Hispanic Ethnicity and Post-traumatic Stress Disorder after a Disaster: Evidence from a General Population Survey after September 11, 2001

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PURPOSE: To assess ethnic differences in the risk of post-traumatic stress disorder (PTSD) after a disaster, and to assess the factors that may explain these differences.

METHODS: We used data from a representative survey of the New York City metropolitan area (n = 2616) conducted 6 months after September 11, 2001. Linear models were fit to assess differences in the prevalence of PTSD between different groups of Hispanics and non-Hispanics and to evaluate potential explanatory variables.

RESULTS: Hispanics of Dominican or Puerto Rican origin (14.3% and 13.2%, respectively) were more likely than other Hispanics (6.1%) and non-Hispanics (5.2%) to report symptoms consistent with probable PTSD after the September 11 terrorist attacks. Dominicans and Puerto Ricans were more likely than persons of other races/ethnicities to have lower incomes, be younger, have lower social support, have had greater exposure to the September 11 attacks, and to have experienced a peri-event panic attack upon hearing of the September 11 attacks; these variables accounted for 60% to 74% of the observed higher prevalence of probable PTSD in these groups.

CONCLUSION: Socio-economic position, event exposures, social support, and peri-event emotional reactions may help explain differences in PTSD risk after disaster between Hispanic subgroups and non-Hispanics.

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KEY WORDS: Post-traumatic Stress Disorder, Race, Ethnicity, Hispanic, Disaster.
Selected Abbreviations and Acronyms

PTSD = post-traumatic stress disorder
SEP = socio-economic position
PR = Puerto Rico
DR = Dominican Republic
WTC = World Trade Center
NWS = National Women’s Study
RDD = random digit dial

Methods

Sample

Data for this analysis were from a random digit dial (RDD) household survey conducted between March 25 and June 25, 2002. The sampling frame for the survey included all adults in the following contiguous geographic areas: New York City and Nassau, Westchester, Suffolk, and Rockland counties in New York State; Hudson, Essex, Bergen, Passaic, Union, Middlesex, Monmouth, Morris, and Somerset counties in New Jersey, and Lower Fairfield County in Connecticut. All interviews were conducted by trained interviewers, in English, Spanish, Mandarin, and Cantonese, using translated and back-translated questionnaires and a computer-assisted telephone interview system. The overall cooperation rate (based on the sum of the number of completed interviews, quota outs, and screen-outs divided by the sum of completed interviews, quota outs, screen outs, refusals, and premature terminations) was 56% and the overall response rate (based on the sum of the number of completed and partial interviews divided by the sum of all numbers that were either eligible as residential telephone numbers or of unknown eligibility) was 34%. We include with this manuscript an Appendix that discusses the implications of response rates in RDD sampling in general and the relevance of these response rates to this study in particular (see Appendix). Sampling weights were developed and applied to our data to correct potential selection bias related to the number of household telephones, persons in the household, and over-sampling. Further discussions of the methods and results from these surveys can also be found elsewhere (12, 13, 16).

Survey Instrument

Respondents were asked questions using a structured interview which assessed the mental health consequences of natural disasters (17). We asked if the respondent was of Hispanic origin and then asked about respondent racial background by asking if respondents considered themselves White, Black, Asian, American Indian/Alaska Native, Native Hawaiian/Pacific Islander. We asked all respondents who considered themselves Hispanic what country their family was primarily from [e.g., Puerto Rico (PR), Dominican Republic (DR), etc.] using a list of predominantly Hispanic countries. Persons could report multiple race/ethnicities. We asked about socio-economic position (SEP) by assessing household income, educational attainment, and employment. We inquired about other demographic characteristics (age, gender, marital status, social support) and about September 11 event-experiences [e.g., proximity to the World Trade Center (WTC) complex during the attacks, if the respondent had witnessed the attacks of September 11, etc.]. We also assessed whether the respondent experienced symptoms consistent with a panic attack in the first few hours after hearing about the September 11 attacks, based on DSM-IV criteria for panic attacks (18). Since surveys were conducted in English, Spanish, Mandarin, and Cantonese, we considered language of survey administration as a potential explanatory variable in the relation between race/ethnicity and PTSD.

We used the National Women’s Study (NWS) PTSD module to assess PTSD symptoms since the September 11 terrorist attacks. The NWS PTSD module is a measure of PTSD that assesses the presence of criterion B, C, and D symptoms and determines content for content-specific symptoms (e.g., content of dreams or nightmares) if symptom presence is endorsed. We assessed probable PTSD that was related to the September 11 attacks based on the presence of necessary PTSD criterion B, C, and D symptoms since September 11. Participants were then required to report at least one re-experiencing symptom specific to the attack, at least three avoidance symptoms (content specific where relevant or since September 11), and two arousal symptoms (since September 11) for a diagnosis of probable PTSD related to the September 11 attacks since September 11.

The NWS PTSD module was validated in a field trial against the PTSD module of the Structured Clinical Interview for DSM-III-R (SCID) (19) administered by mental health professionals. In the field trial, instrument sensitivity
was 99% and specificity was 79% when compared to SCID diagnosis (20). The NWS PTSD module has been used in a number of RDD studies throughout the United States (21–23). Previous research using this measure among persons with a history of specific potentially traumatic events (e.g., rape, physical assault, or crime more generally) has shown that associations of these covariates with PTSD were highly consistent with those reported in other epidemiological studies that carefully assessed both history of events and PTSD, suggesting good construct validity for the NWS PTSD module (23).

### Analyses

We assessed the prevalence of probable PTSD in the following racial/ethnic groups: Asian, Black, White, Hispanic, and Other races and found that the prevalence of probable PTSD was higher among Hispanics than among other racial/ethnic groups, consistent with previous reports (12, 13). Recognizing the potential of this sample to provide a more nuanced analysis of the prevalence of probable PTSD in different Hispanic subgroups, we assessed the prevalence of probable PTSD among Asian, Black, White, Other races, Dominicans, Puerto Ricans, and Other Hispanic reflecting the large number of persons with origins in the DR and PR in our survey. Documenting a comparable prevalence of probable PTSD among other racial/ethnic groups but a higher prevalence of probable PTSD among Puerto Ricans and Dominicans, we collapsed race/ethnicity into four key racial/ethnic groups of interest for the rest of the analyses: Non-Hispanics, Dominicans, Puerto Ricans, and Other Hispanic.

We considered various methods of dealing with missing data in order to present bivariate and multivariable analyses with comparable sample sizes to allow for cross-model comparison. Although there was relatively little missing data for the key variables, income data (important for assessing the role of SEP as a potential confounder) was missing in approximately 18% of cases. As such we conducted the analyses presented here: 1) restricted to persons on whom we had complete data, including income data; 2) using several different data imputation methods, including median, mode, minimum, and maximum imputation for specific demographic groups; and 3) using a dummy variable for missing income data. Parameter estimates of interest in these three different sets of analyses did not differ appreciably, and we present all analyses here restricted to cases with no missing values on all variables but with a dummy variable representing missing income data.

We used two-tailed chi-square testing in bivariate analyses to assess the relations between key covariates and race/ethnicity and between the key covariates and probable PTSD. Covariates that were associated both with race/ethnicity and probable PTSD at the p < 0.1 level were considered to be possible confounders of the relation between race/ethnicity and probable PTSD and were included in multivariate logistic regression analyses.

We used a series of multivariable models to assess the contribution of each of the covariates identified above to the relation between race/ethnicity and probable PTSD. Starting from a bivariate model with race/ethnicity as the only independent variable, we sequentially added each of the potential confounders identified above to the model. We determined the percent change in the race/ethnicity regression parameter in subsequent models to assess the contribution of each added covariate to the race/ethnicity-PTSD relation. In the final multivariable model we assessed interactions between each of the covariates and race/ethnicity. Interactions that were statistically significant at the p < 0.05 level were included in the final model. We present the two interactions that were statistically significant in separate final multivariable models.

### TABLE 1. Demographic characteristics of respondents surveyed 6 months after September 11, 2001 in the NYC metropolitan area compared to anticipated demographic characteristics based on the 2000 US Census (n = 2616)

<table>
<thead>
<tr>
<th>Characteristics</th>
<th>Weighed percent from sample</th>
<th>Percent from 2000 US Census</th>
<th>Chi-square p-value</th>
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<td>38.2</td>
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<td>0.309</td>
</tr>
<tr>
<td>35–54</td>
<td>39.3</td>
<td>39.6</td>
<td></td>
</tr>
<tr>
<td>55 +</td>
<td>22.5</td>
<td>28.3</td>
<td></td>
</tr>
<tr>
<td>Gender</td>
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<td>Male</td>
<td>46.2</td>
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</tr>
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<td>Female</td>
<td>53.8</td>
<td>53.1</td>
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<tr>
<td>Race</td>
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<td></td>
<td></td>
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<tr>
<td>White</td>
<td>53.4</td>
<td>54.8</td>
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<tr>
<td>African American</td>
<td>16.4</td>
<td>16.5</td>
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<tr>
<td>Asian</td>
<td>5.4</td>
<td>7.2</td>
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<tr>
<td>Hispanic</td>
<td>21.0</td>
<td>18.5</td>
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<td>$75,000+</td>
<td>33.3</td>
<td>32.4</td>
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<td>&lt; $20,000</td>
<td>16.9</td>
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<tr>
<td>Never married</td>
<td>34.6</td>
<td>29.0</td>
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### TABLE 2. Bivariate relations between race/ethnicity and socio-economic position, demographic variables, social support, and event exposures in a survey of residents of the New York City metropolitan area conducted 6 months after September 11, 2001

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<th>Characteristics</th>
<th>Other races</th>
<th>Other Hispanic</th>
<th>Puerto Rican</th>
<th>Dominican</th>
<th>p-value*</th>
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<td>N</td>
<td>%</td>
<td>N</td>
<td>%</td>
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<td>155</td>
<td>5.2</td>
<td>17</td>
<td>6.1</td>
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<td>Income categories</td>
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<td>29</td>
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<td>74</td>
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<td>8.7</td>
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(continued)
TABLE 2. Continued

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<th>Other Hispanic</th>
<th>Puerto Rican</th>
<th>Dominican</th>
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<td>%</td>
<td>N</td>
<td>%</td>
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<td>Lost possessions</td>
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<td>Involved in rescue effort</td>
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*pTwo-tailed chi-squared p-value.

RESULTS

Out of 2705 persons eligible for this analysis, we restricted the dataset to 2616 persons on whom we had complete data for all variables except income; we used a dummy variable to represent missing income data to include these persons in all analyses. Table 1 shows the demographic characteristics of the respondents included in these analyses and of residents of the NYC metropolitan area, according to the US Census, to show comparability of the sample. The population sampled is statistically comparable to the anticipated data from the US Census and there were no appreciable differences between the population sampled and the underlying population.

Overall, 459 (21.0%) respondents identified themselves as Hispanic. 3.3% of the sample reported being of multiple race/ethnicities. The prevalence of probable PTSD among Hispanics in the sample was 9.6%, which was higher than that in other racial/ethnic groups (Black 7.0%, White 4.6%, Asian 4.4%, Other 6.2%). Separating the Hispanic group into its three largest constituent groups (persons from the DR, PR, or other countries) showed that the highest prevalence of probable PTSD was in persons from the DR (14.3%) and PR (13.2%); the prevalence of probable PTSD among Hispanics from other countries (6.1%) was not appreciably different than that for persons in other racial groups. The prevalence of probable PTSD in the final grouping that was used for the rest of these analyses was: non-Hispanics (5.2%), Dominicans (14.3%), Puerto Ricans (13.2%), and other Hispanics (6.1%). We note that persons with multiple race/ethnicities who reported that they were Hispanic were analyzed as such (within categories of PR, DR, and other Hispanic), and persons who did not report Hispanicity as one of their ethnicities were considered non-Hispanics.

Table 2 shows bivariate associations between key covariates and race/ethnicity in the final grouping discussed above. Variables that were significantly associated with race/ethnicity were: household income (p < 0.001), educational attainment (p < 0.001), age (p < 0.001), living south of 14th street on September 11 (p < 0.001), social support (p < 0.001), having experienced a peri-event panic attack (p < 0.001), and having had a friend or relative killed on September 11 (p = 0.094). Table 3 shows bivariate associations between key covariates and probable PTSD. Variables that were significantly associated with probable PTSD were: income (p = 0.003), age (p = 0.009), marital status (p = 0.012), living south of 14th street on September 11 (p < 0.001), social support (p = 0.003), prior exposure to traumatic events (p < 0.001), recent exposure to stressors (p < 0.001), having experienced a peri-event panic attack (p < 0.001), having had a friend or relative killed on September 11 (p = 0.001), having lost possessions in the attacks (p < 0.001), and having been involved in the rescue effort (p = 0.003). Language of survey administration was not significantly associated with likelihood of PTSD.

Variables that were associated with both probable PTSD and with race/ethnicity in bivariate analyses and as such were included in the final multivariable model were: income, age, having lived south of 14th street, social support, experiencing a peri-event panic attack, and having had a friend or relative killed in the attacks (Table 4). In the final multivariable model, Dominicans were 1.35 times more likely [95% confidence interval (CI), 0.49–3.77] and Puerto Ricans were 1.51 times more likely (95% CI, 0.59–3.83) to have symptoms consistent with probable PTSD. In the final multivariable model, the covariates added had accounted for 73.5% and 60.2% of the relative increase in PTSD among Dominicans and Puerto Ricans respectively. Interaction terms that were significant in the final multivariable model were the interaction between race/ethnicity and social support and between race/ethnicity and the experience of a peri-event panic attack. Table 5 shows the final multivariable model with the interaction terms added separately. As shown in Model I, among persons who experienced a peri-event panic attack, the odds ratio of probable PTSD was 8.41 for non-Hispanics, 4.57 for Puerto Ricans, and 11.02 for Dominicans, compared to...
### TABLE 3. Bivariate relations between probable PTSD and socio-economic position, demographic variables, social support, and event exposures in a survey of residents of the New York City metropolitan area conducted 6 months after September 11, 2001

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<th>p-value*</th>
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<td>%</td>
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<td>13.6</td>
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(continued)
non-Hispanics who did not experience a peri-event panic attack. As shown in Model II, among persons with low social support, the odds ratio of probable PTSD was 2.59 for non-Hispanics, 0.84 for other Hispanics, 3.90 for Puerto Ricans, and 0.16 for Dominicans compared to non-Hispanics with high social support.

**DISCUSSION**

Using data from a representative population survey of the New York City metropolitan area we found that Dominicans and Puerto Ricans were more likely to report symptoms consistent with probable PTSD than were other Hispanics or non-Hispanics. Variables including SEP, social support, exposure to the September 11 attacks, and peri-event emotional reactions, accounted for approximately 60% to 74% of the observed higher prevalence of probable PTSD in these groups. Experiencing a peri-event panic attack was a strong risk factor for Dominicans while low social support was a strong risk factor for Puerto Ricans.

The underlying finding in this study, namely that Hispanics are more likely to report probable PTSD is consistent with findings from some (3, 24), but not all (2, 7), studies that have explored the issue of racial/ethnic differences in risk for PTSD. A number of covariates partly explained the higher prevalence of probable PTSD among Hispanics in our sample. Adjustment for SEP accounted for up to 17.7% of the higher risk of probable PTSD among Puerto Ricans and up to 14.6% among Dominicans. The fundamental role that SEP may play in confounding the relation between race/ethnicity and health has been well documented although most of the work in this regard has focused on physical rather than mental health (25–27). The differential role of SEP between different minority groups (recognizing that Blacks, typically the racial group with the lowest SEP, did not have a higher prevalence of PTSD than did other groups) suggests that SEP is only partly responsible for the observed ethnic differences in PTSD. In contrast to the NVVRS (3), we found that the extent of exposure to the traumatic event (i.e., combat in the NVVRS, and the September 11 attacks in our study) partly explained the higher prevalence of PTSD reported by Hispanics in our study; this difference was probably attributable to a disproportionately high density of Hispanics living close to the WTC in NYC.

One of the first analyses to address ethnic differences in PTSD risk suggested that low social support and acculturation were responsible for the high risk of PTSD among Hispanics (28). This partly concurs with our observation that social support was a particularly important determinant of differences in PTSD risk between groups of Hispanics. However, we failed to observe a difference in incidence of probable PTSD between Hispanics who were interviewed in English or in Spanish in our study, suggesting that acculturation may play a lesser role in the post-September 11 context. An analysis of NVVRS data including Hispanic and non-Hispanic veterans concurred that acculturation was unlikely to explain the observed higher prevalence of PTSD among Hispanic Vietnam veterans (4). Consistent with our observation that Puerto Ricans (together with Dominicans) had the highest risk of PTSD after September 11, the NVVRS analysis found that Puerto Rican and Mexican veterans had a higher risk for PTSD than other Hispanic veterans (4). There were too few Mexican Hispanics in our sample to reliably assess the prevalence of probable PTSD in that group.

We documented an important role of peri-event emotional reactions in the differential risk of development of PTSD between subgroups of Hispanics; in particular, peri-event panic attacks were more important for persons of Dominican descent than for other Hispanics. Other studies have shown that peri-event panic attacks may play an important etiologic role in the development of PTSD (29,30). "Ataques de nervios," frequently described as a loss of control in emotional expressions, sensations, and actions in Hispanics, have been shown to be associated with panic disorder, post-traumatic stress, and depression (31–35).
TABLE 4. Multivariable models describing the relation between race/ethnicity and probable PTSD in a survey of residents of the New York City metropolitan area conducted 6 months after September 11, 2001 (n = 2616)*

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<tr>
<th>Characteristics</th>
<th>Model I OR 95% CI</th>
<th>Model II OR 95% CI</th>
<th>Model III OR 95% CI</th>
<th>Model IV OR 95% CI</th>
<th>Model V OR 95% CI</th>
<th>Model VI OR 95% CI</th>
<th>Model VII OR 95% CI</th>
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<td>1.00</td>
<td>1.00</td>
<td>1.00</td>
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<td>0.88 0.36–2.18</td>
<td>0.78 0.31–1.99</td>
<td>0.95 0.35–2.59</td>
<td>0.99 0.37–2.68</td>
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<td>2.33 1.17–4.66</td>
<td>2.22 1.12–4.41</td>
<td>2.20 1.10–4.38</td>
<td>2.19 1.09–4.38</td>
<td>1.50 0.62–3.62</td>
<td>1.51 0.59–3.83</td>
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<td>2.37 0.83–6.79</td>
<td>2.06 0.71–6.03</td>
<td>1.35 0.49–3.76</td>
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<td>1.00</td>
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<td>2.46 1.22–4.94</td>
<td>2.16 1.01–4.63</td>
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*Characteristics that were differentially distributed among racial groups and were significantly associated with probable PTSD in bivariate analysis were considered in this analysis. OR, odds ratios; CI, confidence intervals.
TABLE 5. Final models describing the relation between race/ethnicity and probable PTSD in a survey of residents of the New York City metropolitan area conducted 6 months after September 11, 2001 including interaction terms (n = 2616)

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<th>Wald p-value</th>
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<th>OR</th>
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<th>Wald p-value</th>
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*Odds ratios are not given for variables where interactions are present since these odds ratios are uninterpretable. Model parameters must be used to determine relative odds of the relation between specific variables and the outcome. For example, the odds ratio comparing persons from Puerto Rico who had a peri-event panic attack and non-Hispanics who did not would be \( \exp(1.45 + 2.13 - 1.13) = 4.26 \). A few key odds ratios are as follows: Dominicans who had a peri-event panic attack had an odds ratio of 11.02 and non-Hispanics who had a peri-event panic attack had an odds ratio of 8.41, both compared to non-Hispanics who did not report a peri-event panic attack. Persons from Puerto Rico who had low social support had an odds ratio of 3.90, persons from the Dominican Republic with low social support had an odds ratio of 0.16, and non-Hispanics with low social support had an odds ratio of 2.59, both compared to non-Hispanics with high social support.

These findings, taken together, suggest that differences in cultural or subjective factors, or emotional expression, may explain some of the differences we observed in the role of peri-event panic between Dominicans and Puerto Ricans in this sample.

The inferences that can be drawn from our results may be limited by a number of factors. There are several other variables that we did not measure in this survey that could account for some of the racial/ethnic disparities in probable PTSD prevalence. For example, our survey did not include a detailed past psychological history, limiting the conclusions we can draw from respondents’ reports of previous mental or emotional problems. We also did not measure factors such as acculturation, experiences of racism, or place of
birth of respondents, factors that have been considered in the past as potential mediators of the epidemiologic relation between race/ethnicity and PTSD. It is plausible that these and other variables partly account for the relation between race/ethnicity and PTSD in other urban contexts. This necessarily limits the inference that can be drawn from this analysis about the role of race/ethnicity in mental health. Also, the interaction terms discussed here rely on small cell sample sizes and as such, must be interpreted with caution.

It is possible that our sample introduced bias that was not accounted for by our analytic strategy. This is unlikely given the representativeness of the sample and the consistence between many of the analysis previously documented from this sample with those reported in other epidemiologic studies that assessed PTSD after disaster (12, 13, 16). In order to ensure comparability across multivariable models we restricted the sample size for all bivariate and multivariable analyses to the same sub-sample of persons. Although we accounted for missing data on income using multiple analytic techniques and sensitivity analyses did not suggest substantial differences in the conclusions drawn here, it remains possible that non-response bias accounted for some of our observed associations. Some of the conclusions drawn in the manuscript are limited by the modeling technique used. In order to best understand the range of variables that may be important in explaining the race/ethnicity–PTSD relation we included all measured variables that may statistically explain this relation (i.e., variables that were associated with both race/ethnicity and with probable PTSD in bivariate analyses). We note that all the variables included in the final model are important and contribute to the relation between race/ethnicity and probable PTSD in this sample.

In the particular context of Dominicans, we note that the crash of flight 587 bound to the DR and carrying mostly Dominicans, on November 12, 2001, could have had a bearing on some of the observations in this study. Although we did not ask specific questions about the crash of flight 587 in our survey, we asked about other prominent recent stressors and these did not play a role in the higher prevalence of PTSD among Dominicans documented in this study.

CONCLUSIONS
We observed in our study of the general population of the NYC metropolitan area after September 11 that specific socio-demographic variables, event exposures, social support, and peri-event emotional reactions may explain differences in PTSD risk after a disaster between Hispanic subgroups and non-Hispanics and that the latter two factors may be important determinants of differential risk among Hispanic subgroups.

We would like to acknowledge the contributions of Jennifer Ahern to the study discussed here. We would also like to thank Dr. Michael Bucuvalas, Dr. Bruce Link, and two anonymous reviewers for insightful comments that contributed immensely to the final version of this manuscript. We are indebted to those who participated in this study.

APPENDIX
On the Use of Telephone Surveys in Epidemiology
There is growing evidence of the validity of using random digit dial telephone surveys in epidemiologic assessments. One assessment conducted both in-person and telephone interview surveys simultaneously in the same area using the same interview schedule (36). This study, consistent with other work (37, 38), found that telephone respondents tended to be younger, better educated, and more likely to be white than in-person respondents. However, and most importantly, the same assessment showed that there were no substantial differences in accuracy of self-reported conditions or in health utilization questions. In fact, the assessment showed that internal consistency between responses was higher in the telephone surveys than in the in-person surveys. This study also showed that telephone surveys were appreciably cheaper to conduct than in-person surveys. A similar study (37), compared telephone and in-person surveys of physical morbidity using random assignment to one of the two methods. That study found no statistically significant differences between the two interview methods for overall assessment of health status, illnesses reported for the previous four months, or reports of hospitalization. Specific to mental health, several studies have shown that telephone assessment of axis I disorders (including depression and anxiety disorders) produced nearly identical results to in-person assessments using a variety of instruments (39, 40). Another study using both a clinical in-person assessment and a telephone survey found similarity in reporting of trauma and PTSD suggesting appropriateness of telephone methodology for collecting data relevant to this study (41). Importantly, telephone survey methods have become the standard method for the Behavioral Risk Factor Surveillance System (BRFSS), the national risk factor surveillance system. The BRFSS has been invaluable in providing assessments of ongoing trends in self-reported health and risk behaviors and its methods have been emulated in many local public health jurisdictions (42–44). In a recent analysis comparing national estimates of data from the BRFSS and the National Health Interview Survey (NHIS; which obtains information on medical conditions and health risk...
factors via within-household in-person interviews), it was shown that BRFSS estimates were similar to NHIS estimates for 13 of 14 measures examined, suggesting that any effect of telephone versus in-person interview on the quality of the information obtained was negligible (45). This burden of evidence then suggests that random digit dial telephone surveys can be an invaluable addition to the public health and epidemiologic armamentarium. In the context of assessments of large geographic areas such as the New York City metropolitan area, where the organization of representative in-person surveys would be logistically challenging and for most purposes prohibitively expensive, random digit dial surveys provide an efficient method of sampling of the general population.

One of the primary critiques of the use of telephone surveys frequently is that response rates in these assessments are lower than typical in-person epidemiologic assessments. Pinpointing an exact response rate in telephone surveys is often difficult due to the challenge in classifying indeterminate cases. Using standard definitions established by The American Association of Public Opinion Research (46), our study reports survey cooperation (56%) and response rates (34%) that are within the acceptable range for comparable random digit telephone surveys. The BRFSS response rates vary from year to year and from state to state, but the overall decrease in BRFSS response rates is well documented, with median response rates across states falling from a high of 71.4% in 1993 to a low of 48.9% in 2000. Surveys conducted by “official organizations,” such as health departments and the government typically have higher response rates than research surveys conducted by academic or non-profit institutions (47). Importantly, recent analyses of BRFSS data have shown that for a range of response rates for telephone surveys between 30% and 70%, the response rates were at most weakly associated with bias (48). In one analysis, it was shown that although a larger difference in response rate was associated with larger differences in estimates of cigarette smoking prevalence between the BRFSS and the in-person Current Population Survey (CPS), the effects were small with a 45 percentage point difference in response rates having a predicted difference in smoking prevalence estimates of 1.5 percentage points. In analysis comparing data from two identical surveys with response rates of 61% and 36%, very few significant differences were found across 91 comparisons (49). This developing research in the field suggests a weak association between response rates and epidemiologic estimates. Ultimately, further effort to recruit more reluctant non-participants may both be unethical and may itself introduce bias, since the reluctantly recruited participants potentially have reasons for providing false or misleading responses.

REFERENCES


