

Validity and reliability of Resource Utilization Groups (RUG-III) in Finnish long-term care facilities

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Resource Utilization Groups, Version III (RUG-III) is a case-mix system developed in the USA for classification of long-term care residents. This paper examines the validity and reliability of an adapted 22-group version of RUG-III (RUG-III/22) for use in long-term care facilities in Finland. Finnish cost weights for RUG-III/22 groups are calculated and different methods for their computation are evaluated. The study sample (1,964 residents) was collected in 1995–96 from ten long-term care facilities in Finland. RUG-III/22 alone explained 38.2% of the variance of total patient-specific (nursing + auxiliary staff) *per diem* cost. Resource use within RUG groups was relatively homogeneous. Other predictors of resource use included age, gender and length of stay. RUG-III/22 also met the standard for good reliability (i.e. a kappa value of 0.6 or higher) for crucial classification items, such as activities of daily living and high correlation between assessments based on relative cost.

Key words: case-mix, long-term care, nursing homes, patient classification, payment systems, resource use.

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INTRODUCTION

Case-mix classification systems describe resource utilization in healthcare. In acute care hospitals, case-mix measurement is based largely on applications of Diagnosis-Related Groups (DRGs) (1). DRGs have been implemented in the Nordic countries for management and policymaking throughout the 1990s, most recently as a payment method in health districts in southern Finland. In early applications of DRGs in the USA, it became clear that a system predicting cost of care episodes would not apply to long-term care provided to the elderly. In long-term care the variability of length of stay, and thereby of episode cost, is too great to be practical for case-mix classification and payment. Thus, another system, Resource Utilization Groups (RUGs), was developed for measuring case-mix of nursing home residents (2–4). It is based on a measure of *per diem* resource use.

In the USA, RUG Version III (RUG-III) forms the basis for prospective payment of publicly reimbursed nursing home care. In USA Medicaid programs (public health insurance for persons with low income) accounting for about half of all nursing home expenditures, RUG-III is used for payment in one third of US states.

In the federal Medicare program (public health insurance for persons over 65 years), covering nearly 10% of total nursing home costs, RUG-III is being adopted for national prospective payment. RUG classifications have been found to be valid across nations and healthcare systems. RUG-III has been tested successfully in long-term care facilities in England and Wales, Japan and Spain (5–7). An earlier version, RUG Version II, has been validated in Sweden and The Netherlands (8, 9). RUG-III is based on information in the Resident Assessment Instrument (RAI) and its Minimum Data Set (MDS). The RAI is a standardized assessment instrument implemented nationwide in the USA since 1990 for improving care planning and quality of care (10, 11). The RAI and the RUG-III classification are systems currently being tested in all the Nordic countries. In Iceland the RAI has been mandated by the government for use in all nursing homes since 1996 (12).

We assess here the validity and reliability of an adapted, 22-group version of RUG-III (RUG-III/22) in long-term care facilities in Finland. By validity, we mean the criterion-related or predictive validity of the classification system to explain resource use in long-term care. Reliability is measured by the inter-rater

agreement of classifying a resident into a specific RUG-III/22 group. The aim is also to derive Finnish cost weights, case-mix index, for RUG-III/22 to create a basis for a national case-mix measurement, and resource allocation in long-term care, e.g. for determining staffing levels in facilities and for developing payment methods and as an output measure in studies of productivity in long-term care facilities.

In Finland, local municipalities have the primary responsibility for providing institutional long-term care for the elderly. There are two basic settings for institutional care: residential homes and health centre hospitals (13). Residential homes are the most common form of long-term care facility, traditionally owned, managed, and administered through social services of single municipalities. Health centre hospitals provide care for the sickest and frailest long-term residents, although they also provide some acute care. They are owned and managed either by single municipalities or through federations of municipalities.

RESOURCE UTILIZATION GROUPS, VERSION III (RUG-III)

RUG classifications, first developed in the USA in the mid-1980s, were produced to meet statistical, clinical and incentives criteria (2–4). The major statistical criteria included the power of the classification system to explain the cost of caring for long-term care residents, and the homogeneity of cost within the classification groups. The RUG system was also constructed to make clinical sense: that residents within groups had clinical affinity. Patient characteristics used to classify residents, were chosen that could be reliably assessed and which would provide incentives for appropriate care. Where possible, measures of need for service were used rather than the provision of the service itself.

The latest version, RUG-III, was derived based on a sample of 7,658 residents in 203 nursing homes in seven US states (4). The system has been recently validated and slightly revised for the latest version of the MDS (Version 2.0) on an additional sample of over 2,000 residents. RUG-III classifies a resident into one of 44 distinct groups. The system incorporates three dimensions in describing and grouping a resident. The first dimension is represented by seven major clinical categories. These categories are devised as a hierarchy with decreasing cost intensity: 1) special rehabilitation: residents receiving different degrees of physical, occupational, or speech therapy; 2) extensive services: residents with respirator/ventilator care, parenteral feeding suctioning, or tracheostomy; 3) special care: e.g. residents with burns, coma, multiple sclerosis, pressure ulcers stage 3 or 4, quadriplegia, septicaemia, IV medications, or tube feeding; 4) clinically complex:

e.g. residents with aphasia, cerebral palsy, dehydration, hemiplegia, pneumonia, static ulcer, terminal illness, urinary tract infection, dialysis, or four or more physician visits per month; 5) impaired cognition: e.g. residents with impaired decision-making, orientation problems, short-term memory problems; 6) behaviour problems: e.g. residents with physical abuse, verbal abuse, or wandering; and 7) reduced physical functions: residents who do not meet the conditions of earlier categories.

The second dimension, used to subdivide the major categories, is based on an ADL (Activities of Daily Living) Index, a summary measure of functional capability in four ADLs: bed mobility, transfers, eating and toilet use. The ADL Index ranges from 4 to 18, the lowest value (4) indicating independence in all four ADLs, and the highest value total dependency in these same four ADLs. The third dimension forms tertiary splits in the classification and incorporates particular services; rehabilitation provided by nurses, or problems, presence of depression. Depression is used as tertiary splits in the “clinically complex” category, and “nursing rehabilitation” as tertiary splits in “impaired cognition”, “behaviour problems” and “reduced physical functions”. Detailed information of the derivation process and RUG-III definitions are provided elsewhere (4).

The RUG-III/22 model

Early in our empirical analysis it became clear that several groups of the original 44-group RUG-III model would be relatively rare in Finnish patterns of long-term care and the study sample. Thus, we decided to test a reduced model of RUG-III, consisting of only 22 groups. If a reduced model was supported by statistical criteria it would provide more stable estimates of the cost weights. The model we tested reduced the number of groups primarily by collapsing those groups formed by the tertiary splits (depression or nursing rehabilitation). These splits were initially constructed mainly to provide payment incentives, they added only 0.1% to the variance in cost of care explained by RUG-III. The differences in relative cost between two groups formed by tertiary splits were also minimal (4). Collapsing the tertiary splits reduced the number of groups by a total of 13. Given the low numbers of heavy rehabilitation residents, we decided also to collapse the number of rehabilitation groups from 12 to 3. The task of providing special rehabilitation to elderly residents is subject to local variation in Finland, and is not only provided by their long-term care facilities. Residents were classified as “special rehabilitation” if they received a weekly total of 45 min or more, at least four days a week, of physical, occupational or speech

Table 1. Reliability estimates of the RUG-III/22 model

	Number of residents in test group	% agreement RUG-III/22 group	% agreement RUG-III/22 clinical category	Mean kappa RUG-III/22 all items	Mean kappa RUG-III/22 ADL items	Pearson correlation cost weights
Personal nurse vs. personal nurse	41	68%	78%	0.71 ¹	0.73	0.93
Personal nurse vs. outsider evaluator	32	44%	69%	0.59 ²	0.68	0.78

NOTES: Kappa coefficient <0.40 poor agreement, 0.40–0.75 fair to good agreement, >0.75 excellent agreement (Fleiss, 1981).

¹Based on 44% of RUG-III/22 items (prevalence >10%).

²Based on 39% of RUG-III/22 items (prevalence >10%).

category. Informal care time was weighted using the wage-weight for nursing assistant/aide.

Our tests of the validity of the adapted RUG-III/22 model focused mainly on the statistical criteria, since the clinical and “administrative” criteria were regarded to be the same in the USA and Finland. For fitting RUG-III/22 models we used Analysis of Variance (ANOVA), applying regression analysis. Wage-weighted patient specific time served as the dependent variable. Initially, we examined four dependent variables: a wage-weighted time of nursing staff only, and three others adding different categories of auxiliary staff. The results using different compositions of staff care time did not differ greatly. We therefore report here only the results of wage-weighted total patient specific nursing + auxiliary care time. The explanatory power of the system was measured by the percentage of variance in resource use explained by the classification (R-square). Indicator (dummy) variables represented membership in each of the RUG-III/22 groups, the basic independent variables of the analysis. In addition, we evaluated how age, gender, length of stay and ward identifiers increased the fit of the 22-group model. The homogeneity of resource use within groups was measured by each groups’ coefficient of variation and by testing the difference between group means. The cost weights for each RUG-III/22 group were calculated by dividing the group mean by total sample mean of wage-weighted patient specific time. Refined cost weights were calculated adjusting for factors found statistically significant in covariate analysis. The adjusted cost weights were calculated based on parameter estimates of RUG-III/22 dummy variables from regression models where the covariates were included.

RESULTS

Reliability analysis

Overall, the agreement of two evaluators who knew the residents equally well was slightly better than the

agreement of a personal nurse and an outsider evaluator (Table 1). The percentage agreement reported indicates the proportion of residents classified by two evaluators into an identical RUG-III/22 group and main clinical category. For the comparison of two personal nurses, kappa values for each RUG-III/22 item with over 10% prevalence were statistically significant and greater than 0.40. The mean kappa 0.71 indicated good agreement. In the comparison of the personal nurse and outsider evaluator, two RUG-III/22 items in memory/recall ability were lower than 0.40 and the kappas of two additional items were not statistically significant. Overall, the mean kappa (0.59) indicated adequate or good agreement. Kappa values for ADL variables indicated high agreement in both test groups (0.73 and 0.68), ADLs are key items in determining the cost intensity of RUG-III/22 groups.

It should be noted that disagreement on a single item often results in classification into a different RUG-III/22 group, although not necessarily with a very different resource use. For costing purposes, agreement based on cost weights is of more relevance than agreement on clinical group. The correlation coefficient, in comparison of cost weights from assessments performed by the evaluators who knew the residents equally well, indicated high agreement ($r=0.93$). The mean cost weights of each pair of assessments was 0.996 and 0.984, and the difference was not statistically significant. For the other test group, the correlation was somewhat lower ($r=0.78$), but still acceptable. The outsider evaluator tended to grade a lower disability level than the personal nurse. The mean cost weight for the assessments by personal nurses was 0.945, while the corresponding mean by outsider evaluators was 0.864, but again the difference was not statistically significant.

Validity analysis and cost weights

In the sample used to test the validity of the RUG-III/22 model, the median age of the residents was 83 years,

Table 2. Frequency, coefficient of variation (CV), and cost weights by RUG-III/22 groups

RUG-III/22 Group Name	Group code	ADL range	N	CV	Cost weights			
				No adjustments	No adjustments	Adjusted for LOS, gender, age	Adjusted for ward	Adjusted for ward, LOS, gender, age
Special rehabilitation			80	0.63	1.59	1.61	1.58	1.61
REHAB 16–18	RFC	16–18	15	0.54	2.52	2.66	2.57	2.68
REHAB 7–15	RFB	7–15	41	0.50	1.62	1.67	1.64	1.69
REHAB 4–6	RFA	4–6	24	0.47	0.98	0.86	0.86	0.78
Extensive services			38	0.42	1.55	1.63	1.53	1.61
EXTENSIVE 2	SE2		4	0.41	1.40	1.46	1.39	1.43
EXTENSIVE 1	SE1		34	0.43	1.57	1.65	1.55	1.63
Special care			71	0.48	1.59	1.67	1.70	1.77
SPECIAL 17–18	SSC	17–18	37	0.44	1.87	2.00	2.00	2.11
SPECIAL 14–16	SSB	14–16	22	0.42	1.31	1.36	1.42	1.46
SPECIAL 7–13	SSA	7–13	12	0.42	1.26	1.25	1.30	1.28
Clinically complex			829	0.52	1.10	1.11	1.10	1.11
COMPLEX 17–18	CD	17–18	252	0.38	1.31	1.35	1.35	1.39
COMPLEX 11–16	CC	11–16	283	0.43	1.26	1.30	1.25	1.28
COMPLEX 6–10	CB	6–10	129	0.66	0.90	0.88	0.90	0.88
COMPLEX 4–5	CA	4–5	165	0.65	0.68	0.60	0.62	0.56
Impaired cognition			76	0.71	0.57	0.54	0.56	0.53
IMPAIRED 6–10	IB	6–10	30	0.55	0.76	0.74	0.70	0.67
IMPAIRED 4–5	IA	4–5	46	0.79	0.45	0.40	0.47	0.43
Behaviour problems			191	0.78	0.65	0.62	0.67	0.64
BEHAVIOUR 6–10	BB	6–10	78	0.62	0.90	0.90	0.92	0.92
BEHAVIOUR 4–5	BA	4–5	113	0.82	0.48	0.42	0.50	0.45
Reduced physical functions			679	0.67	0.86	0.85	0.85	0.83
PHYSICAL 16–18	PE	16–18	168	0.34	1.29	1.34	1.27	1.31
PHYSICAL 11–15	PD	11–15	152	0.47	1.15	1.19	1.15	1.18
PHYSICAL 9–10	PC	9–10	21	0.46	1.04	1.02	0.99	0.97
PHYSICAL 6–8	PB	6–8	75	0.56	0.77	0.76	0.71	0.69
PHYSICAL 4–5	PA	4–5	263	0.84	0.42	0.34	0.42	0.36
All*			1964	0.65	1.00	1.00	1.00	1.00

*The adjusted cost weights are calibrated the same manner as the non-adjusted, i.e. the mean cost weight of all residents was set to 1.00.

females accounted for 77.5%, and the mean length of stay for the residents was 865 days. The mean staff time caring for residents was 76.4 min per 24-h period, 72.3 min by nursing staff, and 4.1 min by auxiliary staff. The mean informal care time, substituting nursing care, was 8.9 min per 24 h. The case-mix adjusted mean staff time was lower than that of other validation studies. (Case-mix adjusted mean staff time per 24-h period using US cost weights, was 86 min for Finland, 92 min for Japan, 126 min for the USA and 148 min for England and Wales.) The average time for completing the MDS assessment form including only RUG-III items was 14 min.

The distribution of residents into RUG-III/22 main clinical categories and final groups are presented in Table 2. The two largest clinical categories were “clinically complex” (829 residents or 42%) and “reduced physical functions” (679 or 35%). These two categories have similarly been the largest in other validation studies (5–7). Residents were classified into 21 of 22 pos-

sible groups. The 22-group model explained 38.2% of the variance of total wage-weighted patient specific time (Table 3). Our explanatory power was slightly higher than that in the England and Wales study (35.6%), which used a similar sample size, but was lower than the variance explained in the Japanese study (43.8%), and the original US derivation (55.5%). The high variance explanation in the US study has been shown to be partly due to input from rehabilitation staff, the variance explanation of nursing staff cost alone was 41% (2). Adding dummy variables for wards to the RUG-III/22 model increased the explanatory power to 49.9%. Dummy variables for gender, age and length of stay (LOS) were all found statistically significant when added to the basic model. Adjusting for case-mix, residents with LOS of less than two weeks used about 22 min (29%) more care time than others, females used about seven min (10%) more than males, and residents under the age of 75 used six min (8%) more than those older. Interactions between gender,

Table 3. Variance explanation and covariate estimates of RUG-III/22 models

	RUG-III/22	RUG-III/22 + LOS, gender, age	RUG-III/22 + ward	RUG-III/22 + ward LOS, gender, age
Covariate estimates				
Length of stay (LOS) <2 weeks		22.1***		13.8***
Female		7.9***		7.4***
Age <75 years		6.0*		6.1**
R-square	0.382	0.400	0.499	0.508

Parameter estimates for RUG-III/22 groups and ward identifiers are not shown.

* $p < 0.05$.

** $p < 0.01$.

*** $p < 0.001$.

age and LOS were not statistically significant at the 0.05 level.

Table 2 presents four models for calculating cost weights. In the basic non-adjusted model, the cost weights for RUG-III/22 spanned a six-fold range from PA=0.42 (low) to RFC=2.52 (high). The range between the least costly and most costly group increased somewhat in the three adjusted models. In the non-adjusted RUG-III/22 model five groups (PA, BA, IA, CA and CB) had a higher coefficient of variation (CV) of resource use than that of the whole sample (0.65). These less homogeneous groups were also the least costly. A majority (75%) of the group means differed statistically ($p < 0.05$) from each other; when differences were not significant, one group was usually represented by only a few observations. Also, some overlap of cost weights was expected, as groups represent clinically different residents. The mean non-adjusted cost weight for residential home residents was 0.82, and for health centre hospital residents 1.15. Thus, the health centre residents were on average almost 30% more cost intensive than residential home residents. Within each RUG-III/22 clinical category residents of residential homes, not shown, were more concentrated into the lower ADL dependency groups than those of health centres.

DISCUSSION

A simple algorithm of the Resource Utilization Groups, Version III appears to be a valid and reliable system in Finnish long-term care facilities. The inter-rater reliability of RUG-III/22 assessments was high, especially for crucial classification items, such as ADLs, and when comparing cost weights. The low kappa values found in memory/recall ability in the comparison of a personal nurse and outsider evaluator could be due to the fact that the ability to remember

can be highly variable for persons with impaired cognitive skills and may therefore be difficult to assess by an outsider evaluator. Our sample size did not allow assessment of inter-rater reliability of all individual RUG-III/22/MDS variables. Extensive reliability estimates of MDS items have been performed in the USA and in other countries (11, 16).

The adapted RUG-III/22 classification alone explained a considerable amount of variation in cost (38.2%), comparable to that found in other countries. The exclusion of short-term wards may have resulted in a lower variance explanation and a lower proportion of rehabilitation residents. Resource use within groups was relatively homogeneous and the fact that the low cost groups were more heterogeneous also corresponds to results of other derivation and validation studies (4–7). The merging of original RUG-III groups to form the 22-group model did not significantly affect the variance explanation and the homogeneity of groups. The adjustments and exclusion of groups in the “special rehabilitation” category were made based on Finnish practice patterns and the study sample. Whether the current provision of rehabilitation is adequate and effective could not be determined.

The cost weights of the adapted RUG-III/22 followed logical patterns increasing with higher dependency in ADLs. Health centre residents were as expected more costly than residential home residents. The cost weights computed for RUG-III/22 groups were based on relatively large group sizes. In only three groups were the number of observations less than 20. In particular, the cost weight for SE2, with only four observations, must be regarded highly unstable. Overall, the cost weights show patterns similar to those computed in other countries. Consistency of direct mean care time by RUG-III groups between five countries has been previously shown by Carpenter et al. (17). We do not know the potential bias caused by

including two large facilities, and only those in urban areas. It did not appear that staff perceived the project as a "time and motion study" and that this influenced management decisions on staffing levels. Feedback sessions indicated that the amount of training given for collecting the data was sufficient in most, although not all, cases.

The fact that age, gender and length of stay (LOS) were found to be statistically significant can be regarded as reasonable. It is plausible to assume that residents entering a facility receive more care time, controlling for case-mix, during the first few weeks due to greater service needed for "settling in". Younger residents (<75 years) may be more costly because of greater rehabilitation potential. It is not clear why females would be more costly than men, despite the fact that the vast majority of nursing home residents are female. If age, gender and LOS were to be used as separate outputs (e.g. for payment), it would be well founded to use the cost weights that adjust for these three variables. If ward differences express productivity differences, it would be reasonable to use cost weights adjusting for different staff patterns. On the other hand, different staff patterns may have also expressed different quality and outcomes of care, about which we had no information. Overall, the adjusted RUG-III/22 cost weights did not differ greatly from the non-adjusted.

In conclusion, the results of this study support previous evidence on the transferability of the RUG system between healthcare systems. There appears to be good evidence of the feasibility of using the RUG classification in Nordic settings, building upon the work with the earlier version of RUGs (RUG-II) in long-term care facilities in Sweden (8). The RUG-III/22 model presented in this paper could be especially useful in applications such as information systems, as it requires fewer classification items.

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REFERENCES

1. Fetter RB, Brand DA, Gammache D. DRGs: Their Design and Development. Ann Arbor, Michigan: Health Administration Press, 1991.
2. Fries BE, Cooney LM. Resource Utilization Groups: a patient classification for long-term care. *Med Care* 1985; 23: 110–22.
3. Schneider DP, Fries BE, Foley WJ, Desmond M, Gormley W. Case mix for nursing home payment: Resource Utilization Groups, version II. *Health Care Financing Review, Annual Supplement* 1988; 9: 39–52.
4. Fries BE, Schneider DP, Foley WJ, Gavazzi M, Burke R, Cornelius E. Refining a case mix measure for nursing homes: Resource Utilization Groups (RUG-III). *Med Care* 1994; 32: 668–85.
5. Carpenter IG, Main A, Turner GF. Casemix for the elderly inpatient: Resource Utilization Groups (RUGs) validation project. *Age Aging* 1995; 24: 5–13.
6. Ikegami N, Fries BE, Takagi Y, Ikeda S, Ibe T. Applying RUG-III in Japanese long-term care facilities. *Gerontologist* 1994; 34: 628–39.
7. Carrillo E, Garcia-Altes A, Peiro S, Portella E, et al. System for the classification of patients in mid and long-term care facilities: Resource Utilization Groups, version III. Validation in Spain. (in Spanish). *Revista de Gerontologia* 1996; 6: 276–84.
8. Ljunggren G, Fries BE, Winblad U. International validation and reliability testing of a patient classification system for long-term care. *Eur J Gerontol* 1992; 1: 48–59.
9. Frijters D, Van der Kooij C. Resource Utilization Groups for Nursing Home Patients in the Netherlands: SIG, Dutch Centre for Health Care Information, Utrecht, 1991.
10. Morris JN, Hawes C, Fries BE, et al. Designing the National Resident Assessment Instrument for nursing homes. *Gerontologist* 1990; 30: 293–307.
11. Hawes C, Morris JN, Phillips CD, Mor V, Fries BE, Nonemaker S. Reliability estimates for the Minimum Data Set for nursing home resident assessment and care screening (MDS). *Gerontologist* 1995; 35: 172–8.
12. Jónsson PV. Letter from Reykjavik. *Annals of Internal Medicine* 1998; 128: 941–5.
13. Aro S, Noro A, Salinto M. Deinstitutionalization of the elderly in Finland 1981–91. *Scand J Soc Med* 1997; 25: 136–43.
14. Fleiss JL. Statistical methods for rates and proportions. 2nd edn, New York: John Wiley & Sons, 1981.
15. Fleiss JL, Cohen J. The equivalence of weighted kappa and the intraclass correlation coefficient as measures of reliability. *Educ Psychol Meas* 1973; 33: 613–19.
16. Sgadari A, Morris JN, Fries BE, et al. Efforts to establish the reliability of the Resident Assessment Instrument. *Age Aging*, 1997; Suppl 26: 2:27–31.
17. Carpenter GI, Ikegami N, Ljunggren G, Carrillo E, Fries BE. RUG-III and Resource allocation: Comparing the relationship of direct care time with patient characteristics in five countries. *Age Aging* 1997; 26: 61–5.