
96 narrative contrasts with plant species that inhabit meadows, ridges, and slopes, such as *C.*
97 *chalciolepis*. Inhabitants of drier microhabitats may have persisted in high elevation areas that
98 remained free of glaciers and persistent snow throughout glacial cycles (although the persistence
99 of such populations is controversial; Wachter *et al.* 2016), in addition to populations that may
100 have established at lower elevations. As a result, *C. chalciolepis* populations may have remained
101 more interconnected throughout glacial periods. Although the geographic patterning of
102 population structure observed in these two species is consistent with such a narrative (as
103 described in Massatti & Knowles 2014), evaluating the extent to which differences in the
104 patterns of genetic variation can be ascribed to microhabitat differences requires formalization of
105 these narratives into models for statistical testing.

106 To test whether current genetic structure reflects a species' ability/inability to persist within
107 higher elevation microhabitats during glacial periods, we generated genetic expectations under
108 two alternative models – one with glaciated regions as a permeable habitat and another with the
109 regions as a barrier to gene flow – using the integrative distributional, demographic, and
110 coalescent (iDDC) modeling approach (He *et al.* 2013). Specifically, we created demographic
111 models that incorporated spatial and temporal heterogeneity in climatically suitable areas, as

112 informed by present and past ecological niche models, and required that populations either
113 persist within or be excluded from glaciated areas (see also Currat & Excoffier 2004; Wegmann
114 *et al.* 2006). By performing tests of the fit of the empirical data to simulated datasets using
115 approximate Bayesian computation (ABC; see Beaumont *et al.* 2002 for an overview of ABC)
116 and tests of model validation, we evaluated whether the species differ in their support for the
117 alternative models in a manner consistent with the hypothesis that microhabitat differences
118 determine how species respond to shifts in climate. This study is not only a detailed analysis of
119 how species-specific properties may determine whether taxa respond similarly (or dissimilarly)
120 to climate change, but our work also illustrates how general narratives about the processes
121 structuring genetic variation can be formalized into models for statistical testing. We discuss our
122 findings in terms of our ability to generalize the effects of climate change on montane
123 communities as well as acknowledge general caveats with our analyses that leave open questions
124 requiring further analysis.

126 MATERIALS AND METHODS

128 *Genomic data generation and processing*

129 *Carex chalciolepis* ($N = 110$) and *C. nova* ($N = 109$) leaf material was field collected
130 from 11 sampling localities across the full extent of their geographical ranges in the southern
131 Rocky Mountains (Fig. 1; Table S1). Within populations, the collecting distance among
132 individuals was maximized to decrease the probability of sampling related individuals (average
133 distance between samples of 300 m, and a minimum distance of 35 m). Leaf material was stored
134 in silica gel until DNA was extracted with DNeasy Plant Mini Kits (Qiagen, Hilden, Germany)
135 following the manufacturer's protocol. As with previous libraries (see Massatti & Knowles
136 2014), anonymous genomic loci were developed using a restriction associated DNA sequencing
137 (RADseq) approach (for details see Peterson *et al.* 2012); library construction and data
138 processing is described in full detail in the Supporting Information. Briefly, fragments ranging in
139 size from 400 to 500 base pairs were sequenced at The Centre for Applied Genomics (Hospital
140 for Sick Children, Toronto, Canada) to generate 50 base pair, single-end reads. Single nucleotide
141 polymorphisms (SNPs) were identified using a multinomial-based likelihood model that

142 accounts for sequencing error implemented in Stacks v1.25 (Hohenlohe *et al.* 2010; Catchen *et*
143 *al.* 2011; Catchen *et al.* 2013).

144 Five Illumina 2500 sequencing runs were used to generate data for this project. In order
145 to i) maximize the number of unlinked loci, ii) reduce missing data to the fullest extent, iii)
146 maximize the number of individuals per population, and iv) ensure that the subsampled SNP
147 datasets displayed the same genetic patterns among populations as the larger, unfiltered datasets,
148 we employed the following post-processing procedures. Only RADseq loci containing up to
149 three SNPs were retained, and for each RADseq locus, one randomly selected SNP was exported
150 into a STRUCTURE-formatted file if the locus contained less than 50% missing data. Patterns of
151 genome-wide SNP variation among individuals and populations were then visualized for each
152 species with PCAs in the program R (R Core Team 2014) using the ‘adeget’ package (Jombart
153 2008) and the ‘dudi.pca’ function; missing data were replaced by the mean frequency of the
154 corresponding allele. Subsequently, we minimized missing data by manually removing SNPs and
155 individuals containing an excess of missing data and rechecking PCAs to ensure that the
156 subsampling procedure did not alter the major axes of genetic variation among populations (see
157 also Huang & Knowles 2014). Individuals of *C. chalciolepis* and *C. nova* contained an average
158 of 4.9% and 5.3% missing data, respectively (Table S2).

159 Custom scripts were used to convert the STRUCTURE-formatted files into ARLEQUIN-
160 formatted files, which were input into ARLSUMSTAT to extract the empirical summary statistics
161 used in ABC (see below). We also used the empirical ARLEQUIN-formatted files to create masks
162 that were applied to the simulated datasets so the amount and pattern of missing data in the
163 simulated data would match precisely that of the empirical datasets. All custom scripts and
164 genomic data are deposited in Dryad (doi:10.5061/dryad.ng3bv).

165 166 *iDDC approach*

167 Simulated genetic datasets to evaluate the support for the alternative models – that is, (i)
168 glaciers as barriers, and ii) permeable glaciated regions (which allowed for local persistence
169 within and gene flow across glaciated regions) (Fig. 2) – were generated using the iDDC
170 approach (described in detail in He *et al.* 2013). Briefly, simulated genetic datasets were
171 generated using a coalescent model informed by a spatially and temporally explicit demographic
172 model that reflected differences in habitat suitabilities across space and time (based on the

173 ENMs; described below), in addition to either population persistence within or exclusion from
174 glaciated areas (i.e., the hypothesized predictions for the dry and wet-adapted species,
175 respectively). The fit of the empirical genetic data to the two alternative models was evaluated
176 using approximate Bayesian computation (ABC), along with procedures to validate parameter
177 estimates and model quality (Wegmann *et al.* 2010). All scripts for the iDDC analyses are
178 deposited in Dryad (doi:10.5061/dryad.ng3bv).

179
180 Quantification of habitat suitability. Habitat suitability across the southern Rocky Mountains
181 during the present and LGM was estimated for *C. chalciolepis* and *C. nova* with MAXENT v3.3.3e
182 (Phillips *et al.* 2006). Nineteen bioclimatically informative variables for the present (WorldClim
183 v1.4; Hijmans *et al.* 2005) and the LGM (PMIP2-CCSM; Braconnot *et al.* 2007) were used to
184 generate the environmental niche models (ENMs; full details are presented in the Supporting
185 Information). Georeferenced distribution points representative of each species' entire range were
186 collected from personal fieldwork and validated voucher specimens housed at the Rocky
187 Mountain Herbarium (species distribution points are available at doi:10.5061/dryad.ng3bv). To
188 have a computationally tractable number of cells for demographic simulations (detailed below),
189 we statistically downscaled the cell sizes of the ENMs to 0.42 decimal degrees ($\sim 16.5 \text{ km}^2$ per
190 cell) (e.g., Ray *et al.* 2010; He *et al.* 2013). Subsequently, the values of the cells in the LGM and
191 present ENMs denoting the logistic habitat suitability scores (ranging continuously from 0 to 1)
192 determined by MAXENT were reassigned. Specifically, the logistic values were grouped into ten
193 categories using the 'equal interval' clustering method in ArcMap 10.0 (ESRI, Redlands,
194 California, USA) and assigned values ranging from 1-10. Given the extreme similarity of the
195 ENMs between the species (for both the present and past; Fig. S1), an average of the habitat
196 suitability scores of the two species was generated for each time period and used for the
197 demographic simulations (see details below). By using estimates of habitat suitabilities for the
198 present and past based on averages for the two species, we provide a standardized model that
199 avoids the confounding influences of subtle differences in the unique ENMs of the taxa that
200 could contribute to the relative probabilities of the alternative models (as opposed to differences
201 in the permeability of the glaciated areas themselves, which is the central focus of the study). In
202 addition to the present and LGM ENMs that were generated for each species, landscapes of
203 habitat suitabilities specific to the two alternative models were generated: one in which the

204 glaciated areas were a barrier and one in which these regions remained permeable (detailed
205 below), corresponding to the hypothesized effect of differences in microhabitat (i.e., glaciated
206 regions would have been permeable in the dry, but not the wet-adapted species). Note that
207 because microhabitat differences are manifest at a small spatial scale (i.e., meters or less) due to
208 the complex interaction between topography and environmental conditions in montane regions,
209 the “habitat” suitabilities of the taxa based on quantification from the ENMs are similar and only
210 convey that environmental conditions are suitable for the species occurrence.

211
212 Simulated datasets. Demographic and coalescent simulations were performed in SPLATCHE2 (Ray
213 *et al.* 2010) to generate patterns of genetic variation where habitat suitabilities differed across a
214 landscape and through time (see Knowles & Alvarado-Serrano 2010; Brown and Knowles 2012).
215 For each of the two models tested (i.e., glaciers as barriers vs. permeable habitat; Fig. 2), 1 000
216 000 simulated datasets were generated for each species separately (i.e., 4 000 000 total). Uniform
217 priors were used for the demographic parameters (i.e., between-deme migration rate, m ,
218 maximum carrying capacity of a deme, K , and the population sizes of the initial populations,
219 N_{Anc}), and the carrying capacities of demes were scaled proportionally to their habitat suitability
220 values. In other words, data were simulated over a range of maximum K -values, with the highest
221 quality habitat (demes with values of 10) reaching full carrying capacity (100%), while carrying
222 capacities of demes associated with lower quality habitat were decreased proportionally (see
223 Knowles & Alvarado-Serrano 2010; Fig. 2). Each generation, m proportion of the population
224 migrates out of the local deme to adjacent cells (i.e., to the north, south, west, and east). After the
225 exchange of individuals, local demes grow logistically at the rate of 1, regulated by the carrying
226 capacity inferred from the habitat suitability. Demographic simulations were initialized with a
227 southern, central, and northern population (see Fig. 2), each with a population size of N_{Anc} . Note
228 that population trees estimated using empirical SNP data for *C. chalciolepis* and *C. nova*
229 clustered populations sampled from these regions, validating our use of three initial populations.

230 The landscapes that informed the demographic modeling procedure differed over time.
231 Specifically, for both the barrier and permeable models, an initial landscape based on the ENM
232 at the LGM was used to inform the demographic simulations from generations 1-750. The initial
233 landscape was followed by a model-specific landscape (i.e., a landscape representing either (i)
234 glaciated areas as barriers or (ii) permeable glaciated areas) from generations 751-1750, which

235 corresponded to the height of the last Pleistocene glaciation. For the final period from
236 generations 1751-2083, the demographic modeling was informed by a landscape with habitat
237 suitabilities based on an ENM estimated using contemporary climatic conditions for both models
238 (Fig. 2). With a generation time of 3 years for these high elevation species (Körner 2003) and the
239 scaling of the generations by a factor of 15 (to make simulations computationally tractable), the
240 simulations span a time period from the present to the LGM and beyond. Note that modeling was
241 extended beyond the LGM (i.e., the first 750 generations) to provide a sufficient amount of time
242 for all suitable habitat to be occupied over all possible combinations of population demographic
243 parameters. Because of this scaling, any biological interpretation of absolute values of population
244 genetic parameters would need to be adjusted accordingly.

245 As noted above, the initial and final landscapes used to perform the demographic
246 simulations were the same between the models (and species), whereas the landscape
247 corresponding to the height of the LGM for the two models differed in one key aspect: the
248 habitat suitability values of the glaciated areas (which were identified from maps of glacial
249 moraines and glacial till; see Ehlers & Gibbard 2004; Colorado Geological Survey
250 <http://coloradogeologicalsurvey.org/>). Specifically, under the model in which glaciers would
251 have acted as barriers, demes reconstructed as glaciated were assigned a K -value of 0. In
252 contrast, in the alternative model in which such glaciated areas were permeable, the K -values
253 were decreased by 85% of the original value estimated from the LGM ENM, which corresponds
254 to impeded dispersal relative to the surrounding non-glaciated areas (Fig. 2), with a lower bound
255 of $K = 20$. This lower bound was used because of the uncertainty surrounding habitat quality
256 estimates in demes where habitat suitability scores fell below the maximum training sensitivity
257 plus specificity threshold identified by MAXENT. Varying the K -values for demes with highly
258 uncertain habitat suitability could introduce demographic consequences that would have undue
259 influence on the resulting patterns of genetic variation. Preliminary analyses confirmed that gene
260 flow was able to occur among populations separated by glaciers despite reduced carrying
261 capacities. Because *C. chalciolepis* and *C. nova* were predicted to occur in and adjacent to
262 montane habitat during the LGM, but not in the geographically distant lower elevation basins
263 and plains (which were predominantly represented by demes with values of 1), all demes with
264 habitat suitability values of 1 were assumed to be uninhabitable (see Fig 2).

265 A spatially explicit coalescent model informed by the deme-specific demographic
266 parameters was used to simulate genetic data (i.e., genetic variation differed across the landscape
267 depending on the specific combination of m , K , and N_{Anc} ; Excoffier *et al.* 2000; Currat *et al.*
268 2004). An independent coalescent process was run to generate a genealogy for each locus
269 analyzed in the empirical data (i.e., 1142 and 1010 coalescent simulations for *C. chalciolepis* and
270 *C. nova*, respectively) for each of the 4 000 000 datasets simulated with different combinations
271 of demographic parameters (the range of which as specified by the priors). Each simulated
272 dataset consisted of the same number of individuals from the same the corresponding sampling
273 localities, with the same amount of missing data, as in the empirical data.

274 For each of the simulated genetic datasets, as with the empirical data, nine summary
275 statistics were calculated using ARLSUMSTAT v.3.5.2 (Excoffier & Lischer 2010). These included
276 the number of segregating sites (S) for each population and across populations, mean
277 heterozygosity across loci for each population and across populations (H), and pairwise
278 population F_{ST} (Weir & Cockerham 1984), for a total of 83 summary statistic values calculated
279 per simulated dataset.

280
281 Model selection and validation. Approximate Bayesian computation (ABC) was used to select
282 between alternative models, as implemented with ABCestimator in ABCtoolbox (Wegmann *et*
283 *al.* 2010). Rather than using the 83 summary statistic values independently to estimate
284 parameters, partial least squares (PLS) components (Boulesteix & Strimmer 2007) were
285 extracted from the summary statistics using the “PLS” package (Mevik & Wehrens 2007) with
286 boxcox treatment (Box & Cox 1964) in R for the first 20 000 simulations for each model and
287 species. This approach removes the effects of interactions between summary statistics and
288 reduces “the curse of dimensionality” associated with using a large number of summary statistics
289 (Boulesteix & Strimmer 2007). We examined the root mean squared error (RMSE) prediction for
290 each parameter to decide how many PLS components to use for parameter estimation (Fig. S2).
291 Of the million simulated datasets generated for each species under a particular model, 5000
292 simulations (0.5%) whose summary statistics were closest to those calculated from the empirical
293 genomic data were retained and used for parameter estimation and model selection. Post-
294 sampling regression adjustment was applied using the ABC-GLM (general linear model)
295 function (Leuenberger & Wegmann 2010) in R to obtain posterior distributions of the

296 parameters. Bayes factors, which are the ratios between marginal densities of the two models,
297 were used for model selection; a higher ratio indicates more support for the first model (Jeffreys
298 1961).

299 To evaluate whether a model is capable of generating data similar to the empirical data,
300 the likelihood of the empirical data was compared to the likelihoods of the retained simulations
301 under the GLM model. If all the retained simulations have a better likelihood than the observed
302 data (i.e., a low P -value), it would indicate a model is highly unlikely (Wegmann *et al.* 2010). A
303 coefficient of variation (R^2) of each parameter explained by the PLS components was also
304 computed and used as an indicator of the power of estimation (Neuenschwander *et al.* 2008). For
305 the most probable model selected for each species, the accuracy of parameter estimates was
306 validated using 1000 pseudo-observations generated from prior distributions of the parameters. If
307 estimation of the parameters is unbiased, posterior quantiles of the parameters from pseudo runs
308 should be uniformly distributed (Cook *et al.* 2006; Wegmann *et al.* 2010). The posterior
309 quantiles of true parameters for each pseudo run were also calculated based on the posterior
310 distribution of the regression-adjusted 5000 simulations closest to the pseudo-observation.

311

312 RESULTS

313

314 *Empirical genomic dataset*

315 Almost 180 000 000 (average 1 630 280 \pm 719 977 per individual) and 172 000 000
316 (average 1 579 380 \pm 733 329 per individual) reads were generated for the 110 and 109 *C.*
317 *chalciolepis* and *C. nova* individuals, respectively (Table S2). The retention of reads after data
318 processing and assembly with Stacks averaged 86% per individual for both species. The final
319 datasets contained 1142 loci with 1 SNP per locus across 101 individuals of *C. chalciolepis* and
320 1010 loci with 1 SNP per locus across 99 individuals of *C. nova*, after post-processing steps to
321 remove loci and/or individuals because of missing data. The datasets had good representation of
322 all sampled populations for tests with the iDDC approach (Table S3).

323

324 *Model selection and validation*

325 Based on the relative information content contained in the PLS components, the first six
326 PLSs of the predictor variables (i.e., the summary statistics) were retained for ABC analyses (see

327 Fig. S2). Posterior distributions of parameter estimates were distinct from the prior, indicating
328 they contained information relevant to estimating the parameters (Fig. 3). Moreover, comparison
329 of the posterior distributions before and after GLM regression adjustment of the 5000 retained
330 simulations shows a marked improvement of parameter estimates with the GLM regression (Fig.
331 3). The accuracy of the parameter estimates varied (Table 1). Specifically, in addition to a flatter
332 posterior probability (Fig. 3), the lowest power for estimating parameter values was associated
333 with the ancestral population size (N_{Anc}), as indicated by the lowest R^2 values across models
334 (Table 1). Nevertheless, tests of potential bias of parameter estimates show that posterior
335 distributions of N_{Anc} are uniformly distributed in both species (Fig. 4). In contrast, the histograms
336 of the posterior quantiles of m and K did deviate significantly from a uniform distribution for
337 both species based on analyses of 1000 pseudo-observed datasets, suggesting a potential bias
338 (Fig. 4).

339 Based on the marginal densities calculated from the 5000 retained simulations for each
340 model, the model with the best fit to the empirical data (based on Bayes factors) differed
341 between species. Specifically, the model with glaciated areas as barriers was more probable for
342 *C. nova*, whereas the model with permeable glaciated areas was more probable for *C.*
343 *chalciolepis* (although in the latter case, the difference was marginal), which corresponds to the
344 hypothesized effect of microhabitat differences (Table 1). Evaluation of whether the most
345 probable model was capable of producing the empirical data in each species (i.e., consideration
346 of the P -values) indicated that not only were they a good fit, but they were a better fit (i.e., a
347 larger P -value) than the alternative models, suggesting a better correspondence between the
348 empirical data and the simulated data (Table 1).

349

350

DISCUSSION

351

352 Our study highlights the formalization of biologically informed hypotheses into a statistical
353 model-testing framework (see Papadopoulou & Knowles 2016) to evaluate whether species-
354 specific traits may mediate the effects of climate change. Specifically, we show that species'
355 responses to glaciations may be intricately tied to their microhabitat preferences. Below we
356 discuss the implications of our findings when considering the effects of climate change on co-
357 distributed taxa, and especially on species in montane communities.

358 In contrast to studies that rely upon correlative approaches to identify factors that may
359 structure genetic variation (see Massatti & Knowles 2014), here the link between genetic patterns
360 and process is explicit (i.e., data were simulated with either population persistence within or
361 exclusion from glaciated areas – the hypothesized predictions for the dry and wet-adapted
362 species, respectively, based on the differential accumulation of glaciers on ridges and drainages;
363 see Fig. 1) and the fit of empirical data to alternative models was statistically evaluated.
364 Moreover, our study differs from other model-based approaches in which a generic or “naïve”
365 model is applied (i.e., a model that is not informed by biological knowledge of a particular
366 system) (Pelletier & Carstens 2014; Grummer *et al.* 2015; O’Meara *et al.* 2015); we designed our
367 study to target species that differed in one key trait – microhabitat. Minimizing differences
368 between taxa allowed us to explicitly test hypotheses aimed at providing specific insights about
369 the contribution of species-specific traits to patterns of genetic variation (albeit with some
370 caveats; see below). In addition to discussing the utility of the iDDC approach to test if
371 phylogeographic discord reflects deterministic factors, we address both the challenges and
372 limitations with such inferences. Lastly, we highlight the insights that comparative
373 phylogeography can provide regarding the role of biotic factors in structuring genetic variation.

374

375 *Ecological and evolutionary implications for montane taxa*

376 The finding of a correspondence between a model in which glaciated areas were barriers
377 to gene flow and genomic variation in the wet-adapted *C. nova*, as opposed to a model in which
378 glaciated areas were permeable, as with genomic variation in the dry-adapted *C. chalciolepis*
379 (Table 1), has a number of ecological and evolutionary implications. From an ecological
380 perspective, our work highlights how generalizations about the response of montane
381 communities to climate change may be overly simplified (see also Alexander *et al.* 2016).
382 Importantly, the structure of genomic variation of the *Carex* species differs significantly (i.e., the
383 parameterization and fit of the alternative models differed between the taxa; Table 1), despite
384 both sedges being common and co-distributed (Massatti *et al.* 2016) and having many biological
385 similarities, including traits associated with dispersal capabilities. Moreover, the respective fit of
386 the dry-adapted versus wet-adapted species to a model with either population persistence within
387 or exclusion from glaciated areas is in line with expectations derived from knowledge of the
388 interactions of persistent snow and ice with microhabitats in montane ecosystems (Ehlers &

389 Gibbard 2004). As such, this is the first study to directly test causal mechanisms related to the
390 persistence of populations in glaciated regions (Lohse *et al.* 2011; Westergaard *et al.* 2011;
391 Lanier *et al.* 2015; Wachter *et al.* 2016). Specifically, differing support for the most likely model
392 between the species (Table 1) suggests that species adapted to wetter microhabitats were more
393 isolated around the margins of glaciers, whereas species adapted to drier microhabitats persisted
394 within glaciated regions (as well as establishing at lower elevations) and remained relatively
395 connected by gene flow.

396 It is generally accepted that statistical evaluation of alternative processes that might have
397 generated observed patterns of genetic variation (e.g., through model selection procedures, as
398 used here) is critical because similar genetic patterns may result from different demographic
399 processes (Csillery *et al.* 2010; He *et al.* 2013). However, our study is more than an example of
400 model-based statistical phylogeography (Knowles 2009). Specifically, in comparison with other
401 comparative phylogeographic analyses, our study stands out because a deterministic process that
402 is expected to generate discordant genomic variation among taxa is evaluated (i.e., the fit of
403 alternative models is predicted to differ depending on a taxon's microhabitat, corresponding to
404 the hypothesized persistence in or exclusion from previously glaciated areas). This contrasts with
405 the tradition of relying upon phylogeographic concordance for assessing the role of competing
406 processes, and where discord is commonly attributed to the idiosyncrasies of history (reviewed in
407 Papadopoulou & Knowles 2016). Nevertheless, there are some caveats with our approach.
408 Specifically, while we have conducted analyses aimed at addressing the quality of our inference
409 (discussed in the following section), there are additional hypotheses that we cannot rule out. Note
410 that all model-based comparative phylogeographic studies face this issue (Knowles 2009) – that
411 is, this caution is not unique to our study, or specific to the iDDC procedure per se. Moreover,
412 from our perspective, such a discussion is extremely helpful for considering analyses that may be
413 worth pursuing in the future to provide additional insights into the role of biotic factors
414 structuring genetic variation. In this regard, we would argue that the merit of our study, and of
415 comparative phylogeographic studies more generally, is not in identifying “the evolutionary
416 history” of taxa. Instead, it is the insights gained by identifying the probability of one hypothesis
417 relative to others, which is a function of the both the study design and its execution. As noted
418 above, and given that patterns of genetic variation in many species are postulated to reflect the
419 interactions of ecology and climatic oscillations (Avice & Walker 1998; Hewitt 2000; Carstens

420 & Knowles 2007), our results set the stage for additional work that is needed before any broad
421 generalizations can be derived from model-based comparative phylogeographic analyses about
422 the role of species-specific traits in structuring genetic variation (see also Papadopoulou &
423 Knowles 2016).

424

425 *Validation and interpretation of model-based inferences*

426 There are both methodological and conceptual aspects of model-based inferences that
427 must be considered to avoid erroneous conclusions or tests that offer limited insights. For
428 example, there are a number of methodological issues that can be especially challenging with
429 ABC procedures (see Oaks *et al.* 2013) and the iDDC approach we applied. These include issues
430 with approximating the likelihood of models with summary statistics (Pritchard *et al.* 1999;
431 Beaumont *et al.* 2002) as opposed to using all of the data as in full likelihood-based models (Hey
432 & Nielsen 2004, 2007; Kuhner 2006; Nielsen & Beaumont 2009; Hey 2010). Moreover, in
433 addition to difficulties associated with particular steps in ABC procedures (e.g., post-sampling
434 adjustment when the relationship between parameters and summary statistics is extrapolated
435 beyond the region of the observed data set; see Beaumont *et al.* 2002; Leuenberger & Wegmann
436 2010), ABC will always produce a posterior distribution, even if the model is a poor fit to the
437 data (Bertorelle *et al.* 2010). Accordingly, model validation is critical.

438 Several approaches we applied suggest that our results are generally robust. We evaluated
439 the potential bias in parameter estimates (Fig. 4), contrasted the posterior probability of
440 parameter estimates pre and post-GLM (Fig. 3) and relative to the prior, as well as utilized the
441 RMSE of parameter estimates (Fig. S2) to inform decisions about the inclusion of PLS
442 components. While the predictive power of the data for some parameters differed (see R^2 in
443 Table 1) and the posterior quantiles calculated from pseudo observed datasets of both m and K
444 showed a significant departure from a uniform distribution (Cook *et al.* 2006; Wegmann *et al.*
445 2010), we note that for both species, the most probable model provided not only (i) a good fit to
446 the empirical genomic data, but (ii) it was a better fit compared to the less probable model (Table
447 1). Specifically, the Bayes factor indicates strong support for the barrier model in *C. nova*, and
448 this model has a much higher probability of generating simulations with likelihood-values
449 comparable to the empirical data compared to the permeable model (Table 1). In *C. chalciolepis*,
450 the marginal densities of the two models are much more similar. As a consequence, although the

451 empirical data is more probable under the permeable model, the difference in support based on
452 the Bayes factor is not strong. Nevertheless, it is worth noting that, even though some
453 combinations of parameters produce datasets that match the *C. chalciolepis* empirical data under
454 the barrier model, the permeable model has a much wider parameter region that generates data
455 close to the empirical data (as reflected in differences in the *P*-values for the two models; Table
456 1). The fit of the empirical data under these complex models is very encouraging because it can
457 be difficult to capture the complicated nature of a species' history. For example, despite
458 approaches for evaluating complex models, the likelihood of the empirical data under the most
459 probable model may be much lower than any data simulated under such a model (see Excoffier
460 *et al.* 2013).

461 Even though our models are capable of generating the data (Table 1), this does not mean
462 the most probable models for the two species are the “correct” ones. We acknowledge that there
463 could be other models not considered here that might fit the empirical data. However, this does
464 not discount the insights gained with respect to the goal of the study, which was to test whether
465 the empirical data of two species would support alternative models as predicted if the
466 microhabitats of the species mediated their responses to climate change. In this regard,
467 comparison of the estimated parameter values may illuminate possible differences in the
468 population dynamics of the species under climate change scenarios. In particular, we note that *C.*
469 *chalciolepis* tended toward higher values of K_{max} and lower values of m compared to *C. nova*,
470 while the difference in N_{Anc} was more ambiguous because of uncertainty in the estimation of this
471 parameter (Fig. 3). Specifically, the PLSs of the summary statistics were informative for all
472 parameters in both species, but estimates of N_{Anc} are associated with only moderate R^2 values, in
473 contrast to the very high R^2 for the other parameters (Table 1). In the context of glaciations, these
474 combinations of parameter values may intimate that habitat stability for dry-adapted species
475 facilitated larger effective population sizes and lower rates of migration, while continual
476 disturbance within wet microhabitats fostered relatively lower population sizes and higher rates
477 of migration in wet-adapted species (e.g., the reestablishment of populations in disturbed
478 habitats). Our analyses support this proposition, but the detected biases in some parameter
479 estimations (Fig. 4) cautions against interpreting the parameters directly (see Wegmann *et al.*
480 2010).

481 Evaluating potential demographic differences between the species under changing
482 climatic conditions is an important area for future consideration. Such investigations may include
483 tests of whether support for alternative models reflects differences in how estimates of habitat
484 suitability (as informed by ENMs) scale to population demographic parameters. For example, a
485 particular value of habitat suitability may not translate into equivalent predicted carrying
486 capacities between *Carex* species (as modeled here). While there are not pronounced differences
487 in the contemporary abundances of *C. chalciolepis* and *C. nova* that suggest a nonequivalent
488 relationship between the suitability of a habitat and local population sizes, differences in the
489 local stability of the species' populations, or even local adaptation, could potentially contribute
490 to the different patterns of genetic variation in ways not explicitly accounted for in our models.
491 Likewise, we do not directly model microhabitat preference per se, but instead test models with
492 either population persistence or exclusion from glaciated areas based on the hypothesized
493 predictions for the dry and wet-adapted species, respectively, based on the differential
494 accumulation of glaciers on ridges and drainages (see Fig. 1). It is possible that factors other than
495 microhabitat preference might contribute to the persistence or exclusion of taxa from glaciated
496 areas. Hence, it may not be microhabitat preference, but possibly some untested co-varying
497 explanatory variable, that drives the differences in the fit of the taxa to the alternative models.
498 However, the similarity of the taxa makes it difficult to identify other hypothetical, yet realistic,
499 factors. Moreover, any such hypothetical factor would not only have to result in contrasting
500 support for the alternative models between the taxa, but also preserve the directionality of the
501 model fits (e.g., *C. chalciolepis*, but not *C. nova*, must fit the model with persistence within
502 glaciated areas, and vice a versa for the model with exclusion from glaciated areas). We note that
503 other species that are closely related to and co-distributed with *C. chalciolepis* and *C. nova*
504 (Massatti *et al.* 2016) will facilitate the exploration of these intriguing hypotheses, but such tests
505 are beyond the scope of our present analyses.

506

507 *Conclusions*

508 Using intimate knowledge of the interactions between climate and topography within
509 montane ecosystems, as well as utilizing data from other disciplines (e.g., maps of glacial till and
510 glacial moraines; Fig. 1), we tested alternative models to elucidate the potential impact of
511 glaciers on co-distributed species. We evaluated the relative fit of empirical data under a model

512 in which glaciated areas were a barrier versus one in which they were permeable to test the
513 biologically informed hypotheses that differences in microhabitat preferences would result in
514 predictable differences in the responses of the taxa to climate change. Our results supported the
515 hypothesized predictions for the dry and wet-adapted species based on the differential
516 accumulation of glaciers on ridges and drainages (Fig. 1) – the barrier model was the most
517 probable for *C. nova*, whereas the model with permeable glaciated regions was more probable in
518 *C. chalciolepis* (although in the latter case the difference was not strong) (Table 1).

519 The models and approach we apply here go beyond traditional analyses common in
520 phylogeography (e.g., tests of isolation by distance; Slatkin 1993) and comparative
521 phylogeography (e.g., relying on concordant patterns for inferring the role of factors in
522 structuring genetic variation; Papadopoulou & Knowles 2015a). Such approaches are in their
523 infancy and have only been applied in a limited number of studies (e.g., Neuenschwander *et al.*
524 2008; He *et al.* 2013; Martinkova *et al.* 2013). By combining the power that genomic data
525 provide with the proper validation of complex models, approaches such as iDDC provide an
526 exciting opportunity to address ecological and evolutionary principles in a comparative
527 phylogeographic framework that cannot be addressed using traditional methodologies (Excoffier
528 *et al.* 2013; Papadopoulou & Knowles 2016).

529
530

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756 **Data Accessibility**

757 The following data, scripts, and other files will be deposited in Dryad: doi:10.5061/dryad.ng3bv

758 - Species distribution points used in MAXENT

759 - Empirical datasets (composed of >1000 SNPs for each species)

760 - Scripts for editing empirical and simulated SNP datasets

761 - Scripts and settings files used in iDDC analyses

762

763 **Author Contributions**

764 RM collected specimens, conducted lab work, performed modeling, and wrote the manuscript.

765 LLK helped with the development of the experimental design and the modeling scenarios as well

766 as wrote and edited the manuscript.

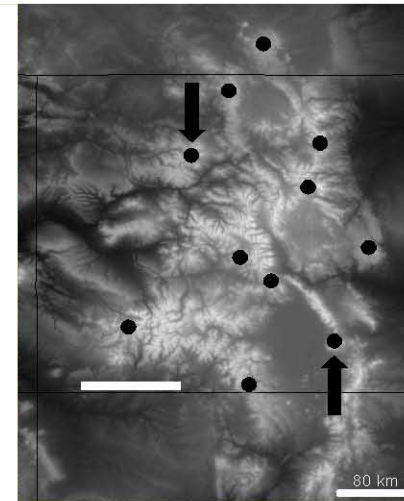
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767 Table 1. Model statistics for each species and modeling scenario. The Bayes factor represents the ratio between the model with the
 768 highest marginal density and the alternative model. Note that the P -value refers to the general fit of the data under a model (i.e., it is
 769 based on the likelihood of the retained simulated datasets relative to the likelihood of the empirical data, where a high P -value
 770 indicates the model is capable of generating the data). Parameters include: K_{max} , the carrying capacity of the deme with the highest
 771 suitability; m , the migration rate per deme per generation; and N_{Anc} , the ancestral population sizes of initial populations before
 772 expansion from refugia. R^2 is the coefficient of determination between a parameter and the six PLSs used herein.

773

Species	Model	Marginal density	P -value	Bayes factor	Parameters	R^2
<i>C. chalciolepis</i>	Barrier	4.87×10^{-5}	0.650	-	K_{max}	0.642
					m	0.966
					N_{Anc}	0.404
	Permeable	1.38×10^{-4}	0.970	2.84	K_{max}	0.698
					m	0.965
					N_{Anc}	0.379
<i>C. nova</i>	Barrier	1.29×10^{-4}	0.844	22.69	K_{max}	0.548
					m	0.961
					N_{Anc}	0.497
	Permeable	5.68×10^{-6}	0.078	-	K_{max}	0.585
					m	0.962
					N_{Anc}	0.479

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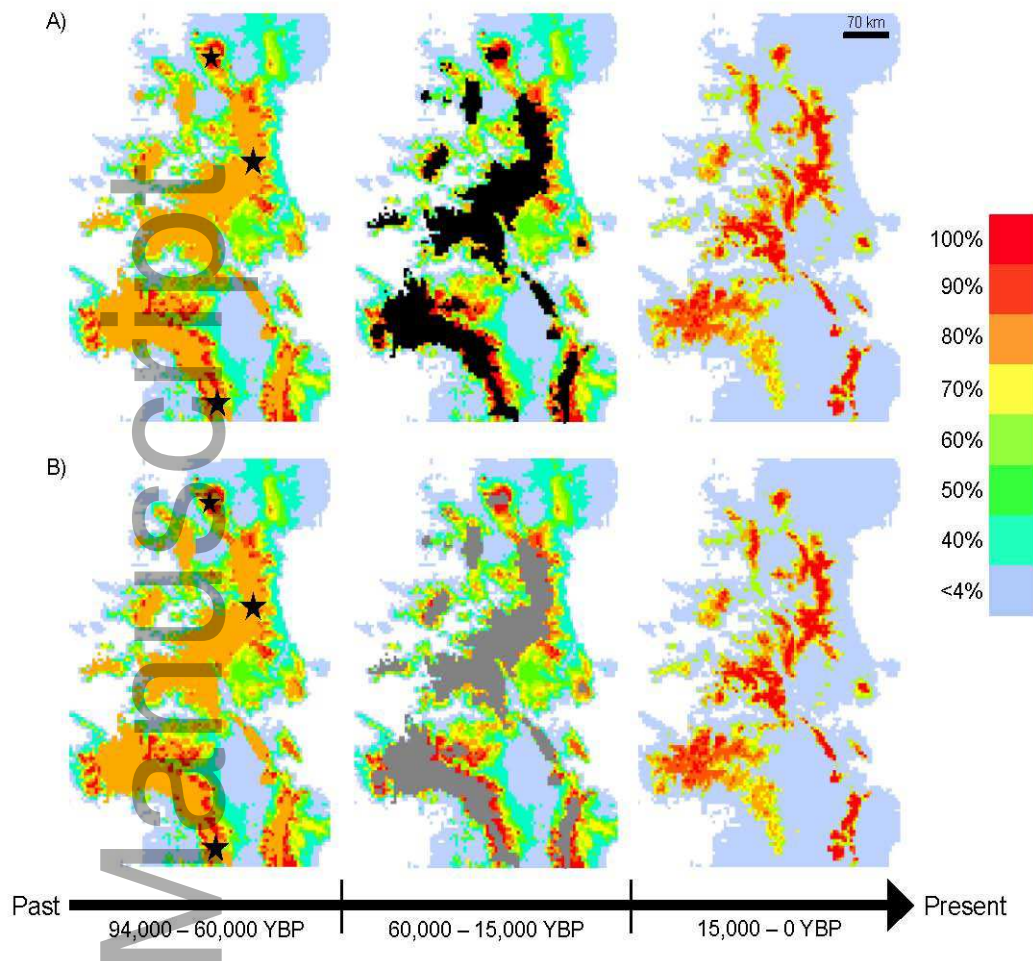
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775 Figure 1. Glacial reconstruction for Colorado during the Last Glacial Maximum based on geologic data such as glacial moraines. Note
776 the differential accumulation of glaciers in drainages versus on ridges. Circles represent collecting localities; arrows identify matching

Microhabitat preference & climate effects

777 localities between the glacial map and the inset (for details, see Table S1). The image is taken from the 'Late Pleistocene glaciers of
778 Colorado' video (Interactive Geology Project, University of Colorado at Boulder, <http://igp.colorado.edu/>).

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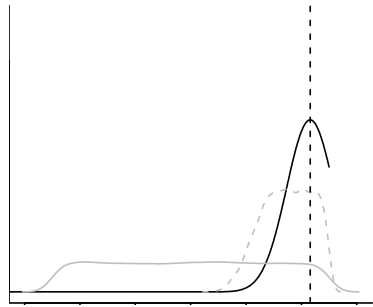


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780 Figure 2. Alternative models used to generate patterns of genetic variation. Both the past and
 781 present landscapes were the same between the models, while the intermediate landscape differed
 782 in how the glaciated areas were modeled. Specifically, under the barrier model (A), the glaciated
 783 areas had a habitat suitability of 0 (denoted by black), whereas under the permeable model (B),
 784 the glaciated areas remained permeable (albeit at a reduced carrying capacity, denoted by grey;
 785 see text for details). The local carrying capacities of demes differed across the landscapes and
 786 were scaled based on habitat suitabilities estimated from ENMs (see colored scale bar). The
 787 black stars in the initial landscape mark the locations of ancestral populations used to initiate the
 788 simulations.

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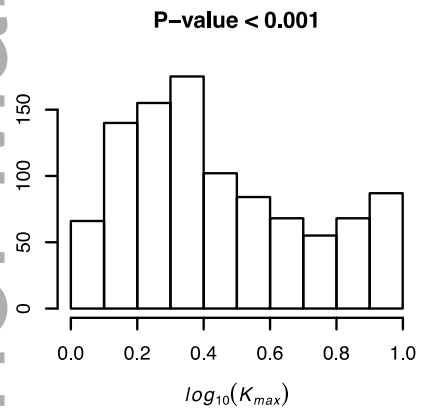
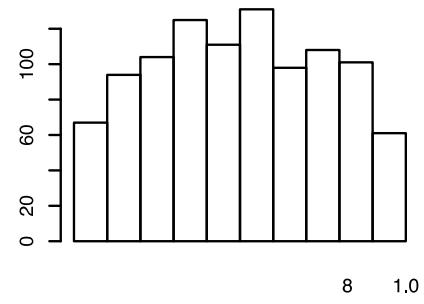
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815 Figure 3. Posterior distribution (black line) and mode (vertical dotted line) of parameter estimates for the most probable model for *C.*
816 *chalciolepis* (permeable model) and *C. nova* (barrier model). Results are based on a GLM regression adjustment of the 5000 retained
817 simulations. The distribution of the retained simulations (dashed line) and the prior (gray line) demonstrate the improvement that the
818 GLM procedure had on parameter estimates and that the data contained information relevant to estimating the parameters.

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846 Figure 4. Distribution of posterior quantiles of parameters for *C. chalciolepis* under the permeable model (A) and *C. nova* under the
847 barrier model (B) to evaluate potential bias in the parameter estimates (results are shown only for the most probable models). Bias is
848 measured by a departure from a uniform distribution using a Kolmogorov-Smirnov test (a P -value <0.05 indicates a non-uniform
849 distribution). Analyses are based on 1000 pseudo-observations (see text for details). Estimation of N_{Anc} is unbiased while the
850 distributions for K and m are too wide for both species.

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