

Meta-Analysis of Association Between Obsessive-Compulsive Disorder and the 3' Region of Neuronal Glutamate Transporter Gene *SLC1A1*

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The neuronal glutamate transporter gene *SLC1A1* is a candidate gene for obsessive-compulsive disorder (OCD) based on linkage studies and convergent evidence implicating glutamate in OCD etiology. The 3' end of *SLC1A1* is the only genomic region with consistently demonstrated OCD association, especially when analyzing male-only probands. However, specific allele associations have not been consistently replicated, and recent OCD genome-wide association and meta-analysis studies have not incorporated all previously associated *SLC1A1* SNPs. To clarify the nature of association between *SLC1A1* and OCD, pooled analysis was performed on all available relevant raw study data, comprising a final sample of 815 trios, 306 cases and 634 controls. This revealed weak association between OCD and one of nine tested *SLC1A1* polymorphisms (rs301443; uncorrected $P=0.046$; non-significant corrected P). Secondary analyses of male-affecteds only ($N=358$ trios and 133 cases) demonstrated modest association between OCD and a different SNP (rs12682807; uncorrected $P=0.012$; non-significant corrected P). Findings of this meta-analysis are consistent with the trend of previous candidate gene studies in psychiatry and do not clarify the putative role of *SLC1A1* in OCD pathophysiology. Nonetheless, it may be important to further examine the potential associations demonstrated in this amalgamated sample, especially since the SNPs with modest associations were not included in the more highly powered recent GWAS or in a past meta-analysis including five *SLC1A1* polymorphisms. This study underscores the need for much larger sample sizes in future genetic association studies and suggests that next-generation sequencing may be beneficial in examining the potential role of rare variants in OCD. © 2013 Wiley Periodicals, Inc.

Key words: obsessive-compulsive disorder; OCD; candidate gene; meta-analysis; *SLC1A1*; glutamate

INTRODUCTION

Obsessive-compulsive disorder (OCD) is a debilitating anxiety disorder characterized by recurrent intrusive thoughts and/or

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rituals intended to mitigate anxiety [American Psychiatric Association, 2000]. OCD is a common psychiatric illness, with estimated lifetime prevalence rates ranging from 1% to 3% worldwide [Karno et al., 1988; Weissman et al., 1994; Kessler et al., 2005]. The genetic architecture of OCD is thought to be complex, [Cavallini et al., 1999; Nestadt et al., 2000a; Hanna et al., 2005a,b] but the specific mechanisms of this etiologic contribution have not yet been established. Despite numerous family, twin, and segregation analyses indicating substantial heritability of OCD, [Inouye, 1965; van Grootheest et al., 2005] most candidate gene findings have not been consistently replicated [Hemmings and Stein, 2006; Grados and Wilcox, 2007; Nestadt et al., 2010]. To date, gene selection for these studies has been guided by pathophysiological and pharmacological knowledge and therefore has focused primarily on the serotonergic and dopaminergic systems. However, among published OCD candidate gene studies, only those focusing on the glutamate transporter gene *SLC1A1* have demonstrated consistent evidence of

association [Hemmings and Stein, 2006; Grados and Wilcox, 2007; Nestadt et al., 2010].

Evidence suggests that abnormal glutamatergic neurotransmission may contribute to OCD pathophysiology [Pittenger et al., 2006, 2011; Wu et al., 2012]. Elevated glutamate concentrations have been found in the cerebrospinal fluid of OCD subjects, [Chakrabarty et al., 2005] and magnetic resonance spectroscopy investigations have revealed altered glutamatergic concentrations in the orbitofrontal white matter of adult patients with OCD [Whiteside et al., 2006] and in the anterior cingulate cortex of both pediatric and adult OCD subjects [Rosenberg et al., 2004; Yucel et al., 2008]. Positron emission tomography studies of OCD have shown increased glucose metabolism and blood flow in the cortico-striatal-thalamo-cortical (CSTC) circuit [Swedo et al., 1989; Rauch et al., 1994; Saxena and Rauch, 2000]. Moreover, two OCD rodent model studies have implicated glutamate-related genes, *SAPAP3* and *SLITRK5* [Welch et al., 2007; Mah, 2010]. Additionally, pharmacological evidence for the role of glutamatergic activity in OCD is emerging. Multiple studies have reported symptom improvement in OCD by augmenting with glutamate-modulating agents such as riluzole [Coric et al., 2005; Grant et al., 2007] or memantine [Aboujaoude et al., 2009; Stewart et al., 2010].

Located on chromosome 9p24, *SLC1A1* encodes the neuronal and epithelial high affinity glutamate transporter EAAC1, also called EAAT3, which is expressed in brain regions implicated in OCD, [Menzies et al., 2008] including the cerebral cortex, striatum, and thalamus [Kanai and Hediger, 2004]. In contrast to the other studied glutamate-related candidate genes such as *GRIK2* [Sampaio et al., 2011] and *BDNF* [Nicolini et al., 2009], *SLC1A1* is also a strong positional candidate. A whole genome linkage scan on 56 individuals from seven multigenerational OCD-affected families by Hanna et al. [2002] identified a linkage peak containing the 9p24 region (LOD 2.25), and a targeted replication study in this region on 50 multiplex or sibling pair pedigrees by Willour et al. [2004b] found suggestive evidence for linkage within 0.5 cM (<350 kb) of the original signal (HLOD 2.26). However, a subsequent full genome linkage scan primarily in sibling pairs did not find evidence for linkage in this region [Shugart et al., 2006b]. Of the known genes within this region, *SLC1A1* is one of only two with demonstrated expression in the brain [Smith et al., 1994].

To date, five candidate gene studies have reported associations between OCD and multiple single nucleotide polymorphisms (SNPs) and/or haplotypes within the 3' region of *SLC1A1* [Arnold et al., 2006; Dickel et al., 2006; Stewart et al., 2007a; Shugart et al., 2009; Wendland et al., 2009] (see Table 1). A recent study by Samuels et al. [2011] also found evidence of association with OCD in the 5' region of *SLC1A1*. No negative candidate gene studies of *SLC1A1* in OCD have been published. Among the five studies examining the 3' region of the gene, some of the analyzed SNPs overlapped, while others were unique across studies. The majority of SNP and haplotype associations reported in individual studies have not been replicated in other samples, and none of the *SLC1A1* polymorphisms has been confirmed as a susceptibility locus for OCD. Of the 18 SNPs analyzed from the 3' region of *SLC1A1* in at least two published studies, only rs3780412 has demonstrated association as a single marker in more than one study sample

(after correction for multiple testing [Dickel et al., 2006; Stewart et al., 2007a]).

Both Dickel et al. [2006] and Stewart et al. [2007a] found evidence of association with rs3780412, although the effects were in opposite directions. Dickel et al. [2006] reported over-transmission while Stewart et al. [2007b] reported under-transmission of the major allele to affected male probands. Incidentally, Kwon et al. [2009] also found weak evidence of support for association between this SNP and atypical antipsychotic-induced OC symptoms (permuted $P=0.07$).

Because the sample sizes of the five OCD 3' *SLC1A1* candidate gene studies were small ($N=66-378$ families), and also given the complex genetics of OCD, these ambiguous results are not surprising. Most recently, the first genome-wide association study (GWAS) of OCD has been published [Stewart et al., 2012], which did not identify significant association with any of the four examined *SLC1A1* SNPs, including rs2072657, rs301430, rs301434, and rs3780412. However, the GWAS notably did not examine several *SLC1A1* SNPs with previous reported association with OCD, and sex-specific analyses were not conducted. Moreover, a recent meta-analysis of OCD genetic association studies did not use raw genotype data to apply standard quality control thresholds across samples, and only examined five *SLC1A1* SNPs in primary analyses [Taylor, 2012].

In order to increase the statistical power to detect association between *SLC1A1* and OCD, and to detect the potential influence of publication bias or issues of multiple testing, the authors aggregated the raw data from the four published family-based association studies with unpublished data from genotyping performed by the OCD Mini-Collaborative, a subgroup of the International OCD Genetics Collaborative, consisting of research sites in Boston, Ann Arbor (Michigan), San Francisco, and Toronto. Analyses comprised an initial pooled analysis of this combined family-based association data, followed by a meta-analysis incorporating these results with the previously published results of a case-control association study. Secondary analyses were also performed to explore potential effects of gender on this association.

METHODS

Identification of Relevant Studies and Samples for Inclusion

Study selection procedure. Published studies examining the association between *SLC1A1* and OCD were originally identified via the MEDLINE/PubMed database using the terms "OCD" and "*SLC1A1*," with a focus on the 3' region of the gene. Additional published studies were identified through examination of the reference lists of the studies identified through MEDLINE. Unpublished datasets and raw genotype data for the pooled analysis were sought by contacting the authors of the published studies.

Study inclusion/exclusion. Studies were included in the pooled and/or meta-analyses if they had a type of control group (either unrelated controls or parents) and an OCD group, if SNPs within the 3' region of the *SLC1A1* gene were examined and if the original authors agreed to provide raw genotype data. Excluded studies involved review articles and association studies that examined the

relationship of *SLC1A1* to non-OCD anxiety disorders. Data from *SLC1A1* SNPs in any published (or unpublished) linkage studies were not included in analyses.

Raw data extraction. Raw data from individual studies that was analyzed included phenotype information, SNP genotypes and minor allele frequencies.

Quality control. In order to obtain the most accurate assessment of data quality, the full datasets provided by the sites were used in the calculation of QC statistics. Genotype data from each site were cleaned using an iterative quality control (QC) process, whereby markers and subjects with excessive missing values (>50%) were dropped before the next round of QC standards were applied. In subsequent steps, SNPs were dropped for excessive missing values (>10%), low minor allele frequencies (<5%), high Mendel error rates (>5%), and apparent deviations from Hardy–Weinberg equilibrium (P -value < 10^{-3}). Subjects were removed for excessive missing values (>10%) and for high Mendel error rates (>5%). All QC steps were performed using PLINK [Purcell et al., 2007] (<http://pngu.mgh.harvard.edu/purcell/plink/>) for the pooled dataset. Additionally, QC checks were performed on the merged final dataset to ensure that Mendel error rates and Hardy–Weinberg P -values remained within acceptable ranges.

Statistical Analyses

We first conducted a pooled analysis of the family based data in PLINK using the transmission disequilibrium test (TDT) [Purcell et al., 2007]. We subsequently identified the common set of genotyped SNPs between family-based studies and case–control studies and then conducted single-marker analyses of these SNPs using PLINK. For the pooled analysis of family-based data, we conducted only single-marker association testing and did not examine haplotypes.

We then performed a meta-analysis using weighted z -scores to combine association results from the pooled family-based analysis and the case–control analysis [see de Bakker et al., 2008, for a review of the method]. The z -score was calculated from the chi-square and odds ratio test statistics provided in the PLINK output:

$$z_i = \sqrt{\chi_1^2} \quad \text{for chi-square testes with 1 df; } z > 0 \\ \text{if odds ratio } > 1, \text{ and } z < 0 \text{ if odds ratio } < 1$$

As the original studies included different groups of SNPs in their analyses, the sample size for each SNP test in this study was essentially unique. Thus, rather than weighting each original study (as is typically done in a meta-analysis) this study weighted each single-marker test. PLINK was used to obtain the number of complete trios (sample size) genotyped for each SNP. The four association tests performed in the case–control sample were given the same weight, as there was only one case–control study. Weighting was established using the Genetic Power Calculator [Purcell et al., 2003]. This online tool was used to obtain the non-centrality parameter (NCP) for each of these samples and to generate a theoretical equivalent “effective” case–control study with a 1:1 case:control ratio. The theoretical case–control sample with the NCP value closest to that of the actual sample was used to weight each single-marker association test as follows:

$$\text{Relative weight} = \sqrt{\frac{N_i}{N_{\text{total}}}}$$

where N_i is the theoretical equivalent case–control sample size and N_{total} is the sum of the theoretical sample sizes of the association tests performed for that SNP, that is the sum of theoretical equivalent weights of the family-based test and case–control test, when a case–control association test had been performed for that SNP.

The tests for the five SNPs genotyped exclusively in family-based studies did not require weighting, as no case–control results were available for meta-analysis. For the other four SNPs, weighted z -scores were summed to produce overall meta-analysis z -scores.

Chi-square values were generated by squaring the z -scores from the meta-analysis, and these chi-square values were then converted into P -values based on the normal distribution with one-degree of freedom. The results of the association tests for each of the nine polymorphisms were indicated as z -scores, P -values, and odds ratios with corresponding 95% confidence intervals. Odds ratios were calculated according to a fixed-effects model [Trikalinos et al., 2008]. To account for multiple testing of the nine examined SNPs, Bonferroni correction was conducted, and the required adjusted P -value was defined at 0.0056.

RESULTS

Study-Sample Descriptions

Samples included in this analysis were incorporated from one case–control study [Wendland et al., 2009] and four family-based studies [Arnold et al., 2006; Dickel et al., 2006; Stewart et al., 2007a; Shugart et al., 2009]. It should be noted that the sample from the study by Samuels et al. [2011] was the same as that by Shugart et al. [2009], but with genotyping of additional SNPs. Authors from each of these five identified studies agreed to provide original raw genotype data from their samples to enable the application of uniform quality control (QC) filters for the purpose of this study. Unpublished family-based data from the OCD Mini-Collaborative, which included samples from Boston, Ann Arbor (Michigan), San Francisco and Toronto, were also entered into the analyses.

In each original study, case or proband status was defined by the presence of a current diagnosis of Obsessive–Compulsive Disorder based on DSM-III-R or DSM-IV (Diagnostic and Statistical Manual of Mental Disorders), [American Psychiatric Association, 2000] criteria which were assessed via established psychiatric interviews. Subjects were recruited primarily from OCD and anxiety clinics, as described previously [Arnold et al., 2006; Dickel et al., 2006; Stewart et al., 2007a; Shugart et al., 2009; Wendland et al., 2009]. Subjects from these original cohorts were primarily Caucasian, although age and, notably, age-of-onset were variable. Only one study [Shugart et al., 2009] and the San Francisco site of the OCD Mini-Collaborative excluded probands with comorbid Tourette disorder, a common comorbidity estimated to occur in 10–30% of OCD cases [Geller et al., 1996; Ivarsson et al., 2008; Langley et al., 2010]. Previously unanalyzed genotype data collected through the OCD Mini-Collaborative were also included. After QC exclusions, the OCD Mini-Collaborative sample contributed 170 new complete

trios as well as 124 complete trios that had been previously included in one of the original five family-based association studies.

Study-Sample Genotyping and Verification

The platforms used for genotyping in the different published studies included TaqMan (TaqMan; Applied Biosystems; Foster City, California), Illumina GoldenGate and BeadArray system (Illumina; San Diego, CA), and Sequenom hME (Sequenom; MassARRAY; San Diego, CA). Genotyping performed by the OCD Mini-Collaborative was conducted on the Sequenom iPLEX platform (Sequenom; MassARRAY; San Diego, CA). In earlier studies performed by Arnold et al. [2006] and Dickel et al. [2006] the primary goal of SNP selection was to provide adequate coverage of *SLC1A1*. Subsequent studies by Stewart et al., Shugart et al., and Wendland et al. were designed, in part, to replicate the positive associations reported by these earlier studies, such that the majority of markers selected had been genotyped previously. In addition, Shugart et al. included SNPs from outside of the *SLC1A1* gene with potential regulatory properties, and Wendland et al. included SNPs predicting gene expression levels.

Some of the original genotyping runs included assays unrelated to *SLC1A1* as well as subjects affected by disorders other than OCD. After QC exclusions were made, the datasets were refined by excluding all data related to OCD-unaffected families (from family-based databases) and non-3' *SLC1A1*-related SNPs. Because the QC thresholds used in this study differed slightly from those applied in the original studies, the post-QC datasets for meta-analysis inclusion were not identical to datasets analyzed in the original studies. To address this fact, attempts were made to replicate the original results by analyzing our post-QC datasets according to the methods—namely programs and settings—described in the published papers [FBAT (<http://www.biostat.harvard.edu/~fbat/default.htm>), [Laird et al., 2000] PBAT (<http://www.biostat.harvard.edu/~clange/default.htm>), [Lange et al., 2004] and TDT in Haploview (<http://www.broadinstitute.org/haploview/haploview>) [Barrett et al., 2005]. For each published study, concordance between original and new results was judged to be adequate (see Supplementary Table I).

Meta-Analysis SNPs

Of the SNPs passing QC thresholds described above, 18 were common to at least two of the published studies and/or the unpublished OCD mini-collaborative data. The SNPs genotyped in the five original studies are presented in Table I. Moreover, SNPs that were included in the recently published OCD GWAS [Stewart et al., 2012], and those included in a combined meta-analysis of multiple OCD candidate genes [Taylor, 2012], are identified. Genomic positions are those described in NCBI (National Center for Biotechnology Information) Build 36dbSNP Build 131/UCSC hg18hg19. Selection of SNPs for inclusion in this meta-analysis was based on whether a significant association, either as a single marker or as part of a haplotype, had been reported previously for the SNP and whether data were available from at least one additional study sample. Accordingly, nine SNPs were analyzed for this study. Data were coded in terms of forward strand orientation and then merged (with the exception of the case–control data) in PLINK [Purcell

et al., 2007] using the default consensus calling setting. Among studies whose participants were later re-genotyped by the OCD Mini-Collaborative, overall concordance rates exceeded 99% in all cases. Additionally, QC checks were performed on the merged final dataset to ensure that Mendel error rates and Hardy–Weinberg *P*-values remained within the acceptable ranges.

The linkage disequilibrium (LD) pattern of the 9 *SLC1A1* loci, which span a 22.5 kb region, was established using genotype data from individuals in this meta-analysis. Haplotype construction was performed in Haploview [Barrett et al., 2005]. The LD structure of these 9 SNPs is illustrated in Figure 1.

Meta-Analysis of Family-Based and Case–Control Single-Marker Associations

Four family-based association studies and one case–control study were included in this meta-analysis. Table II presents characteristics for each of the five studies. The total sample size for this meta-analysis included 815 complete trios and 306 cases with 634 ethnicity- and sex-matched controls. Males comprised 43.9% of OCD-affected probands ($N = 358$) and 43.5% of cases ($N = 133$). The original five studies had reported a combined total of 12 significant single-marker associations between *SLC1A1* polymorphisms and OCD. Among these SNPs, four were found to be associated exclusively in males, and only one was found to be associated in a global sample but not in the corresponding male subsample (see Table I).

In the present meta-analysis, one of the nine examined SNPs (rs301443) had a nominally significant uncorrected *P*-value ($P = 0.046$). This association did not survive Bonferroni correction for nine independent tests (required Bonferroni adjusted $P < 0.0056$) (see Table III).

Secondary Analyses by Gender Stratification

After stratifying the family and case–control datasets by gender of the affected proband, a secondary meta-analysis was performed on the male subsample. Only rs12682807 had a nominally significant uncorrected *P*-value ($P = 0.012$). This association did not survive multiple testing correction (see Table III).

DISCUSSION

To date, *SLC1A1* has been considered the strongest biological candidate gene for OCD based on positional linkage evidence [Hanna et al., 2002; Willour et al., 2004a; Nicolini et al., 2009], functional evidence connecting glutamate and OCD, and relatively consistent genetic association findings across samples. No other OCD candidate gene has demonstrated this level of supportive evidence. In the absence of replication at the polymorphism level, however, the association between *SLC1A1* and OCD has not been established. By integrating all relevant published and unpublished data, this study was able to provide an overall estimate of effect size for each of the nine *SLC1A1* polymorphisms that had been significantly associated with OCD in at least one previous study, and for which data were available from more than one site. Although each of these SNPs had demonstrated association with the OCD phenotype

TABLE I. 3' SLC1A1 Findings From Association, GWAS and Meta-Analysis OCD Studies

3' SLC1A1 SNPs investigated

SNPs	Original study findings included in meta-analysis							Other studies	
	HG18 position	Arnold et al.	Dickel et al.	Stewart et al.	Shugart et al.	Wendland et al.	MiniCollaborative (unpublished)	GWAS	Taylor meta-analysis
rs301430	chr9:4566680	x	Hm $P=0.003$	H perm $P=0.0043$; Hm perm $P=0.033$	x	H $P=0.04$; Hm $P=0.03$; E	x	x	x
rs3780412	chr9:4562480	x	Sm corr $P=0.036$	Sm perm $P=0.045$	x	x	x	x	x
rs3087879	chr9:4576808	S perm $P=0.22$; Sm perm $P=0.03$; H $P=0.003$; Hm $P=0.006$		x	x	H $P=0.04$; Hm $P=0.03$	x		x
rs301434	chr9:4572082	S perm $P=0.006$; Sm perm $P=0.01$; H $P=0.003$; Hm $P=0.006$	x	x	x	x	lost in QC	x	x
rs301979	chr9:4566851	x	Hm $P=0.003$	x	x		x		x
rs301435	chr9:4572843	x		x	x				
rs301443	chr9:4584919		x		S corr $P=0.017$		x		
rs12682807	chr9:4564022			H perm $P=0.0043$; Hm perm $P=0.033$			x		
rs2072657	chr9:4566451			H perm $P=0.0043$; Hm perm $P=0.033$			x		
rs2228622 (formerly s7856209)	chr9:4554432	x		H perm $P=0.0043$; Sm perm $P=0.045$	x		x		
rs7858819	chr9:4549892					H $P=0.04$; Hm $P=0.03$; E			
rs3933331	chr9:4379941					E	x		
rs1980943	chr9:4504246	x	x		x		x		
rs3780415	chr9:4545573	x			x		x		
rs10974625	chr9:4551596				x		x		
rs10814987	chr9:4473303		x				x		
rs10739062	chr9:4492848		x		x		x		
rs10115600	chr9:4509369		x		x				
rs3780413	chr9:4557353								
rs301440	chr9:4580305				x				
rs301437	chr9:4577560				x				

Single nucleotide polymorphisms (SNPs) listed by rs# include those examined in previous OCD studies of 3' SLC1A1. The base pair location for each SNP is listed in the column to the right of the SNP column (HG18 Position). The nine SNPs included in this meta-analysis are listed in the top nine rows. Five columns are provided for previously published study findings, one is provided for previously unpublished findings, one is provided for a published GWAS and one is provided for the only other OCD genotype meta-analysis. Blank cells within rows indicate the absence of that SNP from the related study. Cells with an 'x' indicate that the SNP was genotyped but that the P -value was >0.05 . Cells with a P -value <0.05 are listed according to the type of related association test: S, significant single marker association; Sm, significant single marker association in males; H, part of significant haplotype; Hm, part of significant haplotype in males; corr P , significant single marker association with OCD after correction for multiple testing; perm P , significant single marker association with OCD after permutation testing; E, SLC1A1 eQTL (expression quantitative trait locus).

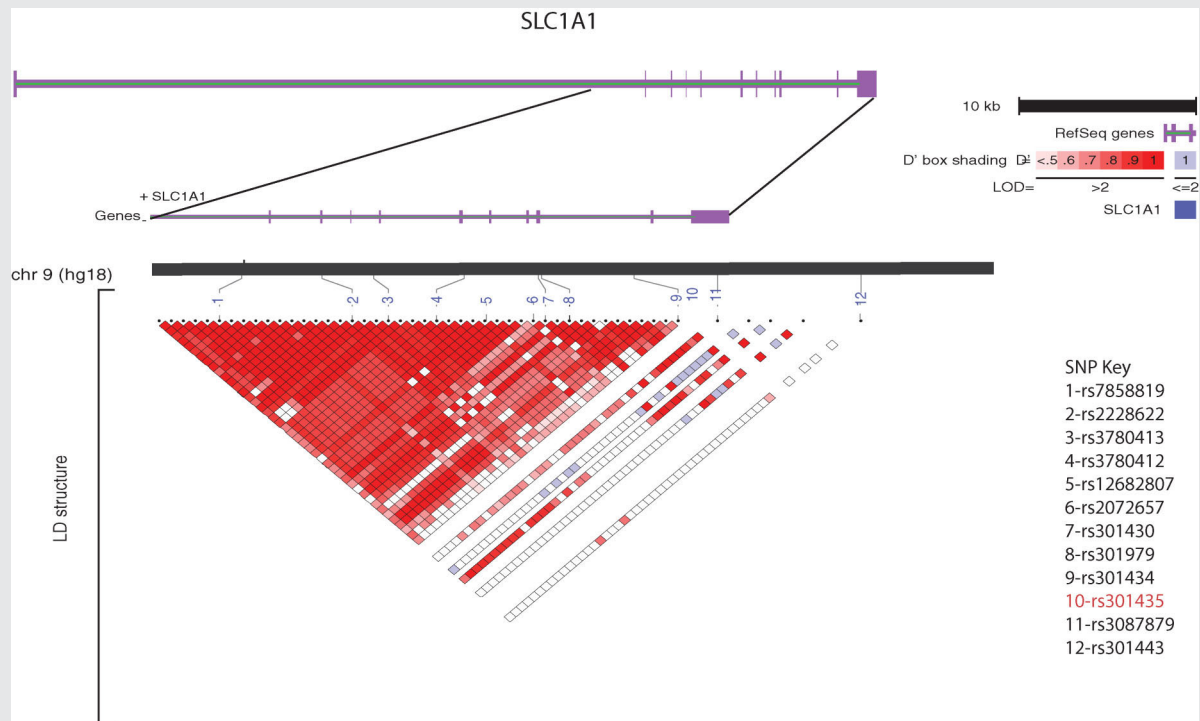


FIG. 1. *SLC1A1* SNPs included for meta-analysis and their previously reported single marker or haplotype OCD associations. **Single marker associations:** SNP 4—Dickel et al., 2006, Stewart et al., 2007; SNP 9—Arnold et al., 2006; SNP 11—Arnold et al., 2006; SNP 12—Shugart et al., 2009. **Haplotype associations:** SNPs 5, 6, 7—Stewart et al., 2007; SNPs 7, 8—Dickel et al., 2006; SNPs 7, 11—Wendland et al., 2009; SNPs 9, 11—Arnold et al., 2006.

in one or more of the original studies, most of these associations were no longer significant when data were combined across studies. This meta-analysis found only two nominally significant associations—rs301443 in the global sample ($P=0.046$, fixed-effects pooled OR = 1.17) and rs12682807 in the male sample ($P=0.012$, fixed-effects pooled OR = 1.65)—among the nine examined *SLC1A1* polymorphisms and OCD.

The greatest strength of this meta-analysis lies in its sample size. It represents a significant increase in power as compared to previous studies. The smallest of the original *SLC1A1* studies contained 66 families [Stewart et al., 2007a] and the largest contained 378 families [Shugart et al., 2009]. The combined sample for this meta-analysis included 815 affected offspring trios and 306 cases with 634 controls. Thus, power to detect an association was markedly increased by combining these samples. The only comparable study, a meta-analysis of polymorphisms from multiple candidate genes in OCD [Taylor, 2012], included five of the *SLC1A1* polymorphisms that were examined by us but excluded data from the large study by Shugart et al. [2009] and also excluded rs3014443 (our major finding). Taylor reported a nominally significant association with the 3' SNP rs3087879, which was not found in our meta-analysis.

Additionally, we were able to conduct a pooled analysis, as authors of all five of the original studies generously agreed to provide their original, pre-QC genotype data. Pooled analyses produce similar results to meta-analyses [Lin and Zeng, 2010],

however they have the advantages of allowing for more consistent quality control and facilitating secondary analyses [Lin and Zeng, 2010; Consortium, 2012]. The previously mentioned meta-analysis by Taylor [2012] was limited by using only data reported in earlier publications, thus necessitating exclusion of studies for which there was insufficient data to compute odds ratios.

Furthermore, our pooled analysis was generally robust against potential bias from population stratification because most of the data were from family-based association studies. Lastly, all sites contributing data to our study used well-validated phenotyping methods based upon standard DSM criteria, decreasing the risk for a dilution effect emerging from misdiagnosis and inappropriate study inclusion.

A secondary analysis was conducted on males to test for a sex-specific association between *SLC1A1* and OCD, given reports of this in all but one of the original published studies, including all four family-based association studies [Arnold et al., 2006; Dickel et al., 2006; Stewart et al., 2007a; Shugart et al., 2009]. The majority of reported single-marker associations in *SLC1A1* were driven by males, which is consistent with the sex differences observed in phenotypic presentations of OCD. The course of the disorder is more chronic in males [Ravizza et al., 1997; Bogetto et al., 1999], who are also more likely to experience onset during childhood. Males are also more likely to present with comorbid conditions such as tic disorder or attention-deficit/hyperactivity disorder [Fischer et al., 1996; Lensi et al., 1996; Bogetto et al., 1999; Zohar et al., 1999;

TABLE II. Characteristics of Previous 3' SLC1A1 Candidate Gene Studies

Study	Design	Number of participants	Race/ethnicity	Number of affected males (%)	Mean OCD onset (SD)	Diagnostic criteria	# of SNPs analyzed	Analysis method
Wendland et al. [2009]	Case-control	325 OCD probands, 662 matched controls	Caucasian	139 (43%)	14 y (9.6)	DSM-IV	6	PLINK
Arnold et al. [2006]	Family-based	157 families	96% Caucasian	60 (38%)	14.4 y (9.20)	DSM-IV	9	Single-locus in FBAT; TDT in Haploview
Stewart et al. [2007]	Family-based	66 Families	Not provided	47 (72.3%)	7.9 y (5.9)	DSM-IV	13	FBAT
Shugart et al. [2009]	Family-based	378 Families	97% Caucasian	325 (36%) (Note: #s are for total study population)	9.3 y	DSM-IV	13	PBAT
Dickel et al. [2006]	Family-based	71 Families	Not provided	25 (35%)	Two groups ascertained: 8.8 y (3.9) 8 y (3.7)	DSM-III-R or DSM-IV	9	Single-locus in FBAT; TDT in Haploview

Matsunaga et al., 2000; Tukul et al., 2004]. Segregation analyses have also revealed evidence of sexual dimorphism in the inheritance of OCD [Nestadt et al., 2000a; Hanna et al., 2005a]. However, the present meta-analysis provided limited evidence of a sex-specific association between *SLC1A1* and OCD. While the total combined OCD-affected sample was less than half (43.8%) male, resulting in a significant reduction in power for the secondary analysis, this meta-analysis is nevertheless the most comprehensive exploration of sex effects on the *SLC1A1*-OCD association to date.

The largely negative results of this meta-analysis are not entirely unexpected, given the findings from previous individual studies. Only three of the nine markers included in this meta-analysis were reported to be significant, either as a single marker or as part of a haplotype, in more than one of the previous studies. The only SNP that had been positively associated with *SLC1A1* as a single marker in more than one study was rs3780412, and this SNP was not significant in either the global or male-specific analysis of the present meta-analysis. This study does not, therefore, clarify the findings from previous association studies involving *SLC1A1* and OCD. The lack of association suggests that previous findings may have been false positive results. However, other limitations of this study may also explain the lack of an association in this meta-analysis. First, despite the relatively large sample size compared to earlier, single site studies, this meta-analysis is admittedly small relative to current guidelines for study size. The Psychiatric GWAS Consortium Coordinating Committee [2009] reported that the majority of recent significant association findings involve alleles with genotypic relative risks of 1.1–1.4, and that detection of such risk alleles may require samples of between 8,000 and 20,000 cases [Cichon et al., 2009].

Second, heterogeneity at both the genetic and phenotypic level may have impeded our ability to detect an association. Sample size considerations are especially relevant in light of the genetic heterogeneity thought to underlie OCD [Liang et al., 2008]. For example, while the first linkage scan for OCD [Hanna et al., 2002; Willour et al., 2004b] reported a linkage signal on chromosome 9p24 that was subsequently replicated [Willour et al., 2004a], subsequent linkage studies using primarily sibling pairs [Shugart et al., 2006a; Hanna et al., 2007; Ross et al., 2011; Mathews et al., 2012] failed to detect this signal. Furthermore, in the linkage study by Willour et al. [2004a] the group noted that only 59% of their pedigrees had demonstrated evidence for linkage with the 9p24 candidate region. Linkage studies are robust to allelic heterogeneity [Risch, 2000] but not to locus heterogeneity, and both forms of heterogeneity will contribute to poor replication across association studies.

Marked heterogeneity of the OCD phenotype has also been consistently identified as a formidable challenge in genetic studies of the disorder. The present meta-analysis combined early-onset OCD cases with adult-onset OCD cases, despite the assertion that genetic etiology of early-onset OCD may be distinct [Walitza et al., 2010]. In a large family study of OCD, Nestadt et al. [2000b] found that earlier age at onset was strongly associated with familiarity and saw no evidence of familiarity when age at onset was older than 17 years. Additionally, a meta-analysis of twin studies in OCD performed by van Grootheest et al. [2005] found dramatically higher heritability in the early-onset form of the disorder, from 0.45 to 0.65, whereas heritability in adult-onset cases ranged from

TABLE III. Meta-analysis Results for Global and Male-only Samples

SNP	MAF		Global sample numbers								Z-scores			OR		Global meta-analysis results				
	Cases	Ctrls	Hap map	Allele	# Trios	# Cases:ctrls	Trio T:U	Trio wt.	c:c wt.	Trios	c:c	Trios	c:c	Trios	c:c	Meta-z	χ^2	P-value	OR	L 95%CI
rs3780412	0.45	0.446	0.425 C	C/T	813	306:634	420: 425	657	412	-0.172	0.251	0.988	1.025	0.021	0.000	0.983	1.411	-31.22	34.04	
rs12682807	0.097	0.088	0.133 C	C/A	747	0	129: 109	604	0	1.297	1.183	0.975	0.983	1.297	1.681	0.195	1.183	0.929	1.437	
rs2072657	0.308	0.313	0.317 G	G/T	728	0	278: 285	589	0	-0.295	0.538	1.016	1.06	-0.295	0.087	0.768	0.975	0.810	1.141	
rs301430	0.285	0.284	0.325 C	C/T	805	306:634	325: 320	651	412	0.197	0.953	0.982	0.837	0.489	0.239	0.625	1.455	-0.048	2.958	
rs301979	0.294	0.313	0.292 G	G/C	798	0	325: 341	645	0	-0.62	0.953	0.982	0.837	-0.62	0.384	0.535	0.953	0.801	1.105	
rs301434	0.457	0.456	0.392 T	T/C	641	306:634	323: 329	518	412	-0.235	-1.795	0.982	0.837	-1.370	1.878	0.171	1.290	0.926	1.654	
rs301435	0.468	0.458	0.347 C	C/T	569	0	291: 296	460	0	-0.206	0.983	0.983	0.983	-0.206	0.043	0.836	0.983	0.821	1.145	
rs3087879	0.337	0.344	0.325 C	C/G	792	306:634	359: 363	640	412	-0.149	-2.217	0.989	0.792	-1.503	2.260	0.133	1.267	0.958	1.575	
rs301443	0.327	0.310	0.258 C	C/G	778	0	351: 300	629	0	1.999	1.17			1.999	3.995	0.046	1.17	1.016	1.324	

SNP	MAF		Male-only sample numbers								Z-scores			OR		Male-only meta-analysis results				
	Cases	Ctrls	Hap map	Allele	# Trios	# Cases:ctrls	Trio T:U	Trio wt.	c:c wt.	Trios	c:c	Trios	c:c	Trios	c:c	Meta-z	χ^2	P-value	OR	L 95%CI
rs3780412	0.436	0.450	0.425 C	C/T	357	133:290	171: 191	289	182	-1.051	-0.033	0.895	0.995	-0.844	0.712	0.399	1.320	0.675	1.964	
rs12682807	0.104	0.088	0.133 C	C/A	326	0	66:40	264	0	2.525	1.65	0.95	0.95	2.525	6.377	0.012	1.65	1.261	2.039	
rs2072657	0.301	0.309	0.317 G	G/T	313	0	114: 120	253	182	-0.392	1.190	1.099	1.212	-0.392	0.154	0.695	0.95	0.694	1.206	
rs301430	0.293	0.28	0.325 C	C/T	350	133:290	156: 142	283	182	0.811	0.940	1.008	0.870	1.377	1.897	0.168	1.616	0.933	2.298	
rs301979	0.297	0.306	0.292 G	G/C	347	0	131: 145	281	182	-0.843	0.903	1.008	0.870	-0.843	0.710	0.399	0.903	0.667	1.140	
rs301434	0.455	0.457	0.392 T	T/C	243	133:290	128: 127	197	182	0.063	-0.940	1.008	0.870	-0.606	0.367	0.544	1.329	0.409	2.250	
rs301435	0.459	0.462	0.347 C	C/T	215	0	113: 116	174	182	-0.198	0.974	0.974	0.974	-0.198	0.039	0.842	0.974	0.715	1.233	
rs3087879	0.332	0.343	0.325 C	C/G	348	133:290	148: 156	282	182	-0.459	-1.563	0.949	0.781	-1.337	1.787	0.181	1.229	0.927	1.531	
rs301443	0.330	0.312	0.258 C	C/G	340	0	163: 134	275	182	1.683	2.832	1.216	1.216	1.683	2.832	0.092	1.216	0.988	1.444	

Single nucleotide polymorphisms (SNPs) listed by rs# included in meta-analysis. In the columns to the right, minor allele frequencies (MAF) are presented for our case samples, for control samples and as reported in HepMap. The minor and major alleles are presented as oriented to the forward strand. Characteristics of the sample size for each SNP are presented including the number of affected offspring (trios) as calculated in PLINK, the number of cases and controls (case-control). The ratio of transmitted to untransmitted minor alleles for trios are presented as the count of transmitted minor alleles: Count of untransmitted minor alleles. Weights for the trio and case-control samples are provided in the columns to the right. Calculated values for the final meta χ^2 and P-values are included, in addition to odds ratios (OR) for trios (as calculated in PLINK), for the case-control samples (c:c) and for the entire sample (meta-OR). The lower and upper limits of the 95% confidence interval (CI) are in the columns to the far right. The rows at the top of the table describe findings for the global sample, and those at the bottom describe findings for the male-only sample.

Bolded rows indicate final meta-analysis $P < 0.05$.

0.27 to 0.47. The results of this meta-analysis also suggest the potential importance of subphenotypic considerations for future genetic studies of OCD. Factor analytic studies have repeatedly demonstrated that OCD is comprised of discrete but overlapping symptom dimensions [Mataix-Cols et al., 2005; Stewart et al., 2007b; Stewart et al., 2008]. It has been proposed that specific symptom dimensions may be more closely tied to underlying genetic etiologic factors than the entire phenotypic spectrum of the disorder, such that different candidate genes may be responsible for phenotypic expression of distinct symptom dimensions [Pato et al., 2002]. Ideally, a meta-analysis of OCD would stratify by symptom dimensions for secondary analyses. However, the present study would not have been adequately powered to detect *SLC1A1* association with putative age-of-onset or symptom-based OCD subphenotypes [Wendland et al., 2009; Samuels et al., 2011].

In addition to the sample size and heterogeneity, a third potential limitation of this study is the particular analytic method selected. Numerous analytic approaches exist for combining case-control and family-based data [Infante-Rivard et al., 2009], including extensions of likelihood models and combinations of distinct estimates. For the latter category, which was employed in this meta-analysis, two separate analyses were first performed; using a chi-square test in the case-control dataset and a TDT in the merged trio dataset. Kazeem and Farrall [2005] described the method of integrating case-control and TDT studies using odds ratios, which can be weighted since the logarithm of the odds ratio is asymptotically normally distributed. Nicodemus [2008] described some of the computational packages designed to carry out such analyses, but for the current meta-analysis we determined that the easiest and most reliable method of combining the estimates of disease-marker association was the weighted *z*-score, as described by de Bakker et al. [2008]. The *z*-score can be easily derived from the chi-square and odds ratio, both of which were provided in the PLINK output. Because the odds ratio was not provided in the FBAT output, PLINK was a much more attractive program than FBAT for analyzing the family-based data. While this study employed a standardized approach across all analyses, a different approach could have been used, which may have altered the results slightly. This slight lack of precision is particularly relevant to the present study, in which markers are demonstrating only modest association to the disease phenotype.

Finally, the current meta-analysis of *SLC1A1* is premised on potentially incorrect assumptions about its underlying genetic architecture. We examined only common polymorphisms, when it may be the case that multiple rare variants within *SLC1A1* contribute to the OCD phenotype. Detecting such rare variants would require a sequencing approach. To date, studies searching for rare variation in *SLC1A1* [Veenstra-VanderWeele et al., 2001, 2012; Wang et al., 2010] have been limited in scope to coding regions and have primarily used capillary-based screening methods, rather than targeting the entire gene via next-generation sequencing methods. Also, if there is substantial allelic heterogeneity for *SLC1A1* and OCD, haplotype [Lambert et al., 2012] or gene-based [Huang et al., 2011; Bacanu, 2012] analytic methods that consider combined effects of SNPs may be more likely to identify associations. Finally, it is possible that insertion/deletion or copy number variants were contributors to the previously identified OCD associations with

SLC1A1. It is interesting in this regard that our strongest findings was for rs301443, which lies within a small deletion detected by Dickel et al. [2006].

In conclusion, the glutamate transporter gene *SLC1A1* has been a very promising candidate gene for OCD, supported by various lines of biological evidence in addition to multiple genetic association studies. However, this meta-analysis does not support a strong role for common *SLC1A1* risk variants in OCD etiology, nor does it convincingly establish a role for gender in the overall OCD risk liability of *SLC1A1* polymorphisms. On the other hand, it is possible that the signals observed in this analysis are legitimate and represent a true association between *SLC1A1* and OCD. Regardless, it is important to note that the results do not undermine the potential contribution of glutamatergic dysregulation to OCD pathology. Investigating the roles of glutamatergic systems in OCD remains warranted and is relevant to both clinical practice and future research. The results from this study demonstrate the need for larger collaborative samples and subphenotyping, and highlight the potential for false-positive findings, or at least the need for cautious interpretation of positive findings, with candidate gene studies. If heterogeneity of the *SLC1A1* gene comprises both common and rare variants, then next-generation sequencing may also be productive in future studies of *SLC1A1* and OCD [Wang et al., 2010; Bailey et al., 2011].

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