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The Informational Role of the Media in Private Lending

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ABSTRACT

We investigate whether a borrower's media coverage influences the syndicated loan origination and participation decisions of informationally disadvantaged lenders, loan syndicate structures, and interest spreads. In syndicated loan deals, information asymmetries can exist between lenders that have a relationship with a borrower and less informed, nonrelationship lenders competing to serve as lead arranger on a syndicated loan, and also between lead arrangers and less informed syndicate participants. Theory

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suggests that the aggressiveness with which less informed lenders compete for a loan deal increases in the sentiment of public information signals about a borrower. We extend this theory to syndicated loans and hypothesize that the likelihood of less informed lenders serving as the lead arranger or joining a loan syndicate is increasing in the sentiment of media-initiated, borrowerspecific articles published prior to loan origination. We find that as media sentiment increases (1) outside, nonrelationship lenders have a higher probability of originating loans; (2) syndicate participants are less likely to have a previous relationship with the borrower or lead bank; (3) lead banks retain a lower percentage of loans; and (4) loan spreads decrease.

JEL codes: G12; G21; M40; M41

Keywords: media; relationship lending; syndicate structure; lead arranger; sentiment

1. Introduction

The business press is an important source of public information about a firm. There is substantial evidence that the business press provides information about firms' fundamentals to equity market participants incremental to that provided by other information intermediaries and by accounting data (e.g., Tetlock, Saar-Tsechansky, and Macskassy [2008]), and that it reduces information asymmetries between equity investors (e.g., Bushee et al. [2010]). However, there is little research examining the role played by the media in private lending markets. In this paper, we investigate the extent to which media coverage of borrowers influences the loan origination and participation decisions of informationally disadvantaged lenders, loan syndicate structures, and interest rate spreads.

A defining feature of equity markets is that investors do not generally have privileged access to firms' confidential information, as securities law prohibits unequal access to such information. This creates scope for the media to serve as an information intermediary for equity investors seeking information from sources external to the firm.¹ In contrast, private debt markets are not subject to securities laws, allowing lenders significant access to borrowers' private information as an integral part of the lending process. Syndicated loans involve multiple lenders where the degree of access to borrowers' inside information can differ substantially across lenders. Such differential access to information creates information asymmetries between lenders at two levels.

First, syndicated lending involves information asymmetry between inside lenders with a prior relationship with a borrower and outside, nonrelationship lenders competing for a mandate to serve as the lead arranger on a

¹Bushee et al. [2010] define an information intermediary as an agent that provides information that is new and useful to other parties because it has either not been publicly released or widely disseminated.

borrower's loan.² Nonrelationship lenders are typically at an information disadvantage relative to relationship lenders. Through sustained engagement, relationship lead banks gain extensive inside knowledge of a borrower's operations and develop private channels of communication with its managers. Borrowers are also inclined to reveal sensitive private information to relationship lenders (e.g., Greenbaum and Thakor [1995], Boot [2000], Bharath et al. [2009]).³ Second, information asymmetries exist between lead banks and other syndicate participants (we use the terms lead arranger and lead bank interchangeably). Although participants can receive some private information from the lead bank or borrower, they are generally at an information disadvantage relative to the lead bank because they tend to maintain an arm's length relationship with the borrower, while the lead bank has the primary due diligence and monitoring responsibilities (e.g., Lee and Mullineaux [2004], Sufi [2007], Ivashina [2009]).

Outside lenders, analogous to equity investors, have incentives to seek credit-relevant information about a borrower to mitigate their information disadvantage relative to more informed inside lenders.⁴ While media is unlikely to fully remove an inside lender's private information advantage, it can reduce *relative* information asymmetries across lenders and influence loan syndicate formation and pricing by transmitting credit-relevant information to less informed lenders that is incremental to information received from other sources. We investigate the informational role of the media in syndicated lending by employing a quantitative measure of media sentiment in business press articles about a borrower. Using data from RavenPack News Analytics, we construct a measure that reflects the average sentiment across all full-size borrower-specific articles published over the six-month period preceding loan origination (Media Sentiment). Raven-Pack's sentiment score reflects assessments of the tone of the news in a given article (i.e., positive or negative news), as well as the strength of the news the article contains.

We first investigate if the media influences competition between inside and outside lenders for the mandate to serve as the lead arranger on a borrower's loan. This analysis builds on the model from Rajan [1992] in which

²A lead arranger establishes a relationship with the borrower and has primary responsibility for information collection, ex ante due diligence, distributing shares of the loan to syndicate participants and ex post monitoring, among other duties. The lead arranger receives a fee for arranging and managing the syndicated loan. We discuss institutional aspects of syndicated lending in more detail in section 2.1.

³ The idea that inside banks have an information advantage is well established in the literature (e.g., Kane and Malkiel [1965], Fama [1985], Greenbaum, Kanatas, and Venezia [1989], Sharpe [1990], Rajan [1992], Petersen and Rajan [1994, 1995], Dell'Ariccia and Marquez [2004]).

⁴For parsimony, we often refer to lead arrangers with a relationship to a borrower and participant banks with a relationship to the lead arranger or a borrower as inside lenders, and other lenders as outside lenders. The premise is that inside lenders are better informed than are outside lenders.

an informed (inside) bank in a current lending relationship with a risky borrower competes with less informed (outside) banks to supply new financing to a borrower. Inside banks exploit their information advantage by opportunistically bidding only for good loans while avoiding bad ones. Because inside banks avoid bad loans, outside lenders are exposed to adverse selection as they face a significant risk of loss if they bid. Outside lenders respond by moderating how aggressively they compete, becoming more reluctant to participate as their assessed probability that a loan is bad increases. Rajan [1992] further considers how public information signals about a borrower shape competitive dynamics, showing that outside lenders' willingness to compete increases in the sentiment of the signal. A more favorable signal increases outside lenders' assessment of loan quality, reducing their risk of getting stuck with a bad loan and increasing the aggressiveness with which they compete against an inside lender. In contrast, a less favorable signal decreases outside lenders' willingness to compete due to heightened adverse selection concerns driven by lower assessed loan quality. The more (less) aggressively outsiders bid, the higher (lower) the probability is of their winning the loan deal.

We extrapolate this economic reasoning to syndicated loans where inside, relationship lenders and outside, nonrelationship lenders compete to be selected by a borrower as the lead arranger of the syndication. We empirically investigate whether media sentiment is associated with the probability that the lead arranger of a loan is an outside lender with no established relationship with a borrower. We hypothesize that the aggressiveness with which outside lenders compete to arrange a loan is increasing in media sentiment, thus raising their probability of winning the loan. We find that the probability of a nonrelationship lender serving as a lead arranger is higher when media sentiment is more positive. A one standard deviation increase in media sentiment increases the probability that a loan is syndicated by a nonrelationship lead arranger by 6.2%.

We next consider whether the media influences the willingness of lenders to participate in a loan syndicate. We classify as more informed, inside participants those lenders who have a previous relationship with either the borrower or the loan's lead arranger (e.g., Sufi [2007], Ivashina [2009]), and as less informed outside participants those without a relationship. We hypothesize that the willingness of less informed lenders to participate in a loan syndicate is increasing in media sentiment. Because a lead arranger has an information advantage relative to syndicate participants and its monitoring effort after a loan closes is unobservable, participants face both adverse selection and moral hazard problems (e.g., Sufi [2007], Ivashina [2009], Mora [2015]). While it is difficult to empirically distinguish adverse selection and moral hazard, we argue that the media can influence lenders' participation through either channel.

To motivate an adverse selection explanation, we extend the economic reasoning from Rajan [1992] to loan syndicates where loan participants compete to become syndicate members. In forming a syndicate, a lead bank sends out invitations to a large set of potential lenders who choose whether or not to bid on joining the syndicate. The number of invitations generally exceeds the number of lenders who ultimately participate (Champagne and Kryzanowski [2007]). As noted by Ivashina and Sun [2011], the syndication process can be viewed as an auction where loan participants submit sealed bids to the lead arranger. The same winner's curse issue developed in Rajan [1992] plausibly operates here to the extent that inside participants pursue good loans, leaving bad loans to less informed bidders. Analogous to Rajan [1992], the media can influence the severity of the winner's curse problem, where more positive media signals increase the willingness of nonrelationship (less informed) lenders to bid for loan participation.

Under moral hazard, syndicate participants are concerned about ex post monitoring efforts of the lead arranger, where monitoring becomes more crucial for troubled borrowers (Sufi [2007]). If the media reveals negative news, outside participants plausibly downgrade their assessment of loan quality, which increases their concern about the lead arranger shirking. Thus, outside participants are less likely to join a syndicate than are relationship participants with superior information about the lead arranger and/or the borrower. Similar to the adverse selection case, outside lenders can either not submit bids or submit bids that are rejected by the lead arranger in favor of superior bids by relationship participants. Based on these arguments, we predict that the probability of nonrelationship syndicate participants joining the syndicate increases in media sentiment. We find evidence consistent with this prediction: a one standard deviation increase in Media Sentiment increases the probability that a participant without a prior relationship with the borrower (lead arranger) joins the syndicate by 7.1% (9.3%).

Next, we examine the proportion of a loan retained by the lead bank. Due to adverse selection and moral hazard concerns, syndicate participants require lead arrangers to put skin in the game by holding a proportion of the loan that is increasing in the severity of the agency problem (e.g., Leland and Pyle [1977], Holmström and Tirole [1997], Sufi [2007]). Because, as argued above, more positive public signals are expected to decrease both adverse selection and moral hazard risks faced by less informed syndicate participants, we hypothesize that the loan share retained by the lead arranger is decreasing in media sentiment. We show that a one standard deviation increase in media sentiment decreases this proportion by 5.3%.

Finally, we investigate whether more positive media sentiment is associated with lower interest rate spreads. We hypothesize that if, as argued earlier, loan attractiveness to nonrelationship lead arrangers and syndicate participants increases when media sentiment is more positive, the expanded set of potential lenders willing to fund the loan will increase the supply of credit available to the borrower and put downward pressure on interest spreads. We find that interest rate spreads are decreasing in media sentiment: a one standard deviation increase in *Media Sentiment* translates into a 6.9 basis point decrease in the spread.

Our previous analyses presume that the media provides new information to less informed lenders by publicly revealing either private information known to inside, but not outside, lenders, or information that is not previously known to either inside or outside lenders. To bolster this claim, we show that our findings with respect to lead arranger choice, syndicate participation, lead arranger share, and interest spreads are robust to controlling for alternative public information sources, including equity analysts, credit rating agencies, and firm-initiated press releases. Further, we split our sample into partitions based on the extent of firms' analyst coverage. Prior research shows that analysts provide useful information to lenders (e.g., Guntav and Hackbarth [2010], Mansi, Maxwell, and Miller [2011]). This suggests that when analyst information is scarce, lenders face more uncertainty about a borrower's prospects and rely more on alternative information sources. We predict that if the media is informative to outside lenders, then the effect of media sentiment on private lending practices will be stronger when analyst coverage is less intensive. Consistently, we find that syndicate structure and loan spreads are more sensitive to media sentiment when analyst coverage is lower. We further find some evidence that the effect of media sentiment is stronger for borrowers that experience an exogenous reduction in analyst coverage preceding loan issuance.

Although our findings are consistent with the media providing incremental information to outside lenders, the media can also influence private lending deals through alternative channels. We explore one such channel by examining the possibility that the media influences consumers' perceptions of a borrower, changing their future purchasing decisions and consequently a borrower's future cash flows. For example, news that a consumer product firm acted irresponsibly toward the environment may cause some consumers to abandon the firm's products. It is plausible that lenders would incorporate the expected impact of changes in consumer behavior on a borrower's future sales and cash flows into loan pricing. We thus hypothesize that the media, operating through a consumer perception channel, impacts the interest spreads of consumer product-oriented borrowers more intensively than it does the spreads of other borrowers. We find strong supporting evidence. This evidence suggests that the media can affect loan pricing via at least two channels-by informing outside lenders with limited access to private information and by changing consumers' perception of the borrower, which is in turn priced by lenders.

While our analyses provide robust evidence consistent with the media being a direct source of information to lenders, we cannot fully rule out the possibility that media sentiment proxies for information from other sources. However, whether the media is a direct source of information or simply mimics information from other sources, our results contribute to the literature by providing substantive evidence on the role of information signal *sentiment* in shaping loan deals. The evidence that lenders' decisions are influenced by media sentiment is consistent with the theory of Rajan [1992] and differentiates our analysis from prior studies that focus primarily on the availability of public information about a borrower without considering the role of its sentiment (e.g., Sufi [2007], Ball, Bushman, and Vasvari [2008], Ivashina [2009], Schenone [2010]). We extend the literature by demonstrating that critical aspects of syndicate lending, such as lead arranger choice, syndicate participation decisions, the loan share retained by the lead arranger, and loan pricing are significantly associated with the sentiment of business press articles, controlling for alternative information sources about the borrower. We also provide novel evidence consistent with media sentiment impacting loan pricing via its influence on consumer perceptions and behavior.

Section 2 presents prior research that motivates our analyses and the hypotheses development. Section 3 describes the sample and data. Section 4 reports our main results and section 5 concludes the paper.

2. Motivation and Related Literature

2.1 INFORMATION ASYMMETRY BETWEEN LENDERS COMPETING TO SERVE AS A LEAD ARRANGER

Private debt contracts rely on the flow of confidential information between lenders and borrowers (e.g., Diamond [1984], Fama [1985]). Confidential information includes hard information, such as timely financial disclosures, covenant compliance information, amendment and waiver requests, financial projections, and plans for acquisitions or dispositions (Standard and Poor's [2011]), as well as softer information, such as observations about a management team's abilities and honesty (Petersen [2004]). Lenders with access to borrowers' confidential information are surely less reliant on public information than are equity investors. However, not all lenders have equivalent access to a borrower's private information, and information asymmetries across lenders create a demand from informationally disadvantaged lenders for public information.

A syndicated loan is a loan issued to a firm jointly by more than one financial institution. Members of a syndicate can be classified as either lead arrangers or participant lenders. The lead arranger initiates a loan, negotiates with a borrower, and then syndicates shares of the loan to participant lenders. The lead arranger's compensation consists of an upfront fee not shared with the rest of the syndicate and interest spread (Ivashina [2009]). This upfront fee increases with the complexity and riskiness of the loan (Standard and Poors [2011]). Lead arrangers also get private benefits from building a relationship with the borrower and gaining access to its future business (Drucker and Puri [2009], Yasuda [2005], Mora [2015]). Before any loan participants are invited to join a syndicate, the borrower first selects the lead arranger from the set of competing lenders. Information asymmetry can affect the competition among lenders competing for a mandate to serve as the lead arranger on a loan. Here, we must

distinguish inside lenders with a prior relationship with a borrower and outside, nonrelationship lenders.

Relationship lenders interact repeatedly with borrowers over an extended time period, and in the process gain extensive inside knowledge of the borrowers' operations and develop private channels of communication with managers. Due to this information advantage, relationships are an important source of an incumbent bank's comparative advantage over new lenders without a relationship (Boot [2000]). Thus, when lenders compete to serve as lead arranger, outside, nonrelationship lenders face adverse selection risk due to their prebidding information disadvantage. This adverse selection issue is the focus of Rajan [1992].

Rajan models competitive dynamics in private lending markets in a setting where a relationship bank competes for a borrower's loan with less informed (outside) banks. The premise is that the inside bank acquires private information through its privileged access to the borrower, while an outside bank lacks access. The model assumes that the inside bank knows for sure whether the firm will succeed or fail, while an outside bank has an assessed probability of success.⁵ Under these conditions, an outside bank is at a disadvantage in bidding against the inside lender. The outside bank understands that if it adopts an aggressive strategy of always bidding, the inside bank will exploit this predictability and always win good loans by offering the borrower slightly better terms than does the outsider in the good state (when the firm will succeed) and leaving outside lenders with bad loans by not bidding in the bad state (when the firm will fail). This exposes the outsider to a winner's curse problem in which it only finances bad projects.

To deal with this problem, an outside lender plays a mixed strategy in which it randomizes bidding and not bidding. This implies that with some probability, the outside lender will not participate in the bidding process. Rajan [1992] shows that outsiders bid less frequently as their assessed probability that a loan is bad increases. Taking as given the outsider's mixed strategy, the insider's equilibrium strategy is to bid for good loans, but not for bad. For good loans, the insider offers a higher interest rate than outsiders do, knowing that when the outsider does not bid, the insider will win the loan and earn high profits, but lose the loan and earn zero profits if the outsider bids. In an important extension, Rajan [1992] shows that public disclosure influences outside lenders' probability of bidding by changing their assessments of loan quality. While a more positive signal increases the adverse selection problem outside lenders face.

⁵ Hauswald and Marquez [2006] demonstrate similar economic consequences of relationship lending in a model where inside banks are imperfectly, but better, informed than are outside banks.

This leads to the result that the probability of an outsider bidding for a loan is higher for signals with higher sentiment.⁶ Further, the greater the probability of outsiders' bidding, the higher their probability is of winning the loan. Building on this intuition, we hypothesize that the probability of an outside, nonrelationship lender serving as the lead arranger of a syndicated loan is increasing in media sentiment. Following Schenone [2010], we classify a loan as syndicated by a relationship lead arranger if the lender has syndicated a majority of the borrower's prior loan deals by volume over the five-year period preceding the loan issuance date, and syndicated by a nonrelationship lender otherwise. We construct a media sentiment variable using the RavenPack database to proxy for the sentiment of public information signals in Rajan [1992].

While Rajan [1992] describes well competition among differentially informed lenders to win a loan deal, the model does not directly consider information asymmetry issues associated with loan syndicate formation. We turn to this issue next.

2.2 INFORMATION ASYMMETRY WITHIN LOAN SYNDICATES

We examine whether the media influences the willingness of lenders to participate as members in a loan syndicate. Participant lenders generally do not directly communicate with the borrower and maintain an arm's length relationship through the lead arranger. The lead arranger establishes and maintains a relationship with the borrower and acts as an agent for the syndicate by collecting and processing information about the borrower. Prior to syndication, the lead bank conducts due diligence and presents an assessment of the borrower's quality to potential participants. Thus, participants face adverse selection risk because the lead bank can have an incentive to syndicate low-quality loans (Sufi [2007], Ivashina [2009]). Lead banks can, for example, originate and then syndicate low-quality loans to obtain upfront fees and capture private benefits, such as future cross-selling opportunities with the borrower (Mora [2015]). After syndication, the lead bank is charged with monitoring the borrower. This introduces a moral hazard risk as lead banks only retain part of the loan, which reduces incentives to monitor.

Most research on syndicate lending does not distinguish between the moral hazard and adverse selection risks (e.g., Dennis and Mullineaux [2000], Jones, Lang, and Nigro [2005], Ball, Bushman, and Vasvari [2008], Ivashina [2009]). Empirically separating these two risks is challenging, as

⁶ While Rajan's [1992] model is analyzed in terms of only a good and bad signal, it is straightforward to generalize this result to a setting with more than two signals. What is required is that the Monotone Likelihood Ratio Property (MLRP) holds such that higher signals indicate a higher likelihood that the loan is good (Rajan's two-signal model implicitly assumes MLRP). In this case, higher signals result in a greater probability of bidding. In the language of our paper, the probability of an outsider bidding for a loan monotonically increases in the sentiment of the information signal.

the consequences of moral hazard and adverse selection are generally observationally equivalent (Cohen and Siegelman [2010]). Although we do not attempt to distinguish between the two, we expect the media to influence lenders' participation through either an adverse selection or moral hazard channel.

The syndication process seeks to match borrowers with the set of lenders that value the loan most highly. To form a syndicate, a lead bank sends out invitations to a large set of potential lenders who choose whether to submit bids to join the syndicate. Invitations to bid generally outnumber acceptances, implying that many invitees choose not to participate (Champagne and Kryzanowski [2007]).⁷ Lenders that accept the invitation submit sealed bids directly to the lead bank. After bids are in, a minimum clearing spread is determined and investors who bid that spread or below are awarded a share of the loan on a pro-rata basis (Ivashina and Sun [2011]).

We extend the adverse selection dynamics reasoning from Rajan [1992] to competition among differentially informed lenders to become syndicate members. We posit that participants with a prior relationship with either the borrower or the lead arranger face less severe information risks than participants without a relationship. Thus, the winner's curse issue developed in Rajan [1992] is likely to play a role in lenders' syndicate participation decisions to the extent that inside participants pursue good loans, leaving bad loans for less informed bidders. In line with Rajan [1992], the media can influence the severity of the winner's curse problem, where more positive media signals increase the willingness of nonrelationship lenders to bid for loan participation. Note that outside lenders that do not join the syndicate can either not submit bids or submit bids that are rejected by the lead arranger in favor of superior bids by relationship participants.

With respect to moral hazard, because the severity of potential losses due to a lack of monitoring effort is likely to be higher for troubled borrowers (Sufi [2007]), negative media news can lower outside participants' assessments of loan quality and thus increase their concerns about lead arranger shirking. This can in turn increase the information disadvantage of nonrelationship participants relative to relationship participants who possess enhanced channels of communication with the borrower or lead arranger. Thus, outside participants can be less willing to bid against relationship lenders or they can demand terms that are less favorable than those sought by relationship participants and therefore are rejected by the lead arranger. Based on these arguments, we hypothesize that the probability of nonrelationship syndicate participants joining the syndicate increases in media sentiment also through the moral hazard channel.

⁷Sample loans have on average 9.6 syndicate participants. Champagne and Kryzanowski [2007] note that the number of invitations sent to potential syndicate participants is generally three times the number of participants that agree to join the syndicate, although this is likely to vary with the tightness of the supply of credit, the loan's purpose, and the complexity of its structure, as well as future cross-selling possibilities, among other things.

We further examine the effect of the media on information asymmetry within syndicates by investigating whether media sentiment is associated with the proportion of the loan retained by the lead bank. Theory posits that a key mechanism for resolving information asymmetry issues and securing potential syndicate lenders willing to join the loan syndicate is to require lead banks to have skin in the game by holding a stake in the loan that increases in the severity of agency problems (e.g., Leland and Pyle [1977], Diamond [1984], Holmström and Tirole [1997]). Empirical research provides evidence consistent with the theory. Sufi [2007] shows that lead banks hold a larger share of the loan when borrowers are more opaque, while Ball, Bushman, and Vasvari [2008] show that the share retained by lead banks is decreasing in the debt contracting value of the borrower's accounting numbers. We extend and innovate on this literature by conjecturing that the share retained by the lead arranger depends not only on the fact that there is public disclosure, but also on the sentiment the disclosure conveys. Because, as argued earlier, a more positive public signal decreases adverse selection and moral hazard risk, we predict that the loan share retained by the lead arranger is decreasing in media sentiment.

2.3 MEDIA SENTIMENT AND LOAN PRICING

We next address whether media sentiment affects the interest rates. Rajan [1992] shows that more positive disclosure increases the aggressiveness of outside lenders in bidding for a loan, which decreases the rents that inside lenders can extract from borrowers and thus results in lower interest rates. Recent empirical research exploits changes in a firm's information environment to identify how the information advantage of inside banks affects loan pricing. Hale and Santos [2009] examine how loan pricing changes following bond initial public offerings (IPOs). They predict that when public information released during the IPO process reveals positive news about a borrower—that it is creditworthy—it increases outside banks' willingness to bid on its loans and consequently reduces inside lenders' information advantage. Consistently, they find that firms revealed as having investment grade credit ratings pay lower spreads on their loans after they undertake a bond IPO.

Hale and Santos [2009] rely on a significant information event to identify how public disclosure reduces inside banks' ability to extract rents from borrowers. We complement Hale and Santos [2009] by examining whether more favorable media sentiment results in lower spreads. In a syndicated loan setting, we expect more favorable media sentiment to increase both the aggressiveness with which outside lenders bid to arrange loans and to participate in loan syndicates, thus expanding the set of lenders willing to fund the loan. This expansion of potential lenders should increase the supply of credit available to the borrower. To the extent that interest spreads are negatively associated with the supply of credit (e.g., Ivashina and Sun [2011]), we predict that the wider range of syndicate participants

and arrangers associated with more positive media sentiment results in lower loan pricing.

2.4 THE ROLE OF THE BUSINESS PRESS IN CAPITAL MARKETS

The main premise of our paper is that the business press plays an important role in the private lending market by mitigating information asymmetries between differentially informed lenders. While there has been little research to date on this topic, a growing body of research demonstrates a substantive role played by the business press in the equity market. Although the information structures and contractual features of private lending differ from those of the equity markets, prior research supports the plausibility of our premise that the media can be informative to less informed outside lenders and influence their behavior.

Miller [2006] finds that the press serves as a watchdog for accounting fraud by rebroadcasting information from other information intermediaries and by undertaking original investigations that provide new information to the markets. Several papers find that media articles pressure firms to alter governance structures and strategic decisions, such as resource allocation, environmental policies, and the structure of executive compensation (e.g., Dyck and Zingales [2002], Bednar, Boive, and Prince [2012], Kuhnen and Niessen [2012], Joe, Louis, and Robinson [2009]).⁸ For lenders without privileged access to a borrower, media articles that reveal fraud or pressure firms to alter governance and strategic choices can plausibly convey new information about the integrity of a borrower's executives, litigation risk, or changes in managerial incentives. These are important aspects of a lender's assessment of a borrower's future prospects and creditworthiness.

Further, there is evidence that the business press provides information about firm fundamentals to equity market participants over and above that provided by other information intermediaries and accounting data (e.g., Fang and Peress [2009], Bushee et al. [2010], Tetlock [2007]). Tetlock, Saar-Tsechansky, and Macskassy [2008] find that the fraction of negative words in firm-specific news stories forecasts low firm earnings and that equity prices quickly respond to the information in negative words. They suggest that linguistic media sentiment captures otherwise hard-to-quantify aspects of firms' fundamentals.⁹ It is thus plausible that the media can also impact private lending by providing information on hard-to-quantify aspects of firms' fundamentals that underpin credit quality. Lenders' assessments of firms' credit risk can be affected by information about products, customers, competitors, industry growth potential, top executive teams, governance, regulation, legal issues, strategic plans, acquisitions, and labor

⁸See Miller and Skinner [2015] for a more detailed discussion of the media literature.

⁹ See also Kothari, Li, and Short [2009], Boudoukh et al. [2013], and Rogers, Skinner, and Zechman [2016]. In contrast, a number of recent papers suggest that media coverage might exacerbate information asymmetry and inefficient trading behavior (e.g., Frankel and Li [2004], Green, Hand, and Penn [2012], Soltes, Solomon, and Sosyura [2014]).

markets, among many other topics (Green, Hand, and Penn [2012]). The information revealed by the media can be new information not available to lenders from other sources, or it could be a signal that confirms, corroborates, or substantiates noisy information that lenders received from a different source. In the latter case, the media could, for example, confirm soft information the lender obtained elsewhere about the borrower, or allow the lender to filter noise from signals received from other sources. While our research design does not allow us to distinguish the specific topics covered in the media articles underpinning media sentiment measure, extant research shows the power of a computational linguistics approach to quantify the language of a wide swath of firms' media coverage rather than focusing on specific categories of news. As noted by Tetlock, Saar-Tsechansky, and Macskassy [2008], by quantifying language, researchers can examine the directional impact of a wide variety of events, rather than focusing on particular event types.

In addition to providing direct information to lenders about a borrower, the media can indirectly influence private lending deals by changing a borrower's circumstances through its impact on the behavior of a borrower's stakeholders. We explore one such channel by examining the possibility that the media influences consumers' perceptions of a borrower, changing their future purchasing decisions and consequently a borrower's future cash flows.¹⁰ The idea is that lenders would incorporate the expected consequences of these changes into the loan terms offered to the borrower.

3. Sample, Data, and Descriptive Statistics

3.1 DATA SOURCES AND SAMPLE SELECTION

We employ the DealScan database provided by the Thomson Reuters Loan Pricing Corporation (TRLPC) to obtain loan-specific characteristics. Media sentiment scores are from RavenPack News Analytics, which covers all news disseminated via Dow Jones Newswires. RavenPack employs a variety of advanced textual analysis techniques to create news sentiment scores for business news stories. We obtain firm characteristics from Compustat. Firms' senior debt ratings, watch list, and outlook data (at the firm level) are retrieved from the S&P historical database. For borrowers with no rating data in the S&P database, we collect the S&P, Moody's, and Fitch ratings from the Internet-based version of the TRLPC.

Table 1 summarizes the sample selection process. For the period from 2000 to 2012, DealScan reports 31,974 facilities issued to U.S. public firms

¹⁰ Previous literature provides evidence that consumer behavior is sensitive to perceptions of a firm. For example, Lev, Petrovits, and Radhakrishnan [2009] find that charitable contributions are significantly positively associated with future revenue, particularly for consumer product–oriented firms. Similarly, Fisman, Heal, and Nair [2006] find that corporate philanthropy and profits are positively related, but only for firms in industries where a firm's image is important to consumers.

Filters	Number of Facilities
Syndicated loans to public U.S. borrowers, in U.S. dollars, issued over the period 2000–2012	31,974
After elimination of the facilities of firms not covered by	24,308
RavenPack	
After elimination of the facilities with missing loan data	12,397
After elimination of the facilities with insufficient firm and	7,244
media data	

TABLE 1Sample Selection

This table presents the sample selection process.

in U.S. dollars. This time period reflects the availability of media data from RavenPack. Merging this sample with RavenPack leaves us with 24,308 facilities. Next, we exclude facilities with insufficient loan data, reducing the sample to 12,397 facilities. We also require borrowers to have sufficient Compustat data to estimate borrower characteristics and sufficient Raven-Pack data to estimate media sentiment prior to loan issuance. We estimate borrower characteristics in the quarter prior to the loan issuance and media sentiment over the 180 days prior to the loan issuance.¹¹ We limit media data to full-size articles, excluding news flashes (news articles composed only of a headline and no body text), news articles composed of a headline and mostly tabular data, and firm-initiated press releases. We further restrict our sample to full-size articles with a relevance score of 75 and above. The relevance score is assigned by RavenPack to indicate when the firm is strongly related to the underlying news story. The scores range from 0 (low relevance) to 100 (high relevance). Our final sample contains 7,244 facilities related to 2,031 firms.

3.2 DESCRIPTIVE STATISTICS

Table 2 presents descriptive statistics for the main variables used in our tests (all variables are described in detail in the appendix). *Media Sentiment* is estimated as the average news sentiment over the 180 days prior to a loan's origination date. We utilize RavenPack's Composite Sentiment Score (CSS), which represents the news sentiment of a given story by combining various textual analysis techniques (see the appendix for a more detailed description of RavenPack's CSS measure). CSS scores range between 0 and 100, with a score above 50 indicating positive news; a score equal to 50, neutral news; and a score below 50, negative news. We apply a linear transformation to the CSS score and define *Media Sentiment* = (CSS - 50)/50, so that the *Media Sentiment* ranges from -1 to 1, with zero being equivalent to

¹¹We focus on the 180 days prior to the loan issuance to allow a sufficient time period prior to the start of the syndication process, which typically takes around three months. Our findings and inferences do not change if we estimate media coverage over 90 or 60 days prior to the loan issuance.

Variables	N	Mean	Median	Std. Dev
Media				
Media Sentiment	6,964	-0.0046	0.0000	0.0386
Primary Loan Characteristics				
Borr-Lead No-Relationship	7,255	0.6379	1.0000	0.4806
Part-Borr No-Relationship	60,082	0.5675	0.0000	0.4954
Part-Lead No-Relationship	60,082	0.1525	1.0000	0.3595
Lead Share (%)	4,613	21.0525	13.3750	20.4714
Spread	6,964	5.0720	5.1648	0.7221
Firm Characteristics				
ROA	6,964	0.0085	0.0102	0.0253
Interest Coverage	6,964	10.5052	2.1228	40.9157
Leverage	6,964	0.2520	0.2373	0.1789
Size	6,964	7.2521	7.2688	1.6609
Z-Score	6,964	2.2403	1.6907	2.2198
MTB	6,964	3.0102	2.0201	3.7461
Loss	6,964	0.1476	0.0000	0.3547
Additional Loan Characteristics				
Amount	6,964	18.9334	19.1138	1.4854
Maturity	6,964	47.9482	57.0000	21.4141
PP	6,964	0.7143	1.0000	0.4519
Collateral	6,964	0.6392	1.0000	0.4802
#Covenants	6,964	2.3109	2.0000	1.0131
Revolver	6,964	0.6229	1.0000	0.4846
Term Loan B	6,964	0.1002	0.0000	0.3006
Additional Information Sources				
Press Release Sentiment	6,964	0.0243	0.0264	0.0287
Rated	6,964	0.5595	1.0000	0.4964
Analyst Coverage	6,964	0.8751	1.0000	0.3307
Returns				
Return	6,964	0.0558	0.0216	0.3121
Negative Return	6,964	0.4275	0.0000	0.4947

 TABLE 2
 2

 Descriptive Statistics
 1

This table provides descriptive statistics (see table 1 for the sample selection procedure). Variables are defined in the appendix.

neutral sentiment. The mean (median) value of *Media Sentiment* is -0.0046 (0.0000).

We examine five primary loan characteristics. Our first measure reflects the previous relationship of the lead arranger and the borrower. The mean value of *Borr-Lead No-Relationship* indicates that 63.8% of loans are issued by nonrelationship lead arrangers. The next two measures are concerned with the previous relationships of the syndicate participants. Sample loan syndicates have on average 9.6 syndicate participants, indicating relatively large syndicates. The interquartile range of the number of participants is 11, which suggests a high variation in syndicate structure for sample loans (untabulated). The mean value of *Part-Borr No-Relationship* indicates that 56.75% of syndicate participants have not participated in a borrower's syndicated loans in the five years preceding a loan's issuance. The mean value

129

of *Part-Lead No-Relationship* indicates that 15.25% of syndicate participants have not participated in syndicates arranged by the loan's lead arranger in the five years preceding the loan's issuance. We also examine the proportion of a loan retained by the lead arranger. For loans with sufficient data, the mean (median) value of *Lead Share* is 21.05% (13.38%). With respect to loan pricing, the average interest rate spread is 159.5 basis points (*Spread* is the logarithm of the interest spread).

With respect to sample firm characteristics, the average ratio of earnings before extraordinary items to total assets (*ROA*) is 0.85%. Sample firms have an average interest coverage ratio (*Interest coverage*) of 10.51, and average *Leverage*, measured as the ratio of total liabilities to total assets, of 0.25. Sample firms are also relatively large, with a mean value of total assets of \$1,411M (*Size* is measured by the logarithm of the total assets). An average Altman's [1968] bankruptcy score (*Z-score*) is 2.24 (a higher score indicates a lower credit risk). The average market-to-book ratio is 3.01. Note that 14.76% of the sample observations relate to the borrowers experiencing losses.

We report statistics for loan characteristics that serve as controls. Sample loans have an average size of \$167.0M (*Amount* is the logarithm of loan amount) and a maturity (*Maturity*) of 47.9 months. The majority of the loans are revolvers (62.3%) and 10% are term loans B and below, which are typically issued to nonbank institutional investors. A total of 71.4% of the sample loans have performance pricing provisions (*PP*), 63.9% are secured (*Collateral*), and loans have on average 2.3 financial covenants (*#Covenants*).

Finally, we provide statistics for three alternative sources of information about borrowers. First, we estimate the sentiment of all borrower-initiated press releases over the 180 days prior to a loan's origination date. To isolate firm-initiated press releases, we impose a more stringent relevance criterion than the one applied to media articles and require a relevance score of 90 or greater. Press releases with a relevance score above 75 and below 90 often relate to cases where the firm is mentioned in the press releases of other firms. We apply a linear transformation to the CSS score and define Press Release Sentiment = (CSS - 50)/50, so that Press Release Sentiment ranges from -1 to 1, with zero being equivalent to neutral sentiment. Press release sentiment is typically positive, with a mean (median) value for Press Release Sentiment of 0.0243 (0.0264). Second, we report on whether the borrower is rated by a credit rating agency. Fifty-six percent of sample loans are issued to rated borrowers. Third, because prior research suggests that equity analysts are informative to lenders (e.g., Guntay and Hackbarth [2010], Mansi, Maxwell, and Miller [2011]), we measure whether a borrower has equity analyst coverage at loan origination. Note that 87.5% of sample loans are issued to borrowers with analyst coverage.

To reflect additional information available about a borrower, we also estimate its abnormal stock returns over the same period during which sentiment is measured. Sample firms experience mean (median) abnormal returns of 5.6% (2.2%) over the 180 days prior to a loan's issuance (*Return*), with 42.75% of firms having a negative return over this period (*Negative Return*).

4. Empirical Results

We organize the discussion of our empirical findings into five subsections. Section 4.1 provides evidence on the association between media sentiment and the probability of a nonrelationship bank leading the loan deal. Section 4.2 examines relations between media sentiment and syndicate structure and section 4.3 considers those between media sentiment and loan pricing. In section 4.4, we provide evidence consistent with the media being a direct source of information to less informed, outside lenders. Finally, in section 4.5 we examine an additional channel through which the media may affect loan pricing.

 $4.1\,$ media sentiment and the probability of a nonrelationship bank leading the loan deal

In this section, we examine whether media sentiment is positively associated with the probability of outside, nonrelationship lenders serving as a lead arranger. We predict that a more positive media sentiment is associated with a higher probability that a lender without an established relationship with the borrower originates the loan. We estimate the following logistic model:

Borr-Lead No-Relationship = $\alpha_0 + \beta_1$ Media Sentiment + β_2 Controls

$$+ \beta_3 Alternative Info Sources$$
, (1)

where *Borr-Lead No-Relationship* is an indicator variable equal to 1 if a loan is issued by a nonrelationship lender, and 0 otherwise.¹² *Media Sentiment* is our main variable of interest; we predict that $\beta_1 > 0$.

We control for a number of factors that prior research suggests are associated with the probability of a nonrelationship lead arranger syndicating the loan (e.g., Gopalan, Udell, and Yerramilli [2011], Bolton et al. [2013]), including whether a firm's previous loan is still outstanding when the current loan is issued (*Outstanding*), the time between the current and previous loan (*Time Between*), and the tightness of the credit supply in the economy (*Tight Credit Supply*). We proxy for credit supply tightness using changes in bank lending standards for mid-sized and large commercial loans, as reported in the Federal Reserve Board's quarterly Senior Loan Officer Opinion Survey on Bank Lending Practices (e.g., Bassett et al. [2012]). The *Tight Credit Supply* variable takes a value of 1 if the change in lending standards

 $^{^{12}}$ To be able to measure this relationship variable, we restrict the estimation of model (1) to the sample of borrowers who issued at least one loan over the five-year period prior to a current loan's origination date.

in the quarter of a loan's origination is in the top quartile of the sample's distribution, and 0 otherwise.

We also control for a borrower's characteristics that reflect its creditworthiness, including profitability (*ROA*), interest coverage ratio (*Interest Coverage*), leverage (*Leverage*), size of total assets (*Size*), z-score (*Z-score*), the market-to-book ratio (*MTB*), and an indicator variable reflecting whether the borrower has experienced losses (*Loss*). In terms of loan-specific characteristics, we control for loan size (*Amount*), maturity (*Maturity*), whether the loan is a revolving line of credit (*Revolver*), and if the loan is a term loan B and below (*Term Loan B*).

Finally, we control for alternative information sources about the borrower, including its press release sentiment (Press Release Sentiment), whether it is rated (*Rated*), and whether it has analysts' coverage (*Analyst Coverage*).¹³ To reflect borrower-related news not captured by additional information sources and firm controls, we also include a borrower's abnormal stock returns (*Return*) and an indicator variable reflecting whether stock returns are negative (Negative Return), measured contemporaneously with Media Sentiment. Although stock returns impound all available information, there is still scope for *Media Sentiment* to play a direct role in informing lenders. The idea here is that the weights placed on the available information signals are tailored to the specific decision context at hand. While stock prices optimally aggregate all available information for valuation purposes, the weights on information signals that are optimal for valuing equity may not be optimal for assessing credit quality. Because a bank cannot generally look at changes in share price and extract the relevant signal from the many factors that move share price, outside banks can utilize relevant information from media stories and weight them appropriately in their credit decisions (Paul [1992], Bushman and Indjejikian [1993]).¹⁴ Finally, we cluster the standard errors at the firm and calendar quarter levels (this applies to all the remaining analyses).¹⁵

We present our findings in table 3. Consistent with the theory of Rajan [1992], we find that the coefficient on *Media Sentiment* is positive and

¹³ In the online appendix, we show that our results are robust to including a more extensive set of measures of alternative information sources in the specification.

¹⁴ Consider a simple example. Assume that equity investors' information set consists of two signals, X_1 (e.g., media sentiment) and X_2 (another signal). This information is impounded into equilibrium price. Assume the equilibrium price applies equal weights to the signals giving: Price = $X_1 + X_2$. Assume that for credit purposes, a lender would optimally weight the signals as: Credit quality = $\gamma_1 X_1 + \gamma_2 X_2$. The lender cannot observe X_1 and X_2 separately, but instead observes X_1 and the stock price. How would lenders use these signals? Given that they want to optimally weight the signals for credit purposes, they would construct the following estimate of credit quality: Credit Quality = $\gamma_1 X_1 + \gamma_2 *$ (Price – X_1) = $\gamma_1 X_1 + \gamma_2 X_2$. That is, even though the price fully impounds X_1 (e.g., media sentiment), lenders find it optimal to use X_1 in addition to price in order to achieve the desired weights on X_1 and X_2 .

¹⁵ Given Greene's [2004] criticism of the inclusion of fixed effects in nonlinear models, we do not include year, industry, and loan purpose fixed effects. The same applies to the probability of a nonrelationship lender syndicate participation model.

Variables	Dependent Variable: Borr-Lead No-Relationshif
Media Sentiment	1.5474**
	(0.6393)
Outstanding	-0.2217^{***}
	(0.0808)
Time-Between	0.0003***
	(0.0001)
Tight Credit Supply	0.0280
	(0.0656)
ROA	0.1141
	(0.8343)
Interest Coverage	0.0017^{**}
	(0.0007)
Leverage	-0.1801
	(0.1812)
Size	-0.0767^{**}
	(0.0305)
Z-score	-0.0177^{*}
	(0.0092)
MTB	0.0097***
	(0.0032)
Loss	0.3560^{***}
	(0.0811)
Amount	-0.1487^{***}
	(0.0323)
Maturity	-0.0068^{***}
	(0.0014)
Revolver	0.1259^{*}
	(0.0664)
Term Loan B	0.0462
	(0.1045)
Press Release Sentiment	-3.1423^{**}
	(1.3519)
Rated	0.0743
	(0.0657)
Analyst Coverage	-0.2312^{***}
	(0.0872)
Return	0.2028^{*}
	(0.1194)
Negative Return	0.0244
-	(0.0707)
Ν	7,255
R ^e	0.0247

 TABLE 3

 Media Sentiment and the Probability of a Nonrelationship Lead Arranger

This table presents the analysis of the effect of media sentiment on the probability of a nonrelationship lender serving as a loan's lead arranger. Standard errors are reported in parentheses. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 levels, respectively. All variables are defined in the appendix.

significant, suggesting that the probability of a nonrelationship lead arranger syndicating the loan is increasing in media sentiment. A one standard deviation increase in media sentiment increases the probability that a loan is syndicated by a nonrelationship lead arranger by 6.2%.

With respect to control variables, similar to Gopalan, Udell, and Yerramilli [2011], we find that the probability of a nonrelationship lender syndicating the loan decreases when the borrower's previous loan is still outstanding when the current loan is issued (*Outstanding*), but increases as the time span between the current and the previous loan increases (*Time-Between*). We do not find that the tightness of the credit supply significantly affects the probability of a nonrelationship lead arranger originating the loan. The probability of a nonrelationship lender syndicating the loan is also affected by a borrower's interest coverage, size, *Z*-score, and marketto-book ratio. When borrowers issue larger loans or loans with a longer maturity, we find that it is less likely for these loans to be arranged by nonrelationship lenders. These results suggest that because nonrelationship lenders have inferior access to a borrower's private information relative to relationship lenders, they are less willing to syndicate larger and longer term credit.

We find a negative coefficient on *Press Release Sentiment*. It is important to note here that press releases are an aspect of a firm's strategic disclosure decisions. For example, firms may issue positive press releases when they are trying to counter negative events or to run up stock prices around mergers (e.g., Davis and Tama-Sweet [2012], Ahern and Sosyura [2014]). Most important for our purposes is that our *Media Sentiment* result is robust to including press release sentiment, which rules out the possibility that media simply echoes firm press releases. We also find a negative coefficient on *Analyst Coverage*; this result is in line with Gopalan, Udell, and Yerramilli [2011].

4.2 THE EFFECT OF MEDIA SENTIMENT ON SYNDICATE PARTICIPATION AND LEAD ARRANGER SHARE

We next investigate the relation between media sentiment and syndicate participation. We predict that the likelihood of less informed lenders with no previous relationship with the borrower or the lead arranger participating in a syndicate is increasing in media sentiment. We follow Sufi [2007] and estimate the following logit model at the syndicate participant level:

$$Participation = \alpha_0 + \beta_1 Media Sentiment + \beta_2 Controls + \beta_3 Alternative Info Sources, \qquad (2)$$

where *Participation* is one of two participant relationship variables. *Part-Borr No-Relationship* is an indicator variable that is equal to 1 if the participant has not participated in a borrower's loans in the five years prior to a loan's issuance date, and 0 otherwise. *Part-Lead No Relationship* is an indicator variable that is equal to 1 if the participant has not participated in a syndicate arranged by the loan's lead arranger in the five years preceding the loan's issuance, and 0 otherwise. *Media Sentiment* is our main variable of interest, for which we predict $\beta_1 > 0$.

As in model (1), we control for borrower characteristics reflecting its creditworthiness. We also extend the set of loan-specific controls. We add controls for a borrower's previous relationship with the lead arranger (*Borr-Lead No-Relationship*) and for whether this is the first deal syndicated by the lead arranger (*First Time Lead*), as both can affect the attractiveness of the loan deal to participants. We also control for whether the loan is secured (*Collateral*), the existence of performance pricing provisions (*PP*), and the number of covenants (#*Covenants*).¹⁶ Similar to model (1), we control for alternative information sources about the borrower and its stock returns.

We present our findings in table 4. In column 1, we report the estimation of model (1) with *Part-Borr No-Relationship* as the dependent variable. Consistent with our predictions, the coefficient on *Media Sentiment* is positive and significant, indicating that the probability of a participant with no relationship with the borrower joining the syndicate increases with media sentiment. This result is also economically significant: a one standard deviation increase in *Media Sentiment* increases this probability by 7.1%.

With respect to controls, in line with lower information asymmetry between the lead arranger and participants when a borrower is more creditworthy, we find positive coefficients on ROA, Interest Coverage, and Z-Score and a negative coefficient on Leverage. We also find a positive coefficient on Borr-Lead No-Relationship, which suggests that the probability of nonrelationship participants joining the syndicate increases when a new lead arranger with no previous relationship with the borrower arranges the loan. The coefficient on *Press Release Sentiment* is insignificant, suggesting that while media sentiment influences less informed participants, the content of firm-initiated press releases does not incrementally affect their participation likelihood. When a firm is rated, the probability of a nonrelationship participant joining the syndicate is significantly higher, consistent with their higher confidence in a loan's quality when creditworthiness is assessed by an independent third party (e.g., Sufi [2007]). Note that the coefficients on Return and Negative Return are difficult to interpret. Stock returns consist of two components: changes in cash flow news and changes in discount rates (e.g., Vuolteenaho [2002]). Stock return controls therefore reflect both news about the firm's performance and the discount rate news associated with a firm's riskiness (our findings do not change when we control for return volatility).

In column 2, we report the results from the estimation of model (1) with *Part-Lead No- Relationship* as the dependent variable and find similar results.

¹⁶We do not control for these three additional loan characteristics in model (1), because they are typically determined during the negotiation process between a lead arranger and the borrower and therefore cannot affect the choice of the lead lender. In any case, in untabulated analyses, we find that our results are robust to the inclusion of these variables.

		Dependent Variable	
	Part-Borr	Part-Lead	
Variable	No-Relationship	No-Relationship	Lead Share
Media Sentiment	1.7654**	2.3114**	-28.7105***
	(0.7714)	(1.1204)	(6.8194)
ROA	3.2091*	-0.9517	-101.6974^{**}
	(1.6573)	(2.3441)	(24.5548)
Interest Coverage	0.0026**	0.0017	0.0013
0	(0.0011)	(0.0014)	(0.0158)
Leverage	-0.4195**	0.4476	-6.2121^{*}
0	(0.1925)	(0.3423)	(3.5286)
Size	-0.1409***	-0.0646	-2.2487**
	(0.0301)	(0.0490)	(0.6833)
Z-score	0.0641***	-0.0190	0.1689
	(0.0192)	(0.0301)	(0.3279)
MTB	-0.0001	-0.0155^{*}	-0.0041
	(0.0061)	(0.0092)	(0.1061)
Loss	0.1358	0.1263	0.5926
1000	(0.0833)	(0.1217)	(1.3263)
Amount	0.0196	-0.0358	-5.9624^{***}
	(0.0256)	(0.0419)	(0.6365)
Maturity	0.0050***	-0.0041^{**}	-0.1084^{***}
maraniy	(0.0011)	(0.0019)	(0.0329)
Spread	0.3591***	0.2945***	0.1706
Spread	(0.0343)	(0.0696)	(1.0446)
Collateral	0.1044*	0.0119	-0.2789
Contaientai	(0.0603)	(0.0896)	(1.1808)
PP	0.2208***	-0.0365	-7.8570^{**}
PP	(0.0543)		(1.7009)
#Comments	0.0377	(0.0896) 0.1479^{***}	-0.9175^{*}
#Covenants			
Revolver	(0.0287)	(0.0424)	(0.5226)
	-0.2142^{***}	-0.1186^{**}	1.0176
	(0.0362)	(0.0605)	(0.9951)
Term Loan B	0.2446***	0.5395***	15.4806***
	(0.0832)	(0.1224)	(2.2275)
Borr-Lead No-Relationship	0.4786***	0.5447***	2.2707***
	(0.0440)	(0.0750)	(0.4899)
First Time Lead	0.1061	1.7979***	4.8580*
	(0.1577)	(0.3378)	(2.7667)
Press Release Sentiment	0.0181	-0.6315	13.6202
	(0.7750)	(1.2702)	(16.4236)
Rated	0.1824^{***}	0.1268	-2.0909^{*}
	(0.0662)	(0.0986)	(1.1038)
Analyst Coverage	0.0117	-0.1215	-5.0446^{**}
	(0.0910)	(0.1569)	(2.1421)
Return	0.3726***	0.3464**	-2.7937^{*}
	(0.1203)	(0.1755)	(1.4571)
Negative Return	0.2131***	0.1245	0.4680
	(0.0616)	(0.0994)	(0.7463)

TABLE 4Media Sentiment and Syndicate Structure

(Continued)

	Dependent Variable		
Variable	Part-Borr No-Relationship	Part-Lead No-Relationship	Lead Share
Model	Logit	Logit	OLS
Fixed effects	None	None	Year/Industry/Purpose
N	60,082	60,082	4,613
R^2	0.0561	0.0620	0.4900

TABLE 4—Continued

This table presents the analyses of the effects of media sentiment on the characteristics of the syndicate's structure. Column 1 (2) presents the analysis of the variable reflecting a participant's prior relationship with the borrower (the lead arranger), while column 3 presents the analysis of the loan share retained by the lead arranger. Standard errors are reported in parentheses. ***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in the appendix.

The coefficient on *Media Sentiment* is positive and significant, suggesting that the probability of a participant without a previous relationship with the lead arranger joining the syndicate increases with media sentiment. A one standard deviation increase in *Media Sentiment* increases this probability by 9.3%. The coefficients on the control variables are generally consistent with those in column 1.

Our results so far indicate that more positive media sentiment extends the range of participants willing to join the syndicate to those without a relationship with the borrower or the lead arranger. Column 3 focuses on the loan share retained by the lead arranger (*Lead Share*). We predict that the share of a loan retained by the lead arranger is decreasing in media sentiment as participants require the lead to have less skin in the game.

We estimate model (2) as an OLS specification with *Lead Share* as the dependent variable. This specification also includes loan purpose,¹⁷ industry, and year fixed effects. In line with our expectations, we find a negative and significant coefficient on *Media Sentiment*. A one standard deviation increase in media sentiment decreases the loan proportion held by the lead arranger by 5.3%. This evidence further supports a significant effect of media sentiment on syndicate structure and suggests that it operates via both channels—syndicate participants' willingness to join the syndicate and the loan share the lead arranger is required to retain.

With respect to control variables, a lead arranger retains a smaller proportion of the loan when the borrower is larger and more profitable. Consistent with lower agency problems within the syndicate when a borrower is more transparent (e.g., Sufi [2007]), the lead arranger retains a smaller proportion of the loan when a borrower is rated and has analyst coverage. We also find that a nonrelationship lead arranger and a first time lead arranger hold a higher loan share, suggesting that participants require a lead arranger to have higher skin in the game in these cases.

¹⁷ It is common in the literature to control for differences in the purpose of the loan (i.e., LBO, M&A, restructuring, refinancing, general corporate purposes, etc.), as the riskiness of a loan and the level of information asymmetry across lenders can vary substantially by loan type.

4.3 THE EFFECT OF MEDIA SENTIMENT ON LOAN PRICING

Our analyses show that the decisions of less informed lenders to arrange and participate in syndicated loans are sensitive to media sentiment, where loan attractiveness is increasing in media sentiment. While it is important to show connections between media sentiment and lenders' decisions, our results thus far do not quantify the implications of media sentiment from the perspective of a borrower seeking financing. We examine this issue in this section.

As we show in our previous analyses, higher media sentiment not only increases a nonrelationship lead arranger's aggressiveness in bidding for a loan, but also attracts less informed syndicate participants to join the syndicate. This expanded set of outside lenders is likely to increase the supply of financing available to fund the borrower's loan, exerting significant downward pressure on the interest rate spread.

To examine this hypothesis, we estimate the following OLS model:

Interest Spread =
$$\alpha_0 + \beta_1$$
 Media Sentiment + β_2 Controls
+ β_3 Alternative Info Sources, (3)

where *Spread* is the logarithm of the interest rate spread in basis points above LIBOR. We predict a negative coefficient on our main variable of interest—*Media Sentiment*. Variables reflecting controls and other information sources about the borrower are defined as previously. Further, the model includes loan purpose, industry, and year fixed effects.

We present our findings in table 5. Consistent with our predictions, we find that the interest spread is inversely associated with media sentiment, with more positive sentiment reducing the interest spread. In terms of economic significance, a one standard deviation increase in *Media Sentiment* translates into a 6.9 basis point decrease in spread. While this effect seems relatively modest, it is similar to the effect of a one standard deviation change in key credit risk measures, such as *ROA* (5.5 basis points) and *Leverage* (12.7 basis points).

With respect to controls, as expected, we report a negative relation between the spread and a borrower's profitability, size, Z-Score, and market-tobook ratio and a positive relation between the spread and leverage and an indicator variable reflecting losses. The coefficients on the loan-level controls are generally consistent with prior research. Larger loans and loans with pricing provisions are associated with lower spreads (Booth [1992], Asquith, Beatty, and Weber [2005]). Due to the endogenous determination of contractual terms, we observe a positive relation between the interest spread and both *Collateral* and *#Covenants* (Berger and Udell [1990], Bradley and Roberts [2004], Costello and Wittenberg Moerman [2011]).¹⁸

¹⁸ Agency theory suggests that there is a tradeoff between the restrictions imposed by the loan contract and the interest spread (Jensen and Meckling [1976], Myers [1977], Smith and

Variable	Dependent Variable: Spread
Media Sentiment	-1.0930***
	(0.2479)
ROA	-1.3283^{***}
	(0.3475)
Interest Coverage	0.0000
5	(0.0000)
Leverage	0.4288^{***}
0	(0.0581)
Size	-0.0396^{***}
	(0.0106)
Z-score	-0.0078^{**}
	(0.0035)
MTB	-0.0076^{**}
	(0.0030)
Loss	0.1287***
	(0.0236)
Amount	-0.0676***
	(0.0078)
Maturity	-0.0004
	(0.0004)
Collateral	0.5123***
	(0.0252)
PP	-0.1143***
	(0.0192)
#Covenants	0.0939***
	(0.0085)
Revolver	-0.0733***
	(0.0157)
Term Loan B	0.1725***
	(0.0258)
Press Release Sentiment	-0.9077***
	(0.2519)
Rated	-0.0021
	(0.0235)
Analyst Coverage	-0.0607***
ajor esterage	(0.0218)
Return	0.0800***
	(0.0294)
Negative Return	0.0314**
	(0.0155)
Fixed effects	Year/Industry/Purpose
N	6,964
R^2	0.6697

TABLE 5 Media Sentiment and Interest Rate Spread

This table presents the analysis of the effects of media sentiment on the interest spread. Standard errors are reported in parentheses. ****, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in the appendix.

While concurrently endogenizing all loan terms is beyond the scope of our paper, in untabulated analyses we estimate the interest rate model excluding loan controls; our findings are unchanged. We find that revolvers (Term Loans B and below) are priced at lower (higher) rates, consistent with prior research (Harjoto, Mullineaux, and Yi [2006], Zhang [2008], Nandy and Shao [2010], Lim, Minton, and Weisbach [2014]). We also show that more positive press release sentiment and analyst coverage are negatively associated with the spread.

The results presented in tables 3–5 are consistent with the media providing incremental information to less informed, nonrelationship lenders. To further support our claim that these lenders learn directly from the media, in an online appendix we provide additional analyses that address alternative information sources available for lenders in more detail. In particular, with respect to information provided by rating agencies, we account for the level of a borrower's credit rating, whether it is on a credit watch list, and its long-term credit outlook at a loan's issuance date as well as over the 180 days period prior to the issuance. We also account for equity analysts' earnings forecasts and recommendations at the time of a loan's issuance and the forecast and recommendation revisions over the 180 days prior to the issuance. We find that our main findings and inferences are robust to controlling for these additional information variables.

4.4 IS THE MEDIA A DIRECT SOURCE OF INFORMATION TO LESS INFORMED, OUTSIDE LENDERS?

To provide further evidence that the media is a direct source of information to lenders, we conduct analyses conditional on the intensity of a borrower's analysts' coverage. We conjecture that if media sentiment is informative to less informed lenders, its effect on syndicate structure will be stronger when a borrower's coverage by other information intermediaries is less intensive. This prediction is motivated by prior evidence that analysts are an important information source to lenders (e.g., Guntay and Hackbarth [2010], Mansi, Maxwell, and Miller [2011]). Therefore, outside lenders will rely more on alternative information sources, such as the media, when analyst following is relatively low. The idea is that investors' posterior expectations will respond more intensively to an information signal of a given precision, as prior uncertainty about a firm's prospects increases (e.g., Verrecchia [2001]). We classify borrowers as having highly intensive analyst coverage if the number of analysts covering the firm falls into the top quintile of the sample distribution, and as less intensive otherwise. We then

Warner [1979]). However, because more risky borrowers are likely to have higher spreads and lenders can simultaneously impose a higher number of covenants and/or require them to provide collateral, empirical tests typically reveal a positive relation between these variables.

estimate models (1)–(3) for the two subsamples—borrowers with highly intensive and less intensive coverage.¹⁹ We report these tests in table 6.

As evidenced from panel A, for the lead arranger analyses, the coefficient on *Media Sentiment* is positive and significant for the less intensive analyst coverage partition, but it is not significant for the highly intensive coverage partition, although the difference in the coefficients between the two partitions is not significant. Economically, for the less intensive analyst coverage partition, a one standard deviation increase in *Media Sentiment* increases the probability that a lender without a prior relationship with the borrower arranges the loan by 7.6%.

We find that, for both the Part-Borr No-Relationship and Part-Lead No-Relationship specifications, the coefficient on Media Sentiment is negative and significant for the less intensive analyst coverage partition, but it is not significant for the highly intensive coverage partition (table 6, panels B and C). The difference in coefficients on Media Sentiment between these partitions is also statistically significant. For the less intensive analyst coverage subsample, a one standard deviation increase in Media Sentiment increases the probability that a syndicate participant without a prior relationship with the borrower (lead arranger) joins the syndicate by 12.1% (12.1%). The results are similar when we examine the lead bank's loan share aspect of the syndicate structure (table 6, panel D). The effect of media sentiment is significant for borrowers with less intensive analyst coverage, but not for those with highly intensive coverage. The difference in coefficients on Media Sentiment between the two partitions is significant at the 10% level. For borrowers with less intensive analyst coverage, a one standard deviation increase in Media Sentiment decreases the proportion held by the lead arranger by 4.6%.

We also predict that the effect of media sentiment on the interest spread will be stronger for the less intensive analyst coverage partition. In table 6, panel E, we report that the coefficient on *Media Sentiment* is negative and significant for the less intensive coverage partition, but not significant for the highly intensive partition. The difference in coefficients on *Media Sentiment* between these partitions is statistically significant. For the less intensive analyst coverage subsample, a one standard deviation increase in *Media Sentiment* translates into a 7.3 basis points decrease in the interest spread.

The results presented in table 6, panels A–E are in line with the media providing more informative public signals to syndicate participants when information production by other information intermediaries is less intensive. However, firms with intensive analyst coverage can potentially differ from firms with less intensive coverage across a number of dimensions. We therefore examine how the effect of media sentiment differs across borrowers that have experienced a reduction in analysts' coverage prior to loan

¹⁹ For all analyses, our results are robust when we define borrowers as having highly intensive analyst coverage if the number of analysts covering the firm falls into the top quartile of the sample distribution.

TABLE 6

Variable	Dependent Variable: Borr-Lead No-Relationship		
	Less Intensive Coverage	Highly Intensive Coverage	
Media Sentiment	1.8891**	0.5731	
	(1.081)	(1.951)	
Model	Logit	Logit	
Controls	Included	Included	
Ν	5,268	1,987	
R^2	0.0402	0.0848	

The Effect of Media Sentiment Conditional on Intensity of Analysts' Coverage

Panel A: The probability of a nonrelationship lead arranger, conditional

Panel B: The syndicate participation of lenders without a prior relationship with the

borrower, conditional on analyst coverage intensity

	Dependent Variable: Part-Borr No-Relationship		
Variable	Less Intensive Coverage	Highly Intensive Coverage	
Media Sentiment	2.956***	-1.3446	
	$(0.908)^{\#}$	(1.693)	
Model	Logit	Logit	
Controls	Included	Included	
Observations	41,760	18,322	
Pseudo-R ²	0.0552	0.0577	

Panel C: The syndicate participation of lenders without a prior relationship with the lead arranger, conditional on analyst coverage intensity

Variable	Dependent Variable: Part-Lead No-Relationship		
	Less Intensive Coverage	Highly Intensive Coverage	
Media Sentiment	2.9520**	0.6063	
	(1.259)##	(1.764)	
Model	Logit	Logit	
Controls	Included	Included	
Observations	41,760	18,322	
Pseudo-R ²	0.0674	0.0529	

Panel D: The loan share retained by the lead arranger, conditional on analyst coverage intensity

	Dependent Variable: Lead Share		
Variable	Less Intensive Coverage	Highly Intensive Coverage	
Media Sentiment	-25.175^{***}	-10.1900	
	$(8.131)^{\#}$	(20.378)	
Model	OLS	OLS	
Controls	Included	Included	
Fixed effects	Year/Industry/Purpose	Year/Industry/Purpose	
Observations	3,108	1,505	
R^2	0.4646	0.6451	

(Continued)

Variable	Dependent Variable: Spread		
	Less Intensive Coverage	Highly Intensive Coverage	
Media Sentiment	-1.1654^{***}	-0.6627	
	$(0.209)^{\#\#}$	(0.449)	
Model	OLS	OLS	
Controls	Included	Included	
Fixed effects	Year/Industry/Purpose	Year/Industry/Purpose	
Observations	5,073	1,891	
R^2	0.6492	0.7580	

TABLE 6-Continued

This table presents the analysis of the effects of media sentiment on the probability of a nonrelationship lender serving as a loan's lead arranger (Borr-Lead No-Relationship), the probability of a participant's prior relationship with the borrower (Part-Borr No-Relationship), the probability of a participant's prior relationship with the lead arranger (Part-Lead No-Relationship), the loan share retained by the lead arranger (Lead Share), and the interest spread (Spread), conditional on equity analysts' coverage intensity. Standard errors are reported in parentheses.

***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. ###, ##, and # indicate that the difference across analyst coverage partitions is significant at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in the appendix.

initiation and those that have not. A reduction in analyst coverage has been shown to be associated with an exogenous (unrelated to changes in firm fundamentals) decrease in public information production about the firm (e.g., Kelly and Ljungqvist [2012], Balakrishnan, Billings, and Ljungqvist [2014]). We hypothesize that media signals will be more informative to nonrelationship lead arrangers and syndicate participants following coverage reductions. We report these analyses in table 7.

We identify that 4.3% of observations in the lead arranger analyses, 6.3% in the syndicate participant analyses, 4.2% in the lender share analyses, and 4.5% in the loan spread analysis relate to borrowers that have experienced a reduction in analyst coverage over the year preceding loan issuance.²⁰ Despite the small size of the coverage reduction partition, we find a significantly higher coefficient on Media Sentiment for this partition in Borr-Lead No-Relationship, Part-Borr No-Relationship, and Spread analyses (panels A, B, and E). For the Lead Share analyses (panel D), the coefficient on Media Sentiment for the analyst coverage reduction partition is higher than the respective coefficient for the no coverage reduction partition, but the difference is not statistically significant, likely due to the extremely small sample size of the coverage reduction partition.²¹

²⁰ When we measure analyst coverage reduction over the 180-day period prior to loan issuance, the proportion of borrowers that have experienced a reduction in analyst coverage becomes even smaller, preventing us from conducting empirical analyses.

²¹ For Part-Lead No-Relationship analysis (panel C), we do not find that the coefficient on Media Sentiment is significantly higher for the analyst reduction partition. This could be due to a lack of power in this small sample setting, as prior research has shown that a previous relationship between a participant and lead arranger is relatively less important than a previous relationship between a participant and the borrower (Sufi [2007]).

TABLE 7 The Effect of Media Sentiment Conditional on Analyst Coverage Reduction

	-
Panel A: The probability of a nonrelationship lead arranger, conditional	
Taner A. The probability of a nonrelationship lead a ranger, conditional	
on analyst coverage reduction	
on unityst coverage reduction	

	Dependent Variable: Borr-Lead No-Relationship				
Variable	No Coverage Reduction	Coverage Reduction			
Media Sentiment	1.3533*	7.9796*			
	(0.957)	$(5.079)^{\#}$			
Model	Logit	Logit			
Controls	Included	Included			
Observations	6,941	314			
Pseudo-R ^e	0.0502	0.0731			

Panel B: The syndicate participation of lenders without a prior relationship with the borrower, conditional on analyst coverage reduction

	Dependent Variable: Part-Borr No-Relationship				
Variable	No Coverage Reduction	Coverage Reduction			
Media Sentiment	1.6771**	4.1932*			
	(0.796)	$(2.640)^{\#}$			
Model	Logit	Logit			
Controls	Included	Included			
Observations	56,274	3,807			
Pseudo-R ²	0.0571	0.0835			

Panel C: The syndicate participation of lenders without a prior relationship with the lead arranger, conditional on analyst coverage reduction

	Dependent Variable: Part-Lead No-Relationship				
Variable	No Coverage Reduction	Coverage Reduction			
Media Sentiment	2.4008**	0.6238			
	(1.155)	(3.920)			
Model	Logit	Logit			
Controls	Included	Included			
Observations	56,274	3,807			
Pseudo-R ²	0.0613	0.0873			

Panel D: The loan share retained by the lead arranger, conditional on analyst coverage reduction

	Dependent Variable: Lead Share				
Variable	No Coverage Reduction	Coverage Reduction			
Media Sentiment	-25.7339^{***}	-35.0953			
	(7.144)	(52.041)			
Model	OLS	OLS			
Controls	Included	Included			
Fixed effects	Year/Industry/Purpose	Year/Industry/Purpose			
Observations	4,417	196			
R^2	0.4890	0.7396			

(Continued)

	Dependent V	ariable: Spread
Variable	No Coverage Reduction	Coverage Reduction
Media Sentiment	-1.1771***	-3.2312**
	(0.206)	$(1.396)^{\#}$
Model	OLS	OLS
Controls	Included	Included
Fixed effects	Year/Industry/Purpose	Year/Industry/Purpose
Observations	6,654	310
R^2	0.6619	0.8496

TABLE 7—Continued

This table presents the analysis of the effects of media sentiment on the probability of a nonrelationship lender serving as a loan's lead arranger (Borr-Lead No-Relationship), the probability of a participant's prior relationship with the borrower (Part-Borr No-Relationship), the probability of a participant's prior relationship with the lead arranger (Part-Lead No-Relationship), the loan share retained by the lead arranger (Lead Share), and the interest spread (Spread), conditional on equity analysts' coverage reduction. Standard errors are reported in parentheses.

***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively. ###, ##, and # indicate that the difference across analyst coverage partitions is significant at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in the appendix.

Overall, the stronger effects of media sentiment on the likelihood of a nonrelationship lead arranger, syndicate structure, and loan pricing for the less intensive analyst coverage and reduction in analyst coverage partitions reinforce our inference that the media serves as a direct source of information to less informed, outside lenders.

4.5 THE SENSITIVITY OF INTEREST SPREADS TO MEDIA SENTIMENT FOR CONSUMER PRODUCT FIRMS

The results in tables 3–7 suggest that the media influences loan pricing by providing incremental information to outside lenders, which increases the supply of credit by attracting nonrelationship lead arrangers and syndicate participants. In this section, we explore an additional channel through which the media can affect interest spreads by examining the possibility that the media influences consumers' perceptions of a borrower, changing their future purchasing decisions and consequently a borrower's future cash flows.

Specifically, we examine whether spreads are more sensitive to media sentiment for consumer product-oriented firms. Consumers' decisions to purchase a firm's products are potentially influenced by positive or negative news about the firm, affecting its future sales and cash flows. Because lenders are likely to incorporate the behavior of a borrower's consumers into loan pricing, we predict that the effect of media sentiment on loan pricing is likely to be more pronounced for more consumer productoriented firms.

Following Fisman, Heal, and Nair [2006], we identify whether a borrower is consumer product-oriented by the relative importance of its advertising expenses. We classify borrowers as having high consumer product intensity if the ratio of advertising expenses to sales revenue is above the sample

	Dependent Variable: Spread				
Variable	Low Intensity	High Intensity -1.0030***			
Media Sentiment	-0.636^{**}				
	(0.327)	$(0.343)^{\#}$			
Model	OLS	OLS			
Controls	Included	Included			
Fixed effects	Year/Industry/Purpose	Year/Industry/Purpose			
Ν	3,357	3,429			
R^2	0.6580	0.7107			

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The Effect of Media Content on Loan Pricing, Conditional on a Borrower's Consumer-Product Intensity

This table presents the analysis of the effects of media sentiment on the interest spread, conditional on a borrower's consumer product intensity. Standard errors are reported in parentheses.

***, **, and * indicate significance at the 0.01, 0.05, and 0.10 level, respectively.

^{###, ##}, and [#] indicate that the difference across consumer product intensity partitions is significant at the 0.01, 0.05, and 0.10 level, respectively. All variables are defined in the appendix.

median. We estimate the interest rate model for the consumer–product intensity partitions and report our findings in table 8. Consistent with our predictions, we find that, although the coefficient on *Media Sentiment* is negative and significant for both partitions, it is significantly larger for the high consumer product intensity partition relative to the low intensity partition.

Overall, our interest spread analyses suggests that media sentiment affects loan pricing both by informing less informed outside lenders and by changing consumers' perception of a borrower, the implications of which are priced by lenders.

5. Summary

This paper examines whether media coverage influences the syndicated loan origination and participation decisions of informationally disadvantaged lenders, loan syndicate structures, and interest spreads. We first investigate if the media influences competition between inside and outside lenders to serve as lead arranger on a borrower's loan. This analysis builds on Rajan [1992], who shows that the aggressiveness with which less informed lenders compete for a loan deal increases in the sentiment of public information signals about a borrower. We find that the probability of a nonrelationship lender serving as a lead arranger is higher when media sentiment is more positive.

We also consider whether media sentiment influences the willingness of lenders to participate as members in a loan syndicate. Syndicate participants face both adverse selection and moral hazard problems (e.g., Sufi [2007], Ivashina [2009]) and the media can influence lenders' participation through either channel. We find that the probability of nonrelationship participants joining the syndicate increases in media sentiment. Consistent with more positive media sentiment reducing the agency problems faced by less informed participants, we also show that the loan share retained by the lead arranger (i.e., the lead's skin in the game) is decreasing in media sentiment. We finally show that more positive media sentiment is associated with a lower interest rate spread. We attribute this finding to the increased supply of credit available to the borrower due to more positive media sentiment attracting nonrelationship lead arrangers and syndicate participants.

Further, we provide evidence consistent with the media being a direct source of information to lenders. First, we show that our results hold after controlling extensively for alternative public information sources, including analysts, rating agencies, and firm-initiated press releases. Second, we show that loan origination and participation decisions, as well as the loan share retained by the lead arranger and loan pricing, are more sensitive to media sentiment when a borrower's analyst coverage is lower and for borrowers that have experienced an exogenous reduction in analyst coverage preceding loan issuance. Finally, we provide evidence that loan spreads are more sensitive to media sentiment for consumer–product oriented borrowers, consistent with lenders accounting for the implications of the news for consumers' perception of a borrower and the resultant impact on the borrower's future sales and cash flows.

By providing evidence that the media serves as an important information intermediary in the private debt market, we extend the growing body of research on the role played by the media in capital markets. Our findings also add to the extensive literature on syndicated lending, as we demonstrate that the media fundamentally alters the information structure and the nature of competition in the loan market. In particular, an important insight gleaned from our work is that not only the availability of public information about a borrower, but also its sentiment, significantly influences the central characteristics of syndicated lending.

variable Definitions		
Variable	Definition	
Amount	The natural logarithm of the loan amount in US dollars (DealScan).	
Analyst Coverage	An indicator variable equal to 1 if there are any analysts following the firm in the month prior to the loan issuance date, 0 otherwise (I/B/E/S).	
Borr-Lead No-Relationship	An indicator variable equal to 1 if a loan's lead arranger has syndicated less than 50% of a borrower's prior loan deals by volume over the five-year period preceding the loan issuance date, 0 otherwise (DealScan).	
Collateral	An indicator equal to 1 if the loan is secured, 0 otherwise (DealScan).	
First Time Lead	An indicator variable equal to 1 if the lender serves as the lead arranger for the first time, 0 otherwise (DealScan).	

APPENDIX Variable Definitions

148	R. M. BUSHMAN,	C. D.	WILLIAMS, ANI) R.	WITTENBERG-MOERMAN

Variable	Definition		
Interest Coverage	Earnings before interest and tax divided by the interest expense,		
	estimated in the quarter preceding a loan's issuance (Compustat).		
Lead Share	The share of the loan held by the lead arranger, expressed in percentages (DealScan).		
Leverage	Total liabilities divided by total assets, estimated in the quarter preceding a loan's issuance (Compustat).		
Loss	An indicator variable equal to 1 if ROA is less than zero, 0 otherwise (Computat).		
Maturity	The number of months to maturity (DealScan).		
Media Sentiment	The average Composite Sentiment Score (CSS) over the 180-day period prior to a loan's issuance date for full-size articles, conditional on the article's relevance score above 75. Firm-initiated press releases are excluded from this estimation		
	CSS combines five sentiment scores (PEQ, BEE, BMQ, BCA, and BAM), while ensuring that there is no sentiment		
	disagreement among these scores. The PEQ score represents		
	the news sentiment of the given news item according to the		
	PEQ classifier, which specializes in identifying positive and		
	negative words and phrases in articles about global equities.		
	The BEE score represents the news sentiment of a given story		
	according to the BEE classifier, which specializes in news storie		
	about earnings evaluations. The BMQ score represents the		
	news sentiment of a given story according to the BMQ classifie		
	which specializes in short commentary and editorials on globa		
	equity markets. The BCA score represents the news sentiment		
	of a given news story according to the BCA classifier, which		
	specializes in reports on corporate action announcements. Th		
	BAM score represents the news sentiment of a given story		
	according to the BAM classifier, which specializes in news		
	stories about mergers, acquisitions, and takeovers.		
	CSS scores range from 0 to 100, with a score above 50 indicating		
	positive news sentiment; equal to 50, neutral news sentiment;		
	and below 50, negative news sentiment. We apply a linear		
	transformation to the CSS score and define <i>Media Content</i> =		
	(CSS - 50)/50, so that the <i>Media Content</i> ranges from -1 to 1, with zero being equivalent to neutral sentiment (RavenPack).		
MTB	The market value of equity divided by the book value of equity,		
	estimated in the quarter preceding a loan's issuance (Compustat).		
Negative Return	An indicator variable equal to 1 if <i>Return</i> is less than zero, 0 otherwise (CRSP).		
Outstanding	An indicator variable equal to 1 if the borrower's previous deals are still outstanding at the current loan's issuance date, 0 otherwise (DealScan).		
Part-Borr No-Relationship	An indicator variable equal to 1 if the participant has not been involved in a deal with the borrower over the five-year period preceding the loan issuance date, 0 otherwise (DealScan).		

Variable	Definition	
Part-Lead No-Relationship	An indicator variable equal to 1 if the participant has not been involved in a deal with the lead arranger over the five-year period preceding the loan issuance date, 0 otherwise (DealScan).	
Press Release Sentiment	The average CSS for firm-initiated press releases with a relevance score greater than 90, estimated over the 180-day period prior to a loan's issuance date (RavenPack).	
PP	An indicator variable equal to 1 if the loan has a performance pricing provision, 0 otherwise (DealScan).	
Rated	An indicator variable equal to 1 if the borrower has a senior debt rating from S&P, Moody's, or Fitch, 0 otherwise (DealScan and S&P historical database).	
Revolver	An indicator variable equal to 1 if the loan is a revolving line of credit, 0 otherwise (DealScan).	
Return	The firm's market-adjusted (value-weighted) cumulative return over the 180-day period prior to a loan's issuance date.	
ROA	Return on assets, defined as earnings before extraordinary items divided by total assets and estimated in the quarter preceding a loan's issuance (Compustat).	
Spread	The natural logarithm of the loan spread over LIBOR (DealScan).	
Size	The natural logarithm of total assets, estimated in the quarter preceding a loan's issuance (Compustat).	
Term Loan B	An indicator variable equal to 1 if the loan type is Term loan B or below (C, D, E, and F), 0 otherwise (DealScan).	
Tight Credit Supply	An indicator variable equal to 1 if the change in bank lending standards for mid-sized and large commercial loans, as reported in the Federal Reserve Board's quarterly Senior Loan Officer Opinion Survey on Bank Lending Practices, in the quarter of a loan's origination is in the top quartile of the sample's distribution, and 0 otherwise.	
Time-Between	The number of days between the loan's issuance date and the previous deal (DealScan).	
Z-Score	Altman's [1968] bankruptcy measure, estimated by the following model: $Z = 1.2X_1 + 1.4X_2 + 3.3X_3 + 0.6X_4 + 0.999X_5$, where X_1 is defined as working capital (total current asset minus total current liabilities) divided by total assets. X_2 is defined as retained earnings divided by total assets. X_3 is defined as earnings before interest and taxes divided by total assets. X_4 is the market value of equity divided by total liabilities. X_5 is total sales divided by total assets. All measures are estimated in the quarter preceding a loan's issuance (Compustat).	
#Covenants	The number of financial covenants (DealScan).	

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